# Aleutian Islands Golden King Crab Model-Based Stock Assessment 

May 2018 Crab SAFE DRAFT REPORT

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## Executive Summary

1. Stock

Golden king crab, Lithodes aequispinus, Aleutian Islands, east of $174^{\circ} \mathrm{W}$ longitude (EAG) and west of $174^{\circ} \mathrm{W}$ longitude (WAG).

## 2. Catches

The Aleutian Islands golden king crab commercial fishery has been prosecuted since 1981/82 and opened every year since then. Retained catch peaked in 1986/87 at 2,686 t (5.922,425 lb) and $3,999 \mathrm{t}(8,816,319 \mathrm{lb})$, respectively, for EAG and WAG, but the retained catch dropped sharply from 1989/90 to 1990/91. The fishery has been managed separately east (EAG) and west (WAG) of $174^{\circ} \mathrm{W}$ longitude since 1996/97 and Guideline Harvest Levels (GHLs) of $1,452 \mathrm{t}(3,200,000 \mathrm{lb})$ for EAG and $1,225 \mathrm{t}(2,700,000 \mathrm{lb})$ for WAG were introduced into management for the first time in 1996/97. The GHL was subsequently reduced to $1,361 \mathrm{t}$ (3,000,000 lb beginning in 1998/99 for EAG. The reduced GHLs remained at $1,361 \mathrm{t}$ (3,000,000 lb) for EAG and $1,225 \mathrm{t}(2,700,000 \mathrm{lb})$ for WAG through 2007/08, but were increased to $1,429 \mathrm{t}(3,150,000 \mathrm{lb})$ for EAG and $1,294 \mathrm{t}(2,835,000 \mathrm{lb})$ for WAG beginning with the 2008/09 fishing season following an Alaska Board of Fisheries (BOF) decision. The acronym changed from GHL to TAC (Total Allowable Catch) since crab rationalization in 2005/06. The TACs were further increased by another BOF decision to $1,501 \mathrm{t}(3,310,000$ lb) for EAG and $1,352 \mathrm{t}(2,980,000 \mathrm{lb})$ for WAG beginning with the $2012 / 13$ fishing season.

Catches have been steady since the introduction of GHL/TAC and the fishery has harvested close to TAC levels since 1996/97. These TAC levels were below the ABCs determined under Tier 5 criteria (considering 1991-1995 mean catch for the whole Aleutian Islands region, $3,145 \mathrm{t}(6,933,822 \mathrm{lb})$, as the limit catch) under the most recent crab management plan. The below par fishery performance in WAG in recent years lead to reduction in TAC to $1,014 \mathrm{t}(2,235,000 \mathrm{lb})$, which reflected a $25 \%$ reduction on the TAC for WAG, while the TAC for EAG was kept at the same level, $1,501 \mathrm{t}(3,310,000 \mathrm{lb})$ for the 2015/16 through 2017/18 fishing seasons. In addition to the retained catch that is allotted as TAC, there was retained catch in a cost-recovery fishery towards a $\$ 300,000$ goal in 2013/14 and 2014/15, and towards a $\$ 500,000$ goal in 2015/16 and 2016/17.

Catch per pot lift (CPUE) of retained legal males decreased from the 1980s into the mid1990s, but increased steadily after 1994/95 and increased markedly at the initiation of the

Crab Rationalization program in 2005/06. Although CPUE for the two areas showed similar trends through 2010/11, during 2011/12-2014/15 CPUE trends have diverged (increasing EAG and decreasing WAG). Total retained catch in 2016/17 was $2,593 \mathrm{t}(5,716,180 \mathrm{lb})$ : $1,578 \mathrm{t}(3,479,529 \mathrm{lb})$ from the EAG fishery, which included cost-recovery catch, $1,015 \mathrm{t}$ $(2,236,651 \mathrm{lb})$ from the WAG fishery. Discarded (non-retained) catch occurs mainly during the directed fishery. Although low levels of discarded catch can occur during other crab fisheries, there have been no such fisheries prosecuted since 2004/05, except as surveys for red king crab conducted under a commissioner's permit (and there were none caught during the cooperative red king crab survey performed by industry and ADF\&G in the Adak area in September 2015 (Hilsinger et al. 2016). Estimates of the bycatch mortality during crab fisheries decreased during 1995/96-2005/06, both in absolute value and relative to the retained catch weight, and stabilized during 2005/06-2014/15. Total estimated bycatch mortality during crab fisheries in 2016/17 was $138 \mathrm{t}(303,832 \mathrm{lb})$ for EAG and $92 \mathrm{t}(202,815$ lb) for WAG. Discarded catch also occurs during fixed-gear and trawl groundfish fisheries, but is small relative to that during the directed fishery and the groundfish fisheries are a minor contributor to total fishery mortality. Estimated bycatch mortality during groundfish fisheries in 2016/17 was $3 \mathrm{t}(6,245 \mathrm{lb})$ for EAG and $3 \mathrm{t}(6,800 \mathrm{lb})$ for WAG. A cooperative golden king crab survey was performed by the Aleutian Islands King Crab Foundation (an industry group) and ADF\&G during the EAG fishery in August 2016, by vessels that were simultaneously fishing. During the survey work, adjustments were made to a portion of the gear so escape mechanisms were no longer functional. However, for the purpose of catch accounting for $2016 / 17$, it was assumed that bycatch mortality that occurred during the survey was accounted for by reported discards for the 2016/17 EAG fishery. The cooperative survey was also conducted in August 2017 during the 2017/18 EAG fishery.

## 3. Stock biomass

Estimated mature male biomass (MMB) for EAG under all scenarios decreased from high levels during the 1990s, then systematically increased during the 2000s and 2010s. Estimated MMB for WAG decreased during the late 1980s and 1990s, systematically increased during the 2000s, and decreased for a number of years since 2009. The low levels of MMB for EAG were observed in 1995-1997 and in 1990s for WAG. Slightly increasing trends in MMB were observed since 2014 in both regions. Stock trends reflected the fishery standardized CPUE trends in both regions.

## 4. Recruitment

The numbers of recruits to the model size groups under all scenarios have fluctuated in both EAG and WAG. For EAG, the model recruitment was high in 1987, 1988, 2008, 2015, 2017, and highest in 2014; and lowest in 1986. An increasing trend in recruitment was observed since the early-1990s in EAG. The model recruitment for WAG was high during 1983 to 1987 and highest in 2015; and lowest in 2011. After 1983 to 1987 peaks, the recruitment trend was low except the 2015 highest recruitment.

## 5. Management performance

The model was accepted at the September 2016 CPT and October 2016 SSC meetings for OFL determination for the 2017/18 fishery cycle. In addition, the CPT in January 2017 and SSC in February 2017 recommended using the Tier 3 method to compute OFL and ABC. The assessment model was first used for setting OFL and ABC for the 2017/18 fishing
season. The CPT in May 2017 and SSC in June 2017 accepted author's recommendation of using scenario 9 (i.e., model using the knife edge maturity to determine MMB) for OFL and ABC calculation. During the May 2017 meeting, the CPT noted that a single OFL and ABC are defined for Aleutian Islands golden king crab (AIGKC). However, separate models are available by area. Following last year's approach, we added OFLs and ABCs by area to calculate OFL and ABC for the entire stock. We could add them together without any modification because the stock status in the two areas after 2016/17 fishery was similar.

Among the six common scenarios for EAG and WAG, we recommend three scenarios (17_0 (base), 17_0d (three catchability and total selectivity), and 17_0e (McAllister and Ianelli method of re-weighting) for consideration and provide the status and catch specifications for the AIGKC stock. Scenario 17_0 is the base scenario with an updated $M$ of $0.21 \mathrm{yr}^{-1}$ and the addition of 2016/17 data. The model formulation is the same as that was accepted in 2017. Scenario 17_0d fits the recent three years' CPUE indices well for EAG, but the OFL and ABC are very low among the three selected scenarios. Scenario 17_0e is an alternative to the base scenario with McAllister and Ianelli method of size composition data weighting instead of Francis' method of reweighting. The OFL and ABC differences between 17_0e and 17_0 are small. The rest of the scenarios have some shortcomings either on adequacy of data or on model diagnostics; hence, are not considered. All scenarios assume the knife-edge maturity selection.

Status and catch specifications (1000 t) of Aleutian Islands golden king crab

| Year | MSST | Biomass <br> (MMB) | TAC | Retained <br> Catch | Total <br> Catch $^{\mathbf{a}}$ | OFL | ABC $^{\mathbf{b}}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $2013 / 14$ | N/A | N/A | 2.853 | 2.894 | 3.192 | 5.69 | 5.12 |
| $2014 / 15$ | N/A | N/A | 2.853 | 2.771 | 3.088 | 5.69 | 4.26 |
| $2015 / 16$ | N/A | N/A | 2.853 | 2.729 | 3.076 | 5.69 | 4.26 |
| $2016 / 17$ | N/A | N/A | 2.515 | 2.593 | 2.947 | 5.69 | 4.26 |
| $2017 / 18$ | 6.044 | 14.205 | 2.515 | 2.585 | 2.942 | 6.048 | 4.536 |
| $2018 / 19^{\text {c }}$ | 6.046 | 17.952 |  |  |  | 5.514 | 4.136 |
| $2018 / 19^{\text {d }}$ | 5.898 | 14.665 |  |  |  | 3.963 | 2.972 |
| $2018 / 19^{\text {e }}$ | 6.107 | 17.793 |  |  |  | 5.581 | 4.186 |

a. Total retained catch plus estimated bycatch mortality of discarded bycatch during crab fisheries and groundfish fisheries.
b. $25 \%$ buffer was applied to total catch OFL to determine ABC.
c. 17_0 base scenario with Francis method of re-weighting
d. $17 \_0 d$ three catchability and total selectivity scenario with Francis method of reweighting
e. $\quad 17$ _0e McAllister and Ianelli method of re-weighting

Status and catch specifications (million lb) of Aleutian Islands golden king crab

| Year | MSST | Biomass <br> (MMB) | TAC | Retained <br> Catch | Total <br> Catch $^{\mathbf{a}}$ | OFL | ABC $^{\mathbf{b}}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $2013 / 14$ | N/A | N/A | 6.290 | 6.38 | 7.038 | 12.54 | 11.28 |
| $2014 / 15$ | N/A | N/A | 6.290 | 6.11 | 6.807 | 12.53 | 9.40 |
| $2015 / 16$ | N/A | N/A | 6.290 | 6.016 | 6.782 | 12.53 | 9.40 |
| $2016 / 17$ | N/A | N/A | 5.545 | 5.716 | 6.497 | 12.53 | 9.40 |
| $2017 / 18$ | 13.325 | 31.315 | 5.545 | 5.699 | 6.487 | 13.333 | 10.000 |
| $2018 / 19^{\text {c }}$ | 13.329 | 39.577 |  |  |  | 12.157 | 9.118 |
| $2018 / 19^{\text {d }}$ | 13.002 | 32.331 |  |  |  | 8.737 | 6.553 |
| $2018 / 19^{\text {e }}$ | 13.464 | 39.227 |  |  |  | 12.305 | 9.228 |

a. Total retained catch plus estimated bycatch mortality of discarded bycatch during crab fisheries and groundfish fisheries.
b. $25 \%$ buffer was applied to total catch OFL to determine ABC.
c. 17_0 base scenario with Francis method of re-weighting
d. 17_0d three catchability and total selectivity scenario with Francis method of reweighting
e. 17_0e McAllister and Ianelli method of re-weighting

Since the $2017 / 18$ total catch of $2,942 \mathbf{t}(6.487$ million lb$)$ is below the OFL catch of $\mathbf{6 , 0 4 8} \mathbf{t}$ ( 13.333 million lb), "overfishing" did not occur in the Aleutian Islands golden king crab fishery in 2017/18.

## 6. Basis for the OFL

The length-based model developed for the Tier 3 analysis estimated MMB on February 15 each year for the period 1986 through 2016 and projected to February 15, 2018 for OFL and ABC determination. The Tier 3 approach uses a constant annual natural mortality $(M)$ and the mean number of recruits for the period 1987 - 2012 for OFL and ABC calculation. An $M$ of $0.21 \mathrm{yr}^{-1}$ derived from the combined data was used.

We provide the OFL and ABC estimates for EAG, WAG, and the two regions pooled together (i.e., for the entire Aleutian Islands, AI) for seven scenarios [17_0, 17_0a, 17_0b, 17_0c, 17_0d, 17_0e, and 17_0f (the last is only for EAG)] in the following six tables. As per September 2017 CPT suggestion, we also provide estimates for May 2017 CPT accepted scenario 9 (modified as $9 * *$ for WAG) in these tables. We treat scenario $17 \_0$ as the base scenario for EAG and WAG. We provide three options of OFL and ABC estimates based on scenarios 17_0, 17_0d, and 17_0e for CPT consideration and selection. Since the OFL and ABC have been set for the entire AI before, we suggest implementing the combined OFL and ABC for AI.

## EAG (Tier 3):

Biomass, total OFL, and ABC for the next fishing season in millions of pounds. Current MMB $=\mathrm{MMB}$ on 15 Feb .2018.
Current MMB for May2017Sc9 =MMB on 15 Feb. 2017.

| Scenario | Tier | MMB ${ }_{35 \%}$ | Current <br> MMB | $\begin{gathered} \text { MMB/ } \\ M M B_{35 \%} \end{gathered}$ | $F_{\text {OFL }}$ | Recruitment Years to define$M M B_{35 \%}$ | OFL |  |  | $\begin{gathered} \hline \mathrm{ABC} \\ (0.75 * \mathrm{OFL}) \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  |  |  | $\begin{gathered} \mathrm{ABC} \\ \left(\mathrm{P}^{*}=0.49\right) \end{gathered}$ |  |  |
|  |  |  |  |  |  |  | $F_{35 \%}$ |  |  |  |
| EAG17_0 | 3a | 15.332 | 25.474 | 1.66 | 0.64 | 1987-2012 | 0.64 | 8.637 | 8.601 | 6.478 |
| EAG17_0a | 3 a | 15.570 | 25.645 | 1.65 | 0.62 | 1987-2012 | 0.62 | 8.729 | 8.683 | 6.547 |
| EAG17_0b | 3a | 14.979 | 22.949 | 1.53 | 0.65 | 1987-2012 | 0.65 | 7.529 | 7.492 | 5.646 |
| EAG17_0c | 3 a | 15.633 | 25.869 | 1.65 | 0.62 | 1987-2012 | 0.62 | 8.920 | 8.872 | 6.690 |
| EAG17_0d | 3a | 14.745 | 17.986 | 1.22 | 0.64 | 1987-2012 | 0.64 | 5.469 | 5.435 | 4.102 |
| EAG17_0e | 3 a | 15.462 | 25.045 | 1.62 | 0.64 | 1987-2012 | 0.64 | 8.761 | 8.725 | 6.570 |
| EAG17_0f | 3 a | 15.312 | 25.340 | 1.65 | 0.64 | 1987-2012 | 0.64 | 8.581 | 8.545 | 6.436 |
| May2017Sc9 | 3 a | 15.539 | 20.515 | 1.32 | 0.75 | 1987-2012 | 0.75 | 9.890 | 9.852 | 7.417 |

Biomass in 1000 t ; total OFL and ABC for the next fishing season in t .

| Scenario | Recruitment |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Tier | MMB ${ }_{35 \%}$ | Current <br> MMB | $\begin{gathered} \text { MMB/ } \\ M M B_{35 \%} \end{gathered}$ | $F_{\text {OFL }}$ | Years to Define $M M B_{35 \%}$ | $F_{35 \%}$ | OFL | $\begin{gathered} \mathrm{ABC} \\ \left(\mathrm{P}^{*}=0.49\right) \end{gathered}$ | $\begin{gathered} \mathrm{ABC} \\ (0.75 * \mathrm{OFL}) \end{gathered}$ |
| EAG17_0 | 3a | 6.954 | 11.555 | 1.66 | 0.64 | 1987-2012 | 0.64 | 3,917.776 | 3,901.317 | 2,938.332 |
| EAG17_0a | 3a | 7.063 | 11.633 | 1.65 | 0.62 | 1987-2012 | 0.62 | 3,959.351 | 3,938.754 | 2,969.513 |
| EAG17_0b | 3a | 6.794 | 10.409 | 1.53 | 0.65 | 1987-2012 | 0.65 | 3,414.981 | 3,398.458 | 2,561.235 |
| EAG17_0c | 3a | 7.091 | 11.734 | 1.65 | 0.62 | 1987-2012 | 0.62 | 4,046.121 | 4,024.483 | 3,034.590 |
| EAG17_0d | 3a | 6.688 | 8.158 | 1.22 | 0.64 | 1987-2012 | 0.64 | 2,480.617 | 2,465.170 | 1,860.463 |
| EAG17_0e | 3a | 7.014 | 11.360 | 1.62 | 0.64 | 1987-2012 | 0.64 | 3,973.77 | 3,957.468 | 2,980.334 |
| EAG17_0f | 3a | 6.946 | 11.494 | 1.65 | 0.64 | 1987-2012 | 0.64 | 3,892.238 | 3,876.174 | 2,919.178 |
| May2017Sc9 | 3a | 7.048 | 9.306 | 1.32 | 0.75 | 1987-2012 | 0.75 | 4,486.052 | 4,468.684 | 3,364.539 |

WAG (Tier 3):
Biomass, total OFL, and ABC for the next fishing season in millions of pounds. Current $\mathrm{MMB}=\mathrm{MMB}$ on 15 Feb. 2018. Current MMB for May2017Sc9 =MMB on 15 Feb. 2017.

| Scenario | Tier | MMB ${ }_{35 \%}$ | Recruitment |  |  |  |  |  |  | $\begin{gathered} \mathrm{ABC} \\ (0.75 * \mathrm{OFL}) \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Current MMB | $\begin{gathered} \mathrm{MMB} / \\ M M B_{35 \%} \end{gathered}$ | $F_{\text {OFL }}$ | Years to Define $M M B_{35 \%}$ | $F_{35 \%}$ | OFL | $\begin{gathered} \mathrm{ABC} \\ \left(\mathrm{P}^{*}=0.49\right) \end{gathered}$ |  |
|  |  |  |  |  |  |  |  |  |  |  |
| WAG17_0 | 3a | 11.327 | 14.103 | 1.25 | 0.60 | 1987-2012 | 0.60 | 3.520 | 3.505 | 2.640 |
| WAG17_0a | 3a | 11.405 | 14.148 | 1.24 | 0.59 | 1987-2012 | 0.59 | 3.503 | 3.489 | 2.627 |
| WAG17_0b | 3a | 11.252 | 13.391 | 1.19 | 0.60 | 1987-2012 | 0.60 | 3.289 | 3.270 | 2.466 |
| WAG17_0c | 3a | 11.294 | 13.947 | 1.23 | 0.60 | 1987-2012 | 0.60 | 3.418 | 3.395 | 2.564 |
| WAG17_0d | 3a | 11.260 | 14.345 | 1.27 | 0.68 | 1987-2012 | 0.68 | 3.268 | 3.248 | 2.451 |
| WAG17_0e | 3a | 11.466 | 14.182 | 1.24 | 0.59 | 1987-2012 | 0.59 | 3.544 | 3.529 | 2.658 |
| May2017Sc9 | 3a | 9.937 | 10.800 | 1.09 | 0.68 | 1993-1997 | 0.68 | 3.443 | 3.428 | 2.582 |

Biomass in 1000 t ; total OFL and ABC for the next fishing season in t .

| Scenario | Tier | MMB ${ }_{35 \%}$ | Current <br> MMB | $\begin{gathered} \text { MMB / } \\ M M B_{35 \%} \end{gathered}$ | $F_{\text {OFL }}$ | Recruitment Years to Define $M M B_{35 \%}$ | $F_{35 \%}$ | OFL | $\begin{gathered} \mathrm{ABC} \\ \left(\mathrm{P}^{*}=0.49\right) \end{gathered}$ | $\begin{gathered} \hline \mathrm{ABC} \\ (0.75 * \mathrm{OFL}) \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| WAG17_0 | 3 a | 5.138 | 6.397 | 1.25 | 0.60 | 1987-2012 | 0.60 | 1,596.535 | 1,589.834 | 1,197.401 |
| WAG17_0a | 3a | 5.173 | 6.417 | 1.24 | 0.59 | 1987-2012 | 0.59 | 1,588.903 | 1,582.813 | 1,191.677 |
| WAG17_0b | 3a | 5.104 | 6.074 | 1.19 | 0.60 | 1987-2012 | 0.60 | 1,491.700 | 1,483.331 | 1,118.775 |
| WAG17_0c | 3 a | 5.123 | 6.326 | 1.23 | 0.60 | 1987-2012 | 0.60 | 1,550.509 | 1,540.027 | 1,162.882 |
| WAG17_0d | 3 a | 5.108 | 6.507 | 1.27 | 0.68 | 1987-2012 | 0.68 | 1,482.383 | 1,473.365 | 1,111.787 |
| WAG17_0e | 3a | 5.201 | 6.433 | 1.24 | 0.59 | 1987-2012 | 0.59 | 1,607.523 | 1,600.637 | 1,205.642 |
| May2017Sc9 | 3 a | 4.507 | 4.899 | $1 . .09$ | 0.68 | 1993-1997 | 0.68 | 1,561.668 | 1,554.794 | 1,171.251 |

Aleutian Islands (AI)
Total OFL and ABC for the next fishing season in millions of pounds.

| Scenario | OFL |  | ABC <br> $\left(\mathrm{P}^{*}=0.49\right)$ | ABC <br> $(0.75 * \mathrm{OFL})$ |
| :---: | ---: | ---: | ---: | ---: |
|  | $17 \_0$ | 12.157 | 12.106 | 9.118 |
|  | $17 \_0 \mathrm{a}$ | 12.232 | 12.172 | 9.174 |
|  | $17 \_0 b$ | 10.818 | 10.762 | 8.112 |
|  | $17 \_0 c$ | 12.338 | 12.267 | 9.254 |
|  | $17 \_0 d$ | 8.737 | 8.683 | 6.553 |
|  | 17_0e | 12.305 | 12.254 | 9.228 |
|  | May2017Sc9 | 13.333 | 13.280 | 9.999 |

Aleutian Islands (AI)
Total OFL and ABC for the next fishing season in t .

| Scenario |  | OFL | $\begin{array}{r} \mathrm{ABC} \\ \left(\mathrm{P}^{*}=0.49\right) \end{array}$ | $\begin{array}{r} \mathrm{ABC} \\ (0.75 * \mathrm{OFL}) \end{array}$ |
| :---: | :---: | :---: | :---: | :---: |
|  | 17_0 | 5,514.311 | 5,491.151 | 4,135.733 |
|  | 17_0a | 5,548.254 | 5,521.567 | 4,161.190 |
|  | 17_0b | 4,906.681 | 4,881.789 | 3,680.010 |
|  | 17_0c | 5,596.630 | 5,564.510 | 4,197.472 |
|  | 17_0d | 3,963.000 | 3,938.535 | 2,972.250 |
|  | 17_0e | 5,581.293 | 5,558.105 | 4,185.976 |
|  | May2017Sc9 | 6,047.720 | 6,023.478 | 4,535.790 |

## 7. Probability density functions of the OFL

Assuming a lognormal distribution of total OFL, we determined the cumulative distributions of OFL and selected the median as the OFL.
8. Basis for the $A B C$ recommendation

An $\mathrm{x} \%$ buffer on the OFL; i.e., $\mathrm{ABC}=(1.0-\mathrm{x} / 100) * \mathrm{OFL}$. We considered $\mathrm{x}=25 \%$.

See also the section G on ABC.
9. A summary of the results of any rebuilding analysis:

Not applicable.

## A. Summary of Major Changes

1. Changes (if any) to management of the fishery

- In 2017, proposed changes to OFL and ABC calculation under model-based Tier 3 assessment were accepted.

2. Changes to input data

- Commercial fisheries data were updated with values from the most recent ADF\&G Area Management report (Leon et al., 2017) and most recent fish ticket data. Fishery data has been updated with the catches during 2016/17: retained catch for the directed fishery and discarded catch estimates for the directed fishery, non-directed crab fisheries, and groundfish fisheries. Thus the time series of data used in the model are: retained catch (1981/82-2016/17), total catch (1990/91-2016/17), and groundfish bycatch (1989/90-2016/17) biomass and size compositions.
- Fish ticket retained CPUE were standardized by the GLM with the lognormal link function for the 1985/86-1998/98 period.
- Observer pot sample legal size crab CPUE data were standardized by the generalized linear model (GLM) with the negative binomial link function with variable selection by R square criterion and CAIC (modified AIC), separately for 1995/96-2004/05 and 2005/06-2015/16 periods.
- For scenario 17_0a, observer data were standardized by VAST. The work is still preliminary.
- For scenario 17_0f for EAG, independent pot survey data from 2015 to 2017 were standardized by GLM and a likelihood component with this set of indices was added.
- Chela height with carapace length data from ADFG (1991) and NMFS (1984) surveys were analyzed outside the assessment model to determine the knife-edge maturity for mature male biomass calculation.


## 3. Changes to assessment methodology

- The equilibrium initial population and Tier $3 M M B_{M S Y}$ reference point estimation procedures used the mean number of recruits for 1987-2012.
- Francis re-weighting method was used to update the input effective sample sizes for length composition data for most scenarios, including $M$ profiling and retrospective analysis except scenario 17_0e in which we applied the McAllister and Ianelli reweighting method (McAllister and Ianelli, 1997; Siddeek et al. 2016c, 2017).
- We also added a stock projection part (Appendix F) to assess the viability of the stock under Tier 3 OFL and ABC control rule and a dynamic B0 analysis part (Appendix $\mathrm{H})$ to assess the biomass dynamics under no fishery.


## 4. Changes to assessment results

Expectedly, addition of one more year data changed the OFL and ABC estimates, but no dramatic changes were observed.

## B. Response to September 2017 CPT comments

Comment 1: The CPT recommended moving forward with the modeling convention adopted by the Groundfish Plan Teams. Naming conventions in groundfish SAFE guidelines: When a model constituting a "major change" from the original version of the base model is introduced, it is given a label of the form "Model $y y . j$," where $y y$ is the year (designated by the last two digits) that the model was introduced, and $j$ is an integer distinguishing this particular "major change" model from other "major change" models introduced in the same year.

When a model constituting only a "minor change" from the original version of the base model is introduced, it is given a label of the form "Model $y y . j x$," where " $x$ " is a letter distinguishing this particular "minor change" model from other "minor change" models derived from the original version of the same base model.

The distinction between "major" and "minor" model changes is determined subjectively by the author on the basis of qualitative differences in model

Response:
We followed this naming convention in labeling model scenarios: 17_0 refers to model was established in 2017 and carried forward to 2018; no major changes occurred in 2018 and remain at the 0-level. 17_0a refers to a minor change to 17_0; for example, CPUE indices were determined by spatio-temporal delta generalized linear mixed model (deltaGLMM) instead of GLM in this case.

## Comment 2: a) Reconsider what crabs are mature vs immature via breakpoint analysis; b)

 Repeat the breakpoint analysis using $\log (\mathrm{CH} / \mathrm{CL})$ vs CL , rather than the $\log \mathrm{CH}$ vs. $\log \mathrm{CL}$; c ) Because it was based on an inappropriate analysis, there is no need to show models with a logistic maturity curve, unless an improved approach can be found.Response:
As suggested by Steve Martel, we used the $\log (\mathrm{CH} / \mathrm{CL})$ vs. CL plot to get a better delineation of points for breakpoint analysis (see Appendix C figures). We used the breakpoint $50 \%$ maturity length for maturity determination in all scenarios. Sizes $\geq 111 \mathrm{~mm}$ CL were treated as mature and below this breakpoint immature.

## Comment 3: It is appropriate to use only the equilibrium abundance as a starting point.

Response:
We used the equilibrium starting point in 1960 in all scenarios.

## Comment 4: Moving forward, do not look at the core data.

Response:
We are not using the core data, but we have analyzed the independent pot survey data to estimate CPUE indices and incorporated them in a separate model scenario (17_0f). In the future we intend to use a spatio-temporal model to analyze the independent pot survey data.

## Comment 5: Continue analysis of spatio-temporal variation of the fishery using a program

 like VAST.
## Response:

We did a preliminary analysis of observer data using a spatio-temporal deltaGLMM (VAST, Thorson et al. 2015) and estimated an additional set of CPUE indices (see Appendix B) for scenario 17_0a. VAST requires spatially explicit catch data and some measure of 'area fished'. This type of information is available from the observer data, which include soak time, lat. and long., and depth. The necessary data for a spatio-temporal detlaGLMM are not available from dock side sampling; therefore, observer data are more suitable (see West coast SSC's March 2017 groundfish subcommittee report on the review of assessment methodologies proposed for use in 2017 groundfish assessments).

However, unlike the open West Coast Sea or Bering Sea, the Aleutian Islands areas provide additional constraints for spatial analysis due to the edge effects from the many islands. More work is needed to improve the use of spatio-temporal models in this region.

## Comment 6: Show a scenario with the McAllister and Ianelli re-weighting for comparison when choosing preferred model.

Response:
We provide scenario 17_0e, which considers McAllister and Ianelli method of re-weighting (see Appendix D for detail).

## Comment 7: Consider interaction terms, specifically area $x$ year interaction for CPUE standardization.

Response:
We standardized the CPUE considering the year: area interaction for scenario 17_0c (see Appendix B for details). The problem with this interaction analysis is that a lot of NAs occurred for many missing factor levels over the years. Anyway, we used the resulting CPUE indices in scenario 17_0c.

Comment 8: Consider scenarios with catchability and/or total selectivity breaking at a third point in 2010 (or a better year).

Response:
We considered scenario 17_0d with different sets of catchability and total selectivity for 1985/86-2004/05; 2005/06-2012/13; and 2013/14-2016/17.

Comment 9: Provide a comparison between the previous CPUE standardization and any new standardization methods that are applied.

Response:



Figure Comm.9. Comparison between May 2017 and May 2018 CPUE indices (top: WAG, bottom: EAG). In 2017 we categorized the area broadly into 10 longitudinally separated regions whereas in 2018 we used individual ADFG coded statistical area. The confidence intervals are $+/-2$ SE. Model estimated additional standard error was added to each input standard error.

## Comment 10: Include last year's model as a scenario for consideration.

Response:
We have included last year's model as scenario May17Sc9 to reflect scenario 9 with knife-edge maturity selectivity, which was accepted last year.

Comment 11: Overall model recommendation for May 2018: base model from last year (equilibrium initial abundance, knife edge maturity, both CPUE analyses with any significant interaction terms).

Response:
Done.

Response to October 2017 SSC comments
Comment 1: The SSC appreciates the CPT's consideration of model number convention and their recommendation to move forward with the modelling convention adopted by the Groundfish Plan Teams.

Response:
Done.

Comment 2: Although the use of chela height-carapace size regression lines has been validated for Chionoecetes crabs (snow, Tanner), the SSC expressed concern that the use of this approach to determine maturity may not be appropriate for lithodid (king) crabs. The SSC recommends that efforts be made to verify this relationship in lab or field experiments, as well as to review the available literature and application of this approach for other non-Chionoecetes species.

Response:
After analyzing a number of lithodid (king) crab stocks for size at maturity, Somerton and Otto (1986) observed that golden king crab provided a better separation of chela height growth at the onset of maturity than either red or blue king crabs (see Appendix C). We have also provided a literature review on king crab maturity determination in Appendix C, which supports the breakpoint type of analysis for male $50 \%$ maturity determination.

Comment 3: The SSC supports the exploration of the VAST geospatial model for investigation of fishery catch rate data, but cautions that the nonrandom nature of fisheries data adds an additional challenge to the standard assumptions of independence between the underlying density and the process of observation beyond that of standard statistically-designed survey programs.

Response:
We did a preliminary run of VAST for observer CPUE standardization and described its advantage and limitation (see response to CPT comment 5).

Comment 4: The SSC encourages the author to explore observer data and to discuss with the participants in the fishery potential changes in fisher behavior that may influence the relationship between fishery catch rates and crab abundance.

Response:
This is an ongoing process. We continue to explore this with the industry input and external experts.

Comment 5:The SSC reiterates previous concerns that this stock assessment relies solely on fishery data, and therefore carries a higher degree of uncertainty than other model-based assessments for crab stocks. The SSC encourages recent and future efforts by the industry to include survey pots in their fishing activity in order to generate additional data to inform this analysis. The SSC extends its appreciation to the industry for their generous cooperative research efforts on this important crab stock.

Response:
We recognized the higher degree of uncertainty in the assessment and therefore set the ABC using $25 \%$ buffer level. For the first time, we used the independent pot survey data in the model even though the time series is short (2015 to 2017).

## C. Introduction

1. Scientific name:

Golden king crab, Lithodes aequispinus J.E. Benedict, 1895.

## 2. Distribution:

General distribution of golden king crab is summarized by NMFS (2004). Golden king crab, also called brown king crab, occur from the Japan Sea to the northern Bering Sea (ca. $61^{\circ} \mathrm{N}$ latitude), around the Aleutian Islands, generally in high-relief habitat such as inter-island passes, on various sea mounts, and as far south as northern British Columbia (Alice Arm) (Jewett et al. 1985). They are typically found on the continental slope at depths of $300-1,000 \mathrm{~m}$ on extremely rough bottom. They are frequently found on coral bottom.

The Aleutian Islands king crab stock boundary is defined by the boundaries of the Aleutian Islands king crab Registration Area O (Figure 2). In this chapter, "Aleutian Islands Area" means the area described by the current definition of Aleutian Islands king crab Registration Area O. Leon et al. (2017) define the boundaries of Aleutian Islands king crab Registration Area O:

The Aleutian Islands king crab management area's eastern boundary is the longitude of Scotch Cap Light ( $164^{\circ} 44.72^{\prime} W$ long), the northern boundary is a line from Cape Sarichef ( $54^{\circ} 36^{\prime} \mathrm{N}$ lat) to $171^{\circ} \mathrm{W}$ long, north to $55^{\circ} 30^{\prime} \mathrm{N}$ lat, and the western boundary the Maritime Boundary Agreement Line as described in the Maritime Boundary Agreement between the United States and the Union of Soviet Socialist Republics signed in Washington, June 1, 1990 (Figure 1-1 in Leon et al. 2017). Area O encompasses
territorial waters of the state of Alaska (0-3 nautical miles) and waters of the Exclusive Economic Zone (3-200 nautical miles).

During 1984/85-1995/96, the Aleutian Islands king crab populations had been managed using the Adak and Dutch Harbor Registration Areas, which were divided at $171^{\circ} \mathrm{W}$ longitude (Figure 3), but from the 1996/97 season to present the fishery has been managed using a division at $174^{\circ} \mathrm{W}$ longitude (Figure 2). In March 1996 the Alaska Board of Fisheries (BOF) replaced the Adak and Dutch Harbor areas with the newly created Aleutian Islands Registration Area O and directed the Alaska Department of Fish and Game (ADF\&G) to manage the golden king crab fishery in the areas east and west of $174^{\circ} \mathrm{W}$ longitude as two distinct stocks. That re-designation of management areas was intended to more accurately reflect golden king crab stock distribution, coherent with the longitudinal pattern in fishery production prior to 1996/97 (Figure 4). The longitudinal pattern in fishery production relative to $174^{\circ} \mathrm{W}$ longitude since 1996/97 is similar to that observed prior to the change in management area definition, although there have been some changes in the longitudinal pattern in fishery production within the areas east and west of $174^{\circ} \mathrm{W}$ longitude (Figure 5).

Commercial fishing for golden king crab in the Aleutian Islands Area typically occurs at depths of 100-275 fathoms (183-503 m). Pots sampled by at-sea fishery observers in 2013/14 were fished at an average depth of 176 fathoms ( $322 \mathrm{~m} ; \mathrm{N}=499$ ) in the area east of $174^{\circ} \mathrm{W}$ longitude and 158 fathoms ( $289 \mathrm{~m} ; \mathrm{N}=1,223$ ) for the area west of $174^{\circ} \mathrm{W}$ longitude (Gaeuman 2014).

## 3. Evidence of stock structure:

Given the expansiveness of the Aleutian Islands Area and the existence of deep ( $>1,000$ m ) canyons between some islands, at least some weak structuring of the stock within the area would be expected. Data for making inferences on stock structure of golden king crab within the Aleutian Islands are largely limited to the geographic distribution of commercial fishery catch and effort. Catch data by statistical area from fish tickets and catch data by location from pots sampled by observers suggest that habitat for legal-sized males may be continuous throughout the waters adjacent to the islands in the Aleutian chain. However, regions of low fishery catch suggest that availability of suitable habitat, in which golden king crab are present at only low densities, may vary longitudinally. Catch has been low in the fishery in the area between $174^{\circ} \mathrm{W}$ longitude and $176^{\circ} \mathrm{W}$ longitude (the Adak Island area, Figures 4 and 5) in comparison to adjacent areas, a pattern that is consistent with low CPUE for golden king crab between $174^{\circ} \mathrm{W}$ longitude and $176^{\circ} \mathrm{W}$ longitude (Figure 6) during the 2002, 2004, 2006, 2010, and 2012 NMFS Aleutian Islands bottom trawl surveys (von Szalay et al. 2011). In addition to longitudinal variation in density, there is also a gap in fishery catch and effort between the Petrel Bank-Petrel Spur area and the Bowers Bank area; both of those areas, which are separated by Bowers Canyon, have reported effort and catch. Recoveries during commercial fisheries of golden king crab tagged during ADF\&G surveys (Blau and Pengilly 1994; Blau et al. 1998; Watson and Gish 2002; Watson 2004, 2007) provided no evidence of substantial movements by crab in the size classes that were tagged (males and females $\geq 90-\mathrm{mm}$ carapace length [CL]). Maximum straight-line distance between
release and recovery location of 90 golden king crab released prior to the 1991/92 fishery and recovered through the 1992/93 fishery was 61.2 km (Blau and Pengilly 1994). Of the 4,567 recoveries reported through 12 April 2016 for the male and female golden king crab tagged and released between $170.5^{\circ} \mathrm{W}$ longitude and $171.5^{\circ} \mathrm{W}$ longitude during the 1991, 1997, 2000, 2003, and 2006 ADF\&G Aleutian Island golden king pot surveys, none of the 3,807 with recovery locations specified by latitude and longitude were recovered west of $173^{\circ} \mathrm{W}$ longitude and only fifteen were recovered west of $172^{\circ} \mathrm{W}$ longitude (V. Vanek, ADF\&G, Kodiak, pers. comm.). Similarly, of 139 recoveries in which only the statistical area of recovery was reported, none were recovered in statistical areas west of $173^{\circ} \mathrm{W}$ longitude and only one was in a statistical area west of $172^{\circ} \mathrm{W}$ longitude.

## 4. Life history characteristics relevant to management:

There is a paucity of information on golden king crab life history characteristics due in part to the deep depth distribution ( $\sim 200-1000 \mathrm{~m}$ ) and the asynchronous nature of life history events (Otto and Cummiskey 1985; Somerton and Otto 1986). The reproductive cycle is thought to last approximately 24 months and at any one time, ovigerous females can be found carrying egg clutches in highly disparate developmental states (Otto and Cummiskey 1985). Females carry large, yolk-rich, eggs, which hatch into lecithotrophic (i.e., the larvae can develop successfully to juvenile crab without eating; Shirley and Zhou 1997) larvae that are negatively phototactic (Adams and Paul 1999). Molting and mating are also asynchronous and protracted (Otto and Cummiskey 1985; Shirley and Zhou 1997) with some indications of seasonality (Hiramoto 1985). Molt increment for large males (adults) in Southeast Alaska is 16.3 mm CL per molt (Koeneman and Buchanan 1985), and was estimated at 14.4 mm CL for legal males in the EAG (Watson et al. 2002). Annual molting probability of males decreases with increasing size, which results in a protracted inter-molt period and creates difficulty in determining annual molt probability (Watson et al. 2002). Male size-at-maturity varies among stocks (Webb 2014) and declines with increasing latitude from about 130 mm CL in the Aleutian Islands to 90 mm CL in Saint Matthew Island section (Somerton and Otto 1986). Along with a lack of annual survey data, limited stock-specific life history stock information prevents development of the standard length-based assessment model.

## 5. Brief summary of management history:

A complete summary of the management history through 2015/16 is provided in Leon et al. (2017, pages 9-14). The first commercial landing of golden king crab in the Aleutian Islands was in 1975/76, but directed fishing did not occur until 1981/82.

The Aleutian Islands golden king crab fishery was restructured beginning in 1996/97 to replace the Adak and Dutch Harbor areas with the newly created Aleutian Islands Registration Area O and golden king crab in the areas east and west of $174^{\circ} \mathrm{W}$ longitude were managed separately as two stocks (ADF\&G 2002). Hereafter, the east of $174^{\circ} \mathrm{W}$ longitude stock segment is referred to as EAG and the west of $174^{\circ} \mathrm{W}$ longitude stock segment is referred to as WAG. Table 1 provides the historical summary of number of vessels, GHL/TAC, harvest, effort, CPUE and average weight in the Aleutian Islands golden king crab fishery.

The fisheries in 1996/97-1997/98 were managed with 1,452 t (3,200,000 lb) for EAG and $1,225 \mathrm{t}(2,700,000 \mathrm{lb})$ for WAG (Table 1). During 1998/99-2004/05 the fisheries were managed with $1,361 \mathrm{t}(3,000,000 \mathrm{lb})$ for EAG and $1,225 \mathrm{t}(2,700,000 \mathrm{lb})$ for WAG. During 2005/06-2007/08 the fisheries were managed with a total allowable catch (TAC) of $1,361 \mathrm{t}(3,000,000 \mathrm{lb})$ for EAG and a TAC of $1,225 \mathrm{t}(2,700,000 \mathrm{lb})$ for WAG. By state regulation (5 AAC 34.612), TAC for the Aleutian Islands golden king crab fishery during 2008/09-2011/12 was $1,429 \mathrm{t}(3,150,000 \mathrm{lb})$ for EAG and $1,286 \mathrm{t}(2,835,000 \mathrm{lb})$ for WAG. In March 2012 the BOF changed 5 AAC 34.612 so that the TAC beginning in $2012 / 13$ would be $1,501 \mathrm{t}(3,310,000 \mathrm{lb})$ for the EAG and $1,352 \mathrm{t}(2,980,000 \mathrm{lb})$ for WAG. Additionally, the BOF added a provision to 5 AAC 34.612 that allows ADF\&G to lower the TAC below the specified level if conservation concerns arise. The TAC for $2016 / 17$ (and 2017/18) was reduced by $25 \%$ for WAG with $1,014 \mathrm{t}(2,235,000 \mathrm{lb})$ while keeping the TAC for EAG at the same level as that in the previous season.

During 1996/97-2016/17 the annual retained catch during commercial fishing (including cost-recovery fishing that occurred during 2013/14-2016/17) has averaged $2 \%$ below the annual GHL/TACs. During 1996/97-2016/17, the retained catch has been as much as $13 \%$ below (1998/99) and as much as $6 \%$ above (2000/01) the GHL/TAC.

A summary of other relevant SOA fishery regulations and management actions pertaining to the Aleutian Islands golden king crab fishery is provided below:

Beginning in 2005/06 the Aleutian Islands golden king crab fishery has been prosecuted under the Crab Rationalization Program. Accompanying the implementation of the Crab Rationalization program was implementation of a community development quota (CDQ) fishery for golden king crab in the eastern Aleutians (i.e., EAG) and the Adak Community Allocation (ACA) fishery for golden king crab in the western Aleutians (i.e., WAG; Hartill 2012). The CDQ fishery in the eastern Aleutians is allocated $10 \%$ of the golden king crab TAC for the area east of $174^{\circ} \mathrm{W}$ longitude and the ACA fishery in the western Aleutians is allocated $10 \%$ of the golden king crab TAC for the area west of $174^{\circ}$ W longitude. The CDQ fishery and the ACA fishery are managed by ADF\&G and prosecuted concurrently with the IFQ fishery.

Golden king crab may be commercially fished only with king crab pots (defined in 5 AAC 34.050). Pots used to fish for golden king crab in the Aleutian Islands Area must be operated from a shellfish longline and, since 1996, must have at least four escape rings of five and one-half inches minimum inside diameter installed on the vertical plane or at least one-third of one vertical surface of the pot composed of not less than nine-inch stretched mesh webbing to permit escapement of undersized golden king crab (5 AAC 34.625 (b)). Prior to the regulation requiring an escape mechanism on pots, some participants in the Aleutian Islands golden king crab fishery voluntarily sewed escape rings (typically 139 mm or 5.5 inches) into their gear or, more rarely, included panels with escape mesh (Beers 1992). With regard to the gear used since the establishment of 5 AAC 34.625 (b) in 1996, Linda Kozak, a representative of the industry, reported in a 19 September 2008 email to the Crab Plan Team that, "... the golden king crab fleet has
modified their gear to allow for small crab sorting," and provided a written statement from Lance Nylander, of Dungeness Gear Works in Seattle, who "believes he makes all the gear for the golden king crab harvesting fleet," saying that, "Since 1999, DGW has installed $9[-\mathrm{inch}]$ escape web on the door of over $95 \%$ of Golden Crab pot orders we manufactured." A study to estimate the contact-selection curve for male golden king crab that was conducted aboard one vessel commercial fishing for golden king crab during the 2012/13 season showed that gear and fishing practices used by that vessel were highly effective in reducing bycatch of sublegal-sized males and females (Vanek et al. 2013). In March 2011 (effective for 2011/12), the BOF amended 5 AAC 34.625 (b) to relax the "biotwine" specification for pots used in the Aleutian Islands golden king crab fishery relative to the requirement in 5 AAC 39.145 that "(1) a sidewall ...of all shellfish and bottomfish pots must contain an opening equal to or exceeding 18 inches in length... The opening must be laced, sewn, or secured together by a single length of untreated, 100 percent cotton twine, no larger than 30 thread." Regulation 5 AAC 34.625 (b)(1) allows the opening described in 5 AAC 39.145 (1) to be "laced, sewn, or secured together by a single length of untreated, 100 percent cotton twine, no larger than 60 [rather than 30] thread."

Regulation (5 AAC 34.610 (b)) sets the commercial fishing season for golden king crab in the Aleutian Islands Area as 1 August through 30 April. That regulatory fishing season became effective in 2015/16 (the commercial fishing season was set in regulation as 15 August through 15 May during 2005/06-2014/15).

Current regulations (5 AAC 39.645 (d)(4)(A)) stipulate that onboard observers are required on catcher vessels during the time that at least $50 \%$ of the retained catch is captured in each of the three trimesters of the 9 -month fishing season. Onboard observers are required on catcher-processors at all times during the fishing season.

Additional management measures include only males of a minimum size may be retained by the commercial golden king crab fishery in the Aleutian Islands Area. By SOA regulation (5 AAC 34.620 (b)), the minimum legal size limit is 6.0 -inches ( 152.4 mm ) carapace width (CW), including spines, which is at least one annual molt increment larger than the $50 \%$ maturity length of 120.8 mm CL for males estimated by Otto and Cummiskey (1985). A carapace length (CL) $\geq 136 \mathrm{~mm}$ is used to identify legal-size males when CW measurements are not available (Table 3-5 in NPFMC 2007b). Note that size limit for golden king crab has been 6 -inches ( 152.4 mm ) CW for the entire Aleutian Islands Area since the 1985/86 season. Prior to the 1985/86 season, the legal size limit was 6.5 -inches ( 165.1 mm ) CW for at least one of the now-defunct Adak or Dutch Harbor Registration Areas.

We re-evaluated the male maturity size using 1991 pot survey measurements of carapace length and chela height in EAG and 1984 NMFS measurements in WAG (Appendix C). Bootstrap analysis of chela height and carapace length data provided the median $50 \%$ male maturity length estimates of 107.02 mm CL in EAG and 107.85 mm CL in WAG. We used a knife-edge $50 \%$ maturity length of 111.0 mm CL, which is the lower limit of the next upper size bin, for mature male biomass (MMB) estimation.

Daily catch and catch-per-unit effort (CPUE) are determined in-season to monitor fishery performance and progress towards the respective TACs. Figures 7 to 9 provide the 1985/86-2016/17 time series of catches, CPUE, and the geographic distribution of catch during the 2016/17 fishing season. Increases in CPUE were observed during the late 1990s through the early 2000s, and with the implementation of crab rationalization in 2005. This is likely due to changes in gear configurations in the late 1990s (crab fishermen, personal communication, July 1, 2008) and, after rationalization, to increased soak time (Siddeek et al. 2015), and decreased competition owing to the reduced number of vessels fishing. Decreased competition could allow crab vessels to target only the most productive fishing areas. Trends in fishery CPUE within the areas EAG and WAG generally paralleled each other during 1985/86-2010/11, but diverged during 2011/122016/17 (an increasing trend in EAG and a decreasing trend in WAG).
6. Brief description of the annual ADF\&G harvest strategy:

The annual TAC is set by state regulation, 5 AAC 34.612 (Harvest Levels for Golden King Crab in Registration Area O), as approved by the BOF in March 2012:
(a) Until the Aleutian Islands golden king crab stock assessment model and a state regulatory harvest strategy are established, the harvest levels for the Registration Area O golden king crab fishery are as follows:
(1) east of $174^{\circ} \mathrm{W}$ long. (EAG): 3.31 million pounds; and
(2) west of $174^{\circ} \mathrm{W}$ long. (WAG): 2.98 million pounds;
(b) The department may reduce the harvest levels based on the best scientific information available and considering the reliability of estimates and performance measures, sources of uncertainty as necessary to avoid overfishing, and any other factors necessary to be consistent with sustained yield principles.

In addition to the retained catch that is limited by the TAC established by ADF\&G under 5 AAC 34.612, ADF\&G also has authority to annually receive receipts of $\$ 500,000$ through cost-recovery fishing on Aleutian Islands golden king crab. The retained catch from that cost-recovery fishing is not counted against attainment of the annuallyestablished TAC.
At the March 2018 meeting, The BOF decided to amend the phrase "may reduce to "may modify" in (b).
7. Summary of the history of the basis and estimates of $M M B_{M S Y}$ or proxy $M M B_{M S Y}$ : We estimated the proxy $M M B_{M S Y}$ as $M M B_{35 \%}$ using the Tier 3 estimation procedure, which is explained in a subsequent section.

## D. Data

1. Summary of new information:
(a) Commercial fishery retained catch by size, estimated total catch by size, groundfish male discard catch by size, observer CPUE index, commercial fishery CPUE index,
and tag-recapture data were updated to include 2016/17 information. The details are given in the pictorial table below.

2. Data presented as time series:
a. Total Catch:

Fish ticket data on retained catch weight, catch numbers, effort (pot lifts), CPUE, and average weight of retained catch for 1981/82-2016/17 (Table 1). Estimated total catch weight for 1990/91-2016/17 (Table 2a).
b. Bycatch and discards:

Retained catch, bycatch mortality (male and female of all sizes included) separated by the crab fishery and groundfish fishery, and total fishery mortality for 1981/822016/17 (Table 2). Crab fishery discards are available after observer sampling was established in 1988/89. Some observer data exists for the 1988/89-1989/90 seasons, but those data are not considered reliable. Table 2 provides crab fishery discards and groundfish fishery bycatch for 1991/92-2016/17 seasons.

## c. Catch-per-unit-effort:

- Pot fishery and observer nominal retained and total CPUE, pot fishery effort, observer sample size, and estimated observer CPUE index delineated by EAG and WAG for 1985/86-2016/17 (Table 3).
- Estimated commercial fishery CPUE index with coefficient of variation (Table 4 for EAG and Table 22 for WAG). The estimation methods, CPUE fits and diagnostic plots are described in Appendixes B and G.


## d. Catch-at-length:

Information on length compositions (Figures 11 to 13 for length compositions for EAG; and 29 to 31 for length compositions for WAG).
e. Survey biomass estimates:

They are not available for the area because no systematic surveys, covering the entire fishing area, have occurred.
f. Survey catch-at-length:

They are not available.
g. Other time series data: None.
3. Data which may be aggregated over time:

- Molt and size transition matrix: Tag release - recapture -time at liberty records from 1991, 1997, 2000, 2003, and 2006 male tag crab releases were aggregated by year at liberty to determine the molt increment and size transition matrix by the integrated model.
- Weight-at-length: Male length-weight relationship: $\mathrm{W}=\mathrm{aL}^{\mathrm{b}}$ where $\mathrm{a}=$ $3.7255^{*} 10^{-4}, \mathrm{~b}=3.0896$ (updated estimates).
- Natural mortality: Model estimated fixed natural mortality value was used in the assessment.

4. Information on any data sources that were available, but were excluded from the assessment:
Data from triennial ADF\&G pot surveys for Aleutian Islands golden king crab in a limited area in EAG (between $170^{\circ} 21^{\prime}$ and $171^{\circ} 33^{\prime} \mathrm{W}$ longitude) that were performed during 1997 (Blau et al. 1998), 2000 (Watson and Gish 2002), 2003 (Watson 2004), and 2006 (Watson 2007) are available, but were not used in this assessment. However, the tag release recapture data from these surveys were used.

## E. Analytic Approach

1. History of modeling approaches for this stock:

A size structured assessment model based on only fisheries data has been under development for several years for the EAG and WAG golden king crab stocks. The model was accepted in 2016 for OFL and ABC setting for the 2017/18 season. The CPT in January 2017 and SSC in February 2017 recommended to using the Tier 3 procedure to set the OFL and ABC. They also suggested to using the maturity data to estimate MMB. We followed these suggestions in this report. This is the second fishing season we are proposing to use the model-based OFL and ABC setting.

## 2. Model Description:

a. Description of overall modeling approach:

The underlying population dynamics model is male-only and length-based (Appendix A). This model combines commercial retained catch, total catch, groundfish fishery discarded catch, standardized observer legal size catch-per-unit-effort (CPUE) indices, fishery retained catch size composition, total catch size composition, and tag recaptures by release-recapture length to estimate stock assessment parameters. The tagging data were used to calculate the size transition matrix. To estimate the male mature biomass (MMB), we used the knife-edge $50 \%$ maturity based on the chela height and carapace length data analysis. To include a long time series of CPUE indices for stock abundance contrast, we also considered the 1985/86-1998/99 legal size standardized CPUE indices as a separate likelihood component in all scenarios (see Table T1). As a first attempt, we used VAST to estimate a separate set of observer CPUE indices for the model scenario 17_0a and also used the 2015-2017 fishery independent pot survey CPUE indices for the model scenario 17_Of.
There were significant changes in fishing practice due to changes in management regulations (e.g., constant TAC since 1996/97 and crab rationalization since 2005/06), pot configuration (escape web on the pot door increased to 9 -inch since 1999), and improved observer recording in Aleutian Islands golden king crab fisheries since 1998. These changes prompted us to consider two sets of catchability and total selectivity parameters with only one set of retention parameters for the periods 1985/86-2004/05 and 2005/06-2016/17. However, in order to respond to the September 2017 CPT comment, we considered three catchabilities, three sets of total selectivity, and one set of retention curves in one scenario (scenario 17_0d).

We fitted the observer and commercial fishery CPUE indices with estimated (by GLM or VAST) standard errors and an additional model estimated constant variance. The assessment model predicted total and retained CPUEs. However, we compared only the predicted retained CPUE with the observer legal size crab CPUE indices in the likelihood function because observer recordings of legal size crabs are reliable.

The data series ranges used for the WAG are the same as those for EAG.
b. Software:

AD Model Builder (Fournier et al. 2012).
c.-f. Details are given in Appendix A.
g. Critical assumptions and consequences of assumption failures:

Because of the lack of an annual stock survey, we relied heavily on standardized CPUE indices (Appendix B) and catch and size composition information to determine the stock abundance trends in both regions. We assumed that the observer and fish ticket CPUE indices are linearly related to exploitable abundance. We kept $M$ constant at $0.21 \mathrm{yr}^{-1}$. The $M$ value was the combined estimates for EAG and WAG (Figure 1). We assumed directed pot fishery discard mortality at $0.20 \mathrm{yr}^{-1}$, overall groundfish fishery mortality at $0.65 \mathrm{yr}^{-1}$ [mean of groundfish pot fishery mortality $\left(0.5 \mathrm{yr}^{-1}\right)$ and groundfish trawl fishery mortality $\left.\left(0.8 \mathrm{yr}^{-1}\right)\right]$, groundfish fishery selectivity at full selection for all length classes (selectivity $=1.0$ ). Any discard of
legal size males in the directed pot fishery was not considered in this analysis. These fixed values invariably reduced the number of model parameters to be estimated and helped in convergence. We assumed different $q$ 's (scaling parameter for standardized CPUE in the model, Equation A. 13 in Appendix A) and logistic selectivity patterns (Equation A. 9 in Appendix A) for different periods for the pot fishery.
h. Changes to any of the above since the previous assessment:

None.
i. Model code has been checked and validated.

The code is available from the authors.

## 3. Model Selection and Evaluation

a. Description of alternative model configurations:

We considered 7 scenarios for EAG and 6 scenarios for WAG (Table T1). We presented OFL and ABC results for all scenarios separately for EAG, WAG, and the entire AI in the executive summary tables. We considered scenario 17_0 as the base scenario. It considers:
i) Initial abundance by the equilibrium condition considering the mean number of recruits for 1987-2012: The equilibrium abundance was determined for 1960, projected forward with only $M$ and annual recruits until 1980, then retained catches removed during 1981-1984 and projected to obtain the initial abundance in 1985 (see Equations A. 4 and A. 5 in Appendix A).
ii) Observer CPUE indices for 1995/96-2016/17.
iii) Fishery CPUE indices for 1985/86-1998/99.
iv) Initial (Stage-1) weighting of effective sample sizes: number of vessel-days for retained and total catch size compositions, and number of fishing trips for groundfish discard size composition (the groundfish size composition was not used in the model fitting); and (Stage-2) iterative re-weighting of effective sample sizes by the Francis and McAllister and Ianelli methods (Appendix D).
v) Two catchability and two sets of logistic total selectivity for the periods 1985/862004/05 and 2005/06-2016/17, and a single set of logistic retention curve parameters.
vi) Full selectivity (selectivity $=1.0$ ) for groundfish (trawl) bycatch.
vii) Knife-edge $50 \%$ maturity size.
viii) Stock dynamics $M=0.21 \mathrm{yr}^{-1}$, pot fishery handling mortality $=0.2 \mathrm{yr}^{-1}$, and mean groundfish bycatch handling mortality $=0.65 \mathrm{yr}^{-1}$.
ix) Size transition matrix using tagging data estimated by the normal probability function with the logistic molt probability sub-model. The tag-recaptures were treated as Bernoulli trials (i.e., Stage-1 weighting).
x) The time period, 1987-2012, was used to determine the mean number of recruits for $M M B_{35 \%}$ (a proxy for $M M B_{M S Y}$ ) estimation under Tier 3.

The salient features and variations from the base scenario of all other scenarios are listed in Table T1. The list of fixed and estimable parameters are provided in Table A1 and
detail weights with coefficient of variations (CVs) assigned to each type of data are listed in Table A2 of Appendix A.

As per CPT and SSC requests, initial parameter values for scenario $17 \_0$ were jittered to confirm model global convergence. The results indicated that global convergence was achieved for almost all the runs (Appendix E).

Table T1. Features of model scenarios. Initial condition was estimated by the equilibrium condition for all scenarios. Changes from scenario 17_0 specifications are highlighted by the light blue shade.

| Scenario | Sizecomposition weighting | Catchability and logistic total selectivity sets | Maturity | Standardized CPUE data type | Treatment of $M$ and Tier 3 $M_{M B}{ }_{M S Y}$ reference points | $\begin{gathered} \text { Natural } \\ \text { mortality }(M \\ \left.y r^{-1}\right) \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0b | Stage1:Number of boat_days/trips Stage-2: Francis method | 2 | Knifeedge, 111 mm CL | Observer from 1995/96-2016/17 \& Fish Ticket from 1985/86-1998/99; <br> GLM variable selection by $R$ square criterion | Estimate a common $M$ using the combined EAG and WAG data without an $M$ prior | $0.2254 ;$ Individual component's estimate: EAG: 0.2142 WAG: 0.2142 |
| 17_0 | Stage1:Number of boat_days/trips Stage-2: Francis method | 2 | $\begin{gathered} \text { Knife- } \\ \text { edge, 111 } \\ \text { mm CL } \end{gathered}$ | Observer from 1995/96-2016/17 \& Fish Ticket from 1985/86-1998/99; <br> GLM variable selection by R square criterion | Single $M$ from combined EAG and WAG data; Tier $3 M M B_{M S Y}$ reference points based on average recruitment from 1987-2012 | 0.21 |
| 17_0a | Stage1 :Number of boat_days/trips Stage-2: Francis method | 2 | Knifeedge, 111 mm CL | Observer CPUE by VAST \& Fish Ticket CPUE by GLM; GLM variable selection by $R$ square criterion | Single $M$ from combined EAG and WAG data; Tier $3 M M B_{M S Y}$ reference points based on average recruitment from 1987-2012 | 0.21 |
| 17_0b | Stage1:Number of boat_days/trips Stage-2: Francis method | 2 | Knifeedge, 111 mm CL | Observer \& Fish Ticket CPUE by GLM; GLM variable selection by CAIC | Single $M$ from combined EAG and WAG data; Tier $3 M M B_{M S Y}$ reference points based on average recruitment from 1987-2012 | 0.21 |


| 17_0c | Stage1:Number of boat_days/trips Stage-2: <br> Francis method | 2 | Knifeedge, 111 mm CL | Observer \& Fish Ticket CPUE standardization considering Year:Area interaction; GLM variable selection by $R$ square criterion | Single $M$ from combined EAG and WAG data; Tier $3 M M B_{M S Y}$ reference points based on average recruitment from 1987-2012 | 0.21 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 17_0d | Stage1:Number of boat_days/trips Stage-2: Francis method | 3 | Knifeedge, 111 mm CL | Observer \& Fish ticket; GLM variable selection by $R$ square criterion | Three different total selectivity curves and catchability coefficients for 1985-2004, 2005-2012, and 2013-2016; single $M$ from combined EAG and WAG data; Tier $3 M M B_{M S Y}$ reference points based on average recruitment from 1987-2012 | 0.21 |
| 17_0e | Stage- <br> 1:Number of boat_days/trips Stage-2: <br> McAllister and Ianelli method | 2 | Knifeedge, 111 mm CL | Observer \& Fish ticket; GLM variable selection by R square criterion | Single $M$ from combined EAG and WAG data; Tier $3 M M B_{M S Y}$ reference points based on average recruitment from 1987-2012 | 0.21 |
| $\begin{aligned} & \text { 17_0f } \\ & \text { (only for } \\ & \text { EAG) } \end{aligned}$ | Stage1:Number of boat_days/trips Stage-2: Francis method | 2 | Knifeedge, 111 mm CL | Observer, Fish ticket, \& fishery independent pot survey (20152016) in EAG; GLM variable selection by $R$ square criterion | Fishery independent pot survey standardized CPUE are considered as a separate likelihood component for EAG; single $M$ from combined EAG and WAG data; Tier $3 M M B_{M S Y}$ reference points based on average recruitment from 1987-2012 | 0.21 |

## b. Progression of results:

The OFL and ABC estimates are similar to those estimated by the 2017 model.
c. Label the approved model from the previous year as model 0:

Following the September CPT suggestion we used the notation 17_0 for the base model which came from the previous assessment.
d. Evidence of search for balance between realistic and simpler models:

Unlike annually surveyed stocks, Aleutian Islands golden king crab stock biomass is difficult to track and several biological parameters are assumed based on knowledge from red king crab (e.g., handling mortality rate of $0.2 \mathrm{yr}^{-1}$ ) due to a lack of species/stock specific information. We fixed a number of model parameters after initially running the model with free parameters to reduce the number of parameters to be estimated (e.g., groundfish bycatch selectivity parameters were fixed). The seven scenarios also considered different configuration of parameters to select parsimonious models. The detailed results of the seven scenarios are provided in tables and figures. The total catch OFLs and the reduction in terminal (2016) MMB from the initial condition (i.e., virgin MMB in 1960) for all scenarios for EAG and WAG are provided in Table 38. We also included the results of the accepted 2017 model scenario, Sc9, in this table for comparison. The reduction in terminal MMB from the initial condition is higher for WAG than EAG.

## e. Convergence status and criteria:

ADMB default convergence criteria were used.
f. Table of the sample sizes assumed for the size compositional data:

We estimated the initial input effective sample sizes (i.e., Stage-1) either as number of vessel-days for retained and total catch compositions and number of fishing trips for groundfish size composition (note: we did not use the groundfish size composition in the model fit) for all scenarios. Then we estimated the Stage-2 effective sample sizes iteratively from Stage-1 input effective sample sizes using the Francis' (2011, 2017) mean length based method and McAllister and Ianelli method (McAllister and Ianelli, 1997) (Appendix D).

We provide the initial input sample sizes (Stage-1) and Stage-2 effective sample sizes for scenarios 17_0 to 17_0f in Tables 5 to 11 for EAG and Tables 23 to 28 for WAG.
g. Provide the basis for data weighting, including whether the input effective sample sizes are tuned and the survey CV adjusted:
Described previously (f) and details are in Appendix D.
h. Do parameter estimates make sense and are they credible?

The estimated parameter values are within the bounds and various plots suggest that the parameter values are reasonable for a fixed $M$ value for the golden king crab stocks.
i. Model selection criteria:

We used a number of diagnostic criteria to select the appropriate models for our recommendation: CPUE fits, observed vs. predicted tag recapture numbers by time at large and release size, retained and total catch, and groundfish bycatch fits. Figures are provided for all scenarios in the Results section.

## j. Residual analysis:

We illustrated residual fits by bubble plots for retained and total catch size composition predictions in various figures in the Results section.
k. Model evaluation:

Only one model with a number of scenarios is presented and the evaluations are presented in the Results section below.

## 4. Results

## 1. List of effective sample sizes and weighting factors:

The Stage-1 and Stage-2 effective sample sizes are listed for various scenarios in Tables 5 to 11 for EAG and Tables 23 to 28 for WAG. The weights for different data sets are provided in Table A2 for various scenarios, respectively, for EAG and WAG (Appendix A). These weights (with the corresponding coefficient of variations) adequately fitted the length compositions and no further changes were examined.

We used weighting factors for catch biomass, recruitment deviation, pot fishery F , and groundfish fishery F. We set the retained catch biomass to a large value (500.0) because retained catches are more reliable than any other data sets. We scaled the total catch biomass in accordance with the observer annual sample sizes with a maximum of 250.0. The total catches were derived from observer nominal total CPUE and effort. In some years, observer sample sizes were low (Tables 3). We chose a small groundfish bycatch weight (0.2) based on the September 2015 CPT suggestion to lower its weight. We used the best fit criteria to choose the lower weight for the groundfish bycatch. Groundfish bycatch of Aleutian Islands golden king crab is very low. We set the CPUE weights to 1.0 for all scenarios. We included a constant (model estimated) variance in addition to input CPUE variance for the CPUE fit. We used the Burnham et al. (1987) suggested formula for $\ln (\mathrm{CPUE})$ [and $\ln (\mathrm{MMB})$ ] variance estimation (Equation A. 14 of Appendix A). However, the estimated additional variance values were small for both observer and fish ticket CPUE indices for the two regions. Nevertheless, the CPUE index variances estimated from the negative binomial and lognormal GLMs were adequate to fit the model, as confirmed by the fit diagnostics (Fox and Weisberg 2011). Parameter estimates are provided in Tables 12 and 13 for EAG and 29 and 30 for WAG for all scenarios. The numbers of estimable parameters are listed in Table A1 of Appendix A. The weights with the corresponding coefficient of variations specifications are detailed in Tables A2 of Appendix A for EAG and WAG.

## 2. Include tables showing differences in likelihood:

Tables 21 and 37 list the total and component negative log likelihood values and their differences between scenarios of similar sample sizes and free parameters for EAG and WAG, respectively.

## 3. Tables of estimates:

a. The parameter estimates with coefficient of variation for all scenarios are summarized respectively in Tables 12 and 13 for EAG and 29 and 30 for WAG. We have also provided the boundaries for parameter searches in those tables. All parameter estimates were within the bounds.
b. All scenarios considered molt probability parameters in addition to the linear growth increment and normally distributed growth variability parameters to determine the size transition matrix.
c. The mature male and legal male abundance time series for all scenarios are summarized in Tables 14 to 20 for EAG and Tables 31 to 36 for WAG.
d. The recruitment estimates for those scenarios are summarized in Tables 14 to 20 for EAG and Tables 31 to 36 for WAG.
e. The negative log-likelihood component values and total negative log-likelihood values for all scenarios are summarized in Table 21 for EAG and Table 37 for WAG. Scenario 17_0d has the minimum total negative log likelihood for EAG whereas scenario 17_0e has the minimum for WAG. Among the scenarios with equal data components (base) and number of free parameters, scenario 17_0e has the lowest total negative log likelihoods for both EAG and WAG. Thus, we chose scenarios 17_0 (base), 17_0d, and 17_0e for OFL and ABC options for consideration.

## 4. Graphs of estimates:

## a. Selectivity:

Total selectivity and retention curves of the pre- and post-rationalization periods for all scenarios are illustrated in Figure 14 for EAG and Figure 32 for WAG. Total selectivity for the pre-rationalization period was used in the tagging model. The groundfish bycatch selectivity appeared flat in the preliminary analysis, indicating that all size groups were vulnerable to the gear. This is also shown in the size compositions of groundfish bycatch (Figures 13 and 31 for EAG and WAG, respectively). Thus, we set the groundfish bycatch selectivity to 1.0 for all length-classes in the subsequent analysis.

## b. Mature male biomass:

The mature male biomass time series for nine (a subset of 11) scenarios are depicted in Figures 28 and 46 for EAG and WAG, respectively. Mature male biomass tracked the CPUE trends well for all scenarios for EAG and WAG. The biomass variance was estimated using Burnham et al. (1987) suggested formula
(Equation A. 14 in Appendix A). We determined the mature male biomass values on 15 February each year and considered the 1987-2012 time series of recruits for estimating mean number of recruits for $M M B_{35 \%}$ calculation under Tier 3 approach.
c. Fishing mortality:

The full selection pot fishery F over time for all scenarios is shown in Figures 27 and 45 for EAG and WAG, respectively. The F peaked in late 1980s and early to mid-1990s and systematically declined in the EAG. On the other hand, the F in the WAG peaked in late 1980s, 1990s and early 2000s, then declined in late 2000s and slightly increased since 2010. The increase in F in recent years may be due to a decline in abundance under constant high harvest allocation to WAG.

## d. F vs. MMB:

We provide these plots for scenarios 17_0 and 17_0d for EAG and WAG in Figure 47.
e. Stock-Recruitment relationship: None.

## f. Recruitment:

The temporal changes in total number of recruits to the modeled population for all scenarios are illustrated in Figure 16 for EAG and in Figure 34 for WAG. The recruitment distribution to the model size group (101-185 mm CL) is shown in Figures 17 and 35 for EAG and WAG, respectively for all scenarios.

## 5. Evaluation of the fit to the data:

## g. Fits to catches:

The fishery retained, total, and groundfish bycatch (observed vs. estimated) plots for all scenarios are illustrated in Figures 19 and 37 for EAG and WAG, respectively. The 1981/82-1984//85 retained catch plots for all scenarios are depicted in Figures 20 and 38 for EAG and WAG, respectively. All predicted fits were very close to observed values, especially for retained catch and groundfish bycatch mortality. However, pre 1995 total catch data did not fit well.
h. Survey data plot:

We did not consider the pot survey data for the analysis.
i. CPUE index data:

The predicted vs. input CPUE indices for all scenarios are shown in Figure 26 for EAG and Figure 44 for WAG. Scenario 17_0d fit the recent three years’ CPUE indices well for EAG; on the other hand, scenario 17_0c did not fit the post rationalization period CPUE indices well for WAG. The CPUE variance was estimated using Burnham et al. (1987) suggested formula (Equation A. 14 in Appendix A).

## j. Tagging data:

The predicted vs. observed tag recaptures by length-class for years 1 to 6 recaptures are depicted in Figure 15 for EAG and Figure 33 for WAG. The predictions appear reasonable. Note that we used the EAG tagging information for size transition matrix estimation for both stocks (EAG and WAG). The size transition matrices estimated using EAG tagging data in the EAG and WAG models were similar.

## k. Molt probability:

The predicted molt probabilities vs. CL for all scenarios are depicted in Figures 18 and 36 for EAG and WAG, respectively. The fits appear to be satisfactory.

1. Fit to catch size compositions:

Retained, total, and groundfish discard length compositions are shown in Figures 11 to 13 for EAG and 29 to 31 for WAG. The retained and total catch size composition fits appear satisfactory. But, the fits to groundfish bycatch size compositions are bad. Note that we did not use the groundfish size composition in any of the model scenario fits.

We illustrate the standardized residual plots as bubble plots of size composition over time for retained catch (Figures 21 and 23 for EAG, and 39 and 41 for WAG) and for total catch (Figures 22 and 24 for EAG, and 40 and 42 for WAG) for two scenarios (17_0 and 17_0d). The retained catch bubble plots appear random for the selected scenarios.
m . Marginal distributions for the fits to the composition data:
We did not provide this plot in this report.
n. Plots of implied versus input effective sample sizes and time series of implied effective sample sizes:
We did not provide the plots, but provided the estimated values in Tables 5 to 11 for EAG and in Tables 23 to 28 for WAG, respectively.
o. Tables of RMSEs for the indices:

We did not provide this table in this report.
p. Quantile-quantile (Q-Q) plots:

We did not provide these plots for model fits in this report. However, we provided these plots in a separate Appendix F for CPUE standardization diagnostic.

## 6. Retrospective and historical analysis:

The retrospective fits for scenarios 17_0 and 17_0d are shown in Figure 25 for EAG and in Figure 43 for WAG. The retrospective fits were prepared for the whole time series 1961 to 2017. The retrospective patterns did not show severe departure when four terminal years' data were removed systematically, especially for WAG and
hence the current formulation of the model appears stable. The Mohn rho values are also given in the figures, which indicate no severe model misspecification (i.e., small rho) (Mohn, 1999; Deroba, 2014). A severe drop in modeled biomass from the initial MMB occurred when the fishery time series started in 1981.
7. Uncertainty and sensitivity analysis:

- The main task was to determine a plausible size transition matrix to project the population over time. In a previous study, we investigated the sensitivity of the model to determining the size transition matrix by using or not using a molt probability function (Siddeek et al. 2016a). The model fit is better when the molt probability model is included. Therefore, we included a molt probability submodel for the size transition matrix calculation in all scenarios.
- We also determined likelihood values at different $M$ and plotted component negative likelihood against $M$ (Figure 1).


## 8. Conduct 'jitter analysis':

We conducted the (random) jitter analysis on scenario 17_0 (base) model fitted parameters. This analysis indicated that the base model achieved the global convergence (details in Appendix E).

## F. Calculation of the OFL

## 1. Specification of the Tier level:

Aleutian Islands golden king crab has been elevated to Tier 3 level in 2017 for OFL and ABC determination. In the following section, we provide the method to determine OFL and ABC

## 2. List of parameter and stock size estimates (or best available proxies thereof) required by limit and target control rules specified in the fishery management plan:

The critical assumptions for $M M B B_{M S Y}$ reference point estimation are:
a. Natural mortality is constant.
b. Growth transition matrix is fixed and estimated using tagging data with the molt probability sub-model.
c. Total fishery selectivity and retention curves are length dependent and the 2005/062016/17 period selectivity estimates are used.
d. Groundfish bycatch fishery selectivity is kept constant at 1.0 for all length groups.
e. Model estimated recruits (in millions of crab) are averaged for the time period 19872012.
f. Model estimated groundfish bycatch mortality values are averaged for the period 2007/08 - 2016/17 (10 years).
g. A knife-edge $50 \%$ maturity size is used for MMB estimation.

## Method:

We simulated the population abundance starting from the model estimated terminal year stock size by length, model estimated parameter values, a fishing mortality value ( F ), and adding a constant number of annual recruits. Once the stock dynamics were stabilized (we used the $99^{\text {th }}$ year estimates) for an F , we calculated the $\mathrm{MMB} / \mathrm{R}$ for that F . We computed the relative $M M B / R$ in percentage, $\left(\frac{M M B}{R}\right)_{x \%}\left(\right.$ where $\mathrm{x} \%=\frac{\frac{M M B_{F}}{R}}{\frac{M M B_{0}}{R}} \times 100$ and $M M B_{0} / R$ is the virgin $M M B / R$ ) for different F values.
$F_{35 \%}$ is the F value that produces the $\mathrm{MMB} / \mathrm{R}$ value equal to $35 \%$ of $M M B_{0} / R$.
$M M B_{35 \%}$ is estimated using the following formula:
$M M B_{35 \%}=\left(\frac{M M B}{R}\right)_{35} \times \bar{R}$, where $\bar{R} \quad$ is the mean number of model estimated recruits for a selected period.

## 3. Specification of the OFL:

## a. Provide the equations (from Amendment 24) on which the OFL is to be based:

$F_{O F L}$ is determined using Equation A. 28 in Appendix A. The OFL is estimated by an iterative procedure accounting for intervening total removals (see Appendix A for the formulas).
b. Basis for projecting MMB to the time of mating:

We followed the NPFMC 2007a guideline.
c. Specification of Fofl, OFL, and other applicable measures (if any) relevant to determining whether the stock is overfished or if overfishing is occurring:
See Management Performance table, below. The OFL and ABC values for 2018/19 in the table below are the recommended values. The TACs for 2013/14-2015/16 in the table below do not include landings towards a cost-recovery fishery goal, but the catches towards cost-recovery fishing in 2013/14-2014/15 are included in the retained and total catch.

Status and catch specifications (1000 t) of Aleutian Islands golden king crab

| Year | MSST | Biomass <br> (MMB) | TAC | Retained <br> Catch | Total <br> Catch $^{\mathbf{a}}$ | OFL | ABC $^{\mathbf{b}}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $2013 / 14$ | N/A | N/A | 2.853 | 2.894 | 3.192 | 5.69 | 5.12 |
| $2014 / 15$ | N/A | N/A | 2.853 | 2.771 | 3.088 | 5.69 | 4.26 |
| $2015 / 16$ | N/A | N/A | 2.853 | 2.729 | 3.076 | 5.69 | 4.26 |
| $2016 / 17$ | N/A | N/A | 2.515 | 2.593 | 2.947 | 5.69 | 4.26 |
| $2017 / 18$ | 6.044 | 14.205 | 2.515 | 2,585 | 2,942 | 6.048 | 4.536 |
| $2018 / 19^{\text {c }}$ | 6.046 | 17.952 |  |  |  | 5.514 | 4.136 |
| $2018 / 19^{\text {d }}$ | 5.898 | 14.665 |  |  |  | 3.963 | 2.972 |
| $2018 / 19^{\text {e }}$ | 6.107 | 17.793 |  |  |  | 5.581 | 4.186 |

a. Total retained catch plus estimated bycatch mortality of discarded bycatch during crab fisheries and groundfish fisheries.
b. $25 \%$ buffer was applied to total catch OFL to determine ABC.
c. 17_0 base scenario with Francis method of re-weighting
d. 17_0d three catchability and total selectivity scenario with Francis method of reweighting
e. 17_0e McAllister and Ianelli method of re-weighting

Status and catch specifications (million lb) of Aleutian Islands golden king crab

| Year | MSST | Biomass <br> (MMB) | TAC | Retained <br> Catch | Total <br> Catch $^{\mathbf{a}}$ | OFL | ABC $^{\mathbf{b}}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $2013 / 14$ | N/A | N/A | 6.290 | 6.38 | 7.038 | 12.54 | 11.28 |
| $2014 / 15$ | N/A | N/A | 6.290 | 6.11 | 6.807 | 12.53 | 9.40 |
| $2015 / 16$ | N/A | N/A | 6.290 | 6.016 | 6.782 | 12.53 | 9.40 |
| $2016 / 17$ | N/A | N/A | 5.545 | 5.716 | 6.497 | 12.53 | 9.40 |
| $2017 / 18$ | 13.325 | 31.315 | 5.545 |  |  | 13.333 | 10.000 |
| $2018 / 19^{\text {c }}$ | 13.329 | 39.577 |  |  |  | 12.157 | 9.118 |
| $2018 / 19^{\text {d }}$ | 13.002 | 32.331 |  |  |  | 8.737 | 6.553 |
| $2018 / 19^{\text {e }}$ | 13.464 | 39.227 |  |  |  | 12.305 | 9.228 |

a. Total retained catch plus estimated bycatch mortality of discarded bycatch during crab fisheries and groundfish fisheries.
b. $25 \%$ buffer was applied to total catch OFL to determine ABC.
c. 17_0 base scenario with Francis method of re-weighting
d. 17_0d three catchability and total selectivity scenario with Francis method of reweighting
e. 17_0e McAllister and Ianelli method of re-weighting

## 4. Specification of the retained portion of the total catch OFL:

The retained catch portion of the total-catch OFL for EAG, WAG, and the entire Aleutian Islands (AI) stock were calculated for the three recommended scenario options (17_0, 17_0d, and 17_0e):

Scenario 17_0:
EAG: $3,756 \mathrm{t}$ ( 8.280 million lb)
WAG: $1,473 \mathrm{t}$ ( 3.248 million lb)
AI: $\quad 5,229 \mathrm{t}$ ( 11.528 million lb).
Scenario 17_0d:
EAG: 2,355 t (5.191 million lb)
WAG: $1,375 \mathrm{t}$ ( 3.031 million lb )
AI: $3,730 \mathrm{t}$ (8.222 million lb).
Scenario 17_0e:
EAG: $3,817 \mathrm{t}$ ( 8.415 million lb )
WAG: $1,484 \mathrm{t}$ ( 3.271 million lb)
AI: $\quad 5,301 \mathrm{t}(11.686$ million lb$)$.

## G. Calculation of ABC

1. We estimated the cumulative probability distribution of OFL assuming a $\log$ normal distribution of OFL. We calculated the OFL at the 0.5 probability and the maximum ABC at the 0.49 probability and considered additional buffer by setting $\mathrm{ABC}=0.75^{*} \mathrm{OFL}$
We provide the ABC estimates with the $25 \%$ buffer for EAG, WAG, and AI considering scenarios 17_0, 17_0d, and 17_0e:

Scenario 17_0:
EAG: $\mathrm{ABC}=2,938 \mathrm{t}(6.478$ million lb)
WAG: $\mathrm{ABC}=1,197 \mathrm{t}(2.640$ million lb$)$
$\mathrm{AI}: \mathrm{ABC}=4,136 \mathrm{t}$ (9.118 million lb).
Scenario 17_0d:
EAG: $\mathrm{ABC}=1,860 \mathrm{t}$ ( 4.102 million lb)
WAG: $\mathrm{ABC}=1,112 \mathrm{t}(2.451$ million lb$)$
$\mathrm{AI}: \mathrm{ABC}=2,972 \mathrm{t}$ (6.553 million lb).
Scenario 17_0e:
EAG: $\mathrm{ABC}=2,980 \mathrm{t}$ ( 6.570 million lb)
WAG: $\mathrm{ABC}=1,206 \mathrm{t}(2.658$ million lb$)$
AI: $\mathrm{ABC}=4,186 \mathrm{t}$ (9.228 million lb).

## 2. List of variables related to scientific uncertainty:

- Model relied largely on fisheries data.
- Observer and fisheries CPUE indices played a major role in the assessment model.
- Natural mortality was estimated in the model and independent estimate is not available.
- The time period to compute the average number of recruits (1987-2012) relative to the assumption that this represents "a time period determined to be representative of the production potential of the stock."
- Fixed bycatch mortality rates were used in each fishery (crab fishery and the groundfish fishery) that discarded golden king crab.
- Discarded catch and bycatch mortality for each fishery that bycatch occurred in during 1981/82-1989/90 were not available.

3. List of additional uncertainties for alternative sigma-b.

We recommended a large buffer of $25 \%$ to account for additional uncertainties.

## 4. Author recommended ABC :

Authors recommended three ABC options based on $25 \%$ buffer on the OFL under scenarios 17_0, 17_0d, and 17_0e.

## H. Rebuilding Analysis

Not applicable. This stock has not been declared overfished.

## I. Data Gaps and Research Priorities

1. The recruit abundances were estimated from commercial catch sampling data. The implicit assumption in the analysis was that the estimated recruits come solely from the same exploited stock through growth and mortality. The current analysis did not consider the possibility that additional recruitment may occur through immigration from neighboring areas and possibly separate sub-stocks. Extensive tagging experiments or resource surveys are needed to investigate stock distributions.
2. We estimated $M$ in the model. However, an independent estimate of $M$ is needed for comparison, which could be achieved with tagging experiments.
3. An extensive tagging study will also provide independent estimates of molting probability and growth. We used the historical tagging data to determine the size transition matrix.
4. An arbitrary $20 \%$ handling mortality rate on discarded males was used, which was obtained from the red king crab literature (Kruse et al. 2000; Siddeek 2002). An experimentally-based independent estimate of handling mortality is needed for Aleutian Islands golden king crab.
5. The Aleutian King Crab Research Foundation recently initiated crab survey programs in the Aleutian Islands. This program needs to be strengthened and continued for golden king crab research to address some of the data gaps and establish a fishery independent data source.
6. We have been using the length-weight relationship established based on late 1990s data for golden king crab. The Aleutian King Crab Research Foundation program can help us to update this relationship by collecting new length weight information.
7. We have recently included male maturity data in the model to determine a maturity curve for MMB estimation. The maturity data available to us were collected in 1984 and 1991. More data and recent data are needed.
8. Morphometric measurements provide morphometric maturity size. Ideally, an experimental study under natural environment condition is needed to collect male size at functional maturity data to determine functional maturity size.

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## K. Literature Cited

Adams, C.F., and A.J. Paul. 1999. Phototaxis and geotaxis of light-adapted zoeae of the golden king crab Lithodes aequispinus (Anomura: Lithodidae) in the laboratory. Journal of Crustacean Biology. 19(1): 106-110.

ADF\&G (Alaska Department of Fish and Game). 2002. Annual management report for the shellfish fisheries of the Westward Region, 2001. Alaska Department of Fish and Game, Division of Commercial Fisheries, Regional Information Report 4K02-54, Kodiak, Alaska.

Barnard, D.R., and R. Burt. 2004. Summary of the 2002 mandatory shellfish observer program database for the general and CDQ fisheries. Alaska Department of Fish and Game, Division of Commercial Fisheries, Regional Information Report 4K04-27, Kodiak, Alaska.

Barnard, D.R., R. Burt, and H. Moore. 2001. Summary of the 2000 mandatory shellfish observer program database for the open access fisheries. Alaska Department of Fish and Game, Division of Commercial Fisheries, Regional Information Report 4K01-39, Kodiak, Alaska.

Beers, D.E. 1992. Annual biological summary of the Westward Region shellfish observer database, 1991. Alaska Department of Fish and game, Division of Commercial Fisheries, Regional Information Report 4K92-33, Kodiak.

Blau, S.F., and D. Pengilly. 1994. Findings from the 1991 Aleutian Islands golden king crab survey in the Dutch Harbor and Adak management areas including analysis of recovered tagged crabs. Alaska Department of Fish and Game, Commercial Fisheries Management and Development Division, Regional Information Report 4K94-35, Kodiak.

Blau, S.F., L.J. Watson, and I. Vining. 1998. The 1997 Aleutian Islands golden king crab survey. Alaska Department of Fish and Game, Commercial Fisheries Management and Development Division, Regional Information Report 4K98-30, Kodiak.

Bowers, F.R., M. Schwenzfeier, S. Coleman, B.J. Failor-Rounds, K. Milani, K. Herring, M. Salmon, and M. Albert. 2008. Annual management report for the commercial and subsistence shellfish fisheries of the Aleutian Islands, Bering Sea, and the Westward Region's shellfish observer program, 2006/07. Alaska Department of Fish and Game, Divisions of Sport Fish and Commercial Fisheries, Fishery Management Report No. 08-02, Anchorage, Alaska.

Bowers, F.R., M. Schwenzfeier, K. Herring, M. Salmon, J. Shaishnikoff, H. Fitch, J. Alas, and B. Baechler. 2011. Annual management report for the commercial and subsistence shellfish fisheries of the Aleutian Islands, Bering Sea, and the Westward Region's shellfish observer program, 2009/10. Alaska Department of Fish and Game, Divisions of Sport Fish and Commercial Fisheries, Fishery Management Report No. 11-05, Anchorage, Alaska.

Burnham, K.P., D.R. Anderson, G.C. White, C. Brownie, and K.H. Pollock. 1987. Design and analysis methods for fish survival experiments based on release-recapture. American Fisheries Society, Monograph 5, 437p.

Deroba, J.J. 2014. Evaluating the consequences of adjusting fish stock assessment estimates of biomass for retrospective patterns using Mohn's rho. North American Journal of Fisheries Management 34:380-390.

Fournier, D.A., H.J. Skaug, J. Ancheta, J. Ianelli, A. Magnusson, M.N. Maunder, A. Nielsen, and J. Sibert. 2012. AD Model Builder: using automatic differentiation for statistical inference of highly parameterized complex nonlinear models. Optim. Methods Softw. 27:233-249.

Fox, J., and S. Weisberg. 2011. An R Companion to Applied Regression. Second edition. Sage Publications, Inc. 449 p.

Francis, R.I.C.C. 2011. Data weighting in statistical fisheries stock assessment models. Canadian Journal of. Fisheries and Aquatic Sciences 68: 1124-1138.

Francis, R.I.C.C. (2017). Revisiting data weighting in fisheries stock assessment models. Fisheries Research 192: 5-15.

Gaeuman, W.B. 2014. Summary of the 2013/2014 Mandatory Crab Observer Program Database for the Bering Sea/Aleutian Islands commercial crab fisheries. Alaska Department of Fish and Game, Fishery Data Series No. 14-49, Anchorage.
Gaeuman, W.B. 2011. Summary of the 2009/10 mandatory crab observer program database for the BSAI commercial crab fisheries. Fishery Data Series No. 11-04. Alaska Department of Fish and Game, Kodiak.

Hartill, T. 2012. Annual management report for the community development quota and Adak Community Allocation crab fisheries in the Bering Sea and Aleutian Islands, 2010/11. Pages 177-194 in Fitch, H., M. Schwenzfeier, B. Baechler, T. Hartill, M. Salmon, M. Deiman, E. Evans, E. Henry, L. Wald, J. Shaishnikoff, and K. Herring. Annual management report for the commercial and subsistence shellfish fisheries of the Aleutian Islands, Bering Sea and the Westward Region's Shellfish Observer Program, 2010/11. Alaska Department of Fish and Game, Fishery Management Report No. 12-22, Anchorage.
Hilsinger, J., C. Siddon and L. Hulbert. 2016. Cooperative red king crab survey in the Adak area, 2015. Anchorage., Alaska Department of Fish and Game, Fishery Data Series No. 16-18.

Hiramoto, K. 1985. Overview of the golden king crab, Lithodes aequispina, fishery and its fishery biology in the Pacific waters of Central Japan. Pages 297-315, In: Proceedings of the International King Crab Symposium. Alaska Sea Grant College Program, AK-SG-85-12, Fairbanks, Alaska.
Jewett, S.C., N.A. Sloan, and D.A. Somerton. 1985. Size at sexual maturity and fecundity of the fjord-dwelling golden king crab Lithodes aequispina Benedict from northern British Columbia. Journal of Crustacean Biology 5: 377-385.

Koeneman, T.M., and D.V. Buchanan. 1985. Growth of the golden king crab, Lithodes aequispina, in Southeast Alaskan waters. Pages 281-297, In: Proceedings of the International King Crab Symposium. Alaska Sea Grant College Program, AK-SG-85-12, Fairbanks, Alaska.

Kruse, G.H., L.C. Byrne, F.C. Funk, S.C. Matulich, and J. Zheng. 2000. Analysis of minimum size limit for the red king crab fishery in Bristol Bay, Alaska. N. Am. J. Fish. Manage. 20:307-319.

Leon, J. M., J. Shaishnikoff, E. Nichols, and M. Westphal. 2017. Annual management report for shellfish fisheries of the Bering Sea-Aleutian Islands management area, 2015/16. Alaska Department of Fish and Game, Fishery Management Report No. 17-10, Anchorage.

Maunder, M.N., and A.E. Punt. 2004. Standardizing catch and effort data: a review of recent approaches. Fisheries Research 70: 141-159.
McAllister, M.K., and J.N. Ianelli. 1997. Bayesian stock assessment using catch-age data and the sampling/importance resampling algorithm. Canadian Journal of. Fisheries and Aquatic Sciences 54: 284-300.

Mohn, R. 1999. The retrospective problem in sequential population analysis: An investigation using cod fishery and simulated data. ICES Journal of Marine Science 56:473-488.

Moore, H., L.C. Byrne, and M.C. Schwenzfeier. 2000. Summary of the 1999 mandatory shellfish observer program database for the open access fisheries. Alaska Department of Fish and Game, Division of Commercial Fisheries, Regional Information Report 4K00-50, Kodiak, Alaska.

Morrison, R., R.K. Gish, and M. Ruccio. 1998. Annual management report for the shellfish fisheries of the Aleutian Islands. Pages 82-139 in ADF\&G. Annual management report for the shellfish fisheries of the Westward Region. Alaska Department of Fish and Game, Division of Commercial Fisheries, Regional Information Report 4K98-39, Kodiak.

National Marine Fisheries Service (NMFS). 2004. Bering Sea Aleutian Islands Crab Fisheries Final Environmental Impact Statement. National Marine Fisheries Service, Alaska Region, Juneau, August 2004.

North Pacific Fishery Management Council (NPFMC). 2007a. Initial Review Draft: Environmental Assessment for proposed Amendment 24 to the Fishery Management Plan for Bering Sea and Aleutian Islands King and Tanner Crabs to Revise Overfishing Definitions. 17 January 2007. North Pacific Fishery Management Council, Anchorage.

North Pacific Fishery Management Council (NPFMC). 2007b. Public Review Draft: Environmental Assessment for proposed Amendment 24 to the Fishery Management Plan for Bering Sea and Aleutian Islands King and Tanner Crabs to Revise Overfishing Definitions. 14 November 2007. North Pacific Fishery Management Council, Anchorage.

Otto, R.S., and P.A. Cummiskey. 1985. Observations on the reproductive biology of golden king crab (Lithodes aequispina) in the Bering Sea and Aleutian Islands. Pages 123-135 In: Proceedings of the International King Crab Symposium. Alaska Sea Grant College Program, AK-SG-85-12, Fairbanks, Alaska.

Pengilly, D. 2016. Aleutian Islands golden king crab - 2016 Tier 5 assessment. 2016 Crab SAFE report chapter (September 2016). North Pacific Fishery Management Council, Anchorage, Alaska.

Punt, A.E., Kennedy, R.B., Frusher, S.D., 1997. Estimating the size-transition matrix for Tasmanian rock lobster, Jasus edwardsii. Mar. Freshw. Res 48, 981-982.

Punt, A.E. (2017). Some insights into data weighting in integrated stock assessments. Fisheries Research 192:52-65.

R Core Team. 2017. R: A language and environment for statistical computing. R Foundation for Statistical Computing, Vienna, Austria. URL http://www.R-project.org/.

Runnebaum, J., Guan, L., Cao, J., O’Brien, L., Chen, Y. 2017. Habitat suitability modeling based on a spatio-temporal model: an example for Cusk in the Gulf of Maine. Canadian Journal of. Fisheries and Aquatic Sciences (in press).

Shirley, T.C., and S. Zhou. 1997. Lecithotrophic development of the golden king crab Lithodes aequispinus (Anomura: Lithodidae). J. Crust. Biol. 17(2):207-216.

Siddeek, M.S.M. 2002. Review of biological reference points used in Bering Sea and Aleutian Islands (king and Tanner) crab management. Alaska Department of Fish and Game, Division of Commercial Fisheries, Regional Information Report 5J02-06, Juneau, Alaska.

Siddeek, M.S.M., D.R. Barnard, L.J. Watson, and R.K. Gish. 2005. A modified catch-length analysis model for golden king crab (Lithodes aequispinus) stock assessment in the eastern Aleutian Islands. Pages 783-805 in Fisheries assessment and management in data limited situations, Alaska Sea Grant College Program, AK-SG-05-02, Fairbanks, Alaska.

Siddeek, M.S.M., J. Zheng, and D. Pengilly. 2015. Aleutian Islands Golden King Crab (Lithodes aequispinus) Model-Based Stock Assessment in Fall 2015. Draft report submitted for the September 2015 Crab Plan Team Meeting. North Pacific Fishery Management Council, Anchorage, Alaska.

Siddeek, M.S.M., J. Zheng, A.E. Punt, and Vicki Vanek. 2016a. Estimation of size-transition matrices with and without moult probability for Alaska golden king crab using tag-recapture data. Fisheries Research 180:161-168.

Siddeek, M.S.M., Jie Zheng, and Doug Pengilly 2016b. Standardizing CPUE from the Aleutian Islands golden king crab observer data. In: T.J. Quinn II, J.L. Armstrong, M.R. Baker, J. Heifetz, and D. Witherell (eds.), Assessing and Managing Data-Limited Fish Stocks. Alaska Sea Grant, University of Alaska Fairbanks, Alaska, USA, pp. 97-116.

Siddeek, M.S.M., J. Zheng, and D. Pengilly. 2016c. Aleutian Islands Golden King Crab (Lithodes aequispinus) Model-Based Stock Assessment in Spring 2016. Draft report submitted for the May 2016 Crab Plan Team Meeting. North Pacific Fishery Management Council, Anchorage, Alaska.
Siddeek, M.S.M., J. Zheng, A.E. Punt, and D. Pengilly 2017. Effect of data weighting on the mature male biomass estimate for Alaskan golden king crab. CAPAM Data weighting Workshop, San Diego, California. Fisheries Research 142: 103-113.

Somerton, D.A., and R.S. Otto. 1986. Distribution and reproductive biology of the golden king crab, Lithodes aequispina, in the Eastern Bering Sea. Fishery Bulletin 81(3): 571-584.

Starr, P.J. 2012. Standardized CPUE analysis exploration: using the rock lobster voluntary logbook and observer catch sampling programmes. New Zealand Fisheries Assessment Report 2012/34, 75 p.
Thorson, J.T., Shelton A.O., Ward, E.J., and Skaug, H.J. 2015. Geostatistical delta-generalized linear mixed models improve precision for estimated abundance indices for West Coast groundfishes. ICES Journal of Marine Sciences 72(5):1297-1310.

Thorson, J.T., and Barnett, L.A.K. 2017. Comparing estimates of abundance trends and distribution shifts using single- and multi-species models of fishes and biogenic habitat. ICES Journal of Marine Sciences 74(5):1311-1321.

Thorson, J.T., Ianelli, J.N., and Kotwicki, S. 2017. Relative influence of temperature and sizestructure on fish distribution shifts. A case study on Walleye Pollock in the Bering Sea. Fish and Fisheries 2017; 1-12.

Vanek, V., Pengilly, D., and Siddeek, M.S.M.. 2013. A study of commercial fishing gear selectivity during the 2012/13 Aleutian Islands Golden King Crab Fishery East of $174^{\circ}$ W

Longitude. Fishery Data Series No. 13-41. Alaska Department of Fish and Game, Division of Sport Fish and Commercial Fisheries, Research and Technical Services, 333 Raspberry Road, Anchorage, Alaska 99518-1565.
Von Szalay, P.G., C.N. Roper, N.W. Raring, and M.H. Martin. 2011. Data report: 2010 Aleutian Islands bottom trawl survey. U.S. Dep. Commerce., NOAA Technical Memorandum NMFS-AFSC-215.

Watson, L.J. 2004. The 2003 triennial Aleutian Islands golden king crab survey and comparisons to the 1997 and 2000 surveys (revised October 17, 2005). Alaska Department of Fish and Game, Division of Commercial Fisheries, Regional Information Report 4K04-42, Kodiak. [Revised 10/17/2005].

Watson, L.J. 2007. The 2006 triennial Aleutian Islands golden king crab survey. Alaska Department of Fish and Game, Fishery Management Report No. 07-07, Anchorage.
Watson, L.J., and R.K. Gish. 2002. The 2000 Aleutian Islands golden king crab survey and recoveries of tagged crabs in the 1997-1999 and 2000-2002 fishing seasons. Alaska Department of Fish and Game, Division of Commercial Fisheries, Regional Information Report 4K02-6, Kodiak.

Watson, L.J., D. Pengilly, and S.F. Blau. 2002. Growth and molting of golden king crabs (Lithodes aequispinus) in the eastern Aleutian Islands, Alaska. Pages 169-187 in Crabs in cold water regions: biology, management, and economics, Alaska Sea Grant College Program, AK-SG-02-01, Fairbanks, Alaska.

Webb. J. 2014. Reproductive ecology of commercially important Lithodid crabs. Pages 285-314 in B.G. Stevens (ed.). King Crabs of the World: Biology and Fisheries Management. CRC Press, Taylor \& Francis Group, New York.

Table 1. Commercial fishery history for the Aleutian Islands golden king crab fishery1981/82-2015/16: number of vessels, guideline harvest level (GHL; established in lb, converted to t) for 1996/97-2004/05, total allowable catch (TAC; established in lb, converted to t) for 2005/06-2016/17, weight of retained catch (Harvest; t), number of retained crab, pot lifts, fishery catch per unit effort (CPUE; retained crab per pot lift), and average weight (kg) of landed crab. The values are separated by EAG and WAG beginning 1996/97.

| Crab <br> Fishing <br> Season | Vessels | GHL/TAC | Harvest ${ }^{\text {a }}$ | Crab ${ }^{\text {b }}$ | Pot Lifts | CPUE ${ }^{\text {b }}$ | Average Weight ${ }^{\mathbf{c}}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1981/82 | 14-20 | - | 599 | 240,458 | 27,533 | 9 | $2.5{ }^{\text {d }}$ |
| 1982/83 | 99-148 | - | 4,169 | 1,737,109 | 179,472 | 10 | $2.4{ }^{\text {d }}$ |
| 1983/84 | 157-204 | - | 4,508 | 1,773,262 | 256,393 | 7 | $2.5{ }^{\text {d }}$ |
| 1984/85 | 38-51 | - | 2,132 | 971,274 | 88,821 | 11 | $2.2{ }^{\text {e }}$ |
| 1985/86 | 53 | - | 5,776 | 2,816,313 | 236,601 | 12 | $2.1{ }^{\text {f }}$ |
| 1986/87 | 64 | - | 6,685 | 3,345,680 | 433,870 | 8 | $2.0{ }^{\text {f }}$ |
| 1987/88 | 66 | - | 4,199 | 2,177,229 | 307,130 | 7 | $1.9{ }^{\text {f }}$ |
| 1988/89 | 76 | - | 4,820 | 2,488,433 | 321,927 | 8 | $1.9{ }^{\text {f }}$ |
| 1989/90 | 68 | - | 5,453 | 2,902,913 | 357,803 | 8 | $1.9{ }^{\text {f }}$ |
| 1990/91 | 24 | - | 3,153 | 1,707,618 | 215,840 | 8 | $1.9{ }^{\text {f }}$ |
| 1991/92 | 20 | - | 3,494 | 1,847,398 | 234,857 | 8 | $1.9{ }^{\text {f }}$ |
| 1992/93 | 22 | - | 2,854 | 1,528,328 | 203,221 | 8 | $1.9{ }^{\text {f }}$ |
| 1993/94 | 21 | - | 2,518 | 1,397,530 | 234,654 | 6 | $1.8{ }^{\text {f }}$ |
| 1994/95 | 35 | - | 3,687 | 1,924,271 | 386,593 | 5 | $1.9{ }^{\text {f }}$ |


| Crab <br> Fishing <br> Season | Vessels |  | GHL/TAC |  | Harvest ${ }^{\text {a }}$ |  | Crab ${ }^{\text {b }}$ |  | Pot Lifts |  | CPUE ${ }^{\text {b }}$ |  | Average Weight ${ }^{\text {c }}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1995/96 | 28 |  | - |  | 3,157 |  | 1,582,333 |  | 293,021 |  | 5 |  | $2.0{ }^{\text {f }}$ |  |
|  | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG |
| 1996/97 | 14 | 13 | 1,452 | 1,225 | 1,493 | 1,145 | 731,909 | 602,968 | 113,460 | 99,267 | 7 | 6 | $2.04{ }^{\text {f }}$ | $1.91{ }^{\text {f }}$ |
| 1997/98 | 13 | 9 | 1,452 | 1,225 | 1,588 | 1,109 | 780,610 | 569,550 | 106,403 | 86,811 | 7 | 7 | $2.04{ }^{\text {f }}$ | $1.95{ }^{\text {f }}$ |
| 1998/99 | 14 | 3 | 1,361 | 1,225 | 1,473 | 768 | 740,011 | 410,018 | 83,378 | 35,975 | 9 | 11 | $2.00^{\text {f }}$ | $1.86{ }^{\text {f }}$ |
| 1999/00 | 15 | 15 | 1,361 | 1,225 | 1,392 | 1,256 | 709,332 | 676,558 | 79,129 | 107,040 | 9 | 6 | $1.95{ }^{\text {f }}$ | $1.86{ }^{\text {f }}$ |
| 2000/01 | 15 | 12 | 1,361 | 1,225 | 1,422 | 1,308 | 704,702 | 705,613 | 71,551 | 101,239 | 10 | 7 | $2.00^{\text {f }}$ | $1.86{ }^{\text {f }}$ |
| 2001/02 | 19 | 9 | 1,361 | 1,225 | 1,442 | 1,243 | 730,030 | 686,738 | 62,639 | 105,512 | 12 | 7 | $2.00^{\text {f }}$ | $1.81{ }^{\text {f }}$ |
| 2002/03 | 19 | 6 | 1,361 | 1,225 | 1,280 | 1,198 | 643,886 | 664,823 | 52,042 | 78,979 | 12 | 8 | $2.00^{\text {f }}$ | $1.81{ }^{\text {f }}$ |
| 2003/04 | 18 | 6 | 1,361 | 1,225 | 1,350 | 1,220 | 643,074 | 676,633 | 58,883 | 66,236 | 11 | 10 | $2.09{ }^{\text {f }}$ | $1.81{ }^{\text {f }}$ |
| 2004/05 | 19 | 6 | 1,361 | 1,225 | 1,309 | 1,219 | 637,536 | 685,465 | 34,848 | 56,846 | 18 | 12 | $2.04{ }^{\text {f }}$ | $1.77{ }^{\text {f }}$ |
| 2005/06 | 7 | 3 | 1,361 | 1,225 | 1,300 | 1,204 | 623,971 | 639,368 | 24,569 | 30,116 | 25 | 21 | $2.09{ }^{\text {f }}$ | $1.91{ }^{\text {f }}$ |
| 2006/07 | 6 | 4 | 1,361 | 1,225 | 1,357 | 1,030 | 650,587 | 527,734 | 26,195 | 26,870 | 25 | 20 | $2.09^{\text {f }}$ | $1.95{ }^{\text {f }}$ |
| 2007/08 | 4 | 3 | 1,361 | 1,225 | 1,356 | 1,142 | 633,253 | 600,595 | 22,653 | 29,950 | 28 | 20 | $2.13{ }^{\text {f }}$ | $1.91{ }^{\text {f }}$ |
| 2008/09 | 3 | 3 | 1,361 | 1,286 | 1,426 | 1,150 | 666,946 | 587,661 | 24,466 | 26,200 | 27 | 22 | $2.13{ }^{\text {f }}$ | $1.95{ }^{\text {f }}$ |
| 2009/10 | 3 | 3 | 1,429 | 1,286 | 1,429 | 1,253 | 679,886 | 628,332 | 29,298 | 26,489 | 26 | 24 | $2.09^{\text {f }}$ | $2.00^{\text {f }}$ |
| 2010/11 | 3 | 3 | 1,429 | 1,286 | 1,428 | 1,279 | 670,983 | 626,246 | 25,851 | 29,994 | 26 | 21 | $2.13{ }^{\text {f }}$ | $2.04{ }^{\text {f }}$ |

$\qquad$

| Crab <br> Fishing <br> Season | Vessels |  | GHL/TAC |  | Harvest ${ }^{\text {a }}$ |  | Crab ${ }^{\text {b }}$ |  | Pot Lifts |  | CPUE ${ }^{\text {b }}$ |  | Average Weight ${ }^{\text {c }}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG |
| 2011/12 | 3 | 3 | 1,429 | 1,286 | 1,429 | 1,276 | 668,828 | 616,118 | 17,915 | 26,326 | 37 | 23 | $2.13{ }^{\text {f }}$ | $2.09^{\text {f }}$ |
| 2012/13 | 3 | 3 | 1,501 | 1,352 | 1,504 | 1,339 | 687,666 | 672,916 | 20,827 | 32,716 | 33 | 21 | $2.18{ }^{\text {f }}$ | $2.00^{\text {f }}$ |
| 2013/14 | 3 | 3 | 1,501 | 1,352 | 1,546 | 1,347 | 720,220 | 686,883 | 21,388 | 41,835 | 34 | 16 | $2.13{ }^{\text {f }}$ | $1.95{ }^{\text {f }}$ |
| 2014/15 | 3 | 2 | 1,501 | 1,352 | 1,554 | 1,217 | 719,064 | 635,312 | 17,002 | 41,548 | 42 | 15 | $2.18{ }^{\text {f }}$ | $1.91{ }^{\text {f }}$ |
| 2015/16 | 3 | 2 | 1,501 | 1,352 | 1,590 | 1,139 | 763,604 | 615,355 | 19,376 | 41,108 | 39 | 15 | $2.09^{\text {f }}$ | $1.85{ }^{\text {f }}$ |
| 2016/17 | 3 | 3 | 1,501 | 1,014 | 1,578 | 1,015 | 793,983 | 543,796 | 24,470 | 38,118 | 32 | 14 | $1.99{ }^{\text {f }}$ | $1.87{ }^{\text {f }}$ |

Note:
a. Includes deadloss.
b. Number of crab per pot lift.
c. Average weight of landed crab, including deadloss.
d. Managed with $6.5^{\prime \prime}$ carapace width (CW) minimum size limit.
e. Managed with $6.5^{\prime \prime} \mathrm{CW}$ minimum size limit west of $171^{\circ} \mathrm{W}$ longitude and $6.0^{\prime \prime}$ minimum size limit east of $171^{\circ} \mathrm{W}$ longitude.
f. Managed with 6.0" minimum size limit.

Catch and effort data include cost recovery fishery.

Table 2. Annual weight of total fishery mortality to Aleutian Islands golden king crab, 1981/82 2016/17, partitioned by source of mortality: retained catch, bycatch mortality during crab fisheries, and bycatch mortality during groundfish fisheries. For bycatch in the federal groundfish fisheries, historical data (1991-2008) are not available for areas east and west of 174 W , and are listed for federal groundfish reporting areas 541,542 , and 543 combined. The 2009- present data are available by separate EAG and WAG fisheries and are listed as such. A mortality rate of $20 \%$ was applied for crab fisheries bycatch, and a mortality rate of $50 \%$ for groundfish pot fisheries and $80 \%$ for the trawl fisheries were applied.

| Season | Retained Catch (t) |  | Bycatch Mortality by Fishery Type (t) |  |  |  | Total Fishery Mortality (t) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Crab |  | Grou | dfish |  |  |  |
|  | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | $\begin{aligned} & \text { Entire } \\ & \text { AI } \end{aligned}$ |
| 1981/82 | 490 | 95 |  |  |  |  |  |  | 585 |
| 1982/83 | 1,260 | 2,655 |  |  |  |  |  |  | 3,914 |
| 1983/84 | 1,554 | 2,991 |  |  |  |  |  |  | 4,545 |
| 1984/85 | 1,839 | 424 |  |  |  |  |  |  | 2,263 |
| 1985/86 | 2,677 | 1,996 |  |  |  |  |  |  | 4,673 |
| 1986/87 | 2,798 | 4,200 |  |  |  |  |  |  | 6,998 |
| 1987/88 | 1,882 | 2,496 |  |  |  |  |  |  | 4,379 |
| 1988/89 | 2,382 | 2,441 |  |  |  |  |  |  | 4,823 |
| 1989/90 | 2,738 | 3,028 |  |  |  |  |  |  | 5,766 |
| 1990/91 | 1,623 | 1,621 |  |  |  |  |  |  | 3,244 |
| 1991/92 | 2,035 | 1,397 | 515 | 344 |  | 0 |  |  | 4,291 |
| 1992/93 | 2,112 | 1,025 | 1,206 | 373 |  | 0 |  |  | 4,716 |
| 1993/94 | 1,439 | 686 | 383 | 258 |  | 4 |  |  | 2,770 |
| 1994/95 | 2,044 | 1,540 | 687 | 823 |  | 1 |  |  | 5,095 |
| 1995/96 | 2,259 | 1,203 | 725 | 530 |  | 2 |  |  | 4,719 |
| 1996/97 | 1,738 | 1,259 | 485 | 439 |  | 5 |  |  | 3,926 |
| 1997/98 | 1,588 | 1,083 | 441 | 343 |  | 1 |  |  | 3,455 |
| 1998/99 | 1,473 | 955 | 434 | 285 |  | 1 |  |  | 3,149 |
| 1999/00 | 1,392 | 1,222 | 313 | 385 |  | 3 |  |  | 3,316 |
| 2000/01 | 1,422 | 1,342 | 82 | 437 |  | 2 |  |  | 3,285 |
| 2001/02 | 1,442 | 1,243 | 74 | 387 |  | 0 |  |  | 3,146 |
| 2002/03 | 1,280 | 1,198 | 52 | 303 |  | 18 |  |  | 2,850 |
| 2003/04 | 1,350 | 1,220 | 53 | 148 |  | 20 |  |  | 2,792 |
| 2004/05 | 1,309 | 1,219 | 41 | 143 |  | 1 |  |  | 2,715 |
| 2005/06 | 1,300 | 1,204 | 22 | 73 |  | 2 |  |  | 2,601 |
| 2006/07 | 1,357 | 1,022 | 28 | 81 |  | 18 |  |  | 2,506 |
| 2007/08 | 1,356 | 1,142 | 24 | 114 |  | 59 |  |  | 2,695 |
| 2008/09 | 1,426 | 1,150 | 61 | 102 |  | 33 |  |  | 2,772 |
| 2009/10 | 1,429 | 1,253 | 111 | 108 | 18 | 5 | 1,558 | 1,366 | 2,923 |
| 2010/11 | 1,428 | 1,279 | 123 | 124 | 49 | 3 | 1,600 | 1,407 | 3,006 |
| 2011/12 | 1,429 | 1,276 | 106 | 117 | 25 | 4 | 1,560 | 1,398 | 2,957 |
| 2012/13 | 1,504 | 1,339 | 118 | 145 | 9 | 6 | 1,631 | 1,491 | 3,122 |


| $2013 / 14$ | 1,546 | 1,347 | 113 | 174 | 5 | 7 | 1,665 | 1,528 | 3,192 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $2014 / 15$ | 1,554 | 1,217 | 127 | 175 | 9 | 5 | 1,691 | 1,397 | 3,088 |
| $2015 / 16$ | 1,590 | 1,139 | 165 | 157 | 23 | 2 | 1,778 | 1,298 | 3,076 |
| $2016 / 17$ | 1,578 | 1,015 | 203 | 145 | 3 | 3 | 1,785 | 1,163 | 2,947 |
| $2017 / 18$ | 1,571 | 1,014 | 219 | 126 | 10 | 2 | 1,801 | 1,142 | 2,942 |

Table 2a. Time series of estimated total male catch (weight of crabs on the deck without applying any handling mortality) for the EAG and WAG golden king crab stocks (1990/912016/17). The crab weights are for the size range $\geq 101 \mathrm{~mm}$ CL and Length-Weight formula was used to predict weight at the mid-point of each size bin. NA: no observer sampling to compute catch.

| Year | Total Catch <br> Biomass $(\mathbf{t})$ <br> EAG | Total Catch <br> Biomass $(\mathbf{t})$ <br> WAG |
| :---: | :---: | :---: |
| $1990 / 91$ | 3,672 | 3,736 |
| $1991 / 92$ | 3,946 | 2,275 |
| $1992 / 93$ | 5,570 | 1,500 |
| $1993 / 94$ | NA | 2,800 |
| $1994 / 95$ | 2,020 | 4,945 |
| $1995 / 96$ | 3,724 | 2,125 |
| $1996 / 97$ | 2,035 | 1,766 |
| $1997 / 98$ | 2,534 | 1,794 |
| $1998 / 99$ | 2,797 | 1,083 |
| $1999 / 00$ | 2,272 | 2,085 |
| $2000 / 01$ | 2,551 | 2,225 |
| $2001 / 02$ | 2,107 | 2,131 |
| $2002 / 03$ | 1,796 | 1,889 |
| $2003 / 04$ | 1,819 | 1,853 |
| $2004 / 05$ | 1,618 | 1,873 |
| $2005 / 06$ | 1,713 | 1,786 |
| $2006 / 07$ | 1,621 | 1,542 |
| $2007 / 08$ | 1,790 | 1,602 |
| $2008 / 09$ | 1,796 | 1,719 |
| $2009 / 10$ | 1,750 | 1,667 |
| $2010 / 11$ | 1,719 | 1,580 |
| $2011 / 12$ | 1,736 | 1,504 |
| $2012 / 13$ | 1,927 | 1,811 |
| $2013 / 14$ | 1,818 | 1,890 |
| $2014 / 15$ | 1,939 | 1,583 |
| $2015 / 16$ | 2,104 | 1,547 |
| $2016 / 17$ | 2,104 | 1,425 |
|  |  |  |

Table 3. Time series of nominal annual pot fishery retained, observer retained, and observer total catch-per-unit-effort (CPUE, number of crabs per pot lift), total pot fishing effort (number of pot lifts), observer sample size (number of sampled pots), and GLM estimated observer CPUE Index for the EAG and WAG golden king crab stocks, 1985/86-2016/17. Observer retained CPUE includes retained and non-retained legal size crabs.

| Year | Pot Fishery Nominal Retained CPUE |  | Obs. Nominal Retained CPUE |  | Obs. Nominal Total CPUE |  | Pot Fishery Effort (no.pot lifts) |  | Obs. Sample <br> Size (no.pot lifts) |  | Obs. CPUE Index |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG |
| 1985/86 | 11.90 | 11.90 |  |  |  |  | 117,718 | 118,563 |  |  |  |  |
| 1986/87 | 8.42 | 7.32 |  |  |  |  | 155,240 | 277,780 |  |  |  |  |
| 1987/88 | 7.03 | 7.15 |  |  |  |  | 146,501 | 160,229 |  |  |  |  |
| 1988/89 | 7.52 | 7.93 |  |  |  |  | 155,518 | 166,409 |  |  |  |  |
| 1989/90 | 8.49 | 7.83 |  |  |  |  | 155,262 | 202,541 |  |  |  |  |
| 1990/91 | 8.90 | 7.00 | 6.84 | 8.00 | 33.60 | 27.04 | 106,281 | 108,533 | 138 | 340 |  |  |
| 1991/92 | 8.20 | 7.40 | 8.11 | 6.83 | 24.69 | 17.01 | 133,428 | 101,429 | 377 | 857 |  |  |
| 1992/93 | 8.40 | 5.90 | 10.42 | 6.35 | 38.46 | 16.64 | 133,778 | 69,443 | 199 | 690 |  |  |
| 1993/94 | 7.80 | 4.40 | 5.07 | 6.51 | 20.81 | 17.14 | 106,890 | 127,764 | 31 | 174 |  |  |
| 1994/95 | 5.90 | 4.10 | 2.54 | 6.71 | 12.91 | 19.25 | 191,455 | 195,138 | 127 | 1,270 |  |  |
| 1995/96 | 5.90 | 4.70 | 5.03 | 4.96 | 16.94 | 14.26 | 177,773 | 115,248 | 6,388 | 5,598 | 0.75 | 1.14 |
| 1996/97 | 6.50 | 6.10 | 5.11 | 5.43 | 13.65 | 13.56 | 113,460 | 99,267 | 8,360 | 7,194 | 0.77 | 0.99 |
| 1997/98 | 7.30 | 6.60 | 7.11 | 6.53 | 18.15 | 15.03 | 106,403 | 86,811 | 4,670 | 3,985 | 0.79 | 1.01 |
| 1998/99 | 8.90 | 11.40 | 9.10 | 9.41 | 25.76 | 23.05 | 83,378 | 35,975 | 3,616 | 1,876 | 0.96 | 1.05 |
| 1999/00 | 9.00 | 6.30 | 9.21 | 5.92 | 20.70 | 14.47 | 79,129 | 107,040 | 3,851 | 4,523 | 0.91 | 0.91 |
| 2000/01 | 9.90 | 7.00 | 9.90 | 6.39 | 25.35 | 16.63 | 71,551 | 101,239 | 5,043 | 4,740 | 0.91 | 0.89 |
| 2001/02 | 11.70 | 6.50 | 11.19 | 5.99 | 22.59 | 14.64 | 62,639 | 105,512 | 4,626 | 4,454 | 1.15 | 0.86 |
| 2002/03 | 12.40 | 8.40 | 11.94 | 7.47 | 22.54 | 17.37 | 52,042 | 78,979 | 3,980 | 2,509 | 1.21 | 0.93 |
| 2003/04 | 10.90 | 10.20 | 11.03 | 9.28 | 19.46 | 18.15 | 58,883 | 66,236 | 3,960 | 3,334 | 1.11 | 1.10 |
| 2004/05 | 18.30 | 12.10 | 17.71 | 11.13 | 28.47 | 22.43 | 34,848 | 56,846 | 2,206 | 2,619 | 1.78 | 1.19 |
| 2005/06 | 25.40 | 21.20 | 29.44 | 23.89 | 38.47 | 36.23 | 24,569 | 30,116 | 1,193 | 1,365 | 1.01 | 1.19 |
| 2006/07 | 24.80 | 19.60 | 25.21 | 24.01 | 33.52 | 33.47 | 26,195 | 26,870 | 1,098 | 1,183 | 0.82 | 1.16 |
| 2007/08 | 28.00 | 20.00 | 31.09 | 21.07 | 40.37 | 32.48 | 22,653 | 29,950 | 998 | 1,082 | 0.95 | 1.06 |
| 2008/09 | 27.30 | 22.40 | 29.92 | 24.54 | 38.36 | 38.12 | 24,466 | 26,200 | 613 | 979 | 0.92 | 1.15 |
| 2009/10 | 25.90 | 23.70 | 26.64 | 26.54 | 35.89 | 34.07 | 26,298 | 26,489 | 408 | 892 | 0.77 | 1.22 |
| 2010/11 | 26.00 | 20.90 | 26.05 | 22.35 | 36.76 | 29.05 | 25,851 | 29,994 | 436 | 867 | 0.77 | 1.06 |
| 2011/12 | 37.30 | 23.40 | 38.79 | 23.76 | 51.69 | 31.09 | 17,915 | 26,326 | 361 | 837 | 1.13 | 1.10 |
| 2012/13 | 33.02 | 20.57 | 38.00 | 22.81 | 47.74 | 30.73 | 20,827 | 32,716 | 438 | 1,109 | 1.08 | 1.06 |
| 2013/14 | 33.67 | 16.42 | 35.83 | 16.93 | 46.16 | 24.95 | 21,388 | 41,835 | 499 | 1,223 | 1.04 | 0.83 |
| 2014/15 | 42.29 | 15.29 | 46.96 | 15.28 | 60.00 | 22.67 | 17,002 | 41,548 | 376 | 1,137 | 1.34 | 0.71 |
| 2015/16 | 39.41 | 14.97 | 43.17 | 15.75 | 58.81 | 22.13 | 19,376 | 41,108 | 478 | 1,296 | 1.28 | 0.77 |
| 2016/17 | 32.45 | 14.29 | 37.01 | 16.63 | 52.78 | 24.25 | 24,470 | 38,118 | 617 | 1,060 | 1.09 | 0.87 |

Table 4. Time series of GLM estimated CPUE indices and coefficient of variations (CV) for the fish ticket based retained catch-per-pot lift for the EAG golden king crab stock. The GLM was fitted to the 1985/86 to 1998/99 time series of data. GLM predictor variables selected by R square criteria.

| Year | CPUE <br> Index | CV |
| :--- | :---: | :---: |
| $1985 / 86$ | 1.63 | 0.05 |
| $1986 / 87$ | 1.20 | 0.05 |
| $1987 / 88$ | 0.93 | 0.06 |
| $1988 / 89$ | 1.02 | 0.05 |
| $1989 / 90$ | 1.05 | 0.04 |
| $1990 / 91$ | 0.85 | 0.06 |
| $1991 / 92$ | 0.87 | 0.06 |
| $1992 / 93$ | 0.94 | 0.06 |
| $1993 / 94$ | 0.89 | 0.06 |
| $1994 / 95$ | 0.80 | 0.06 |
| $1995 / 96$ | 0.77 | 0.07 |
| $1996 / 97$ | 0.82 | 0.07 |
| $1997 / 98$ | 1.19 | 0.05 |
| $1998 / 99$ | 1.39 | 0.05 |

Table 5. The initial input number of vessel-days/trips and Stage-2 effective sample sizes iteratively estimated by Francis method for retained, total, and groundfish discard catch size compositions of golden king crab for scenario 17_0 model fit to EAG data. NA: not available.

| Year | Initial <br> Input Retained VesselDays Sample Size (no) | Stage-2 <br> Retained <br> Effective <br> Sample <br> Size (no) | Initial <br> Input <br> Total <br> VesselDays Sample Size (no) | Stage-2 <br> Total Effective Sample Size (no) | Initial Input Groundfish Trip Sample Size (no) | Stage-2 Groundfish Effective Sample Size (no) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1985/86 | 57 | 48 |  |  |  |  |
| 1986/87 | 11 | 9 |  |  |  |  |
| 1987/88 | 61 | 51 |  |  |  |  |
| 1988/89 | 352 | 293 |  |  |  |  |
| 1989/90 | 792 | 660 |  |  | 9 | 4 |
| 1990/91 | 163 | 136 | 22 | 11 | 13 | 6 |
| 1991/92 | 140 | 117 | 48 | 24 | NA | NA |
| 1992/93 | 49 | 41 | 41 | 21 | 2 | 1 |
| 1993/94 | 340 | 283 | NA | NA | 2 | 1 |
| 1994/95 | 319 | 266 | 34 | 17 | 4 | 2 |
| 1995/96 | 879 | 733 | 1,117 | 568 | 5 | 2 |
| 1996/97 | 547 | 456 | 509 | 259 | 4 | 2 |
| 1997/98 | 538 | 449 | 711 | 362 | 8 | 4 |
| 1998/99 | 541 | 451 | 574 | 292 | 15 | 7 |
| 1999/00 | 463 | 386 | 607 | 309 | 14 | 6 |
| 2000/01 | 436 | 363 | 495 | 252 | 16 | 7 |
| 2001/02 | 488 | 407 | 510 | 259 | 13 | 6 |
| 2002/03 | 406 | 338 | 438 | 223 | 15 | 7 |
| 2003/04 | 405 | 338 | 416 | 212 | 17 | 8 |
| 2004/05 | 280 | 233 | 299 | 152 | 10 | 4 |
| 2005/06 | 266 | 222 | 232 | 118 | 12 | 5 |
| 2006/07 | 234 | 195 | 143 | 73 | 14 | 6 |
| 2007/08 | 199 | 166 | 134 | 68 | 17 | 8 |
| 2008/09 | 197 | 164 | 113 | 57 | 15 | 7 |
| 2009/10 | 170 | 142 | 95 | 48 | 16 | 7 |
| 2010/11 | 183 | 153 | 108 | 55 | 26 | 12 |
| 2011/12 | 160 | 133 | 107 | 54 | 13 | 6 |
| 2012/13 | 187 | 156 | 99 | 50 | 18 | 8 |
| 2013/14 | 193 | 161 | 122 | 62 | 17 | 8 |
| 2014/15 | 168 | 140 | 99 | 50 | 16 | 7 |
| 2015/16 | 190 | 158 | 125 | 64 | 10 | 4 |
| 2016/17 | 223 | 186 | 155 | 79 | 12 | 5 |

Table 6. The initial input number of vessel-days/trips and Stage-2 effective sample sizes iteratively estimated by Francis method for retained, total, and groundfish discard catch size compositions of golden king crab for scenario 17_0a model fit to EAG data. NA: not available.

| Year | Initial <br> Input <br> Retained <br> Vessel- <br> Days | Stage-2 <br> Retained <br> Effective <br> Sample <br> Size (no) | Initial <br> Input <br> Total <br> Vessel- <br> Days <br> Sample | Stage-2 <br> Total <br> Effective <br> Sample <br> Size (no) | Initial <br> Input <br> Groundfish <br> Trip <br> Sample <br> Size (no) | Stage-2 <br> Groundfish <br> Effective |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Sample <br> Size (no) |  |  |  |  |  |  |
|  | Size (no) |  |  |  |  |  |

Table 7. The initial input number of vessel-days/trips and Stage-2 effective sample sizes iteratively estimated by Francis method for retained, total, and groundfish discard catch size compositions of golden king crab for scenario 17_0b model fit to EAG data. NA: not available.

| Year | Initial <br> Input <br> Retained <br> Vessel- <br> Days | Stage-2 <br> Retained <br> Effective <br> Sample <br> Size (no) | Initial <br> Input <br> Total <br> Vessel- <br> Days <br> Sample | Stage-2 <br> Total <br> Effective <br> Sample <br> Size (no) | Initial <br> Input <br> Groundfish <br> Trip <br> Sample <br> Size (no) | Stage-2 <br> Groundfish <br> Effective |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Sample <br> Size (no) |  |  |  |  |  |  |
|  | Size (no) |  |  |  |  |  |

Table 8. The initial input number of vessel-days/trips and Stage-2 effective sample sizes iteratively estimated by Francis method for retained, total, and groundfish discard catch size compositions of golden king crab for scenario 17_0c model fit to EAG data. NA: not available.

| Year | Initial <br> Input Retained VesselDays Sample Size (no) | Stage-2 <br> Retained <br> Effective <br> Sample <br> Size (no) | Initial <br> Input <br> Total <br> Vessel- <br> Days <br> Sample <br> Size <br> (no) | Stage-2 <br> Total Effective <br> Sample <br> Size (no) | Initial Input Groundfish Trip Sample Size (no) | Stage-2 <br> Groundfish Effective Sample Size (no) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1985/86 | 57 | 47 |  |  |  |  |
| 1986/87 | 11 | 9 |  |  |  |  |
| 1987/88 | 61 | 50 |  |  |  |  |
| 1988/89 | 352 | 288 |  |  |  |  |
| 1989/90 | 792 | 648 |  |  | 9 | 4 |
| 1990/91 | 163 | 133 | 22 | 12 | 13 | 6 |
| 1991/92 | 140 | 115 | 48 | 26 | NA | NA |
| 1992/93 | 49 | 40 | 41 | 22 | 2 | 1 |
| 1993/94 | 340 | 278 | NA | NA | 2 | 1 |
| 1994/95 | 319 | 261 | 34 | 18 | 4 | 2 |
| 1995/96 | 879 | 719 | 1,117 | 602 | 5 | 2 |
| 1996/97 | 547 | 447 | 509 | 274 | 4 | 2 |
| 1997/98 | 538 | 440 | 711 | 383 | 8 | 4 |
| 1998/99 | 541 | 443 | 574 | 309 | 15 | 7 |
| 1999/00 | 463 | 379 | 607 | 327 | 14 | 6 |
| 2000/01 | 436 | 357 | 495 | 267 | 16 | 7 |
| 2001/02 | 488 | 399 | 510 | 275 | 13 | 6 |
| 2002/03 | 406 | 332 | 438 | 236 | 15 | 7 |
| 2003/04 | 405 | 331 | 416 | 224 | 17 | 8 |
| 2004/05 | 280 | 229 | 299 | 161 | 10 | 4 |
| 2005/06 | 266 | 218 | 232 | 125 | 12 | 5 |
| 2006/07 | 234 | 191 | 143 | 77 | 14 | 6 |
| 2007/08 | 199 | 163 | 134 | 72 | 17 | 8 |
| 2008/09 | 197 | 161 | 113 | 61 | 15 | 7 |
| 2009/10 | 170 | 139 | 95 | 51 | 16 | 7 |
| 2010/11 | 183 | 150 | 108 | 58 | 26 | 12 |
| 2011/12 | 160 | 131 | 107 | 58 | 13 | 6 |
| 2012/13 | 187 | 153 | 99 | 53 | 18 | 8 |
| 2013/14 | 193 | 158 | 122 | 66 | 17 | 8 |
| 2014/15 | 168 | 137 | 99 | 53 | 16 | 7 |
| 2015/16 | 190 | 155 | 125 | 67 | 10 | 4 |
| 2016/17 | 223 | 182 | 155 | 84 | 12 | 5 |

Table 9. The initial input number of vessel-days/trips and Stage-2 effective sample sizes iteratively estimated by Francis method for retained, total, and groundfish discard catch size compositions of golden king crab for scenario 17_0d model fit to EAG data. NA: not available.

| Year | Initial <br> Input Retained VesselDays Sample Size (no) | Stage-2 <br> Retained <br> Effective <br> Sample <br> Size (no) | Initial <br> Input <br> Total VesselDays Sample Size (no) | Stage-2 <br> Total Effective Sample Size (no) | Initial Input Groundfish Trip Sample Size (no) | Stage-2 <br> Groundfish Effective Sample Size (no) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1985/86 | 57 | 48 |  |  |  |  |
| 1986/87 | 11 | 9 |  |  |  |  |
| 1987/88 | 61 | 52 |  |  |  |  |
| 1988/89 | 352 | 298 |  |  |  |  |
| 1989/90 | 792 | 669 |  |  | 9 | 4 |
| 1990/91 | 163 | 138 | 22 | 12 | 13 | 6 |
| 1991/92 | 140 | 118 | 48 | 25 | NA | NA |
| 1992/93 | 49 | 41 | 41 | 22 | 2 | 1 |
| 1993/94 | 340 | 287 | NA | NA | 2 | 1 |
| 1994/95 | 319 | 270 | 34 | 18 | 4 | 2 |
| 1995/96 | 879 | 743 | 1,117 | 593 | 5 | 2 |
| 1996/97 | 547 | 462 | 509 | 270 | 4 | 2 |
| 1997/98 | 538 | 455 | 711 | 378 | 8 | 4 |
| 1998/99 | 541 | 457 | 574 | 305 | 15 | 7 |
| 1999/00 | 463 | 391 | 607 | 322 | 14 | 6 |
| 2000/01 | 436 | 369 | 495 | 263 | 16 | 7 |
| 2001/02 | 488 | 412 | 510 | 271 | 13 | 6 |
| 2002/03 | 406 | 343 | 438 | 233 | 15 | 7 |
| 2003/04 | 405 | 342 | 416 | 221 | 17 | 8 |
| 2004/05 | 280 | 237 | 299 | 159 | 10 | 4 |
| 2005/06 | 266 | 225 | 232 | 123 | 12 | 5 |
| 2006/07 | 234 | 198 | 143 | 76 | 14 | 6 |
| 2007/08 | 199 | 168 | 134 | 71 | 17 | 8 |
| 2008/09 | 197 | 167 | 113 | 60 | 15 | 7 |
| 2009/10 | 170 | 144 | 95 | 50 | 16 | 7 |
| 2010/11 | 183 | 155 | 108 | 57 | 26 | 12 |
| 2011/12 | 160 | 135 | 107 | 57 | 13 | 6 |
| 2012/13 | 187 | 158 | 99 | 53 | 18 | 8 |
| 2013/14 | 193 | 163 | 122 | 65 | 17 | 8 |
| 2014/15 | 168 | 142 | 99 | 53 | 16 | 7 |
| 2015/16 | 190 | 161 | 125 | 66 | 10 | 4 |
| 2016/17 | 223 | 188 | 155 | 82 | 12 | 5 |

Table 10. The initial input number of vessel-days/trips and Stage-2 effective sample sizes iteratively estimated by McAllister and Ianelli method for retained, total, and groundfish discard catch size compositions of golden king crab for scenario 17_0e model fit to EAG data. NA: not available.

| Year | Initial <br> Input Retained VesselDays Sample Size (no) | Stage-2 <br> Retained <br> Effective <br> Sample <br> Size (no) | Initial <br> Input <br> Total <br> Vessel- <br> Days <br> Sample <br> Size <br> (no) | Stage-2 <br> Total <br> Effective <br> Sample <br> Size (no) | Initial Input Groundfish Trip Sample Size (no) | Stage-2 Groundfish Effective Sample Size (no) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1985/86 | 57 | 72 |  |  |  |  |
| 1986/87 | 11 | 14 |  |  |  |  |
| 1987/88 | 61 | 77 |  |  |  |  |
| 1988/89 | 352 | 443 |  |  |  |  |
| 1989/90 | 792 | 997 |  |  | 9 | 7 |
| 1990/91 | 163 | 205 | 22 | 8 | 13 | 10 |
| 1991/92 | 140 | 176 | 48 | 18 | NA | NA |
| 1992/93 | 49 | 62 | 41 | 16 | 2 | 1 |
| 1993/94 | 340 | 428 | NA | NA | 2 | 1 |
| 1994/95 | 319 | 402 | 34 | 13 | 4 | 3 |
| 1995/96 | 879 | 1,106 | 1,117 | 424 | 5 | 4 |
| 1996/97 | 547 | 689 | 509 | 193 | 4 | 3 |
| 1997/98 | 538 | 677 | 711 | 270 | 8 | 6 |
| 1998/99 | 541 | 681 | 574 | 218 | 15 | 11 |
| 1999/00 | 463 | 583 | 607 | 230 | 14 | 10 |
| 2000/01 | 436 | 549 | 495 | 188 | 16 | 12 |
| 2001/02 | 488 | 614 | 510 | 194 | 13 | 10 |
| 2002/03 | 406 | 511 | 438 | 166 | 15 | 11 |
| 2003/04 | 405 | 510 | 416 | 158 | 17 | 13 |
| 2004/05 | 280 | 352 | 299 | 113 | 10 | 7 |
| 2005/06 | 266 | 335 | 232 | 88 | 12 | 9 |
| 2006/07 | 234 | 295 | 143 | 54 | 14 | 10 |
| 2007/08 | 199 | 250 | 134 | 51 | 17 | 13 |
| 2008/09 | 197 | 248 | 113 | 43 | 15 | 11 |
| 2009/10 | 170 | 214 | 95 | 36 | 16 | 12 |
| 2010/11 | 183 | 230 | 108 | 41 | 26 | 19 |
| 2011/12 | 160 | 201 | 107 | 41 | 13 | 10 |
| 2012/13 | 187 | 235 | 99 | 38 | 18 | 13 |
| 2013/14 | 193 | 243 | 122 | 46 | 17 | 13 |
| 2014/15 | 168 | 211 | 99 | 38 | 16 | 12 |
| 2015/16 | 190 | 239 | 125 | 47 | 10 | 7 |
| 2016/17 | 223 | 281 | 155 | 59 | 12 | 9 |

Table 11. The initial input number of vessel-days/trips and Stage-2 effective sample sizes iteratively estimated by Francis method for retained, total, and groundfish discard catch size compositions of golden king crab for scenario 17_0f model fit to EAG data. NA: not available.

| Year | Initial <br> Input Retained VesselDays Sample Size (no) | Stage-2 <br> Retained <br> Effective <br> Sample <br> Size (no) | Initial <br> Input <br> Total <br> Vessel- <br> Days <br> Sample <br> Size <br> (no) | Stage-2 <br> Total Effective <br> Sample <br> Size (no) | Initial Input Groundfish Trip Sample Size (no) | Stage-2 <br> Groundfish Effective Sample Size (no) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1985/86 | 57 | 48 |  |  |  |  |
| 1986/87 | 11 | 9 |  |  |  |  |
| 1987/88 | 61 | 51 |  |  |  |  |
| 1988/89 | 352 | 294 |  |  |  |  |
| 1989/90 | 792 | 661 |  |  | 9 | 4 |
| 1990/91 | 163 | 136 | 22 | 11 | 13 | 6 |
| 1991/92 | 140 | 117 | 48 | 24 | NA | NA |
| 1992/93 | 49 | 41 | 41 | 21 | 2 | 1 |
| 1993/94 | 340 | 284 | NA | NA | 2 | 1 |
| 1994/95 | 319 | 266 | 34 | 17 | 4 | 2 |
| 1995/96 | 879 | 734 | 1,117 | 569 | 5 | 2 |
| 1996/97 | 547 | 457 | 509 | 259 | 4 | 2 |
| 1997/98 | 538 | 449 | 711 | 362 | 8 | 4 |
| 1998/99 | 541 | 452 | 574 | 292 | 15 | 7 |
| 1999/00 | 463 | 386 | 607 | 309 | 14 | 6 |
| 2000/01 | 436 | 364 | 495 | 252 | 16 | 7 |
| 2001/02 | 488 | 407 | 510 | 260 | 13 | 6 |
| 2002/03 | 406 | 339 | 438 | 223 | 15 | 7 |
| 2003/04 | 405 | 338 | 416 | 212 | 17 | 8 |
| 2004/05 | 280 | 234 | 299 | 152 | 10 | 4 |
| 2005/06 | 266 | 222 | 232 | 118 | 12 | 5 |
| 2006/07 | 234 | 195 | 143 | 73 | 14 | 6 |
| 2007/08 | 199 | 166 | 134 | 68 | 17 | 8 |
| 2008/09 | 197 | 164 | 113 | 58 | 15 | 7 |
| 2009/10 | 170 | 142 | 95 | 48 | 16 | 7 |
| 2010/11 | 183 | 153 | 108 | 55 | 26 | 12 |
| 2011/12 | 160 | 134 | 107 | 55 | 13 | 6 |
| 2012/13 | 187 | 156 | 99 | 50 | 18 | 8 |
| 2013/14 | 193 | 161 | 122 | 62 | 17 | 8 |
| 2014/15 | 168 | 140 | 99 | 50 | 16 | 7 |
| 2015/16 | 190 | 159 | 125 | 64 | 10 | 4 |
| 2016/17 | 223 | 186 | 155 | 79 | 12 | 5 |

Table 12. Parameter estimates and coefficient of variations (CV) with the 2016 MMB (MMB on 15 Feb 2017) for scenarios 17_0, 17_0a, 17_0b, and 17_0c for the golden king crab data from the EAG, 1985/86-2016/17. Recruitment and fishing mortality deviations and initial size frequency determination parameters were omitted from this list.

|  | Scenario 17_0 |  | Scenario 17_0a |  | Scenario 17_0b |  | Scenario 17_0c |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Parameter | Estimate | CV | Estimate | CV | Estimate | CV | Estimate | CV | Limits |
| $\log _{\sim} \omega_{1}$ ( growth incr. intercept) | 2.54 | 0.006 | 2.54 | 0.006 | 2.54 | 0.006 | 2.54 | 0.006 | 1.0, 4.5 |
| $\omega_{2}$ ( growth incr. slope) | -8.20 | 0.21 | -8.22 | 0.21 | -8.22 | 0.21 | -8.26 | 0.21 | -12.0,-5.0 |
| $\log _{\text {_ }} \mathrm{a}$ (molt prob. slope) | -2.50 | 0.02 | -2.48 | 0.02 | -2.50 | 0.02 | -2.48 | 0.02 | -4.61,-1.39 |
| $\log _{-} \mathrm{b}$ (molt prob. L50) | 4.95 | 0.001 | 4.95 | 0.00 | 4.95 | 0.001 | 4.95 | 0.001 | 3.869,5.05 |
| $\sigma$ (growth variability std) | 3.68 | 0.03 | 3.68 | 0.03 | 3.68 | 0.03 | 3.68 | 0.03 | 0.1,12.0 |
| log_total sel delta $\theta$, 1985-04 | 3.38 | 0.020 | 3.38 | 0.02 | 3.37 | 0.020 | 3.38 | 0.019 | 0.,4.4 |
| $\log _{-}$total sel delta $\theta$, 2005-16 | 2.97 | 0.030 | 2.93 | 0.030 | 2.98 | 0.030 | 2.92 | 0.031 | 0.,4.4 |
| $\log _{-}$ret. sel delta0, 1985-16 | 1.85 | 0.023 | 1.85 | 0.023 | 1.85 | 0.0234 | 1.85 | 0.0233 | 0.,4.4 |
| log_tot sel $\theta_{50}, 1985-04$ | 4.83 | 0.003 | 4.84 | 0.003 | 4.83 | 0.003 | 4.84 | 0.003 | 4.0,5.0 |
| log_tot sel $\theta_{50}, 2005-16$ | 4.92 | 0.002 | 4.91 | 0.002 | 4.92 | 0.0021 | 4.91 | 0.0019 | 4.0,5.0 |
| log_ret. sel $\theta_{50}, 1985-16$ | 4.91 | 0.0003 | 4.91 | 0.00 | 4.91 | 0.0003 | 4.91 | 0.0003 | 4.0,5.0 |
| $\log _{\_} \beta_{\mathrm{r}}$ (rec.distribution par.) | -1.09 | 0.18 | -1.08 | 0.18 | -1.09 | 0.18 | -1.06 | 0.18 | -12.0, 12.0 |
| $\operatorname{logq} 2$ (catchability 1995-04) | -0.59 | 0.12 | -0.61 | 0.13 | -0.57 | 0.13 | -0.69 | 0.15 | -9.0, 2.25 |
| $\operatorname{logq} 3$ (catchability 2005-16) | -0.97 | 0.13 | -1.06 | 0.13 | -0.89 | 0.15 | -1.09 | 0.13 | -9.0, 2.25 |
| log_mean_rec (mean rec.) | 0.874 | 0.05 | 0.890 | 0.05 | 0.855 | 0.05 | 0.893 | 0.06 | 0.01, 5.0 |
| log_mean_Fpot (Pot fishery F) | -1.060 | 0.06 | -1.108 | 0.06 | -1.032 | 0.07 | -1.119 | 0.07 | -15.0, -0.01 |
| log_mean_Fground (GF byc. F) | -9.244 | 0.09 | -9.278 | 0.09 | -9.210 | 0.09 | -9.289 | 0.09 | -15.0, -1.6 |
| $\sigma_{e}^{2}$ (observer CPUE additional var) | 0.018 | 0.37 | 0.029 | 0.43 | 0.032 | 0.39 | 0.031 | 0.47 | 0.0, 0.15 |
| $\sigma_{e}^{2}$ (fishery CPUE additional var) | 0.051 | 0.43 | 0.051 | 0.44 | 0.040 | 0.432 | 0.173 | 0.58 | 0.0,1.0 |
| 2016 MMB | 13,455 | 0.17 | 13,579 | 0.20 | 11,842 | 0.19 | 13,767 | 0.21 |  |

Table 13. Parameter estimates and coefficient of variations (CV) with the 2016 MMB (MMB on 15 Feb 2017) for scenarios 17_0d, 17_0e, and 17_0f for the golden king crab data from the EAG, 1985/86-2016/17. Recruitment and fishing mortality deviations and initial size frequency determination parameters were omitted from this list.

|  | Scenario 17_0d |  | Scenario 17_0e |  | Scenario 17_0f |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Parameter | Estimate | CV | Estimate | CV | Estimate | CV | Limits |
| $\log _{-} \omega_{1}$ ( growth incr. intercept) | 2.54 | 0.006 | 2.54 | 0.006 | 2.54 | 0.006 | 1.0, 4.5 |
| $\omega_{2}$ (growth incr. slope) | -8.24 | 0.21 | -7.94 | 0.21 | -8.20 | 0.21 | -12.0,-5.0 |
| $\log _{\text {_ }}$ ( (molt prob. slope) | -2.50 | 0.02 | -2.51 | 0.02 | -2.50 | 0.02 | -4.61,-1.39 |
| $\log _{-} \mathrm{b}$ (molt prob. L50) | 4.95 | 0.001 | 4.95 | 0.001 | 4.95 | 0.00 | 3.869,5.05 |
| $\sigma$ (growth variability std) | 3.68 | 0.03 | 3.68 | 0.03 | 3.68 | 0.03 | 0.1,12.0 |
| log_total sel delta 0 , 1985-04 | 3.38 | 0.02 | 3.34 | 0.02 | 3.38 | 0.02 | 0.,4.4 |
| $\log _{-}$total sel delta $\theta$, 2005-12 | 2.93 | 0.04 |  |  | 2.97 | 0.03 | 0.,4.4 |
| $\log _{-}$total sel delta, , 2013-16 or 2005-16 | 3.02 | 0.05 | 2.96 | 0.03 | 1.85 | 0.02 | 0.,4.4 |
| $\log _{-}$ret. sel delta $\theta, 1985-16$ | 1.85 | 0.02 | 1.85 | 0.02 | 4.83 | 0.003 | 0.,4.4 |
| log_tot sel $\theta_{50}, 1985-04$ | 4.83 | 0.002 | 4.83 | 0.002 | 4.92 | 0.002 | 4.0,5.0 |
| log_tot sel $\theta_{50}, 2005-12$ | 4.92 | 0.002 |  |  | 4.91 | 0.0003 | 4.0,5.0 |
| $\log _{\text {_tot sel }} \theta_{50}, 2013-16$ or 2005-16 | 4.92 | 0.004 | 4.92 | 0.002 | -1.09 | 0.18 | 4.0,5.0 |
| log_ret. sel $\theta_{50}, 1985-16$ | 4.91 | 0.0003 | 4.91 | 0.0003 | -0.59 | 0.12 | 4.0,5.0 |
| $\log _{\_} \beta_{\mathrm{r}}$ (rec.distribution par.) | -1.08 | 0.18 | -1.15 | 0.17 | -0.97 | 0.13 | -12.0, 12.0 |
| Logq1 (catchability 1985-04) | -0.60 | 0.12 | -0.60 | 0.12 | 0.873 | 0.05 | -9.0, 2.25 |
| Logq3 (catchability 2005-12) | -0.99 | 0.11 |  |  | -1.060 | 0.06 | -9.0, 2.25 |
| Logq2 (catchability 2013-16 or 2005-16) | -0.57 | 0.36 | -1.01 | 0.12 | -9.242 | 0.09 | -9.0, 2.25 |
| log_mean_rec (mean rec.) | 0.83 | 0.06 | 0.872 | 0.05 | 0.018 | 0.38 | 0.01, 5.0 |
| log_mean_Fpot (Pot fishery F) | -1.02 | 0.07 | -1.080 | 0.06 | 0.051 | 0.43 | -15.0, -0.01 |
| log_mean_Fground (GF byc. F) | -9.18 | 0.09 | -9.259 | 0.09 | 2.54 | 0.006 | -15.0, -1.6 |
| $\sigma_{e}^{2}$ (observer CPUE additional var) | 0.02 | 0.36 | 0.018 | 0.37 | -8.20 | 0.21 | 0.0, 0.15 |
| $\sigma_{e}^{2}$ (fishery CPUE additional var) | 0.05 | 0.43 | 0.052 | 0.42 | -2.50 | 0.02 | 0.0,1.0 |
| $\sigma_{e}^{2}$ (survey CPUE additional var) |  |  |  |  | 0.0000003 | 1001.0 | 0.0,1.0 |
| 2016 MMB | 8,833 | 0.23 | 13,440 | 0.17 | 13,368 | 0.17 |  |

Table 14. Annual abundance estimates of model recruits (millions of crabs), legal male biomass ( $t$ ) with coefficient of variations (CV), and mature male biomass ( $t$ ) with CV for scenario 17_0 for golden king crab in the EAG. Legal male biomass was estimated on July 1 (start of fishing year) of fishing year y. Mature male biomass for fishing year y was estimated on February 15 of year $y+1$, after the year y fishery total catch removal. Recruits estimates for 1961 to 2017 are restricted to 1985-2017. Equilibrium $\mathrm{MMB}_{\mathrm{eq}}$ and $\mathrm{MMB}_{35 \%}$ are also listed.

| Year | Recruits to the Model ( $\geq$ 101 mm CL) | $\begin{gathered} \text { Mature Male } \\ \text { Biomass } \\ (\geq 111 \mathrm{~mm} \mathrm{CL}) \end{gathered}$ | CV | Legal Size Male <br> Biomass ( $\geq \mathbf{1 3 6}$ <br> mm CL) | CV |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{aligned} & \mathrm{MMB}_{\mathrm{eq}}=23,950 \\ & M M B_{35 \%}=6,954 \end{aligned}$ |  |  |  |
| 1985 | 1.67 |  |  | 9,618 | 0.05 |
| 1986 | 1.00 | 9,534 | 0.04 | 8,147 | 0.04 |
| 1987 | 4.12 | 7,286 | 0.04 | 6,353 | 0.04 |
| 1988 | 3.77 | 6,652 | 0.05 | 5,274 | 0.05 |
| 1989 | 2.20 | 6,706 | 0.05 | 4,698 | 0.07 |
| 1990 | 2.71 | 5,973 | 0.06 | 4,287 | 0.07 |
| 1991 | 3.52 | 6,078 | 0.05 | 4,647 | 0.06 |
| 1992 | 2.27 | 6,116 | 0.04 | 4,466 | 0.05 |
| 1993 | 2.13 | 6,058 | 0.04 | 4,471 | 0.05 |
| 1994 | 2.45 | 6,195 | 0.03 | 4,889 | 0.04 |
| 1995 | 2.29 | 5,716 | 0.04 | 4,442 | 0.04 |
| 1996 | 2.25 | 5,139 | 0.04 | 3,850 | 0.04 |
| 1997 | 3.03 | 5,253 | 0.04 | 3,987 | 0.05 |
| 1998 | 2.78 | 5,529 | 0.05 | 4,100 | 0.05 |
| 1999 | 2.96 | 6,118 | 0.05 | 4,542 | 0.05 |
| 2000 | 2.78 | 6,811 | 0.05 | 5,202 | 0.06 |
| 2001 | 2.11 | 7,463 | 0.06 | 5,847 | 0.06 |
| 2002 | 2.70 | 7,848 | 0.06 | 6,414 | 0.06 |
| 2003 | 2.26 | 8,179 | 0.06 | 6,787 | 0.07 |
| 2004 | 1.95 | 8,507 | 0.07 | 7,089 | 0.07 |
| 2005 | 2.95 | 8,577 | 0.07 | 7,304 | 0.07 |
| 2006 | 2.25 | 8,649 | 0.07 | 7,233 | 0.08 |
| 2007 | 2.17 | 8,903 | 0.08 | 7,390 | 0.08 |
| 2008 | 3.52 | 8,910 | 0.08 | 7,543 | 0.08 |
| 2009 | 2.39 | 9,127 | 0.08 | 7,509 | 0.09 |
| 2010 | 2.19 | 9,630 | 0.08 | 7,914 | 0.09 |
| 2011 | 2.82 | 9,685 | 0.08 | 8,235 | 0.08 |
| 2012 | 2.74 | 9,708 | 0.08 | 8,229 | 0.09 |
| 2013 | 2.36 | 9,885 | 0.09 | 8,273 | 0.09 |
| 2014 | 5.63 | 9,913 | 0.10 | 8,368 | 0.10 |
| 2015 | 4.76 | 10,626 | 0.11 | 8,432 | 0.11 |
| 2016 | 2.59 | 12,484 | 0.14 | 9,623 | 0.13 |
| 2017 | 4.70 | 13,455 | 0.17 |  |  |

Table 15. Annual abundance estimates of model recruits (millions of crabs), legal male biomass ( $t$ ) with coefficient of variations (CV), and mature male biomass ( t ) with CV for scenario 17_0a for golden king crab in the EAG. Legal male biomass was estimated on July 1 (start of fishing year) of fishing year y. Mature male biomass for fishing year y was estimated on February 15 of year $y+1$, after the year y fishery total catch removal. Recruits estimates for 1961 to 2017 are restricted to 1985-2017. Equilibrium $\mathrm{MMB}_{\mathrm{eq}}$ and $\mathrm{MMB}_{35 \%}$ are also listed.

| Year | Recruits to the Model ( $\geq \mathbf{1 0 1}$ mm CL) | $\begin{gathered} \text { Mature Male } \\ \text { Biomass } \\ (\geq 111 \mathrm{~mm} \mathrm{CL}) \end{gathered}$ | CV | Legal Size Male <br> Biomass ( $\geq$ 136 mm CL ) | CV |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{aligned} & \mathrm{MMB}_{\mathrm{eq}}=24,335 \\ & M M B_{35 \%}=7,063 \end{aligned}$ |  |  |  |
| 1985 | 1.67 |  |  | 9,713 | 0.05 |
| 1986 | 1.01 | 9,594 | 0.04 | 8,206 | 0.04 |
| 1987 | 4.14 | 7,331 | 0.04 | 6,389 | 0.04 |
| 1988 | 3.75 | 6,696 | 0.05 | 5,303 | 0.05 |
| 1989 | 2.20 | 6,751 | 0.05 | 4,732 | 0.07 |
| 1990 | 2.70 | 6,013 | 0.06 | 4,321 | 0.07 |
| 1991 | 3.51 | 6,116 | 0.05 | 4,678 | 0.06 |
| 1992 | 2.23 | 6,146 | 0.04 | 4,492 | 0.05 |
| 1993 | 2.11 | 6,069 | 0.04 | 4,488 | 0.05 |
| 1994 | 2.49 | 6,179 | 0.03 | 4,881 | 0.04 |
| 1995 | 2.36 | 5,695 | 0.04 | 4,414 | 0.04 |
| 1996 | 2.32 | 5,152 | 0.04 | 3,836 | 0.04 |
| 1997 | 3.17 | 5,321 | 0.05 | 4,016 | 0.05 |
| 1998 | 2.94 | 5,673 | 0.05 | 4,185 | 0.05 |
| 1999 | 3.12 | 6,369 | 0.05 | 4,711 | 0.06 |
| 2000 | 2.97 | 7,180 | 0.06 | 5,476 | 0.06 |
| 2001 | 2.26 | 7,954 | 0.06 | 6,231 | 0.06 |
| 2002 | 2.83 | 8,457 | 0.07 | 6,914 | 0.07 |
| 2003 | 2.35 | 8,870 | 0.07 | 7,381 | 0.07 |
| 2004 | 2.02 | 9,248 | 0.08 | 7,747 | 0.08 |
| 2005 | 2.96 | 9,328 | 0.08 | 7,988 | 0.08 |
| 2006 | 2.41 | 9,369 | 0.08 | 7,907 | 0.09 |
| 2007 | 2.32 | 9,594 | 0.08 | 8,018 | 0.09 |
| 2008 | 3.46 | 9,635 | 0.08 | 8,168 | 0.09 |
| 2009 | 2.27 | 9,840 | 0.09 | 8,162 | 0.09 |
| 2010 | 2.23 | 10,218 | 0.09 | 8,513 | 0.09 |
| 2011 | 2.81 | 10,143 | 0.09 | 8,698 | 0.10 |
| 2012 | 2.68 | 10,095 | 0.10 | 8,594 | 0.10 |
| 2013 | 2.42 | 10,186 | 0.11 | 8,567 | 0.11 |
| 2014 | 5.58 | 10,143 | 0.13 | 8,582 | 0.13 |
| 2015 | 4.78 | 10,821 | 0.15 | 8,606 | 0.15 |
| 2016 | 2.63 | 12,627 | 0.18 | 9,754 | 0.18 |
| 2017 | 4.70 | 13,579 | 0.20 |  |  |

Table 16. Annual abundance estimates of model recruits (millions of crabs), legal male biomass (t) with coefficient of variations (CV), and mature male biomass ( t ) with CV for scenario 17_0b for golden king crab in the EAG. Legal male biomass was estimated on July 1 (start of fishing year) of fishing year y. Mature male biomass for fishing year y was estimated on February 15 of year $y+1$, after the year y fishery total catch removal. Recruits estimates for 1961 to 2017 are restricted to 1985-2017. Equilibrium MMBeq and MMB35\% are also listed.

| Year | Recruits to the Model ( $\geq 101$ mm CL) | $\begin{gathered} \text { Mature Male } \\ \text { Biomass } \\ (\geq 111 \mathrm{~mm} \mathrm{CL}) \\ \hline \end{gathered}$ | CV | Legal Size Male Biomass ( $\geq 136$ mm CL) | CV |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{aligned} & \text { MMBeq }=23,449 \\ & \text { MMB35\% }=6,794 \end{aligned}$ |  |  |  |
| 1985 | 1.68 |  |  | 9,583 | 0.06 |
| 1986 | 1.00 | 9,536 | 0.04 | 8,138 | 0.04 |
| 1987 | 4.17 | 7,300 | 0.04 | 6,364 | 0.04 |
| 1988 | 3.73 | 6,683 | 0.05 | 5,295 | 0.05 |
| 1989 | 2.13 | 6,751 | 0.05 | 4,738 | 0.07 |
| 1990 | 2.72 | 5,991 | 0.06 | 4,325 | 0.07 |
| 1991 | 3.54 | 6,066 | 0.05 | 4,652 | 0.06 |
| 1992 | 2.28 | 6,107 | 0.04 | 4,455 | 0.05 |
| 1993 | 2.15 | 6,060 | 0.04 | 4,466 | 0.05 |
| 1994 | 2.43 | 6,209 | 0.03 | 4,896 | 0.04 |
| 1995 | 2.27 | 5,735 | 0.04 | 4,462 | 0.04 |
| 1996 | 2.23 | 5,145 | 0.04 | 3,864 | 0.04 |
| 1997 | 3.01 | 5,242 | 0.04 | 3,987 | 0.05 |
| 1998 | 2.74 | 5,498 | 0.05 | 4,082 | 0.05 |
| 1999 | 2.90 | 6,064 | 0.05 | 4,506 | 0.05 |
| 2000 | 2.71 | 6,720 | 0.06 | 5,140 | 0.06 |
| 2001 | 2.05 | 7,327 | 0.06 | 5,747 | 0.06 |
| 2002 | 2.60 | 7,667 | 0.07 | 6,270 | 0.07 |
| 2003 | 2.24 | 7,946 | 0.07 | 6,601 | 0.07 |
| 2004 | 1.92 | 8,231 | 0.07 | 6,852 | 0.08 |
| 2005 | 2.92 | 8,294 | 0.08 | 7,044 | 0.08 |
| 2006 | 2.27 | 8,364 | 0.08 | 6,971 | 0.08 |
| 2007 | 2.17 | 8,633 | 0.08 | 7,133 | 0.09 |
| 2008 | 3.32 | 8,676 | 0.08 | 7,313 | 0.09 |
| 2009 | 2.20 | 8,873 | 0.08 | 7,306 | 0.09 |
| 2010 | 2.08 | 9,261 | 0.08 | 7,650 | 0.09 |
| 2011 | 2.59 | 9,207 | 0.08 | 7,855 | 0.08 |
| 2012 | 2.45 | 9,126 | 0.08 | 7,753 | 0.09 |
| 2013 | 2.18 | 9,143 | 0.09 | 7,678 | 0.09 |
| 2014 | 4.97 | 9,019 | 0.11 | 7,618 | 0.11 |
| 2015 | 4.39 | 9,511 | 0.13 | 7,548 | 0.12 |
| 2016 | 2.54 | 11,020 | 0.16 | 8,450 | 0.16 |
| 2017 | 4.70 | 11,842 | 0.19 |  |  |

Table 17. Annual abundance estimates of model recruits (millions of crabs), legal male biomass ( t ) with coefficient of variations (CV), and mature male biomass ( t ) with coefficient of variation (CV) for scenario 17_0c for golden king crab in the EAG. Legal male biomass was estimated on July 1 (start of fishing year) of fishing year y. Mature male biomass for fishing year y was estimated on February 15 of year $\mathrm{y}+1$, after the year y fishery total catch removal. Recruits estimates for 1961 to 2017 are restricted to 1985-2017. Equilibrium $\mathrm{MMB}_{\text {eq }}$ and $\mathrm{MMB}_{35 \%}$ are also listed.

| Year | Recruits to the Model ( $\geq$ 101 mm CL) | Mature Male Biomass $(\geq 111 \mathrm{~mm} \mathrm{CL})$ | CV | Legal Size Male <br> Biomass ( $\geq 136$ <br> mm CL) | CV |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{aligned} & \mathrm{MMB}_{\mathrm{eq}}=24,526 \\ & M M B_{35 \%}=7,091 \end{aligned}$ |  |  |  |
| 1985 | 1.59 |  |  | 9,750 | 0.06 |
| 1986 | 0.98 | 9,598 | 0.04 | 8,236 | 0.04 |
| 1987 | 3.98 | 7,282 | 0.04 | 6,372 | 0.04 |
| 1988 | 3.99 | 6,587 | 0.05 | 5,241 | 0.05 |
| 1989 | 2.18 | 6,625 | 0.06 | 4,601 | 0.07 |
| 1990 | 2.73 | 5,964 | 0.06 | 4,239 | 0.07 |
| 1991 | 3.52 | 6,080 | 0.05 | 4,634 | 0.06 |
| 1992 | 2.22 | 6,136 | 0.04 | 4,474 | 0.05 |
| 1993 | 2.09 | 6,069 | 0.04 | 4,487 | 0.05 |
| 1994 | 2.48 | 6,175 | 0.03 | 4,884 | 0.04 |
| 1995 | 2.35 | 5,682 | 0.04 | 4,409 | 0.04 |
| 1996 | 2.34 | 5,129 | 0.04 | 3,820 | 0.05 |
| 1997 | 3.24 | 5,300 | 0.05 | 3,995 | 0.05 |
| 1998 | 3.02 | 5,684 | 0.05 | 4,176 | 0.06 |
| 1999 | 3.15 | 6,445 | 0.06 | 4,749 | 0.06 |
| 2000 | 3.00 | 7,314 | 0.06 | 5,577 | 0.07 |
| 2001 | 2.27 | 8,117 | 0.07 | 6,375 | 0.07 |
| 2002 | 2.85 | 8,630 | 0.07 | 7,075 | 0.08 |
| 2003 | 2.38 | 9,044 | 0.08 | 7,545 | 0.08 |
| 2004 | 2.04 | 9,420 | 0.08 | 7,908 | 0.09 |
| 2005 | 3.03 | 9,502 | 0.09 | 8,147 | 0.09 |
| 2006 | 2.38 | 9,561 | 0.09 | 8,070 | 0.09 |
| 2007 | 2.30 | 9,795 | 0.09 | 8,202 | 0.10 |
| 2008 | 3.52 | 9,803 | 0.09 | 8,342 | 0.10 |
| 2009 | 2.37 | 9,997 | 0.09 | 8,307 | 0.10 |
| 2010 | 2.22 | 10,407 | 0.10 | 8,662 | 0.10 |
| 2011 | 2.81 | 10,360 | 0.10 | 8,886 | 0.10 |
| 2012 | 2.72 | 10,295 | 0.11 | 8,788 | 0.11 |
| 2013 | 2.41 | 10,377 | 0.12 | 8,745 | 0.12 |
| 2014 | 5.59 | 10,332 | 0.14 | 8,758 | 0.14 |
| 2015 | 4.85 | 10,995 | 0.16 | 8,773 | 0.16 |
| 2016 | 2.65 | 12,801 | 0.18 | 9,910 | 0.18 |
| 2017 | 4.70 | 13,767 | 0.21 |  |  |

Table 18. Annual abundance estimates of model recruits (millions of crabs), legal male biomass ( $t$ ) with coefficient of variations (CV), and mature male biomass ( t ) with CV for scenario 17_0d for golden king crab in the EAG. Legal male biomass was estimated on July 1 (start of fishing year) of fishing year y. Mature male biomass for fishing year y was estimated on February 15 of year $y+1$, after the year y fishery total catch removal. Recruits estimates for 1961 to 2017 are restricted to 1985-2017. Equilibrium $\mathrm{MMB}_{\text {eq }}$ and $\mathrm{MMB}_{35 \%}$ are also listed.

| Year | Recruits to the Model ( $\geq$ 101 mm CL) | Mature Male Biomass $(\geq 111 \mathrm{~mm} \mathrm{CL})$ | CV | Legal Size Male <br> Biomass ( $\geq 136$ mm CL) | CV |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{aligned} & \mathrm{MMB}_{\mathrm{eq}}=23,043 \\ & M M B_{35 \%}=6,688 \end{aligned}$ |  |  |  |
| 1985 | 1.66 |  |  | 9,599 | 0.06 |
| 1986 | 0.99 | 9,538 | 0.04 | 8,146 | 0.04 |
| 1987 | 4.14 | 7,286 | 0.04 | 6,356 | 0.04 |
| 1988 | 3.79 | 6,653 | 0.05 | 5,272 | 0.05 |
| 1989 | 2.18 | 6,719 | 0.05 | 4,698 | 0.07 |
| 1990 | 2.70 | 5,989 | 0.06 | 4,296 | 0.07 |
| 1991 | 3.53 | 6,087 | 0.05 | 4,654 | 0.06 |
| 1992 | 2.26 | 6,126 | 0.04 | 4,470 | 0.05 |
| 1993 | 2.13 | 6,067 | 0.04 | 4,475 | 0.05 |
| 1994 | 2.46 | 6,200 | 0.03 | 4,890 | 0.04 |
| 1995 | 2.30 | 5,721 | 0.04 | 4,442 | 0.04 |
| 1996 | 2.26 | 5,150 | 0.04 | 3,853 | 0.04 |
| 1997 | 3.04 | 5,271 | 0.04 | 3,997 | 0.05 |
| 1998 | 2.79 | 5,553 | 0.05 | 4,115 | 0.05 |
| 1999 | 2.97 | 6,149 | 0.05 | 4,562 | 0.05 |
| 2000 | 2.79 | 6,849 | 0.05 | 5,229 | 0.06 |
| 2001 | 2.12 | 7,504 | 0.06 | 5,879 | 0.06 |
| 2002 | 2.69 | 7,894 | 0.06 | 6,449 | 0.06 |
| 2003 | 2.23 | 8,224 | 0.06 | 6,824 | 0.07 |
| 2004 | 1.92 | 8,539 | 0.07 | 7,123 | 0.07 |
| 2005 | 2.89 | 8,584 | 0.07 | 7,319 | 0.07 |
| 2006 | 2.16 | 8,619 | 0.07 | 7,221 | 0.07 |
| 2007 | 2.06 | 8,813 | 0.07 | 7,334 | 0.08 |
| 2008 | 3.21 | 8,743 | 0.07 | 7,424 | 0.08 |
| 2009 | 2.04 | 8,824 | 0.07 | 7,307 | 0.08 |
| 2010 | 1.78 | 9,087 | 0.08 | 7,538 | 0.08 |
| 2011 | 2.15 | 8,861 | 0.08 | 7,617 | 0.08 |
| 2012 | 2.11 | 8,518 | 0.09 | 7,326 | 0.09 |
| 2013 | 1.82 | 8,233 | 0.10 | 6,976 | 0.11 |
| 2014 | 3.58 | 7,848 | 0.12 | 6,651 | 0.12 |
| 2015 | 3.43 | 7,832 | 0.15 | 6,312 | 0.15 |
| 2016 | 2.42 | 8,486 | 0.19 | 6,546 | 0.19 |
| 2017 | 4.70 | 8,833 | 0.23 |  |  |

Table 19. Annual abundance estimates of model recruits (millions of crabs), legal male biomass ( $t$ ) with coefficient of variations (CV), and mature male biomass ( t ) with CV for scenario 17_0e for golden king crab in the EAG. Legal male biomass was estimated on July 1 (start of fishing year) of fishing year y. Mature male biomass for fishing year y was estimated on February 15 of year $y+1$, after the year y fishery total catch removal. Recruits estimates for 1961 to 2017 are restricted to 1985-2017. Equilibrium $\mathrm{MMB}_{\text {eq }}$ and $\mathrm{MMB}_{35 \%}$ are also listed.

| Year | Recruits to the Model ( $\geq$ 101 mm CL) | $\begin{gathered} \text { Mature Male } \\ \text { Biomass } \\ (\geq 111 \mathrm{~mm} \mathrm{CL}) \end{gathered}$ | CV | Legal Size Male Biomass ( $\geq 136$ mm CL) | CV |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{aligned} & \mathrm{MMB}_{\mathrm{eq}}=24,217 \\ & M M B_{35 \%}=7,014 \end{aligned}$ |  |  |  |
| 1985 | 1.75 |  |  | 9,489 | 0.05 |
| 1986 | 0.99 | 9,601 | 0.04 | 8,137 | 0.04 |
| 1987 | 4.25 | 7,417 | 0.04 | 6,449 | 0.04 |
| 1988 | 3.36 | 6,806 | 0.04 | 5,406 | 0.04 |
| 1989 | 2.38 | 6,813 | 0.05 | 4,854 | 0.06 |
| 1990 | 2.63 | 5,974 | 0.05 | 4,325 | 0.06 |
| 1991 | 3.69 | 6,109 | 0.04 | 4,662 | 0.06 |
| 1992 | 2.26 | 6,148 | 0.04 | 4,471 | 0.04 |
| 1993 | 2.07 | 6,158 | 0.04 | 4,522 | 0.04 |
| 1994 | 2.37 | 6,286 | 0.03 | 4,977 | 0.03 |
| 1995 | 2.28 | 5,760 | 0.03 | 4,514 | 0.03 |
| 1996 | 2.22 | 5,130 | 0.04 | 3,860 | 0.04 |
| 1997 | 3.05 | 5,218 | 0.04 | 3,961 | 0.04 |
| 1998 | 2.69 | 5,473 | 0.05 | 4,048 | 0.05 |
| 1999 | 2.99 | 6,042 | 0.05 | 4,479 | 0.05 |
| 2000 | 2.88 | 6,701 | 0.05 | 5,105 | 0.06 |
| 2001 | 2.06 | 7,391 | 0.06 | 5,747 | 0.06 |
| 2002 | 2.87 | 7,821 | 0.06 | 6,364 | 0.06 |
| 2003 | 2.41 | 8,181 | 0.06 | 6,759 | 0.07 |
| 2004 | 1.92 | 8,631 | 0.07 | 7,136 | 0.07 |
| 2005 | 3.12 | 8,776 | 0.07 | 7,458 | 0.07 |
| 2006 | 2.40 | 8,879 | 0.07 | 7,426 | 0.08 |
| 2007 | 2.11 | 9,240 | 0.07 | 7,645 | 0.08 |
| 2008 | 3.84 | 9,295 | 0.08 | 7,886 | 0.08 |
| 2009 | 2.17 | 9,547 | 0.08 | 7,862 | 0.08 |
| 2010 | 2.26 | 10,114 | 0.08 | 8,353 | 0.08 |
| 2011 | 2.99 | 10,069 | 0.08 | 8,635 | 0.08 |
| 2012 | 2.77 | 10,108 | 0.08 | 8,581 | 0.09 |
| 2013 | 2.27 | 10,327 | 0.09 | 8,653 | 0.09 |
| 2014 | 5.75 | 10,317 | 0.10 | 8,764 | 0.10 |
| 2015 | 4.33 | 10,957 | 0.11 | 8,763 | 0.11 |
| 2016 | 2.55 | 12,709 | 0.14 | 9,910 | 0.13 |
| 2017 | 2.39 | 13,440 | 0.17 |  |  |

Table 20. Annual abundance estimates of model recruits (millions of crabs), legal male biomass ( t ) with coefficient of variations (CV), and mature male biomass (t) with CV for scenario 17_0f for golden king crab in the EAG. Legal male biomass was estimated on July 1 (start of fishing year) of fishing year y. Mature male biomass for fishing year y was estimated on February 15 of year $y+1$, after the year y fishery total catch removal. Recruits estimates for 1961 to 2017 are restricted to 1985-2017. Equilibrium $\mathrm{MMB}_{\text {eq }}$ and $\mathrm{MMB}_{35 \%}$ are also listed.

| Year | Recruits to the Model ( $\geq \mathbf{1 0 1}$ mm CL) | $\begin{gathered} \text { Mature Male } \\ \text { Biomass } \\ (\geq 111 \mathrm{~mm} \mathrm{CL}) \end{gathered}$ | CV | Legal Size Male <br> Biomass ( $\geq 136$ mm CL) | CV |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{aligned} & \mathrm{MMB}_{\mathrm{eq}}=23,924 \\ & M M B_{35 \%}=6,946 \end{aligned}$ |  |  |  |
| 1985 | 1.67 |  |  | 9,618 | 0.05 |
| 1986 | 1.00 | 9,534 | 0.04 | 8,146 | 0.04 |
| 1987 | 4.12 | 7,286 | 0.04 | 6,353 | 0.04 |
| 1988 | 3.77 | 6,652 | 0.05 | 5,274 | 0.05 |
| 1989 | 2.20 | 6,706 | 0.05 | 4,698 | 0.07 |
| 1990 | 2.71 | 5,973 | 0.06 | 4,287 | 0.07 |
| 1991 | 3.52 | 6,078 | 0.05 | 4,647 | 0.06 |
| 1992 | 2.27 | 6,116 | 0.04 | 4,466 | 0.05 |
| 1993 | 2.13 | 6,058 | 0.04 | 4,471 | 0.05 |
| 1994 | 2.45 | 6,195 | 0.03 | 4,889 | 0.04 |
| 1995 | 2.29 | 5,716 | 0.04 | 4,442 | 0.04 |
| 1996 | 2.25 | 5,139 | 0.04 | 3,849 | 0.04 |
| 1997 | 3.04 | 5,253 | 0.04 | 3,987 | 0.05 |
| 1998 | 2.78 | 5,530 | 0.05 | 4,100 | 0.05 |
| 1999 | 2.96 | 6,120 | 0.05 | 4,543 | 0.05 |
| 2000 | 2.78 | 6,814 | 0.05 | 5,204 | 0.06 |
| 2001 | 2.11 | 7,466 | 0.06 | 5,850 | 0.06 |
| 2002 | 2.70 | 7,853 | 0.06 | 6,417 | 0.06 |
| 2003 | 2.26 | 8,183 | 0.06 | 6,791 | 0.07 |
| 2004 | 1.95 | 8,511 | 0.07 | 7,093 | 0.07 |
| 2005 | 2.95 | 8,581 | 0.07 | 7,307 | 0.07 |
| 2006 | 2.25 | 8,653 | 0.07 | 7,237 | 0.08 |
| 2007 | 2.17 | 8,907 | 0.08 | 7,394 | 0.08 |
| 2008 | 3.51 | 8,914 | 0.08 | 7,547 | 0.08 |
| 2009 | 2.38 | 9,128 | 0.08 | 7,513 | 0.09 |
| 2010 | 2.18 | 9,625 | 0.08 | 7,913 | 0.09 |
| 2011 | 2.80 | 9,670 | 0.08 | 8,226 | 0.08 |
| 2012 | 2.70 | 9,683 | 0.08 | 8,212 | 0.09 |
| 2013 | 2.34 | 9,841 | 0.09 | 8,243 | 0.09 |
| 2014 | 5.66 | 9,845 | 0.10 | 8,317 | 0.10 |
| 2015 | 4.71 | 10,553 | 0.11 | 8,362 | 0.11 |
| 2016 | 2.59 | 12,415 | 0.13 | 9,560 | 0.13 |
| 2017 | 4.70 | 13,368 | 0.17 |  |  |

Table 21. Negative log-likelihood values of the fits for scenarios (Sc) 17_0 (base), 17_0a (observer CPUE by VAST), 17_0b (observer and fishtick CPUE variable selection by CAIC), 17_0c (Year:Area interaction for observer and fishtick CPUE), 17_0d (three total selectivity and catchability for 1985-04, 2005-12, and 2013-16 time periods), 17_0e (Stage 2 effective sample sizes by McAllister and Ianelli method), and 17_0f (independent pot survey CPUE as an additional likelihood component) for golden king crab in the EAG. Differences in likelihood values are given for scenarios with the same number of data points (base) and free parameters. Likelihood components with zero entry in the entire rows are omitted. Retdcatch $\mathrm{B}=$ retained catch biomass.

| Likelihood Component | Sc 17_0 | $\begin{gathered} \text { Sc } \\ \text { 17_0a } \end{gathered}$ | $\begin{gathered} \text { Sc } \\ \text { 17_0b } \end{gathered}$ | $\begin{gathered} \text { Sc } \\ \text { 17_0c } \end{gathered}$ | $\begin{gathered} \text { Sc } \\ \text { 17_0d } \end{gathered}$ | $\begin{gathered} \text { Sc } \\ \text { 17_0e } \end{gathered}$ | $\begin{gathered} \text { Sc } \\ \text { 17_0f } \end{gathered}$ | $\begin{gathered} \text { Sc17_0a- } \\ \text { Sc 17_0 } \end{gathered}$ | $\begin{gathered} \hline \text { Sc 17_0b - } \\ \text { Sc 17_0 } \end{gathered}$ | $\begin{gathered} \hline \text { Sc 17_0c } \\ - \\ \text { Sc 17_0 } \\ \hline \end{gathered}$ | $\begin{gathered} \text { Sc 17_0e- } \\ \text { Sc 17_0 } \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Number of free parameters | 140 | 140 | 140 | 140 | 143 | 140 | 141 |  |  |  |  |
| Data | Base | Base | Base | Base | Base | Base |  |  |  |  |  |
| Retlencomp | -1177.540 | -1177.110 | -1178.030 | -1174.470 | -1180.060 | -1235.080 | -1177.740 | 0.43 | -0.490 | 3.070 | -57.540 |
| Totallencomp | -1249.120 | -1260.300 | -1248.190 | -1261.890 | -1258.200 | -1192.770 | -1249.490 | -11.18 | 0.930 | -12.770 | 56.350 |
| Observer cpue | -12.551 | -5.466 | -6.545 | -3.945 | -12.776 | -12.429 | -12.364 | 7.085 | 6.006 | 8.606 | 0.122 |
| RetdcatchB | 7.502 | 8.109 | 7.283 | 8.009 | 7.581 | 7.034 | 7.501 | 0.607 | -0.219 | 0.507 | -0.468 |
| TotalcatchB | 18.260 | 18.609 | 18.199 | 18.611 | 18.419 | 17.723 | 18.267 | 0.349 | -0.061 | 0.351 | -0.537 |
| GdiscdcatchB | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0 | 0 | 0 | 0 |
| Rec_dev | 7.571 | 7.435 | 6.880 | 7.804 | 5.937 | 7.966 | 7.552 | -0.136 | -0.691 | 0.233 | 0.395 |
| Pot F_dev | 0.013 | 0.014 | 0.013 | 0.015 | 0.013 | 0.013 | 0.013 | 0.001 | 0 | 0.002 | 0 |
| Gbyc_F_dev | 0.026 | 0.026 | 0.026 | 0.026 | 0.028 | 0.026 | 0.026 | 0 | 0 | 0 | 0 |
| Tag | 2692.200 | 2691.860 | 2692.350 | 2691.730 | 2692.220 | 2692.450 | $2692.200$ | -0.34 | 0.150 | -0.470 | 0.250 |
| Fishery cpue | -0.460 | -0.565 | -2.206 | 10.74300 | -0.461 | $-0.347$ | $-0.463$ | -0.105 | -1.745 | 11.203 | 0.113 |
| RetcatchN | 0.007999 | 0.007584 | 0.007019 | 0.007569 | 0.005034 | 0.010917 | 0.0079 | -0.00042 | -0.00098 | -0.00043 | 0.002918 |
| Total | 285.910 | 282.618 | 289.789 | 296.634 | 272.703 | 284.602 | 285.765 | -3.292 | 3.879 | 10.724 | -1.308 |

Table 22. Time series of GLM estimated CPUE indices and coefficient of variations (CV) for the fish ticket based retained catch-per-pot lift for the WAG golden king crab stock. The GLM was fitted to the 1985/86 to 1998/99 time series of data. GLM predictor variables selected by R square criteria.

| Year | CPUE <br> Index | CV |
| :---: | :---: | :---: |
| $1985 / 86$ | 1.87 | 0.03 |
| $1986 / 87$ | 1.68 | 0.03 |
| $1987 / 88$ | 1.26 | 0.04 |
| $1988 / 89$ | 1.37 | 0.03 |
| $1989 / 90$ | 1.10 | 0.03 |
| $1990 / 91$ | 0.84 | 0.04 |
| $1991 / 92$ | 0.73 | 0.06 |
| $1992 / 93$ | 0.70 | 0.06 |
| $1993 / 94$ | 0.67 | 0.08 |
| $1994 / 95$ | 0.84 | 0.05 |
| $1995 / 96$ | 0.87 | 0.05 |
| $1996 / 97$ | 0.85 | 0.04 |
| $1997 / 98$ | 0.84 | 0.04 |
| $1998 / 99$ | 1.12 | 0.03 |

Table 23. The initial input number of vessel-days/trips and Stage-2 effective sample sizes iteratively estimated by Francis method for retained, total, and groundfish discard catch size compositions of golden king crab for scenario 17_0 model fit to WAG data. NA: not available.

| Year | Initial <br> Input <br> Retained <br> Vessel- | Stage-2 <br> Retained <br> Effective <br> Sample <br> Dize (no) | Initial <br> Input <br> Total <br> Vessel- <br> Days <br> Sample | Stage-2 <br> Total <br> Effective <br> Sample <br> Size (no) | Initial <br> Input <br> Groundfish <br> Trip <br> Sample <br> Saze (no) | Stage-2 <br> Groundfish <br> Effective |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Sample <br> Size (no) |  |  |  |  |  |  |
|  | Size (no) |  |  |  |  |  |

Table 24. The initial input number of vessel-days/trips and Stage-2 effective sample sizes iteratively estimated by Francis method for retained, total, and groundfish discard catch size compositions of golden king crab for scenario 17_0a model fit to WAG data. NA: not available.

| Year | Initial <br> Input <br> Retained <br> Vessel- <br> Days | Stage-2 <br> Retained <br> Effective <br> Sample <br> Size (no) | Initial <br> Input <br> Total <br> Vessel- <br> Days <br> Sample | Stage-2 <br> Total <br> Effective <br> Sample <br> Size (no) | Initial <br> Input <br> Groundfish <br> Trip <br> Sample <br> Size (no) | Stage-2 <br> Groundfish <br> Effective <br> Sample <br> Size (no) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Size (no) |  | Size <br> (no) |  |  |  |
| 1985/86 | 45 | 23 |  |  |  |  |
| $1986 / 87$ | 23 | 12 |  |  |  |  |
| 1987/88 | 8 | 4 |  |  |  |  |
| $1988 / 89$ | 286 | 148 |  |  |  |  |
| $1989 / 90$ | 513 | 266 |  |  | 7 | 5 |
| $1990 / 91$ | 205 | 106 | 190 | 89 | 6 | 5 |
| $1991 / 92$ | 102 | 53 | 104 | 49 | 1 | 1 |
| $1992 / 93$ | 76 | 39 | 94 | 44 | 3 | 2 |
| $1993 / 94$ | 378 | 196 | 62 | 29 | NA | NA |
| $1994 / 95$ | 367 | 190 | 119 | 56 | 2 | 2 |
| $1995 / 96$ | 705 | 365 | 907 | 427 | 5 | 4 |
| $1996 / 97$ | 817 | 423 | 1,061 | 499 | 8 | 6 |
| $1997 / 98$ | 984 | 510 | 1,116 | 525 | 6 | 5 |
| $1998 / 99$ | 613 | 318 | 638 | 300 | 14 | 11 |
| $1999 / 00$ | 915 | 474 | 1,155 | 543 | 18 | 14 |
| $2000 / 01$ | 1,029 | 533 | 1,205 | 567 | 11 | 8 |
| $2001 / 02$ | 898 | 465 | 975 | 459 | 11 | 8 |
| $2002 / 03$ | 628 | 325 | 675 | 318 | 16 | 12 |
| $2003 / 04$ | 688 | 357 | 700 | 329 | 8 | 6 |
| $2004 / 05$ | 449 | 233 | 488 | 230 | 9 | 7 |
| $2005 / 06$ | 337 | 175 | 220 | 104 | 6 | 5 |
| $2006 / 07$ | 337 | 175 | 321 | 151 | 14 | 11 |
| $2007 / 08$ | 276 | 143 | 257 | 121 | 17 | 13 |
| $2008 / 09$ | 318 | 165 | 258 | 121 | 19 | 14 |
| $2009 / 10$ | 362 | 188 | 292 | 137 | 24 | 18 |
| $2010 / 11$ | 328 | 170 | 222 | 104 | 13 | 10 |
| $2011 / 12$ | 295 | 153 | 252 | 119 | 14 | 11 |
| $2012 / 13$ | 288 | 149 | 241 | 113 | 18 | 14 |
| $2013 / 14$ | 327 | 169 | 236 | 111 | 17 | 13 |
| $2014 / 15$ | 305 | 158 | 219 | 103 | 18 | 14 |
| $2015 / 16$ | 287 | 149 | 243 | 114 | 10 | 8 |
| $2016 / 17$ | 392 | 203 | 253 | 119 | 12 | 9 |
|  |  |  |  |  |  |  |

Table 25. The initial input number of vessel-days/trips and Stage-2 effective sample sizes iteratively estimated by Francis method for retained, total, and groundfish discard catch size compositions of golden king crab for scenario 17_0b model fit to WAG data. NA: not available.

| Year | Initial <br> Input Retained VesselDays Sample Size (no) | Stage-2 <br> Retained <br> Effective <br> Sample <br> Size (no) | Initial <br> Input <br> Total <br> Vessel- <br> Days <br> Sample <br> Size (no) | Stage-2 <br> Total <br> Effective <br> Sample <br> Size (no) | Initial Input Groundfish Trip Sample Size (no) | Stage-2 Groundfish Effective Sample Size (no) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1985/86 | 45 | 23 |  |  |  |  |
| 1986/87 | 23 | 12 |  |  |  |  |
| 1987/88 | 8 | 4 |  |  |  |  |
| 1988/89 | 286 | 145 |  |  |  |  |
| 1989/90 | 513 | 261 |  |  | 7 | 5 |
| 1990/91 | 205 | 104 | 190 | 92 | 6 | 5 |
| 1991/92 | 102 | 52 | 104 | 50 | 1 | 1 |
| 1992/93 | 76 | 39 | 94 | 45 | 3 | 2 |
| 1993/94 | 378 | 192 | 62 | 30 | NA | NA |
| 1994/95 | 367 | 187 | 119 | 57 | 2 | 2 |
| 1995/96 | 705 | 358 | 907 | 438 | 5 | 4 |
| 1996/97 | 817 | 415 | 1,061 | 513 | 8 | 6 |
| 1997/98 | 984 | 500 | 1,116 | 539 | 6 | 5 |
| 1998/99 | 613 | 312 | 638 | 308 | 14 | 11 |
| 1999/00 | 915 | 465 | 1,155 | 558 | 18 | 14 |
| 2000/01 | 1,029 | 523 | 1,205 | 582 | 11 | 8 |
| 2001/02 | 898 | 456 | 975 | 471 | 11 | 8 |
| 2002/03 | 628 | 319 | 675 | 326 | 16 | 12 |
| 2003/04 | 688 | 350 | 700 | 338 | 8 | 6 |
| 2004/05 | 449 | 228 | 488 | 236 | 9 | 7 |
| 2005/06 | 337 | 171 | 220 | 106 | 6 | 5 |
| 2006/07 | 337 | 171 | 321 | 155 | 14 | 11 |
| 2007/08 | 276 | 140 | 257 | 124 | 17 | 13 |
| 2008/09 | 318 | 162 | 258 | 125 | 19 | 15 |
| 2009/10 | 362 | 184 | 292 | 141 | 24 | 18 |
| 2010/11 | 328 | 167 | 222 | 107 | 13 | 10 |
| 2011/12 | 295 | 150 | 252 | 122 | 14 | 11 |
| 2012/13 | 288 | 146 | 241 | 116 | 18 | 14 |
| 2013/14 | 327 | 166 | 236 | 114 | 17 | 13 |
| 2014/15 | 305 | 155 | 219 | 106 | 18 | 14 |
| 2015/16 | 287 | 146 | 243 | 117 | 10 | 8 |
| 2016/17 | 392 | 199 | 253 | 122 | 12 | 9 |

Table 26. The initial input number of vessel-days/trips and Stage-2 effective sample sizes iteratively estimated by Francis method for retained, total, and groundfish discard catch size compositions of golden king crab for scenario 17_0c model fit to WAG data. NA: not available.

| Year | Initial <br> Input Retained VesselDays Sample Size (no) | Stage-2 <br> Retained <br> Effective <br> Sample <br> Size (no) | Initial <br> Input <br> Total <br> Vessel- <br> Days <br> Sample <br> Size <br> (no) | Stage-2 <br> Total Effective Sample Size (no) | Initial Input Groundfish Trip Sample Size (no) | Stage-2 <br> Groundfish Effective Sample Size (no) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1985/86 | 45 | 22 |  |  |  |  |
| 1986/87 | 23 | 11 |  |  |  |  |
| 1987/88 | 8 | 4 |  |  |  |  |
| 1988/89 | 286 | 142 |  |  |  |  |
| 1989/90 | 513 | 255 |  |  | 7 | 5 |
| 1990/91 | 205 | 102 | 190 | 91 | 6 | 5 |
| 1991/92 | 102 | 51 | 104 | 50 | 1 | 1 |
| 1992/93 | 76 | 38 | 94 | 45 | 3 | 2 |
| 1993/94 | 378 | 188 | 62 | 30 | NA | NA |
| 1994/95 | 367 | 183 | 119 | 57 | 2 | 2 |
| 1995/96 | 705 | 351 | 907 | 433 | 5 | 4 |
| 1996/97 | 817 | 407 | 1,061 | 506 | 8 | 6 |
| 1997/98 | 984 | 490 | 1,116 | 533 | 6 | 5 |
| 1998/99 | 613 | 305 | 638 | 305 | 14 | 11 |
| 1999/00 | 915 | 456 | 1,155 | 551 | 18 | 14 |
| 2000/01 | 1,029 | 512 | 1,205 | 575 | 11 | 8 |
| 2001/02 | 898 | 447 | 975 | 465 | 11 | 8 |
| 2002/03 | 628 | 313 | 675 | 322 | 16 | 12 |
| 2003/04 | 688 | 343 | 700 | 334 | 8 | 6 |
| 2004/05 | 449 | 224 | 488 | 233 | 9 | 7 |
| 2005/06 | 337 | 168 | 220 | 105 | 6 | 5 |
| 2006/07 | 337 | 168 | 321 | 153 | 14 | 11 |
| 2007/08 | 276 | 137 | 257 | 123 | 17 | 13 |
| 2008/09 | 318 | 158 | 258 | 123 | 19 | 14 |
| 2009/10 | 362 | 180 | 292 | 139 | 24 | 18 |
| 2010/11 | 328 | 163 | 222 | 106 | 13 | 10 |
| 2011/12 | 295 | 147 | 252 | 120 | 14 | 11 |
| 2012/13 | 288 | 143 | 241 | 115 | 18 | 14 |
| 2013/14 | 327 | 163 | 236 | 113 | 17 | 13 |
| 2014/15 | 305 | 152 | 219 | 105 | 18 | 14 |
| 2015/16 | 287 | 143 | 243 | 116 | 10 | 8 |
| 2016/17 | 392 | 195 | 253 | 121 | 12 | 9 |

Table 27. The initial input number of vessel-days/trips and Stage-2 effective sample sizes iteratively estimated by Francis method for retained, total, and groundfish discard catch size compositions of golden king crab for scenario 17_0d model fit to WAG data. NA: not available.

| Year | Initial <br> Input Retained VesselDays Sample Size (no) | Stage-2 <br> Retained <br> Effective <br> Sample <br> Size (no) | Initial <br> Input <br> Total <br> Vessel- <br> Days <br> Sample <br> Size <br> (no) | Stage-2 <br> Total Effective Sample Size (no) | Initial Input Groundfish Trip Sample Size (no) | Stage-2 <br> Groundfish Effective Sample Size (no) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1985/86 | 45 | 25 |  |  |  |  |
| 1986/87 | 23 | 13 |  |  |  |  |
| 1987/88 | 8 | 4 |  |  |  |  |
| 1988/89 | 286 | 160 |  |  |  |  |
| 1989/90 | 513 | 286 |  |  | 7 | 5 |
| 1990/91 | 205 | 114 | 190 | 92 | 6 | 5 |
| 1991/92 | 102 | 57 | 104 | 50 | 1 | 1 |
| 1992/93 | 76 | 42 | 94 | 45 | 3 | 2 |
| 1993/94 | 378 | 211 | 62 | 30 | NA | NA |
| 1994/95 | 367 | 205 | 119 | 57 | 2 | 2 |
| 1995/96 | 705 | 393 | 907 | 438 | 5 | 4 |
| 1996/97 | 817 | 456 | 1,061 | 512 | 8 | 6 |
| 1997/98 | 984 | 549 | 1,116 | 539 | 6 | 5 |
| 1998/99 | 613 | 342 | 638 | 308 | 14 | 11 |
| 1999/00 | 915 | 510 | 1,155 | 557 | 18 | 14 |
| 2000/01 | 1,029 | 574 | 1,205 | 582 | 11 | 8 |
| 2001/02 | 898 | 501 | 975 | 471 | 11 | 8 |
| 2002/03 | 628 | 350 | 675 | 326 | 16 | 12 |
| 2003/04 | 688 | 384 | 700 | 338 | 8 | 6 |
| 2004/05 | 449 | 250 | 488 | 236 | 9 | 7 |
| 2005/06 | 337 | 188 | 220 | 106 | 6 | 5 |
| 2006/07 | 337 | 188 | 321 | 155 | 14 | 11 |
| 2007/08 | 276 | 154 | 257 | 124 | 17 | 13 |
| 2008/09 | 318 | 177 | 258 | 125 | 19 | 14 |
| 2009/10 | 362 | 202 | 292 | 141 | 24 | 18 |
| 2010/11 | 328 | 183 | 222 | 107 | 13 | 10 |
| 2011/12 | 295 | 165 | 252 | 122 | 14 | 11 |
| 2012/13 | 288 | 161 | 241 | 116 | 18 | 14 |
| 2013/14 | 327 | 182 | 236 | 114 | 17 | 13 |
| 2014/15 | 305 | 170 | 219 | 106 | 18 | 14 |
| 2015/16 | 287 | 160 | 243 | 117 | 10 | 8 |
| 2016/17 | 392 | 219 | 253 | 122 | 12 | 9 |

Table 28. The initial input number of vessel-days/trips and Stage-2 effective sample sizes iteratively estimated by McAllister and Ianelli method for retained, total, and groundfish discard catch size compositions of golden king crab for scenario 17_0e model fit to WAG data. NA: not available.
$\left.\begin{array}{ccccccc}\hline \text { Year } & \begin{array}{c}\text { Initial } \\ \text { Input } \\ \text { Retained } \\ \text { Vessel- } \\ \text { Days }\end{array} & \begin{array}{c}\text { Stage-2 } \\ \text { Retained } \\ \text { Effective } \\ \text { Sample } \\ \text { Size (no) }\end{array} & \begin{array}{c}\text { Initial } \\ \text { Input } \\ \text { Total } \\ \text { Vessel- } \\ \text { Days } \\ \text { Sample }\end{array} & \begin{array}{c}\text { Stage-2 } \\ \text { Total } \\ \text { Effective } \\ \text { Sample } \\ \text { Size (no) }\end{array} & \begin{array}{c}\text { Initial } \\ \text { Input } \\ \text { Groundfish } \\ \text { Trip } \\ \text { Sample } \\ \text { Size (no) }\end{array} & \begin{array}{c}\text { Stage-2 } \\ \text { Groundfish } \\ \text { Effective } \\ \text { Sample Size }\end{array} \\ & \begin{array}{ccc}\text { (no) } \\ \text { Size (no) }\end{array} & & & \\ \text { Size (no) }\end{array}\right]$

Table 29. Parameter estimates and coefficient of variations (CV) with the 2016 MMB (MMB on 15 Feb 2017) for scenarios 17_0, 17_0a, 17_0b, and 17_0c for the golden king crab data from the WAG, 1985/86-2016/17. Recruitment and fishing mortality deviations and initial size frequency determination parameters were omitted from this list.

|  | Scenario 17_0 |  | Scenario 17_0a |  | Scenario 17_0b |  | Scenario 17_0c |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Parameter | Estimate | CV | Estimate | CV | Estimate | CV | Estimate | CV | Limits |
| $\log _{\_} \omega_{1}$ ( growth incr. intercept) | 2.54 | 0.006 | 2.54 | 0.006 | 2.54 | 0.006 | 2.54 | 0.006 | 1.0, 4.5 |
| $\omega_{2}$ (growth incr. slope) | -7.81 | 0.22 | -7.84 | 0.22 | -7.74 | 0.22 | -7.74 | 0.22 | -12.0,-5.0 |
| $\log _{-} \mathrm{a}$ (molt prob. slope) | -2.61 | 0.03 | -2.61 | 0.03 | -2.61 | 0.03 | -2.61 | 0.03 | -4.61,-1.39 |
| log_b (molt prob. L50) | 4.95 | 0.001 | 4.95 | 0.00 | 4.95 | 0.001 | 4.95 | 0.001 | 3.869,5.05 |
| $\sigma$ (growth variability std) | 3.69 | 0.03 | 3.68 | 0.03 | 3.69 | 0.03 | 3.69 | 0.03 | 0.1,12.0 |
| log_total sel deltae, 1985-04 | 3.40 | 0.02 | 3.40 | 0.02 | 3.40 | 0.01 | 3.40 | 0.01 | 0.,4.4 |
| $\log _{-}$total sel delta $\theta, 2005-16$ | 2.90 | 0.02 | 2.89 | 0.02 | 2.89 | 0.02 | 2.89 | 0.02 | 0.,4.4 |
| $\log _{-}$ret. sel delta $\theta$, 1985-16 | 1.78 | 0.02 | 1.77 | 0.02 | 1.78 | 0.02 | 1.78 | 0.02 | 0.,4.4 |
| log_tot sel $\theta_{50}$, 1985-04 | 4.86 | 0.002 | 4.86 | 0.002 | 4.87 | 0.002 | 4.87 | 0.002 | 4.0,5.0 |
| log_tot sel $\theta_{50}, 2005-16$ | 4.90 | 0.002 | 4.90 | 0.002 | 4.90 | 0.002 | 4.90 | 0.002 | 4.0,5.0 |
| log_ret. sel $\theta_{50}, 1985-16$ | 4.92 | 0.0002 | 4.92 | 0.00 | 4.92 | 0.0002 | 4.92 | 0.0002 | 4.0,5.0 |
| $\log _{\_} \beta_{\mathrm{r}}$ (rec.distribution par.) | -1.05 | 0.16 | -1.06 | 0.16 | -1.05 | 0.16 | -1.05 | 0.16 | -12.0, 12.0 |
| $\operatorname{logq} 2$ (catchability 1995-04) | -0.06 | 1.18 | -0.06 | 1.16 | -0.09 | 0.75 | -0.09 | 0.75 | -9.0, 2.25 |
| logq3 (catchability 2005-16) | -0.38 | 0.24 | -0.39 | 0.22 | -0.37 | 0.29 | -0.37 | 0.29 | -9.0, 2.25 |
| log_mean_rec (mean rec.) | 0.725 | 0.06 | 0.727 | 0.06 | 0.720 | 0.06 | 0.720 | 0.06 | 0.01, 5.0 |
| log_mean_Fpot (Pot fishery F) | -0.699 | 0.09 | -0.709 | 0.09 | -0.692 | 0.09 | -0.692 | 0.09 | -15.0, -0.01 |
| log_mean_Fground (GF byc. F) | -8.371 | 0.10 | -8.376 | 0.10 | -8.364 | 0.10 | -8.364 | 0.10 | -15.0, -1.6 |
| $\sigma_{e}^{2}$ (observer CPUE additional var) | 0.019 | 0.38 | 0.012 | 0.47 | 0.054 | 0.34 | 0.054 | 0.34 | 0.0, 0.15 |
| $\sigma_{e}^{2}$ (fishery CPUE additional var) | 0.025 | 0.60 | 0.025 | 0.62 | 0.013 | 0.58 | 0.013 | 0.58 | 0.0,1.0 |
| 2016 MMB | 6,269 | 0.17 | 6,280 | 0.16 | 5,884 | 0.22 | 5,884 | 0.22 |  |

Table 30. Parameter estimates and coefficient of variations (CV) with the 2016 MMB (MMB on 15 Feb 2017) for scenarios 17_0d and 17_0e for the golden king crab data from the WAG, 1985/86-2016/17. Recruitment and fishing mortality deviations and initial size frequency determination parameters were omitted from this list.

|  | Scenario 17_0d |  | Scenario 17_0e |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Parameter | Estimate | CV | Estimate | CV | Limits |
| $\log _{-} \omega_{1}$ ( growth incr. intercept) | 2.54 | 0.006 | 2.54 | 0.006 | 1.0, 4.5 |
| $\omega_{2}$ ( growth incr. slope) | -7.74 | 0.22 | -7.29 | 0.23 | -12.0,-5.0 |
| $\log _{\_} \mathrm{a}$ (molt prob. slope) | -2.62 | 0.03 | -2.67 | 0.02 | -4.61,-1.39 |
| $\log _{-} \mathrm{b}$ (molt prob. L50) | 4.95 | 0.001 | 4.95 | 0.00 | 3.869,5.05 |
| $\sigma$ (growth variability std) | 3.68 | 0.03 | 3.69 | 0.03 | 0.1,12.0 |
| log_total sel deltay, 1985-04 | 3.39 | 0.01 | 3.36 | 0.01 | 0.,4.4 |
| $\log _{-}$total sel delta $\theta$, 2005-12 | 2.90 | 0.03 |  |  | 0.,4.4 |
| $\log _{-}$total sel delta $\theta, 2013-16$ or 2005-16 | 2.92 | 0.03 | 2.89 | 0.02 | 0.,4.4 |
| $\log _{-}$ret. sel delta $\theta, 1985-16$ | 1.78 | 0.02 | 1.78 | 0.02 | 0.,4.4 |
| log_tot sel $\theta_{50}, 1985-04$ | 4.87 | 0.002 | 4.87 | 0.002 | 4.0,5.0 |
| log_tot sel $\theta_{50}, 2005-12$ | 4.89 | 0.002 |  |  | 4.0,5.0 |
| log_tot sel $\theta_{50}, 2013-16$ or 2005-16 | 4.92 | 0.003 | 4.90 | 0.002 | 4.0,5.0 |
| log_ret. sel $\theta_{50}, 1985-16$ | 4.92 | 0.00 | 4.92 | 0.00 | 4.0,5.0 |
| $\log _{\_} \beta_{\mathrm{r}}$ (rec.distribution par.) | -1.06 | 0.15 | -1.10 | 0.15 | -12.0, 12.0 |
| Logq1 (catchability 1985-04) | -0.067 | 1.02 | -0.04 | 1.62 | -9.0, 2.25 |
| Logq3 (catchability 2005-12) | -0.424 | 0.21 |  |  | -9.0, 2.25 |
| Logq2 (catchability 2013-16 or 2005-16) | -0.098 | 1.80 | -0.41 | 0.20 | -9.0, 2.25 |
| log_mean_rec (mean rec.) | 0.719 | 0.06 | 0.717 | 0.06 | 0.01, 5.0 |
| log_mean_Fpot (Pot fishery F) | -0.681 | 0.09 | -0.710 | 0.08 | -15.0, -0.01 |
| log_mean_Fground (GF byc. F) | -8.364 | 0.10 | -8.390 | 0.10 | -15.0, -1.6 |
| $\sigma_{e}^{2}$ (observer CPUE additional var) | 0.023 | 0.38 | 0.020 | 0.39 | 0.0, 0.15 |
| $\sigma_{e}^{2}$ (fishery CPUE additional var) | 0.026 | 0.57 | 0.037 | 0.53 | 0.0,1.0 |
| 2016 MMB | 6,136 | 0.23 | 6,355 | 0.17 |  |

Table 31. Annual abundance estimates of model recruits (millions of crabs), legal male biomass ( t ) with coefficient of variations (CV), and mature male biomass ( t ) with CV for scenario 17_0 for golden king crab in the WAG. Legal male biomass was estimated on July 1 (start of fishing year) of fishing year y. Mature male biomass for fishing year y was estimated on February 15 of year $y+1$, after the year y fishery total catch removal. Recruits estimates for 1961 to 2017 are restricted to 1985-2017. Equilibrium $\mathrm{MMB}_{\text {eq }}$ and $\mathrm{MMB}_{35 \%}$ are also listed.

| Year | Recruits to the <br> Model ( $\geq \mathbf{1 0 1}$ mm CL) | $\begin{gathered} \text { Mature Male } \\ \text { Biomass } \\ (\geq 111 \mathrm{~mm} \mathrm{CL}) \end{gathered}$ | CV | Legal Size Male Biomass ( $\geq 136$ mm CL) | CV |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{aligned} & \mathrm{MMB}_{\mathrm{eq}}=17,827 \\ & M M B_{35 \%}=5,138 \end{aligned}$ |  |  |  |
| 1985 | 3.75 |  |  | 8,812 | 0.11 |
| 1986 | 3.41 | 10,641 | 0.05 | 8,387 | 0.08 |
| 1987 | 2.69 | 8,164 | 0.05 | 5,971 | 0.06 |
| 1988 | 1.92 | 7,496 | 0.04 | 5,553 | 0.05 |
| 1989 | 2.55 | 6,432 | 0.04 | 4,896 | 0.04 |
| 1990 | 1.85 | 4,468 | 0.05 | 3,106 | 0.06 |
| 1991 | 1.56 | 4,172 | 0.05 | 2,870 | 0.05 |
| 1992 | 2.07 | 3,906 | 0.05 | 2,810 | 0.05 |
| 1993 | 1.60 | 4,025 | 0.04 | 2,923 | 0.05 |
| 1994 | 1.96 | 4,613 | 0.03 | 3,493 | 0.03 |
| 1995 | 1.88 | 3,924 | 0.03 | 2,833 | 0.04 |
| 1996 | 1.72 | 3,925 | 0.04 | 2,785 | 0.04 |
| 1997 | 1.84 | 3,934 | 0.04 | 2,828 | 0.04 |
| 1998 | 1.90 | 4,002 | 0.04 | 2,909 | 0.04 |
| 1999 | 2.23 | 4,318 | 0.04 | 3,184 | 0.04 |
| 2000 | 2.49 | 4,351 | 0.04 | 3,122 | 0.04 |
| 2001 | 2.54 | 4,507 | 0.04 | 3,129 | 0.04 |
| 2002 | 2.48 | 4,943 | 0.05 | 3,451 | 0.05 |
| 2003 | 1.78 | 5,489 | 0.05 | 3,961 | 0.05 |
| 2004 | 2.27 | 5,810 | 0.06 | 4,442 | 0.06 |
| 2005 | 2.29 | 5,913 | 0.06 | 4,626 | 0.06 |
| 2006 | 2.41 | 6,194 | 0.06 | 4,797 | 0.06 |
| 2007 | 1.71 | 6,698 | 0.06 | 5,224 | 0.06 |
| 2008 | 1.48 | 6,863 | 0.05 | 5,502 | 0.06 |
| 2009 | 1.89 | 6,658 | 0.05 | 5,539 | 0.05 |
| 2010 | 1.59 | 6,263 | 0.05 | 5,173 | 0.05 |
| 2011 | 1.14 | 5,972 | 0.05 | 4,864 | 0.05 |
| 2012 | 1.80 | 5,465 | 0.05 | 4,521 | 0.05 |
| 2013 | 2.29 | 4,850 | 0.05 | 3,903 | 0.05 |
| 2014 | 1.59 | 4,627 | 0.07 | 3,421 | 0.07 |
| 2015 | 3.63 | 4,719 | 0.09 | 3,491 | 0.08 |
| 2016 | 2.23 | 5,204 | 0.13 | 3,650 | 0.12 |
| 2017 | 2.06 | 6,269 | 0.17 |  |  |

Table 32. Annual abundance estimates of model recruits (millions of crabs), legal male biomass ( t ) with coefficient of variations (CV), and mature male biomass ( t ) CV for scenario 17_0a for golden king crab in the WAG. Legal male biomass was estimated on July 1 (start of fishing year) of fishing year y. Mature male biomass for fishing year y was estimated on February 15 of year $y+1$, after the year y fishery total catch removal. Recruits estimates for 1961 to 2017 are restricted to 1985-2017. Equilibrium $\mathrm{MMB}_{\mathrm{eq}}$ and $\mathrm{MMB}_{35 \%}$ are also listed.

| Year | Recruits to <br> the Model ( $\geq$ <br> 101 mm CL) | Mature Male <br> Biomass <br> $(\geq \mathbf{1 1 1 ~ m m ~ C L ) ~}$ | CV | Legal Size Male <br> Biomass ( $\geq \mathbf{1 3 6}$ <br> mm CL) | CV |
| :--- | :---: | :---: | :---: | :---: | :---: |
|  | MMBeq $=17,862$ <br> MMB $35 \%=5,173$ |  |  |  |  |
| 1985 | 3.76 | 10,646 | 0.05 |  |  |
| 1986 | 3.41 | 8,170 | 0.05 | 8,388 | 0.11 |
| 1987 | 2.69 | 7,499 | 0.04 | 5,556 | 0.08 |
| 1988 | 1.92 | 6,435 | 0.04 | 4,898 | 0.06 |
| 1989 | 2.55 | 4,471 | 0.04 | 3,108 | 0.05 |
| 1990 | 1.85 | 4,175 | 0.05 | 2,873 | 0.04 |
| 1991 | 1.56 | 3,908 | 0.05 | 2,812 | 0.06 |
| 1992 | 2.06 | 4,022 | 0.04 | 2,924 | 0.05 |
| 1993 | 1.60 | 4,605 | 0.03 | 3,487 | 0.05 |
| 1994 | 1.99 | 3,921 | 0.03 | 2,826 | 0.05 |
| 1995 | 1.89 | 3,940 | 0.04 | 2,791 | 0.04 |
| 1996 | 1.72 | 3,957 | 0.04 | 2,846 | 0.04 |
| 1997 | 1.85 | 4,026 | 0.04 | 2,931 | 0.04 |
| 1998 | 1.91 | 4,344 | 0.04 | 3,206 | 0.04 |
| 1999 | 2.23 | 4,379 | 0.04 | 3,147 | 0.04 |
| 2000 | 2.54 | 4,544 | 0.04 | 3,154 | 0.04 |
| 2001 | 2.59 | 5,013 | 0.05 | 3,495 | 0.04 |
| 2002 | 2.50 | 5,592 | 0.05 | 4,038 | 0.05 |
| 2003 | 1.81 | 5,932 | 0.05 | 4,543 | 0.05 |
| 2004 | 2.26 | 6,044 | 0.06 | 4,744 | 0.05 |
| 2005 | 2.21 | 6,295 | 0.06 | 4,913 | 0.06 |
| 2006 | 2.42 | 6,755 | 0.05 | 5,299 | 0.06 |
| 2007 | 1.69 | 6,898 | 0.05 | 5,545 | 0.06 |
| 2008 | 1.48 | 6,668 | 0.05 | 5,557 | 0.06 |
| 2009 | 1.91 | 6,268 | 0.05 | 5,175 | 0.05 |
| 2010 | 1.61 | 5,987 | 0.05 | 4,867 | 0.05 |
| 2011 | 1.13 | 5,488 | 0.05 | 4,536 | 0.05 |
| 2012 | 1.81 | 4,872 | 0.05 | 3,922 | 0.05 |
| 2013 | 2.28 | 4,650 | 0.06 | 3,443 | 0.05 |
| 2014 | 1.59 | 4,739 | 0.08 | 3,510 | 0.06 |
| 2015 | 3.62 | 5,218 | 0.11 | 3,666 | 0.08 |
| 2016 | 2.23 | 6,280 | 0.16 |  | 0.10 |
| 2017 | 2.07 |  |  |  |  |
|  |  |  |  |  |  |

Table 33. Annual abundance estimates of model recruits (millions of crabs), legal male biomass ( $t$ ) with coefficient of variations (CV), and mature male biomass ( t ) with CV for scenario 17_0b for golden king crab in the WAG. Legal male biomass was estimated on July 1 (start of fishing year) of fishing year y. Mature male biomass for fishing year y was estimated on February 15 of year $y+1$, after the year y fishery total catch removal. Recruits estimates for 1961 to 2017 are restricted to 1985-2017. Equilibrium $\mathrm{MMB}_{\text {eq }}$ and $\mathrm{MMB}_{35 \%}$ are also listed.

| Year | Recruits to the Model ( $\geq$ 101 mm CL) | Mature Male Biomass ( $\geq 111$ mm CL) | CV | Legal Size Male <br> Biomass ( $\geq 136$ mm CL) | CV |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{gathered} \text { MMBeq }=17,730 \\ M M B_{35 \%}=5,104 \end{gathered}$ |  |  |  |
| 1985 | 3.89 |  |  | 8,932 | 0.09 |
| 1986 | 3.57 | 10,650 | 0.05 | 8,419 | 0.07 |
| 1987 | 2.65 | 8,254 | 0.05 | 5,995 | 0.06 |
| 1988 | 1.80 | 7,644 | 0.04 | 5,650 | 0.04 |
| 1989 | 2.36 | 6,540 | 0.04 | 5,019 | 0.04 |
| 1990 | 1.84 | 4,474 | 0.04 | 3,175 | 0.05 |
| 1991 | 1.65 | 4,091 | 0.05 | 2,841 | 0.05 |
| 1992 | 2.08 | 3,828 | 0.05 | 2,725 | 0.05 |
| 1993 | 1.56 | 3,985 | 0.04 | 2,857 | 0.05 |
| 1994 | 1.97 | 4,575 | 0.03 | 3,451 | 0.03 |
| 1995 | 1.87 | 3,879 | 0.03 | 2,792 | 0.03 |
| 1996 | 1.73 | 3,885 | 0.03 | 2,745 | 0.03 |
| 1997 | 1.85 | 3,895 | 0.04 | 2,787 | 0.04 |
| 1998 | 1.91 | 3,974 | 0.04 | 2,874 | 0.04 |
| 1999 | 2.25 | 4,301 | 0.04 | 3,158 | 0.04 |
| 2000 | 2.51 | 4,346 | 0.04 | 3,107 | 0.04 |
| 2001 | 2.55 | 4,519 | 0.04 | 3,126 | 0.04 |
| 2002 | 2.48 | 4,971 | 0.05 | 3,463 | 0.05 |
| 2003 | 1.76 | 5,525 | 0.05 | 3,985 | 0.05 |
| 2004 | 2.29 | 5,840 | 0.06 | 4,468 | 0.06 |
| 2005 | 2.33 | 5,937 | 0.06 | 4,645 | 0.06 |
| 2006 | 2.42 | 6,235 | 0.06 | 4,818 | 0.07 |
| 2007 | 1.70 | 6,758 | 0.06 | 5,264 | 0.06 |
| 2008 | 1.47 | 6,918 | 0.05 | 5,551 | 0.06 |
| 2009 | 1.85 | 6,698 | 0.05 | 5,581 | 0.05 |
| 2010 | 1.58 | 6,282 | 0.05 | 5,201 | 0.05 |
| 2011 | 1.13 | 5,965 | 0.05 | 4,867 | 0.05 |
| 2012 | 1.78 | 5,444 | 0.05 | 4,503 | 0.05 |
| 2013 | 2.16 | 4,817 | 0.06 | 3,876 | 0.06 |
| 2014 | 1.48 | 4,546 | 0.08 | 3,377 | 0.08 |
| 2015 | 3.43 | 4,547 | 0.12 | 3,382 | 0.11 |
| 2016 | 2.19 | 4,921 | 0.17 | 3,455 | 0.16 |
| 2017 | 2.05 | 5,884 | 0.22 |  |  |

Table 34. Annual abundance estimates of model recruits (millions of crabs), legal male biomass ( t ) with coefficient of variations (CV), and mature male biomass ( t ) CV for scenario 17_0c for golden king crab in the WAG. Legal male biomass was estimated on July 1 (start of fishing year) of fishing year y. Mature male biomass for fishing year y was estimated on February 15 of year $y+1$, after the year y fishery total catch removal. Recruits estimates for 1961 to 2017 are restricted to 1985-2017. Equilibrium $\mathrm{MMB}_{\text {eq }}$ and $\mathrm{MMB}_{35 \%}$ are also listed.

| Year | Recruits to the Model ( $\geq$ 101 mm CL) | Mature Male Biomass ( $\geq 111$ mm CL) | CV | Legal Size Male <br> Biomass ( $\geq 136$ mm CL) | CV |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{aligned} & \mathrm{MMB}_{\mathrm{eq}}=17,720 \\ & M M B_{35 \%}=5,123 \end{aligned}$ |  |  |  |
| 1985 | 3.03 |  |  | 8,932 | 0.09 |
| 1986 | 3.64 | 10,650 | 0.05 | 8,419 | 0.07 |
| 1987 | 2.56 | 8,254 | 0.05 | 5,995 | 0.06 |
| 1988 | 1.87 | 7,644 | 0.04 | 5,650 | 0.04 |
| 1989 | 2.59 | 6,540 | 0.04 | 5,019 | 0.04 |
| 1990 | 1.87 | 4,474 | 0.04 | 3,175 | 0.05 |
| 1991 | 1.57 | 4,091 | 0.05 | 2,841 | 0.05 |
| 1992 | 1.86 | 3,828 | 0.05 | 2,725 | 0.05 |
| 1993 | 1.57 | 3,985 | 0.04 | 2,857 | 0.05 |
| 1994 | 1.97 | 4,575 | 0.03 | 3,451 | 0.03 |
| 1995 | 1.85 | 3,879 | 0.03 | 2,792 | 0.03 |
| 1996 | 1.71 | 3,885 | 0.03 | 2,745 | 0.03 |
| 1997 | 1.87 | 3,895 | 0.04 | 2,787 | 0.04 |
| 1998 | 1.89 | 3,974 | 0.04 | 2,874 | 0.04 |
| 1999 | 2.23 | 4,301 | 0.04 | 3,158 | 0.04 |
| 2000 | 2.48 | 4,346 | 0.04 | 3,107 | 0.04 |
| 2001 | 2.52 | 4,519 | 0.04 | 3,126 | 0.04 |
| 2002 | 2.45 | 4,971 | 0.05 | 3,463 | 0.05 |
| 2003 | 1.75 | 5,525 | 0.05 | 3,985 | 0.05 |
| 2004 | 2.32 | 5,840 | 0.06 | 4,468 | 0.06 |
| 2005 | 2.40 | 5,937 | 0.06 | 4,645 | 0.06 |
| 2006 | 2.37 | 6,235 | 0.06 | 4,818 | 0.07 |
| 2007 | 1.71 | 6,758 | 0.06 | 5,264 | 0.06 |
| 2008 | 1.49 | 6,918 | 0.05 | 5,551 | 0.06 |
| 2009 | 1.84 | 6,698 | 0.05 | 5,581 | 0.05 |
| 2010 | 1.61 | 6,282 | 0.05 | 5,201 | 0.05 |
| 2011 | 1.18 | 5,965 | 0.05 | 4,867 | 0.05 |
| 2012 | 1.80 | 5,444 | 0.05 | 4,503 | 0.05 |
| 2013 | 2.20 | 4,817 | 0.06 | 3,876 | 0.06 |
| 2014 | 1.55 | 4,546 | 0.08 | 3,377 | 0.08 |
| 2015 | 3.60 | 4,547 | 0.12 | 3,382 | 0.11 |
| 2016 | 2.23 | 4,921 | 0.17 | 3,455 | 0.16 |
| 2017 | 2.08 | 5,884 | 0.22 |  |  |

Table 35. Annual abundance estimates of model recruits (millions of crabs), legal male biomass ( t ) with coefficient of variations (CV), and mature male biomass ( t ) with CV for scenario 17_0d for golden king crab in the WAG. Legal male biomass was estimated on July 1 (start of fishing year) of fishing year y. Mature male biomass for fishing year y was estimated on February 15 of year $y+1$, after the year y fishery total catch removal. Recruits estimates for 1961 to 2017 are restricted to 1985-2017. Equilibrium $\mathrm{MMB}_{\text {eq }}$ and $\mathrm{MMB}_{35 \%}$ are also listed.

| Year | Recruits to the Model ( $\geq$ 101 mm CL) | Mature Male Biomass ( $\geq 111$ mm CL) | CV | Legal Size Male <br> Biomass ( $\geq 136$ <br> mm CL) | CV |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{gathered} \text { MMBeq }=17,710 \\ M M B_{35 \%}=5,108 \end{gathered}$ |  |  |  |
| 1985 | 3.68 |  |  | 8,888 | 0.10 |
| 1986 | 3.43 | 10,707 | 0.05 | 8,462 | 0.07 |
| 1987 | 2.68 | 8,202 | 0.05 | 6,010 | 0.06 |
| 1988 | 1.91 | 7,530 | 0.04 | 5,574 | 0.05 |
| 1989 | 2.56 | 6,457 | 0.04 | 4,911 | 0.04 |
| 1990 | 1.85 | 4,489 | 0.04 | 3,116 | 0.06 |
| 1991 | 1.56 | 4,198 | 0.04 | 2,882 | 0.05 |
| 1992 | 2.05 | 3,934 | 0.05 | 2,826 | 0.05 |
| 1993 | 1.58 | 4,047 | 0.04 | 2,941 | 0.05 |
| 1994 | 1.97 | 4,619 | 0.03 | 3,501 | 0.03 |
| 1995 | 1.89 | 3,920 | 0.03 | 2,828 | 0.03 |
| 1996 | 1.74 | 3,922 | 0.04 | 2,774 | 0.04 |
| 1997 | 1.86 | 3,936 | 0.04 | 2,818 | 0.04 |
| 1998 | 1.91 | 4,012 | 0.04 | 2,906 | 0.04 |
| 1999 | 2.26 | 4,336 | 0.04 | 3,188 | 0.04 |
| 2000 | 2.54 | 4,382 | 0.04 | 3,135 | 0.04 |
| 2001 | 2.62 | 4,564 | 0.04 | 3,156 | 0.04 |
| 2002 | 2.60 | 5,042 | 0.05 | 3,506 | 0.05 |
| 2003 | 1.83 | 5,654 | 0.05 | 4,061 | 0.05 |
| 2004 | 2.30 | 6,040 | 0.06 | 4,608 | 0.06 |
| 2005 | 2.21 | 6,171 | 0.06 | 4,843 | 0.06 |
| 2006 | 2.40 | 6,428 | 0.06 | 5,029 | 0.06 |
| 2007 | 1.64 | 6,871 | 0.05 | 5,411 | 0.06 |
| 2008 | 1.39 | 6,979 | 0.05 | 5,634 | 0.05 |
| 2009 | 1.71 | 6,688 | 0.05 | 5,606 | 0.05 |
| 2010 | 1.37 | 6,176 | 0.04 | 5,154 | 0.05 |
| 2011 | 1.08 | 5,723 | 0.04 | 4,720 | 0.05 |
| 2012 | 1.80 | 5,082 | 0.05 | 4,224 | 0.05 |
| 2013 | 2.09 | 4,424 | 0.06 | 3,513 | 0.06 |
| 2014 | 1.57 | 4,169 | 0.08 | 3,014 | 0.08 |
| 2015 | 4.05 | 4,211 | 0.11 | 3,039 | 0.11 |
| 2016 | 2.26 | 4,829 | 0.18 | 3,195 | 0.16 |
| 2017 | 2.05 | 6,136 | 0.23 |  |  |

Table 36. Annual abundance estimates of model recruits (millions of crabs), legal male biomass ( $t$ ) with coefficient of variations (CV), and mature male biomass ( t ) with CV for scenario 17_0e for golden king crab in the WAG. Legal male biomass was estimated on July 1 (start of fishing year) of fishing year y. Mature male biomass for fishing year y was estimated on February 15 of year $y+1$, after the year y fishery total catch removal. Recruits estimates for 1961 to 2017 are restricted to 1985-2017. Equilibrium $\mathrm{MMB}_{\text {eq }}$ and $\mathrm{MMB}_{35 \%}$ are also listed.

| Year | Recruits to the Model ( $\geq$ 101 mm CL) | Mature Male Biomass $(\geq 111 \mathrm{~mm} \mathrm{CL})$ | CV | Legal Size Male Biomass ( $\geq \mathbf{1 3 6}$ mm CL) | CV |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{gathered} \text { MMBeq }=18,001 \\ M M B_{35 \%}=5,201 \end{gathered}$ |  |  |  |
| 1985 | 3.33 |  |  | 9,215 | 0.08 |
| 1986 | 3.56 | 10,884 | 0.04 | 8,762 | 0.06 |
| 1987 | 2.62 | 8,250 | 0.04 | 6,106 | 0.04 |
| 1988 | 1.91 | 7,581 | 0.04 | 5,590 | 0.04 |
| 1989 | 2.68 | 6,476 | 0.04 | 4,903 | 0.04 |
| 1990 | 1.89 | 4,534 | 0.04 | 3,106 | 0.05 |
| 1991 | 1.54 | 4,296 | 0.04 | 2,908 | 0.05 |
| 1992 | 2.00 | 4,046 | 0.04 | 2,895 | 0.05 |
| 1993 | 1.54 | 4,139 | 0.04 | 3,022 | 0.04 |
| 1994 | 1.90 | 4,675 | 0.03 | 3,558 | 0.03 |
| 1995 | 1.86 | 3,931 | 0.03 | 2,848 | 0.03 |
| 1996 | 1.86 | 3,884 | 0.03 | 2,743 | 0.03 |
| 1997 | 1.77 | 3,906 | 0.04 | 2,753 | 0.03 |
| 1998 | 1.88 | 4,003 | 0.03 | 2,865 | 0.03 |
| 1999 | 2.20 | 4,285 | 0.03 | 3,140 | 0.03 |
| 2000 | 2.51 | 4,297 | 0.04 | 3,057 | 0.04 |
| 2001 | 2.67 | 4,437 | 0.04 | 3,035 | 0.04 |
| 2002 | 2.76 | 4,905 | 0.05 | 3,347 | 0.05 |
| 2003 | 1.95 | 5,573 | 0.05 | 3,907 | 0.05 |
| 2004 | 2.34 | 6,071 | 0.05 | 4,531 | 0.06 |
| 2005 | 2.25 | 6,289 | 0.05 | 4,875 | 0.06 |
| 2006 | 2.30 | 6,598 | 0.05 | 5,137 | 0.06 |
| 2007 | 1.65 | 7,039 | 0.05 | 5,561 | 0.05 |
| 2008 | 1.44 | 7,104 | 0.05 | 5,758 | 0.05 |
| 2009 | 1.86 | 6,820 | 0.04 | 5,710 | 0.05 |
| 2010 | 1.66 | 6,363 | 0.04 | 5,277 | 0.05 |
| 2011 | 1.02 | 6,048 | 0.04 | 4,914 | 0.04 |
| 2012 | 1.90 | 5,522 | 0.04 | 4,563 | 0.04 |
| 2013 | 2.48 | 4,868 | 0.05 | 3,910 | 0.05 |
| 2014 | 1.58 | 4,714 | 0.06 | 3,426 | 0.06 |
| 2015 | 3.58 | 4,878 | 0.09 | 3,561 | 0.08 |
| 2016 | 2.21 | 5,336 | 0.13 | 3,756 | 0.12 |
| 2017 | 2.05 | 6,355 | 0.17 |  |  |

Table 37. Negative log-likelihood values of the fits for scenarios (Sc) 17_0 (base), 17_0a (observer CPUE by VAST), 17_0b (observer and fishtick CPUE variable selection by CAIC), 17_0c (Year:Area interaction for observer and fishtick CPUE), 17_0d (three total selectivity and catchability for 1985-04, 2005-12, and 2013-16 time periods), and 17_0e (Stage 2 effective sample sizes by McAllister and Ianelli method) for golden king crab in the WAG. Differences in likelihood values are given for scenarios with the same number of data points (base) and free parameters. Likelihood components with zero entry in the entire rows are omitted. RetdcatchB= retained catch biomass.

| Likelihood Component | Sc 17_0 | $\begin{gathered} \text { Sc } \\ \text { 17_0a } \end{gathered}$ | $\begin{gathered} \text { Sc } \\ \text { 17_0b } \end{gathered}$ | $\begin{gathered} \text { Sc } \\ \text { 17_0c } \end{gathered}$ | $\begin{gathered} \text { Sc } \\ \text { 17_0d } \end{gathered}$ | $\begin{gathered} \text { Sc } \\ \text { 17_0e } \end{gathered}$ | $\begin{gathered} \text { Sc17_0a- } \\ \text { Sc 17_0 } \end{gathered}$ | $\begin{gathered} \hline \text { Sc 17_0b - } \\ \text { Sc 17_0 } \end{gathered}$ | $\begin{gathered} \hline \text { Sc 17_0c - } \\ \text { Sc 17_0 } \end{gathered}$ | $\begin{gathered} \hline \text { Sc 17_0e - } \\ \text { Sc 17_0 } \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Number of free parameters | 140 | 140 | 140 | 140 | 143 | 140 |  |  |  |  |
| Data | Base | Base | Base | Base | Base | Base |  |  |  |  |
| Retlencomp | -1146.700 | -1147.140 | -1143.350 | -1142.310 | -1161.250 | -1243.980 | -0.440 | 3.350 | 4.390 | -97.280 |
| Totallencomp | -1389.720 | -1389.680 | -1395.850 | -1396.210 | -1396.220 | -1370.230 | 0.040 | -6.130 | -6.490 | 19.490 |
| Observer cpue | -11.773 | -14.747 | -0.680 | 15.078 | -10.040 | -11.199 | -2.974 | 11.093 | 26.851 | 0.574 |
| RetdcatchB | 4.721 | 4.854 | 4.853 | 5.858 | $4.846$ | $4.956$ | 0.133 | 0.132 | 1.137 | 0.235 |
| TotalcatchB | 43.783 | 43.745 | 43.936 | 44.348 | 43.849 | 47.086 | -0.038 | 0.153 | 0.565 | 3.303 |
| GdiscdcatchB | 0.001 | 0.001 | 0.001 | 0.001 | 0.001 | 0.001 | 0.000 | 0.000 | 0.000 | 0.000 |
| Rec_dev | 5.243 | 5.248 | 5.254 | 4.797 | 6.091 | 6.103 | 0.005 | 0.011 | -0.446 | 0.860 |
| Pot F_dev | 0.026 | 0.026 | 0.026 | 0.027 | 0.027 | 0.026 | 0.000 | 0.000 | 0.001 | 0.000 |
| Gbyc_F_dev | 0.037 | 0.037 | 0.037 | 0.037 | 0.038 | 0.037 | 0.000 | 0.000 | 0.000 | 0.000 |
| Tag | 2693.630 | 2693.450 | 2693.710 | 2693.780 | 2693.910 | 2695.840 | -0.180 | 0.080 | 0.150 | 2.210 |
| Fishery cpue | -5.155 | -5.207 | -9.456 | 17.685 | -5.004 | -2.783 | -0.052 | -4.301 | 22.840 | 2.371 |
| RetcatchN | 0.002129 | 0.002068 | 0.001757 | 0.000874 | 0.002098 | 0.005553 | -0.000061 | -0.000372 | -0.001255 | 0.003424 |
| Total | 194.090 | 190.591 | 198.490 | 243.086 | 176.255 | 125.863 | -3.499 | 4.400 | 48.996 | -68.227 |

Table 38. Predicted total catch OFL ( t ), $M M B_{35 \%}$, and terminal MMB ratio for various scenarios for EAG and WAG, respectively. Sc $=$ scenario; $\mathrm{MMB}_{2016} / \mathrm{MMB}_{\text {initial }}=$ ratio of terminal MMB relative to initial MMB $\left(=\right.$ MMB $\left._{1960}\right)$. Note: $\mathrm{MMB}_{2016}$ is estimated on Feb 15, 2017.

|  | EAG |  |  | WAG |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Sc | $\begin{gathered} \text { Tier } 3 \\ \text { Total } \\ \text { Catch } \\ \text { OFL }(\mathbf{t}) \end{gathered}$ | $\text { MMB }_{35 \%}$ <br> (t) | $\begin{gathered} \text { MMB }_{2016} \\ / \\ \text { MMB }_{\text {initial }} \end{gathered}$ | Tier 3 <br> Total Catch OFL (t) | $\text { MMB }_{35 \%}$ <br> (t) | $\begin{aligned} & \hline \text { MMB }_{2016} / \\ & \text { MMB }_{\text {initial }} \end{aligned}$ | M $\mathbf{y r}^{-1}$ | Remarks |
| 17_0 | 3,918 | 6,954 | 0.68 | 1,597 | 5,138 | 0.42 | 0.21 | Base scenario: 1960 equilibrium initial size composition, 1995/96-2016/17 observer CPUE, 1985/86-1998/99 <br> Fishery CPUE, time period for mean R calculation for equilibrium initial abundance and $M M B_{M S Y}$ reference point calculations 1987-2012, knife-edge maturity $\geq 111 \mathrm{~mm}$ CL, Francis re-weighting, |
| 17_0a | 3,959 | 7,063 | 0.67 | 1,589 | 5,173 | 0.42 | 0.21 | Observer CPUE standardization by VAST |
| 17_0b | 3,415 | 6,794 | 0.61 | 1,492 | 5,104 | 0.40 | 0.21 | Variable selection for CPUE standardization by CAIC |
| 17_0c | 4,046 | 7,091 | 0.67 | 1,551 | 5,123 | 0.40 | 0.21 | Year:Area interaction for CPUE standardization Three catchability and asymptotic total selectivity for |
| 17_0d | 2,481 | 6,688 | 0.46 | 1,482 | 5,108 | 0.42 | 0.21 | 1985/86-2004/05, 2005/06-2012/13, and 2013/14-2016/17 |
| 17_0e | 3,974 | 7,014 | 0.67 | 1,608 | 5,201 | 0.43 | 0.21 | McAllister and Ianelli method of re-weighting EAG fishery independent pot survey (2015/16-2016/17) |
| 17_0f <br> May | 3,892 | 6,946 | 0.67 |  |  |  | 0.21 | CPUE indices as an additional likelihood component. |
| 2017 Sc9 | 4,486 | 7,048 | 0.60 | 1,562 | 4,507 | 0.34 | 0.224 | 2017 assessment. Knife-edge maturity $\geq 111 \mathrm{~mm}$ CL |



Figure 1. Total and components negative log-likelihoods vs. $M$ for scenario $\mathbf{0 b}$ model fit for EAG and WAG combined data. The $M$ estimate was obtained without any $M$ penalty. The $M$ estimate was $0.2254 \mathrm{yr}^{-1}\left( \pm 0.0199 \mathrm{yr}^{-1}\right)$. The negative log likelihood values were estimated for fixed proportions of estimated $M$ without using an $M$ penalty and they were zero adjusted. The $M$ profile indicates an $M$ value of $0.2142 \mathrm{yr}^{-1}$ at the minima of negative total likelihood for combined data as well as individual date sets. Hence an $M$ value of $0.21 \mathrm{yr}^{-1}$ was used in all scenarios.


Figure 2. Aleutian Islands, Area O, red and golden king crab management area (from Leon et al. 2017).


Figure 3. Adak (Area R) and Dutch Harbor (Area O) king crab registration area and districts, 1984/85-1995/96 seasons (Leon et al., 2017).


Figure 4. Percent of total 1981/82-1995/96 golden king crab retained catch weight (harvest) from one-degree longitude intervals in the Aleutian Islands, with dotted line denoting the border at $171^{\circ} \mathrm{W}$ longitude used during the 1984/85-1995/96 seasons to divide fishery management between the Dutch Harbor Area (east of $171^{\circ} \mathrm{W}$ longitude) and the Adak Area (west of $171^{\circ} \mathrm{W}$ longitude) and solid line denoting the border at $174^{\circ} \mathrm{W}$ longitude used since the 1996/97 season to manage crab east and west of $174^{\circ} \mathrm{W}$ longitude (adapted from Figure 4-2 in Morrison et al. 1998).


Figure 5. Retained catch ( t ) of golden king crab within one-degree longitude intervals in the Aleutian Islands during the 2000/01 through 2016/17commercial fishery seasons; solid line denotes the border at $174^{\circ} \mathrm{W}$ longitude that has been used since the 1996/97 season to manage Aleutian Island golden king crab as separate stocks east and west of $174^{\circ} \mathrm{W}$ longitude and dashed line denotes the border at $171^{\circ} \mathrm{W}$ longitude used during the 1984/85-1995/96 seasons to divide fishery management between the Dutch Harbor Area (east of $171^{\circ} \mathrm{W}$ longitude) and the Adak Area (west of $171^{\circ} \mathrm{W}$ longitude).


Figure 6. Average golden king crab CPUE ( $\mathrm{kg} / \mathrm{nm} 2$ ) for tows, number of tows, and average depth of tows from one-degree longitude intervals during the 2002, 2004, 2006, 2010, and 2012 NMFS Aleutian Islands bottom trawl surveys; preliminary summary of data obtained on 1 April 2013 from http://www.afsc.noaa.gov/RACE/groundfish/survey_data/default.htm.


Figure 7. Historical commercial harvest (from fish tickets; metric tons) and catch-per-unit effort (CPUE, number of crabs per pot lift) of golden king crab in the EAG, 1985/86-2016/17 fisheries (note: 1985 refers to the 1985/86 fishing year).


Figure 8. Historical commercial harvest (from fish tickets; metric tons) and catch-per-unit effort (CPUE, number of crabs per pot lift) of golden king crab in the WAG, 1985/86-2016/17 fisheries (note: 1985 refers to the 1985/86 fishing year).


Figure 9. Catch distribution by statistical area.in 2016/17.


Figure 10. Standard deviation of recruit_dev plot for EAG and WAG. The mean recruit for years with standard deviation less than 0.7 sigma R was used to initialize model. We selected the 1987-2012 period for mean recruit estimation.


Figure 11. Predicted (line) vs. observed (bar) retained catch relative length frequency distributions under scenarios 17_0 (black line), 17_0a (orange line), 17_0b (red line), 17_0c (blue line), 17_0d (violet line), 17_0e (dark green line), and 17_0f (green line) for golden king crab in the EAG, 1985/86 to 2016/17. This color scheme is used in all other graphs.


Figure 12. Predicted (line) vs. observed (bar) total catch relative length frequency distributions under scenarios 17_0 to 17_0f for golden king crab in the EAG, 1990/91 to 2016/17.


Figure 13. Predicted (line) vs. observed (bar) groundfish (or trawl) discarded bycatch relative length frequency distributions under scenarios $17 \_0$ to $17 \_0 f$ for golden king crab in the EAG, 1989/90 to 2016/17. Note that this data set was not used in the model fitting.






Figure 14. Estimated total (black solid line) and retained selectivity (red dotted line) for pre- and post- rationalization periods under scenarios 17_0 to May 2017 Sc9 model fits to golden king crab data in the EAG.


Figure 15. Observed (open circles) vs. predicted (solid line) tag recaptures by size bin for years 1 to 6 recaptures under scenario 17_0 for EAG golden king crab.


Figure 16. Estimated number of male recruits (crab size $\geq 101 \mathrm{~mm} \mathrm{CL}$ ) to the assessment model under scenarios (Sc) 17_0 to May 2017 Sc9 for EAG golden king crab data, 1961-2017. Top left: scenarios 17_0 and 17_0a; top right: scenarios 17_0, 17_0b, and 17_0c; bottom left: scenarios 17_0, 17_0d, and 17_0e; and bottom right: scenarios 17_0, 17_0f, and May 2017 Sc9. This grouping scheme was used in a number of subsequent figures. The number of recruits are centralized using (R-mean R )/mean R for comparing different scenarios' results.


Figure 17. Recruit size distribution to the assessment model under scenarios (Sc) 17_0 to May 2017 Sc9 for EAG golden king crab.


Figure 18. Estimated molt probability vs. carapace length of golden king crab for scenarios 17_0 to May 2017 Sc9 in the EAG.



Figure 19. Observed (open circle) vs. predicted (solid line) retained catch (top left in each scenario set), total catch (top right in each scenario set), and groundfish bycatch (bottom left in each scenario set) of golden king crab for scenarios 17_0 to May 2017 Sc9, in EAG, 1981/822016/17.


Figure 20. Observed (open circle) vs. predicted (solid line) retained catch of golden king crab for scenarios 17_0 to May 2017 Sc9 fits in the EAG, 1981/82-1984/85. Note: Input retained catches to the model during pre-1985 fishery period were in number of crabs.

EAG 17_0 Retained Catch Size Composition Standardized Residuals


Figure 21. Bubble plot of standardized residuals of retained catch length composition for scenario 17_0 fit for EAG golden king crab, 1985/86-2016/17. Green circles are the positive and pink circles are the negative standardized residuals. The area of the circle is the relative magnitude of the residual.

EAG 17_0 Total Catch Size Composition Standardized Residuals


Figure 22. Bubble plot of standardized residuals of total catch length composition for scenario $17 \_0$ fit for EAG golden king crab, 1990/91-2016/17. Green circles are the positive and pink circles are the negative standardized residuals. The area of the circle is the relative magnitude of the residual.


Figure 23. Bubble plot of standardized residuals of retained catch length composition for scenario 17_0d fit for EAG golden king crab, 1985/86-2015/16. Blue circles are the positive and
pink circles are the negative standardized residuals. The area of the circle is the relative magnitude of the residual.


Figure 24. Bubble plot of standardized residuals of total catch length composition for scenario 9 fit for EAG golden king crab, 1990/91-2015/16. Blue circles are the positive and pink circles are the negative standardized residuals. The area of the circle is the relative magnitude of the residual.


Figure 25. Retrospective fits of MMB by the model following removal of terminal year data under scenarios 17_0 (top) and 17_0d (bottom) for golden king crab in the EAG, 1960/612016/17.

Mohn rho ( $\rho$ ) formula (modified by Deroba, 2014) is as follows:

$$
\text { Mohn } \rho=\frac{\sum_{n=1}^{x} \frac{\left[\widehat{M M B}_{y=T-n, T-n}-\widehat{M M B}_{y=T-n, T}\right]}{\widehat{M M B}_{y=T-n, T}}}{x}
$$

where, $\widehat{M M B}_{y=T-n, T-n}$ is the MMB estimated for year T-n (left subscript) using data up to T-n years (right subscript), T is the terminal year of the entire data, x is the total number of peels, most recent year's data is "peeled off" recursively n times, where $\mathrm{n}=1,2,3 \ldots$. x .
We used four peels ( $x=4$ ) and our $T=2016$.


Figure 26. Comparison of input CPUE indices (open circles with $+/-2$ SE) with predicted CPUE indices (colored solid lines) under scenarios 17_0 to May 2017 Sc9 for EAG golden king crab data, 1985/86-2016/17. Model estimated additional standard error was added to each input standard error.


Figure 27. Trends in pot fishery full selection total fishing mortality of golden king crab for scenarios 17_0 to May 2017 Sc9 model fits in the EAG, 1981/82-2016/17.


Figure 28. Trends in golden king crab mature male biomass for scenarios $17 \_0$ to May 2017 Sc9 fits in the EAG, 1960/61-2016/17. Scenario 17_0 estimates have two standard errors confidence limits.


Figure 29. Predicted (line) vs. observed (bar) retained catch relative length frequency distributions under scenarios 17_0 (black line), 17_0a (orange line), 17_0b (red line), 17_0c (blue line), 17_0d (violet line), and 17_0e (dark green line) for golden king crab in the WAG, 1985/86 to 2016/17. This color scheme is used in all other graphs.


Figure 30. Predicted (line) vs. observed (bar) total catch relative length frequency distributions under scenarios 17_0 to 17_0e for golden king crab in the WAG, 1990/91 to 2016/17.


Figure 31. Predicted (line) vs. observed (bar) groundfish (or trawl) discarded bycatch relative length frequency distributions under scenarios $17 \_0$ to $17 \_0 \mathrm{e}$ for golden king crab in the WAG, 1989/90 to 2016/17. Note that this data set was not used in the model fitting.



Figure 32. Estimated total (black solid line) and retained selectivity (red dotted line) for pre- and post- rationalization periods under scenarios 17_0 to May 2017 Sc9 fits to golden king crab data in the WAG.


Figure 33. Observed (open circles) vs. predicted (solid line) tag recaptures by size bin for years 1 to 6 recaptures under scenario 17_0 for WAG golden king crab.


Figure 34. Estimated number of male recruits (crab size $\geq 101 \mathrm{~mm} \mathrm{CL}$ ) to the assessment model under scenarios 17_0 to May 2017 Sc9 for WAG golden king crab data, 1961-2017. Top left: scenarios 17_0 and 17_0a; top right: scenarios 17_0, 17_0b, and 17_0c; and bottom left: scenarios 17_0, 17_0d, and 17_0e and May 2017 Sc9. The number of recruits are centralized using ( R -mean R )/mean R for comparing different scenarios' results.


Figure 35. Recruit size distribution to the assessment model under scenarios (Sc) 17_0 to May 2017 Sc9 for WAG golden king crab.


Figure 36. Estimated molt probability vs. carapace length of golden king crab for scenarios 17_0 to May 2017 Sc9 in the WAG.



Figure 37. Observed (open circle) vs. predicted (solid line) retained catch (top left in each scenario set), total catch (top right in each scenario set), and groundfish bycatch (bottom left in each scenario set) of golden king crab for scenarios 17_0 to May 2017 Sc9 fits in the WAG, 1981/82-2016/17.


Figure 38. Observed (open circle) vs. predicted (solid line) retained catch of golden king crab for scenarios 17_0 to May 2017 Sc9 fits in the WAG, 1981/82-1984/85. Note: Input retained catches to the model during pre-1985 fishery period were in number of crabs.

WAG 17_0 Retained Catch Size Composition Standardized Residuals


Figure 39. Bubble plot of standardized residuals of retained catch length composition for scenario 17_0 fit for WAG golden king crab, 1985/86-2016/17. Green circles are the positive and pink circles are the negative standardized residuals. The area of the circle is the relative magnitude of the residual.

WAG 17_0 Total Catch Size Composition Standardized Residuals


Figure 40. Bubble plot of standardized residuals of total catch length composition for scenario 17_0 fit for WAG golden king crab, 1990/91-2016/17. Green circles are the positive and pink circles are the negative standardized residuals. The area of the circle is the relative magnitude of the residual.


Figure 41. Bubble plot of standardized residuals of retained catch length composition for scenario 17_0d fit for WAG golden king crab, 1985/86-2016/17. Blue circles are the positive and pink circles are the negative standardized residuals. The area of the circle is the relative magnitude of the residual.

WAG 17_0d Total Catch Size Composition Standardized Residuals


Figure 42. Bubble plot of standardized residuals of total catch length composition for scenario 17_0d fit for WAG golden king crab, 1990/91-2016/17. Blue circles are the positive and pink circles are the negative standardized residuals. The area of the circle is the relative magnitude of the residual.


Figure 43. Retrospective fits of MMB by the model following removal of terminal year data under scenarios 17_0 (top) and 17_0d (bottom) for golden king crab in the WAG, 1960/612016/17.


Figure 44. Comparison of input CPUE indices (open circles with +/- 2 SE) with predicted CPUE indices (colored solid lines) under scenarios 17_0 to May 2017 Sc9 for WAG golden king crab data, 1985/86-2016/17. Model estimated additional standard error was added to each input standard error.


Figure 45. Trends in pot fishery full selection total fishing mortality of golden king crab for scenarios 17_0 to May 2017 Sc9 model fits in the WAG, 1981/82-2016/17.


Figure 46. Trends in golden king crab mature male biomass for scenarios 17_0 to May 2017 Sc9 model fits in the WAG, 1960/61-2016/17. Scenario 17_0 estimates have two standard errors confidence limits.


Figure 47. Relationships between full fishing mortalities for the directed pot fishery and mature male biomass on Feb. 15 during 1985/86-2016/17 under scenarios 17_0 and 17_0d for EAG and WAG. Average recruitment from 1987 to 2012 was used to estimate $M_{3}{ }_{35 \%}$. Pot and groundfish handling mortality rates were assumed to be 0.2 and 0.65 , respectively.

## Appendix A: Integrated model

Aleutian Islands Golden King Crab (Lithodes aequispinus) Stock Assessment Model Development- east of $174^{\circ} \mathrm{W}$ (EAG) and west of $174^{\circ} \mathrm{W}$ (WAG) Aleutian Island stocks

## Basic population dynamics

The annual [male] abundances by size are modeled using the equation:
$N_{t+1, j}=\sum_{i=1}^{j}\left[N_{t, i} e^{-M}-\left(\hat{C}_{t, i}+\widehat{D}_{t, i}+\widehat{\operatorname{Tr}}_{t, i}\right) e^{\left(y_{t}-1\right) M}\right] X_{i, j}+R_{t+1, j}$
where $N_{t, i}$ is the number of [male] crab in length class i on 1 July (start of fishing year) of year $\mathrm{t} ; \hat{C}_{t, i}, \hat{D}_{t, i}$, and $\hat{T} r_{t, i}$ are respectively the predicted fishery retained, pot fishery discard dead, and groundfish fishery discard dead catches in length class $i$ during year $t ; \widehat{D}_{t, i}$ is estimated from the intermediate total ( $\hat{T}_{t, i \text { temp }}$ ) catch and the retained ( $\hat{C}_{t, i}$ ) catch by Equation A.2c. ${ }^{X_{i, j}}$ is the probability of length-class $i$ growing into length-class $j$ during the year; $y_{t}$ is elapsed time period from 1 July to the mid -point of fishing period in year $t ; M$ is instantaneous rate of natural mortality; and $R_{t+1, j}$ recruitment to length class $j$ in year $t+1$.

The catches are predicted using the equations
$\hat{T}_{t, j, t e m p}=\frac{F_{t} s_{t, j}^{T}}{Z_{t, j}} N_{t, j} e^{-y_{t} M}\left(1-e^{-Z_{t, j}}\right)$
$\hat{C}_{t, j}=\frac{F_{t} s_{t, j}^{T} s_{t, j}^{r}}{z_{t, j}} N_{t, j} e^{-y_{t} M}\left(1-e^{-Z_{t, j}}\right)$
$\widehat{D}_{t, j}=0.2\left(\widehat{T}_{t, j, t e m p}-\hat{C}_{t, j}\right)$
$\widehat{T r}_{t, j}=0.65 \frac{F_{t}^{T r} s_{j}^{T r}}{z_{t, j}} N_{t, j} e^{-y_{t} M}\left(1-e^{-Z_{t, j}}\right)$
$\widehat{T}_{t, j}=\hat{C}_{t, j}+\widehat{D}_{t, j}$
where $Z_{t, j}$ is total fishery-related mortality on animals in length-class $j$ during year $t$ :

$$
\begin{equation*}
Z_{t, j}=F_{t} s_{t, j}^{T} s_{t, j}^{r}+0.2 F_{t} s_{t, j}^{T}\left(1-s_{t, j}^{r}\right)+0.65 F_{t}^{T r} s_{j}^{T r} \tag{A.3}
\end{equation*}
$$

$F_{t}$ is the full selection fishing mortality in the pot fishery, $F_{t}^{T r}$ is the full selection fishing mortality in the trawl fishery, $s_{t, j}^{T}$ is the total selectivity for animals in length-class $j$ by the pot fishery during year $t, s_{j}^{T r}$ is the selectivity for animals in length-class $j$ by the trawl fishery, $s_{t, j}^{r}$ is the probability of retention for animals in length-class $j$ by the pot fishery during year $t$. Pot
bycatch mortality of 0.2 and groundfish bycatch mortality of 0.65 (average of trawl (0.8) and fish pot ( 0.5 ) mortality) were assumed.

## Initial abundance

The initial conditions are computed as the equilibrium initial condition using the following relations:

The equilibrium stock abundance is
$N=X . S . N+R$
The equilibrium abundance in $1960, N_{1960}$, is
$\underline{N}_{1960}=(I-X S)^{-1} \underline{R}$
where $X$ is the growth matrix, $S$ is a matrix with diagonal elements given by $e^{-M}, I$ is the identity matrix, and $\underline{R}$ is the product of average recruitment and relative proportion of total recruitment to each size-class.

We used the mean number of recruits from 1987 to 2012 in equation (A.5) to obtain the equilibrium solution under only natural mortality in year 1960, and then projected the equilibrium abundance under natural mortality with recruitment estimated for each year after 1960 up to 1985 with removal of retained catches during 1981/82 to 1984/85.

## Growth Matrix

The growth matrix $X$ is modeled as follows:
$X_{i, j}= \begin{cases}0 & \text { if } j<i \\ P_{i, j}+\left(1-m_{i}\right) & \text { if } j=i \\ P_{i, j} & \text { if } j>i\end{cases}$
where:

$$
P_{i, j}=m_{i}\left\{\begin{array}{rr}
\int_{-\infty}^{j_{2}-L_{i}} N\left(x \mid \mu_{i}, \sigma^{2}\right) d x & \text { if } j=i \\
\int_{j_{1}-L_{i}}^{j_{2}-L_{i}} N\left(x \mid \mu_{i}, \sigma^{2}\right) d x & \text { if } i<j<n  \tag{A.7}\\
\int_{j_{1}-L_{i}}^{\infty} N\left(x \mid \mu_{i}, \sigma^{2}\right) d x & \text { if } i=n
\end{array}, \begin{array}{c} 
\\
N\left(x \mid \mu_{i}, \sigma^{2}\right)=\frac{1}{\sqrt{2 \pi \sigma^{2}}} e^{-\left(\frac{x-\mu_{i}}{\sqrt{2} \sigma}\right)^{2}}, \text { and }
\end{array}\right.
$$

$\mu_{i}$ is the mean growth increment for crab in size-class $i$ :
$\mu_{i}=\omega_{1}+\omega_{2} * \bar{L}_{i}$.
$\omega_{1} \quad, \omega_{2}, \quad$ and $\sigma$ are estimable parameters, and $j_{1}$ and $j_{2}$ are the lower and upper limits of the receiving length-class $j$ (in mm CL), and $\bar{L}_{i}$ is the mid-point of the contributing length interval $i$. The quantity $m_{i}$ is the molt probability for size-class $i$ :
$m_{i}=\frac{1}{1+e^{c\left(\tau_{i}-d\right)}}$
where $\tau_{i}$ is the mid-length of the $i$-th length-class, $c$ and $d$ are parameters.

## Selectivity and retention

Selectivity and retention are both assumed to be logistic functions of length. Selectivity depends on the fishing period for the pot fishery:
$S_{i}=\frac{1}{1+e^{\left[-\ln \left(199 \frac{\tau_{i}-\theta_{50}}{\left.\theta_{95}-\theta_{50}\right]}\right.\right.}}$
where $\theta_{95}$ and $\theta_{50}$ are the parameters of the selectivity/ retention pattern (Mark Maunder, unpublished generic crab model). In the program, we re-parameterized the denominator ( $\theta_{95}$ $\left.\theta_{50}\right)$ to $\log (\operatorname{delta\theta })$ so that the difference is always positive and transformed $\theta_{50}$ to $\log \left(\theta_{50}\right)$ to keep the estimate always positive.

Recruitment
Recruitment to length-class i during year $t$ is modeled as $R_{t, i}=\bar{R} e^{\epsilon_{i}} \Omega_{i}$ where $\Omega_{i}$ is a normalized gamma function

$$
\begin{equation*}
\operatorname{gamma}\left(x \mid \alpha_{r}, \beta_{r}\right)=\frac{x^{\alpha_{r}-1} e^{\frac{x}{\beta_{r}}}}{\beta_{r}{ }^{\alpha_{r}} \Gamma_{\left(\alpha_{r}\right)}} \tag{A.10}
\end{equation*}
$$

with $\alpha_{r}$ and $\beta_{r}$ (restricted to the first five length classes).

## Parameter estimation

Table A1 lists the parameters of the model indicating which are estimated and which are prespecified. The objective function includes contributions related to the fit of the model to the available data and penalties (priors on various parameters).

Tables A2 lists parameter values (with the corresponding coefficient of variations in parentheses) used to weight the components of the objective functions for EAG and WAG.

## Likelihood components

## Catches

The contribution of the catch data (retained, total, and groundfish discarded) to the objective function is given by:

$$
\begin{align*}
& L L_{r}^{\text {catch }}=\lambda_{r} \sum_{t}\left\{\ln \left(\sum_{j} \hat{C}_{t, j} w_{j}+c\right)-\ln \left(\sum_{j} C_{t, j} w_{j}+c\right)\right\}^{2}  \tag{A.11a}\\
& L L_{T}^{\text {catch }}=\lambda_{T} \sum_{t}\left\{\ln \left(\sum_{j} \widehat{T}_{t, j} w_{j}+c\right)-\ln \left(\sum_{j} T_{t, j} w_{j}+c\right)\right\}^{2}  \tag{A.11b}\\
& L L_{G D}^{\text {catch }}=\lambda_{G D} \sum_{t}\left\{\ln \left(\sum_{j} \widehat{T r}_{t, j} w_{j}+c\right)-\ln \left(\sum_{j} T r_{t, j} w_{j}+c\right)\right\}^{2} \tag{A.11c}
\end{align*}
$$

where $\lambda_{r}, \lambda_{T}$, and $\lambda_{G D}$ are weights assigned to likelihood components for the retained, pot total, and groundfish discard catches; ${ }_{j}$ is the average mass of a crab is length-class $j ;{ }^{C_{t, j}}, T_{t, j}$, and $T r_{t, j}$ are, respectively, the observed numbers of crab in size class $j$ for retained, pot total, and groundfish fishery discarded crab during year $t$, and $c$ is a small constant value. We assumed $c=$ 0.001 .

An additional retained catch likelihood (using Equation A.11a without w) for the retained catch in number of crabs during 1981/82 to 1984/85 was also considered in all scenarios.

## Catch-rate indices

The catch-rate indices are assumed to be lognormally distributed about the model prediction. Account is taken of variation in addition to that related to sampling variation:
$L L_{r}^{C P U E}=\lambda_{r, C P U E}\left\{0.5 \sum_{t} \ln \left[2 \pi\left(\sigma_{r, t}^{2}+\sigma_{e}^{2}\right)\right]+\sum_{t} \frac{\left(\ln \left(C P U E_{t}^{r}+c\right)-\ln \left(C \widehat{P U E}_{t}^{r}+c\right)\right)^{2}}{2\left(\sigma_{r, t}^{2}+\sigma_{e}^{2}\right)}\right\}$
where ${ }^{C P U E} E_{t}^{r}$ is the standardized retain catch-rate index for year $t,{ }^{\sigma_{r, t}}$ is standard error of the logarithm of $C P U E_{t}^{r}$, and $\widehat{C U E} r t$ is the model-estimate of $C P U E_{t}^{r}$ :

$$
\begin{equation*}
\widehat{C P U E}{ }_{t}^{r}=q_{k} \sum_{j} S_{j}^{T} S_{j}^{r}\left(N_{t, j}-0.5\left[\widehat{C_{t, j}}+\widehat{D_{t, j}}+\widehat{T r_{t, j}}\right]\right) e^{-y_{t} M} \tag{A.13}
\end{equation*}
$$

in which $q_{k}$ is the catchability coefficient during the $k$-th time period (e.g., pre- and postrationalization time periods), $\sigma_{e}$ is the extent of over-dispersion, $c$ is a small constant to prevent zero values (we assumed $c=0.001$ ), and $\lambda_{r, C P U E}$ is the weight assigned to the catch-rate data. We used the same likelihood formula (A.12) for fish ticket retained catch rate indices.

Following Burnham et al. (1987), we computed the $\ln (C P U E)$ variance by:

$$
\begin{equation*}
\sigma_{\mathrm{r}, \mathrm{t}}^{2}=\ln \left(1+\mathrm{CV}_{\mathrm{r}, \mathrm{t}}^{2}\right) \tag{A.14}
\end{equation*}
$$

## Length-composition data

The length-composition data are included in the likelihood function using the robust normal for proportions likelihood, i.e., generically:

$$
\begin{equation*}
L L_{r}^{L F}=0.5 \sum_{t} \sum_{j} \ln \left(2 \pi \sigma_{t, j}^{2}\right)-\sum_{t} \sum_{j} \ln \left[\exp \left(-\frac{\left(P_{t, j}-\hat{P}_{t, j}\right)^{2}}{2 \sigma_{t, j}^{2}}\right)+0.01\right] \tag{A.15}
\end{equation*}
$$

where $P_{t, j}$ is the observed proportion of crabs in length-class j in the catch during year $\mathrm{t},{ }^{\boldsymbol{P}_{t, j}}$ is the model-estimate corresponding to ${ }^{P_{t, j}}$, i.e.:
$\hat{\mathrm{L}}_{\mathrm{t}, \mathrm{j}}^{\mathrm{r}}=\frac{\widehat{\mathrm{C}}_{\mathrm{t}, \mathrm{j}}}{\sum_{\mathrm{j}}^{\mathrm{n}} \mathrm{C}_{\mathrm{t}, \mathrm{j}}}$
$\hat{\mathrm{L}}_{\mathrm{t}, \mathrm{j}}^{\mathrm{T}}=\frac{\widehat{\mathrm{T}}_{\mathrm{t}, \mathrm{j}}}{\sum_{\mathrm{j}}^{\mathrm{n}} \mathrm{T}_{\mathrm{t}, \mathrm{j}}}$
$\hat{\mathrm{L}}_{\mathrm{t}, \mathrm{j}}^{\mathrm{GF}}=\frac{\widehat{\mathrm{Tr}}_{\mathrm{t}, \mathrm{j}}}{\sum_{\mathrm{j}}^{\mathrm{n}} \widehat{T r}_{\mathrm{t}, \mathrm{j}}}$
$\sigma_{t, j}^{2}$ is the variance of $\boldsymbol{P}_{t, j}$ :
$\sigma_{t, j}^{2}=\left[\left(1-P_{t, j}\right) P_{t, j}+\frac{0.1}{n}\right] / S_{t}$
and $S_{t}$ is the effective sample size for year $t$ and $n$ is the number of size classes.

Note: The likelihood calculation for retained length composition starts from length-class 6 (mid length 128 mm CL ) because the length-classes 1 to 5 mostly contain zero data.

## Tagging data

Let $V_{j, t, y}$ be the number of tagged male crab that were released during year $t$ that were in sizeclass $j$ when they were released and were recaptured after $y$ years, and $\underline{\rho}_{j, t, y}$ be the vector of recaptures by size-class from the males that were released in year $t$ that were in size-class $j$ when they were released and were recaptured after $y$ years. The log-likelihood corresponding to the multinomial distribution for the tagging data is then:
$\ln L=\lambda_{y, t a g} \sum_{j} \sum_{t} \sum_{y} \sum_{i} \rho_{j, t, y, i} \ln \hat{\rho}_{j, t, y, i}$
where $\lambda_{y, t a g}$ is the weight assigned to the tagging data for recapture year $y, \hat{\rho}_{j, t, y, i}$ is the proportion in size-class $i$ of the recaptures of males that were released during year $t$ that were in size-class $j$ when they were released and were recaptured after $y$ years:
$\underline{\hat{\rho}}_{j, t, y} \propto \underline{s}^{T}[\mathbf{X}]^{y} \underline{Z}^{(j)}$
where $Z^{(j)}$ is a vector with $V_{j, t, y}$ at element $j$ and 0 otherwise, and $S^{T}$ is the vector of total selectivities for tagged male crab by the pot fishery. This log-likelihood function is predicated on the assumption that all recaptures are in the pot fishery and the reporting rate is independent of the size of crab.

## Penalties

Penalties are imposed on the deviations of annual pot fishing mortality about mean pot fishing mortality, annual trawl fishing mortality about mean trawl fishing mortality, recruitment about mean recruitment, and the posfunction (fpen):

$$
\begin{align*}
& P_{1}=\lambda_{F} \sum_{t}\left(\ln F_{t}-\ln \bar{F}\right)^{2}  \tag{A.20}\\
& P_{2}=\lambda_{F^{T r}} \sum_{t}\left(\ln F_{t}^{T r}-\ln \bar{F}^{T r}\right)^{2}
\end{align*}
$$

(A.21)

$$
\begin{align*}
& P_{3}=\lambda_{R} \sum_{t}\left(\ln \varepsilon_{t}\right)^{2}  \tag{A.22}\\
& \mathrm{P}_{5}=\lambda_{\text {posfn }} * \text { fpen } \tag{A.23}
\end{align*}
$$

## Standardized Residual of Length Composition

$$
\begin{equation*}
\text { Std. } \operatorname{Res}_{\mathrm{t}, \mathrm{j}}=\frac{\mathrm{P}_{\mathrm{t}, \mathrm{j}}-\stackrel{P}{\mathrm{t}, \mathrm{j}}}{\sqrt{2 \sigma_{\mathrm{t}, \mathrm{j}}^{2}}} \tag{A.24}
\end{equation*}
$$

## Output Quantities

## Harvest rate

Total pot fishery harvest rate:

$$
\begin{equation*}
E_{t}=\frac{\sum_{j=1}^{n}\left(\widehat{\mathrm{C}}_{\mathrm{j}, t}+\widehat{\mathrm{D}}_{\mathrm{j}, t}\right)}{\sum_{j=1}^{\mathrm{N}} \mathrm{~N}_{\mathrm{j}, \mathrm{t}}} \tag{A.25}
\end{equation*}
$$

Exploited legal male biomass at the start of year $t$ :
$L M B_{t}=\sum_{j=\text { legal size }}^{n} s_{j}^{T} s_{j}^{r} N_{j, t} w_{j}$
where $w_{j}$ is the weight of an animal in length-class $j$.
Mature male biomass on 15 February spawning time (NPFMC 2007) in the following year:
$\mathrm{MMB}_{\mathrm{t}}=\sum_{\mathrm{j}=\text { mature size }}^{\mathrm{n}}\left\{\mathrm{N}_{\mathrm{j}, \mathrm{t}} \mathrm{e}^{-\mathrm{y}^{\prime} \mathrm{M}}-\left(\widehat{\mathrm{C}}_{\mathrm{j}, \mathrm{t}}+\widehat{\mathrm{D}}_{\mathrm{j}, \mathrm{t}}+\widehat{\operatorname{Tr}}_{\mathrm{j}, \mathrm{t}}\right) \mathrm{e}^{\left(\mathrm{y}_{\mathrm{t}}-\mathrm{y}^{\prime}\right) \mathrm{M}}\right\} \mathrm{w}_{\mathrm{j}}$
where $y^{\prime}$ is the elapsed time from 1 July to 15 February in the following year.
For estimating the next year limit harvest levels from current year stock abundances, a $F_{\text {OFL }}$ value is needed. Current crab management plan specifies five different Tier formulas for different stocks depending on the strength of information available for a stock, for computing $F_{O F L}$ (NPFMC 2007). For the golden king crab, the following Tier 3 formula is applied to compute $F_{O F L}$ :

$$
\begin{align*}
& \text { If, } \\
& M M B_{\text {current }}>M M B_{35 \%}, F_{O F L}=F_{35 \%} \\
& \text { If, } \\
& M M B_{\text {current }} \leq M M B_{35 \%} \text { and } M M B_{\text {current }}>0.25 M M B_{35 \%}, \\
& F_{O F L}=F_{35 \%} \frac{\left(\frac{M M B_{\text {current }}}{M M B_{35 \%}}-\alpha\right)}{(1-\alpha)} \tag{A.28}
\end{align*}
$$

If,
$M M B_{\text {current }} \leq 0.25 M M B_{35 \%}$,
$F_{O F L}=0$.
where $\alpha$ is a parameter, $\mathrm{MMB}_{\text {current }}$ is the mature male biomass in the current year and $M M B_{35 \%}$ is the proxy $M M B_{M S Y}$ for Tier 3 stocks. We assumed $\alpha=0.1$.
Because projected $\mathrm{MMB}_{\mathrm{t}}$ (i.e., $\mathrm{MMB}_{\text {current }}$ ) depends on the intervening retained and discard catch (i.e., $\mathrm{MMB}_{\mathrm{t}}$ is estimated after the fishery), an iterative procedure is applied using Equations A. 27 and A. 28 with retained and discard catch predicted from Equations A.2b-d. The next year limit harvest catch is estimated using Equations A.2b-d with the estimated $F_{O F L}$ value.

Table A1. Pre-specified and estimated parameters of the population dynamics model
Parameter Number of parameters

Initial conditions:
Length specific equilibrium abundance
17 (estimated)

## Fishing mortalities:

Pot fishery, ${ }_{t}$
Mean pot fishery fishing mortality, $\bar{F}$
Groundfish fishery, $\boldsymbol{F}_{t}{ }^{T r}$

Mean groundfish fishery fishing mortality, $\bar{F}^{T r}$
1981-2016 (estimated)
1 (estimated)
1989-2016 (the mean F for 1989 to 1994 was used to estimate groundfish discards back to 1981 (estimated)
1 (estimated)

Selectivity and retention:
Pot fishery total selectivity, $\theta_{50}^{\mathrm{T}} \quad 2$ (1981-2004; 2005+) or 3 (1981-2004,
Pot fishery total selectivity difference, delta $\theta^{T}$
Pot fishery retention, $\theta_{50}^{\mathrm{r}}$
Pot fishery retention selectivity difference, delta $\theta^{r}$
Groundfish fishery selectivity 2005-2012, 2013+) (estimated)
2 (1981-2004; 2005+) or 3 (1981-2004;
2005-2012; 2013+) (estimated)
1 (1981+) (estimated)
1 (1981+) (estimated)
fixed at 1 for all size-classes
Growth:
Expected growth increment, $\omega_{1}, \omega_{2}$
Variability in growth increment, $\sigma$
Molt probability (size transition matrix with tag data), a
Molt probability (size transition matrix with tag data), b
Natural mortality, M
2 (estimated)

Recruitment:
Number of recruiting length-classes
Mean recruit length
Distribution to length-class, $\beta_{\mathrm{r}}$
Median recruitment, $\overline{\mathrm{R}}$
Recruitment deviations, $\mathcal{E}_{t}$

1 (estimated)
1 (estimated)
1 (estimated)
1 (pre-specified, $0.21 \mathrm{yr}^{-1}$ )

5 (pre-specified)
1 (pre-specified, 110 mmCL )
1 (estimated)
1 (estimated)
57 (1961-2017) (estimated)

Fishery catchability, q
2 (1985-2004; 2005+) or 3 (1981-2004;
2005-2012; 2013+) (estimated)
Additional CPUE indices standard deviation, $\sigma_{\mathrm{e}}$ Likelihood weights (coefficient of variation)

1 (estimated)
Pre-specified, varies by scenario

Table A2. Specifications for the weights with corresponding coefficient of variations* in parentheses for each scenario for EAG and WAG. select. phase = selectivity phase. Scenario 17_Of is for the independent survey and applicable only to EAG.

| Weight | Value |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Scenario 17_0 | Scenario 17 0a | Scenario $170 b$ | Scenario 17 0c | Scenario 17 0d | Scenario <br> 17 0e | Scenario 17 Of |
| Catch: |  |  |  |  |  |  |  |
| Retained catch for 1981- | 500 (0.032) | 500 | 500 | 500 | 500 | 500 | 500 |
| 1984 and/or 1985-2016, $\lambda_{r}$ Total catch for 1990-2016, $\lambda_{T}$ | $\begin{array}{lr} \text { Number } & \text { of } \\ \text { sampled } & \text { pots } \\ \text { scaled to a } & \max \\ 250 \end{array}$ | $\begin{array}{lr} \text { Number } & \text { of } \\ \text { sampled } & \text { pots } \\ \text { scaled to a } \\ 250 \end{array}$ | $\begin{aligned} & \text { Number of } \\ & \text { sampled pots } \\ & \text { scaled to a max } \\ & 250 \end{aligned}$ | Numberof <br> sampled pots <br> scaled to a max <br> 250 <br> 02 | Numberof <br> sampled <br> pots <br> scaled to a <br> 250$\quad$max <br> 00 | Numberof <br> sampled <br> pots <br> scaled to a max <br> 250 | Number of sampled pots scaled to a max 250 |
| Groundfish bycatch for $0.2(3.344)$ 0.2 0.2 0.2 0.2 0.2 <br> $1989-2016, ~$       <br> Catch-rate:       |  |  |  |  |  |  |  |
| $\lambda_{r, \text { CPUE }}$ | 1(0.805) | 1 | 1 | 1 | 1 | 1 | 1 |
| Independent survey catchrate for 2015-2017, $\lambda_{r, \text { CPUE }}$ |  |  |  |  |  |  | 1(0.805) |
| Fish ticket retained crab catch-rate for 1985-1998 , $\lambda_{r, \text { CPUE }}$ | 1(0.805) | 1 | 1 | 1 | 1 | 1 | 1 |
| Penalty weights: |  |  |  |  |  |  |  |
| Pot fishing mortality dev, $\lambda_{F}$ | Initially 1000 , relaxed to 0.001 at phases $\geq$ select. phase | Initially 1000 , relaxed to 0.001 at phases $\geq$ select. phase | Initially 1000, relaxed to 0.001 at phases $\geq$ select. phase | Initially 1000 , relaxed to 0.001 at phases $\geq$ select. phase | Initially 1000 , relaxed to 0.001 at phases $\geq$ select. phase | Initially 1000 , relaxed to 0.001 at phases $\geq$ select. phase | Initially 1000 , relaxed to 0.001 at phases $\geq$ select. phase |
| Groundfish fishing mortality dev, $\boldsymbol{\lambda}_{F^{r r}}$ | Initially 1000 , relaxed to 0.001 at phases $\geq$ select. phase 2 (0.533) | Initially 1000 , relaxed to 0.001 at phases $\geq$ select. phase 2 | Initially 1000, relaxed to 0.001 at phases $\geq$ select. phase 2 | Initially 1000 , relaxed to 0.001 at phases $\geq$ select. phase 2 | Initially 1000 , relaxed to 0.001 at phases $\geq$ select. phase 2 | Initially 1000 , relaxed to 0.001 at phases $\geq$ select. phase 2 | Initially 1000 , relaxed to 0.001 at phases $\geq$ select. phase 2 |
| Recruitment, Posfunction (to keep abundance estimates always positive), $\lambda_{\text {posfn }}$ | $2(0.533)$ $1000(0.022)$ | 2 1000 | 2 1000 | 2 1000 | 2 1000 | 2 1000 | 2 1000 |


| Tagging likelihood | EAG individual <br> tag returns | EAG tag data | EAG tag data | EAG tag data |
| :--- | :--- | :--- | :--- | :--- | EAG tag data $\quad$ EAG tag data EAG tag data

* Coefficient of Variation, $C V=\sqrt{\exp \left[\frac{1}{2 W}\right]-1}, \quad w=$ weight


## Appendix B: Catch and CPUE data

The commercial catch and length frequency distribution were estimated from ADF\&G landing records and dockside sampling (Bowers et al. 2008, 2011). The annual retained catch, total catch, and groundfish (or trawl) discarded mortality are provided in Tables 1, 2, and $2 b$ for EAG and WAG. The weighted length frequency data were used to distribute the catch into $5-\mathrm{mm}$ size intervals. The length frequency data for a year were weighted by each sampled vessel's catch as follows. The $i$-th length-class frequency was estimated as:

$$
\begin{equation*}
\sum_{j=1}^{k} C_{j} \frac{L F_{j, i}}{\sum_{i=1}^{n} L F_{j, i}} \tag{B.1}
\end{equation*}
$$

where $k=$ number of sampled vessels in a year, $L F_{j, i}=$ number of crabs in the $i$-th lengthclass in the sample from $j$-th vessel, $\mathrm{n}=$ number of size classes, $C_{j}=$ number of crabs caught by $j$-th vessel. Then the relative frequency for the year was calculated and applied to the annual retained catch (in number of crabs) to obtain retained catch by length-class.

The annual total catch (in number of crabs) was estimated by the observer nominal (unstandardized) total CPUE considering all vessels multiplied by the total fishing effort (number of pot lifts). The weighted length frequency of the observer samples across the fleet was estimated using Equation B.1. Observer measurement of crab ranged from 20 to 220 mm CL. To restrict the total number of crabs to the model assumed size range (101-185+ mm CL ), the proportion of observer total relative length frequency corresponding to this size range was multiplied by the total catch (number of crabs). This total number of crabs was distributed into length-classes using the weighted relative length frequency. Thus, crab sizes < 101 mm CL were excluded from the model. In addition, all crab > 185 mm CL were pooled into a plus length class. Note that the total crab catch by size that went into the model did not consider retained and discard components separately. However, once the model estimated the annual total catch, then retained catch was deducted from this total and multiplied by handling mortality [we used a $20 \%$ handling mortality (Siddeek et al. 2005) to obtain the directed fishery discarded (dead) catch].

Observer data have been collected since 1988 (Moore et al. 2000; Barnard et al. 2001; Barnard and Burt 2004; Gaeuman 2011), but data were not comprehensive in the initial years, so a shorter time series of data for the period 1990/91-2016/17 was selected for this analysis. During 1990/91-1994/95, observers were only deployed on catcher-processor vessels. During 1995/96-2004/05, observers were deployed on all fishing vessels during fishing activity. Observers have been deployed on all fishing vessels since 2005/06, but catcher-only vessels are only required to carry observers for a minimum of $50 \%$ of their fishing activity during a season; catcher-processor vessels are still required to carry observers during all fishing activity. Onboard observers sample seven pots per day (it can be different number of pots per string) and count and measure all crabs caught and categorize catch as females, sublegal males, retained legal males, and non-retained legal males in a sampled pot. Prior to the 2009/10 season, depending on season, area, and type of fishing vessel, observers were also instructed to sample additional pots in which all crab were only counted and categorized as females, sublegal males, retained legal males, and non-retained legal males, but were not measured. Annual mean nominal CPUEs of retained and total crabs were estimated considering all sampled pots within each season (Table 3). The observer CPUE
data collection improved over the years and the data since 1995/96 are more reliable. Thus, for model fitting, the observer CPUE time series was restricted to 1995/96-2016/17. The 1990/91-2016/17 observer database consists of 112,510 records and that of 1995/96-2016/17 contains 108,231 records, For CPUE standardization, these data were further reduced by $5 \%$ cutoff of Soak time and $1 \%$ cutoff of Depth on both ends of the variable range to remove unreliable data or data from dysfunctional pot operations, and restricting to vessels which have made five trips per year for at least three years during 1985/86-2016/17.

Length-specific CPUE data collected by observers provides information on a wider size range of the stock than did the commercial catch length frequency data obtained from mostly legal-sized landed males.

There were significant changes in fishing practice due to changes in management regulations (e.g., since 1996/97 constant TAC and since 2005/06 crab rationalization), pot configuration (escape web on the pot door increased to 9 " since 1999), and improved observer recording in Aleutian Islands golden king crab fisheries since 1998. These changes prompted us to consider two separate observer CPUE time series, 1995/96-2004/05 and 2005/06-2016/17, to estimate CPUE indices for model input.

To include a long time series of CPUE indices for stock abundance contrast, we also considered the 1985/86-1998/99 legal size standardized CPUE as a separate likelihood component in all scenarios. Because of the lack of soak time data previous to 1990, we estimated the CPUE index considering a limited set of explanatory variables (e.g., vessel, captain, area, month) and fitting the lognormal GLM to fish ticket data (Tables 4 and 26).

When using CPUE indices in the model fit, we compared the predicted with the observed legal male CPUE in the observer CPUE likelihoods because legal male (retained plus nonretained) data are more reliable than total in the observer samples.

Most scenarios used CPUE indices estimated by the GLM method. One scenario (17_0a) used the deltaGLMM spatio-temporal method (VAST, Thorson et al., 2015) to estimate observer CPUE indices. We describe both below:

## a. Observer CPUE index by GLM:

The CPUE standardization followed the GLM fitting procedure (Maunder and Punt 2004; Starr 2012; Siddeek et al. 2016b). We considered the negative binomial GLM on positive and zero catches to select the explanatory variables. The response variable CPUE is the observer sample catch record for a pot haul. The negative binomial model uses the log link function for the GLM fit. Therefore, we assumed the null model to be

$$
\begin{equation*}
\ln \left(\text { CPUE }_{i}\right)=\text { Year }_{y_{i}} \tag{B.2}
\end{equation*}
$$

where Year is a factorial variable.
The maximum set of model terms offered to the stepwise selection procedure was:

$$
\begin{align*}
& \ln \left(\text { CPUE }_{I}\right)=\text { Year }_{y_{i}}+\text { ns }\left(\text { Soak }_{\text {si }}, \text { df }\right)+\text { Month }_{m_{i}}+\text { Vessel }_{\text {vi }}+\text { Captain }_{\text {ci }}+\text { Area }_{\text {ai }}+ \\
& \operatorname{Gear}_{\mathrm{gi}}+\mathrm{ns}\left(\text { Depth }_{\text {di }}, \text { df }\right)+\text { ns }\left(\text { VesSoak }_{\text {vsi }}, \text { df }\right), \tag{B.3}
\end{align*}
$$

where Soak is in unit of days and is numeric; Month, Area code, Vessel code, Captain code, and Gear code are factorial variables; Depth in fathom is a numeric variable; VesSoak is a numeric variable computed as annual number of vessels times annual mean soak days (to account for other vessels' effect on CPUE); ns=cubic spline, and $\mathrm{df}=$ degree of freedom.

We used a log link function and a dispersion parameter ( $\theta$ ) in the GLM fitting process. We used the $\mathrm{R}^{2}$ criterion for predictor variable selection (Siddeek et al. 2016b).

The $\mathrm{R}^{2}$ formula for explanatory variable selection is as follows:
$R^{2}=\frac{(n u l l \text { model deviance-added parameter model deviance) }}{\text { null model deviance }}$
An arbitrary $\mathrm{R}^{2}$ minimum increment of 0.01 was set to select the model terms.
First we determined the dispersion parameter $(\theta)$ by a grid search method (Fox and Weisberg, 2011). The best $\theta$ value was obtained at the minimum AIC:

Table B.1. Dispersion parameter search.

|  | Time Period | $\boldsymbol{\theta}$ | AIC |
| :--- | :--- | :--- | :--- |
| EAG | $1995 / 96-2004 / 05$ | 1.37 | 223,933 |
|  | $2005 / 06-2016 / 17$ | 2.30 | 59,284 |
|  |  |  |  |
| WAG | $1995 / 96-2004 / 05$ | 1.00 | 196,290 |
|  | $2005 / 06-2016 / 17$ | 1.17 | 94,190 |

Then we used the optimized dispersion parameter value in the GLM model for individual predictor variable fit to determine appropriate df value based on the minimum AIC:

Table B.2. Predictor variable degree of freedom search.

|  | Time Period | Predictor Variable | Df | AIC |
| :---: | :---: | :---: | :---: | :---: |
| EAG | 1995/96-2004/05 | Soak | 4 | 235,222 |
|  |  | Depth | 2 | 237,098 |
|  |  | VesSoak | 9 | 232,152 |
|  | 2005/06-2016/17 | Soak | 11 | 59,988 |
|  |  | Depth | 6 | 60,215 |
|  |  | VesSoak | 4 | 59,982 |
| WAG | 1995/96-2004/05 | Soak | 10 | 201,755 |
|  |  | Depth | 9 | 205,398 |
|  |  | VesSoak | 7 | 204,841 |
|  | 2005/06-2016/17 | Soak | 5 | 95,181 |
|  |  | Depth | 2 | 95,202 |
|  |  | VesSoak | 4 | 94,954 |

We also used the "stepAIC" package (R Core Team, 2018) for forward selection of predictor variables for CPUE standardization for scenarios EAG17_0b and WAG17_0b, respectively. Instead of using the traditional AIC (-2log_likelihood+2p) we used CAIC \{$2 \log$ _likelihood $+[\ln (\mathrm{n})+1] \mathrm{p}\}$ for variable selection, where $\mathrm{n}=$ number of observations and $\mathrm{p}=$ number of parameters to be estimated.

The final main effect models for EAG were:
Under R square selection criteria:
$\ln ($ CPUE $)=$ Year + Gear + Captain + Area + ns(Soak, 4)
for the 1995/96-2004/05 period $\left[\theta=1.37, \mathrm{R}^{2}=0.2473\right.$, $\left.\mathrm{AIC}=223,933\right]$
Under CAIC selection criteria:
$\ln ($ CPUE $)=$ Year + Gear + Captain + ns $($ Soak, 4$)+$ Month + Area
for the 1995/96-2004/05 period $\left[\theta=1.37, \mathrm{R}^{2}=0.2563\right.$, $\left.\mathrm{AIC}=224,707\right]$
Under R square selection criteria:
$\ln ($ CPUE $)=$ Year + Captain + Gear + ns $($ Soak, 11)
for the 2005/06-2016/17 period $\left(\theta=2.30, \mathrm{R}^{2}=0.1177\right.$, $\left.\mathrm{AIC}=59,284\right)$.
Under CAIC selection criteria:
$\ln ($ CPUE $)=$ Year + Vessel + Gear + ns(Soak, 11)
for the 2005/06-2016/17 period $\left(\theta=2.30, R^{2}=0.1143, \mathrm{AIC}=59,610\right)$.
The final models for WAG were:

Under R square selection criteria:
$\ln ($ CPUE $)=$ Year + Captain + Gear + ns (Soak, 10) + Area
for the 1995/96-2004/05 period $\left[\theta=1.00, \mathrm{R}^{2}=0.2031\right.$, $\left.\mathrm{AIC}=196,290\right]$
Under CAIC selection criteria:

$$
\begin{equation*}
\ln (\text { CPUE })=\text { Year }+ \text { Captain }+ \text { Gear }+ \text { ns }(\text { Soak, } 10)+\text { Month }+ \tag{B.10}
\end{equation*}
$$

Vessel + ns(Depth, 9)
for the 1995/96-2004/05 period $\left[\theta=1.00, R^{2}=0.1948, \mathrm{AIC}=197,640\right]$
Under R square selection criteria:
$\ln ($ CPUE $)=$ Year + Area + Gear $+\mathrm{ns}($ Soak, 5$)$
for the 2005/06-2016/17 period
$\left[\theta=1.17, \mathrm{R}^{2}=0.0831\right.$, AIC $=94,190$ with $\mathrm{ns}($ Soak, 5$)$ forced in$]$
Under CAIC selection criteria:
$\ln ($ CPUE $)=$ Year + Gear + Vessel + ns $($ Depth, 2$)+$ Month
+ns(Soak, 5)
for the 2005/06-2016/17 period
$\left[\theta=1.17, R^{2}=0.0684\right.$, AIC $=94,699$ with ns(Soak, 5$)$ forced in]

The final model after adding the Year:Area interaction term in the scope of variables for EAG were:

$$
\begin{equation*}
\ln (\text { CPUE })=\text { Year }+ \text { Gear }+ \text { Captain }+ \text { Area }+ \text { Year: Area }+ \text { ns }(\text { Soak }, 4) \tag{B.13}
\end{equation*}
$$

for the 1995/96-2004/05 period
$\left[\theta=1.37, \mathrm{R}^{2}=0.2684, \mathrm{AIC}=223,164\right.$ with $\mathrm{ns}($ Soak, 4$)$ forced in ]
Note: A number of indeterminate parameter values for interaction factors were observed. However, as per January 2018 CPT request, we used the resulting CPUE indices in scenario EAG17_0c.
$\ln ($ CPUE $)=$ Year + Captain + Gear + ns $($ Soak, 11)
for the 2005/06-2016/17 period $\left[\theta=2.30, \mathrm{R}^{2}=0.1177\right.$, $\left.\mathrm{AIC}=59,284\right]$.
Note: The Year:Area interaction term was not selected.

The final model after adding Year:Area interaction term in the scope of variables for WAG were:
$\ln ($ CPUE $)=$ Year + Captain + Gear + ns $($ Soak, 10 $)+$ Area
for the 1995/96-2004/05 period $\left[\theta=1.00, \mathrm{R}^{2}=0.2031\right.$, $\left.\mathrm{AIC}=196,290\right]$
Note: The Year:Area interaction term was not selected.
$\ln ($ CPUE $)=$ Year + Area + Year: Area $+\mathrm{ns}($ Soak, 5$)$
for the 2005/06-2016/17 period
$\left[\theta=1.17, R^{2}=0.1356\right.$, AIC $=94,273$ with $\mathrm{ns}($ Soak, 5$)$ forced in]
Note: A number of indeterminate parameter values for interaction factors were observed. However, as per January 2018 CPT request, we used the resulting CPUE indices in scenario WAG17_0c.

Figures B. 1 to B. 4 depict the trends in nominal and standardized CPUE indices for the two CPUE time series for EAG and WAG, respectively.

## b. Fishery independent survey CPUE index by GLM:

The fishing industry and ADF\&G cooperative fishery independent surveys have been conducted during the first month of each fishing season (i.e., August) for the last three years, 2015-2017 in the EAG, and this project is expected to continue. The sampling procedure is different from the observer sampling design. Fishing operations are conducted in a randomly selected grids ( 2 km X 2 km ) and five pots per string are sampled for fishery and biological data collection (e.g., date, vessel, captain, soak time, depth, Lat. Long., pot number, string number, species, sex, size, legal status, catch, etc.). There are 7294 records for EAG golden
king crab. For CPUE standardization, these data were further reduced by $5 \%$ cutoff of Soak time and $1 \%$ cutoff of Depth on both ends of the variable range to remove unreliable data or data from dysfunctional pot operations.

The GLM followed the same procedure as that for observer data for standardizing CPUE. Only $R^{2}$ criterion was used for variable selection. The null model was
$\ln \left(\right.$ CPUE $\left._{i}\right)=$ Year $_{y_{i}}$
where Year is a factorial variable.
The maximum set of model terms offered to the stepwise selection procedure was:

$$
\begin{aligned}
& \ln \left(\mathrm{CPUE}_{\mathrm{i}}\right)=\text { Year }_{\mathrm{y}_{\mathrm{i}}}+\mathrm{ns}\left(\text { Soak }_{\text {si }}, \text { df }\right)+\text { Vessel }_{\mathrm{vi}}+\text { Captain }_{\mathrm{ci}}+\text { VesStrPot }_{\mathrm{Si}}+\text { Lat }_{i}+ \\
& \text { LongLong }_{i}+\mathrm{ns}\left(\text { Depth }_{\mathrm{di}}, \text { df }\right) \\
& \text { (B. 18) }
\end{aligned}
$$

where Soak is in unit of days and is numeric; Depth in fathom is a numeric variable; Vessel code, Captain code, VesStrPot , Lat, and Long are factorial variables; ns=cubic spline, and df $=$ degree of freedom. To make a unique factor level for vessel, string, and pot, we concatenated the Vessel code, string ID, and PotID (VesStrPot)..

The final model was
$\ln ($ CPUE $)=$ Year + VesStrPot + Lat + ns $($ Soak, 11)
for the 2015/16-2017/18 period
$\left[\theta=2.30, R^{2}=0.55695, \operatorname{AIC}=30,481\right.$ with $n s($ Soak, 11) forced in $]$.
Because the assessment model considered fisheries data up to 2016/17, we used CPUE indices for 2015/16-2016/17 in the fitting of scenario EAG17_0f.

Figure B. 5 shows the trends in nominal and standardized CPUE indices for the two CPUE time series for the independent survey in EAG.
c. Observer CPUE index by VAST:

We used a spatio-temporal deltaGLMM (Thorson et al., 2015; Thorson et al., 2017; Thorson and Barnett, 2017) to develop separate sets of CPUE indices based on the observer data for the pre-(1995/96-2004/05) and post- rationalization (2005/06-2016/17) periods. This is a two-stage model that first estimates the probability of presence (B.20) then estimates positive catch rates in the second stage (B.21). To account for the spatial dependence of crab density within the model, spatial and spatio-temporal autocorrelation are incorporated into the model as random effects. Positive catch rates in the model are a function of area fished. Since area swept is difficult to define for a pot gear, we used soak time as the area fished proxy. The number of knots is user defined and derived over the spatial domain based on the relative sampling density. Based on the fishing locations recorded during 1995/96-2016/17, one hundred knots were selected for each of EAG and WAG (Thorson et al., 2015; Runnebaum et al., 2017) (Figure B.6).

The final models applied to each period for EAG and WAG data are:
$P_{i}=\operatorname{logit}^{-1}\left[d_{T_{(i)}}^{(p)}+r_{v_{i}}^{(p)}+\omega_{J_{(i)}}^{(p)}+\varepsilon_{J_{(i),}, T_{(i)}}^{(p)}\right]$
$\lambda_{i}=w_{i} \exp \left[d_{T_{(i)}}^{(\lambda)}+r_{v_{i}}^{(\lambda)}+\omega_{J_{(i)}}^{(\lambda)}+\varepsilon_{J_{(i)}, T_{(i)}}^{(\lambda)}\right]$
where
$P_{i}$ and $\lambda_{i}$ are the expected probabilities of an occupied habitat and positive catches given occupied habitat for sample i at a given location; $d_{T_{(i)}}$ is the average annual density in year $T_{(i)}$, $J_{i}$ is the nearest knot to sample $\mathrm{i} ; w_{i}$ is the soak time for sample $\mathrm{i} ; \omega_{J_{(i)}}$ is a random field accounting for spatially correlated variability at knot $J_{(i)}$ that is persistent among years; $\varepsilon_{J_{(i)}, T_{(i)}}$ is the random field accounting for spatio-temporal correlation at knot $J_{(i)}$ in year $T_{(i)}$; and $r_{v_{i}}$ is a random effect accounting for differences in catch between vessels.

Figure B. 7 compares the CPUE index trends between GLM and deltaGLMM estimates for EAG and WAG. The CPUE trends are similar, in particular during the post rationalization period. The confidence intervals for deltaGLMM estimated CPUE indices are wider than those of GLM estimated CPUE indices. Spatio-temporal models have been shown to provide more precision compared to design sampling of stock distribution because they are able to account for the spatial variation in density, thereby minimizing unexplained variability (Thorson et al., 2015). However, this was not the case when using a spatio-temporal deltaGLMM for golden king crab along the Aleutian Islands. There are likely two contributing factors to the increased variability in CPUE estimates. 1) Currently the VAST modeling framework is not able to account for an 'edge effect' when extrapolating species density to a given grid cell. In essence, there is no recognition of a land existing between density distributions, it appears density estimates from a given knot are being extrapolated over land. 2) Standard abundance surveys use a pre-designed grid and sample all grid cells consistently; on the other hand, commercial fishery samples the stock area opportunistically. Consequently, there are some years where there are large gaps in coverage, resulting in large areas being assigned density estimates with no direct observations for that area. This is leading to likely uncertainty in the spatial variability in density estimates. These are the two likely causes of increased variability in standard error estimates when using the deltaGLMM for Aleutian Islands golden king crab.

## Fish Ticket CPUE index:

We also fitted the lognormal GLM for the fish ticket retained CPUE time series 1985/861998/99 offering Year, Month, Vessel, Captain, and Area as explanatory variables. The fitting procedure was similar to that followed for observer data analysis. There were 20,435 records for 1985/86-2016/17. The number of records was reduced by considering only those for 1985/86 1998/99, positive catches, and Vessels with five trips per year for at least three years.

The final model for EAG was:
Under R square selection criteria:
$\ln ($ CPUE $)=$ Year + Captain + Area + Vessel + Month, $R^{2}=0.5037$, AIC $=4,957$
Under CAIC selection criteria:
$\ln ($ CPUE $)=$ Year + Vessel + Month, $\mathrm{R}^{2}=0.3700$, AIC $=5,345$
and those for WAG was:
Under R square selection criteria:
$\ln ($ CPUE $)=$ Year + Captain + Vessel + Area, $R^{2}=0.4971$, AIC $=9,923$
Under CAIC selection criteria:
$\ln ($ CPUE $)=$ Year + Vessel, $R^{2}=0.3679$, AIC $=10,670$
The final model after adding the Year:Area interaction term in the scope of variables for EAG were:

Under R square selection criteria:
$\ln ($ CPUE $)=$ Year + Captain + Area + Vessel + Month + Year: Area, $R^{2}=0.6086$, AIC $=$ 4,783

The final model after adding the Year:Area interaction term in the scope of variables for WAG were:

Under R square selection criteria:
$\ln ($ CPUE $)=$ Year + Captain + Vessel + Area + Year: Area, $R^{2}=0.6105$, AIC $=9,802$

Note:

1. A number of indeterminate parameter values for Year:Area interaction factors were observed. However, as per January 2018 CPT request, we used the resulting CPUE indices in scenarios EAG17_0c and WAG17_0c.
2. The $\mathrm{R}^{2}$ values for the fish ticket data fits are much higher compared to that for observer data fits.

Figures B. 8 and B. 9 depict the trends in nominal and standardized CPUE indices for the fish ticket CPUE time series for EAG and WAG, respectively.

Note: For brevity we did not present the diagnostic figures for the fits in this document. They are available with the first author.


Figure B.1. Trends in non-standardized [arithmetic (nominal)] and standardized (negative binomial GLM) CPUE indices with +/-2 SE for Aleutian Islands golden king crab observer data from EAG (east of $174{ }^{\circ} \mathrm{W}$ longitude). Top panel: 1995/96-2004/05, and bottom panel: 2005/06-2016/17. Standardized indices: black line and non-standardized indices: red line. Variable selection by $\mathrm{R}^{2}$ criteria.


Figure B.2. Trends in non-standardized [arithmetic (nominal)] and standardized (negative binomial GLM) CPUE indices with +/- 2 SE for Aleutian Islands golden king crab observer data from EAG (east of $174^{\circ} \mathrm{W}$ longitude). Top panel: 1995/96-2004/05, and bottom panel: 2005/06-2016/17. Standardized indices: black line and non-standardized indices: red line. Variable selection by CAIC criteria.


Figure B.3. Trends in non-standardized [arithmetic (nominal)] and standardized (negative binomial GLM) CPUE indices with +/- 2 SE for Aleutian Islands golden king crab observer data from WAG (east of $174^{\circ}$ W longitude). Top panel: 1995/96-2004/05, and bottom panel: 2005/06-2016/17. Standardized indices: black line and non-standardized indices: red line. Variable selection by $\mathrm{R}^{2}$ criteria.


Figure B.4. Trends in non-standardized [arithmetic (nominal)] and standardized (negative binomial GLM) CPUE indices with +/- 2 SE for Aleutian Islands golden king crab observer data from WAG (east of $174^{\circ}$ W longitude). Top panel: 1995/96-2004/05, and bottom panel: 2005/06-2016/17. Standardized indices: black line and non-standardized indices: red line. Variable selection by CAIC criteria.


Figure B.5. Trends in non-standardized [arithmetic (nominal)] and standardized (negative binomial GLM) CPUE indices with +/- 2 SE for Aleutian Islands golden king crab independent survey data from EAG (east of $174^{\circ} \mathrm{W}$ longitude) during 2015-2017. Standardized indices: black line and non-standardized indices: red line. Variable selection by $\mathrm{R}^{2}$ criteria.


Figure B6. One hundred knots selected each for EAG (left panel) and WAG (right panel) for spatio-temporal delta GLMM model fitting for CPUE indices estimation.


Figure B.7. Comparison of GLM (black) and VAST (green) estimated CPUE indices with +/- 2 SE for Aleutian Islands golden king crab in EAG (top panel) and WAG (bottom panel) for 1995/96-2016/17. GLM variable selection by $\mathrm{R}^{2}$ criteria.


Figure B.8. Trends in non-standardized [arithmetic (nominal)] and standardized (lognormal GLM) CPUE indices with +/- 2 SE for Aleutian Islands golden king crab from EAG. The 1985/86-1998/99 fish ticket data set was used. Standardized indices: black line and nonstandardized indices: red line. Top panel: variable selection by $\mathrm{R}^{2}$ criteria; bottom panel: variable selection by CAIC square criteria.


Figure B.9. Trends in non-standardized [arithmetic (nominal)] and standardized (lognormal GLM) CPUE indices with $+/-2$ SE for Aleutian Islands golden king crab from WAG. The 1985/86-1998/99 fish ticket data set was used. Standardized indices: black line and nonstandardized indices: red line. Top panel: variable selection by $\mathrm{R}^{2}$ criteria; bottom panel: variable selection by CAIC criteria.

## Appendix C: Male maturity

## Male maturity:

## Method:

We used the 1991 EAG pot survey collected 2457 carapace length (mm CL) and chela height (up to one-tenth of a mm CH ) measurements in the EAG and NMFS survey collected 508 same measurements in Bowers Ridge, WAG for male $50 \%$ maturity length determination. We determined the $50 \%$ maturity length outside the assessment model using the 'segmented regression' package available in R ( R Core Team 2017). We used the $50 \%$ maturity length as the break point for categorizing immature and mature crab for mature male biomass (MMB) determination for EAG and WAG.

First we fitted a linear regression model to the data pair using the R package as follows:
$\ln (C H / C L)=\beta_{0}+\beta_{1} C L$
where $\beta_{0}$ and $\beta_{1}$ are regression parameters
The procedure of 'segmented regression' uses maximum likelihood to fit a somewhat different parameterization of the linear model. It can be approximated as $\ln (C H / C L)=\beta_{0}+\beta_{1} C L+\beta_{2}[\mathrm{CL}-c]+\gamma I[\mathrm{CL}>c]$
where $\beta_{2}$ is a regression parameter and c is the break point. $\gamma I[\mathrm{CL}>c]$ is a dummy variable. When $\mathrm{CL}<\mathrm{c}$, the model reduces to,
$\ln (C H / C L)=\beta_{0}+\beta_{1} C L+\beta_{2}[C L-c]$
The $\gamma$ term is a measure of the distance between the end of the first segment and the beginning of the next. The model converges when $\gamma$ is minimized, thus this method constrains the segments to be (nearly) continuous.

Results:
Table C1. Breakpoint analysis results for EAG:

| Breakpoint <br> Estimate, CL: | 107.015 | Standard <br> (SE): | Error 1.916 |
| :--- | :--- | :--- | :--- |

Meaningful coefficients of the linear terms:

|  | Estimate | SE | t value | $\operatorname{Pr}(>\|t\|)$ |
| :--- | :--- | :--- | :--- | :---: |
| Intercept | -1.60175 | 0.02286 | -70.05 | $<2 \mathrm{e}-16^{* * *}$ |
| CL | 0.00070 | 0.00026 | 2.72 | $0.00657^{* *}$ |
| U1.CL | 0.00424 | 0.00029 | 14.45 | NA |

Signif. codes: 0 '***' $0.001{ }^{\text {'**' }} 0.01^{\prime *}{ }^{*} 0.05{ }^{\prime} .{ }^{\prime} 0.1^{\prime}{ }^{\prime} 1$
Adjusted R-squared: $0.4551, \mathrm{df}=2453$
Thus, the break point estimate of male CL (i.e., $50 \%$ maturity length) $=107.015 \mathrm{~mm}$ CL.

Table C2. Breakpoint analysis results for WAG:

| Breakpoint <br> Estimate, CL: | 107.482 | Standard <br> (SE): | Error | 2.747 |  |
| :--- | :---: | :---: | :---: | :--- | :---: |
| Meaningful coefficients of the linear terms: |  |  |  |  |  |
|  | Estimate | SE | t value | $\operatorname{Pr}(>\mid t)$ |  |
| Intercept | -1.63672 | 0.05592 | -29.271 | $<2 \mathrm{e}-16^{* * *}$ |  |
| CL | 0.00086 | 0.00059 | 1.446 | 0.149 |  |
| U1.CL | 0.00441 | 0.00063 | 7.035 | NA |  |

Signif. codes: 0 '***' $0.001^{\text {'**' }} 0.01^{\text {'*' }} 0.05^{\prime} .{ }^{\prime} 0.1^{\text {' }} 1$
Adjusted R-squared: $0.7389, \mathrm{df}=504$
Thus, the break point estimate of male CL (i.e., $50 \%$ maturity length) $=107.482 \mathrm{~mm}$ CL.

Figures C. 1 and C. 2 provide the segment regression fit to the $\log (\mathrm{CH} / \mathrm{CL})$ vs. CL data pair for EAG and WAG, respectively:


Figure C.1. Segmented linear regression fit to $\ln (\mathrm{CH} / \mathrm{CL})$ vs. CL data of male golden king crab in EAG.


Figure C.2. Segmented linear regression fit to $\ln (\mathrm{CH} / \mathrm{CL})$ vs. CL data of male golden king crab in WAG.

Bootstrap estimate of breakpoint with $95 \%$ confidence limits:
We created 1000 bootstrap samples of the $\ln (\mathrm{CH} / \mathrm{CL})$ and CL pair and fitted the segmented regression to each sample $[\ln (\mathrm{CH} / \mathrm{CL})$ vs CL$]$ and estimated the median and the $95 \%$ confidence interval ( $2.5 \%$ and $97.5 \%$ percentiles of CL of the breakpoints) for EAG and WAG.

Table C.3. Median and $95 \%$ confidence limits of 1000 bootstrap estimates of male maturity by breakpoint analysis of chela height and carapace length data of golden king crab in EAG (1991 data) and WAG (1984 data).

| Males | Median | Lower 95\% <br> Limit | Upper 95\% Limit |
| :--- | :--- | :--- | :--- |
| EAG <br> Maturity Breakpoint <br> (mm CL) | 107.02 | 85.12 | 111.02 |
| WAG |  |  |  |
| Maturity Breakpoint <br> (mm CL) | 107.85 | 103.46 | 126.03 |

We considered one bin above the median maturity size falling bin as the knife edge breakpoint of maturity. Thus all sizes equal and above 111 mmCL were considered to be fully mature and below this size immature for MMB calculation.

```
Essential R steps:
# Segmented regression:
# fit a single linear regression first then apply segmented
    library(segmented)
    singleline.mod<- lm(log(CH/CL)~CL)
    segmented.mod<- segmented(singleline.mod,seg.Z=~CL)
```


## Review of king crab male maturity:

Chelae allometry has been used to determine morphometric male size-at-maturity among a number of king crab (Lithodidae) stocks. Golden king crab provides a better discrimination of chelae height against size at the onset of maturity than other king crab stocks (Somerton and Otto, 1986). Table C. 4 lists the literature reported estimates of size-at-maturity of males and females of different king crab stocks in the northern hemisphere including golden king crab. Breakpoint analysis has been used to estimate maturity on king crabs in majority of cases (Table C. 4 and Webb, 2014).

Table C.4. Review of estimates of male and female size-at-maturity of golden (Lithodes aequispins), blue ( Paralithodes platypus), and red (Paralithodes camtschatica) king crabs by area and stocks. Numbers in parentheses are standard deviations estimated by the bootstrap sampling method.

| Species | Sex | Size-at- <br> Maturity <br> (mm CL) | Method | Area | Sources |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Lithodes aequispins | Male | 114 (11.4) | Breakpoint analysis on $\log$ (chela height) vs. $\log$ (carapace length) | British Columbia, Canada | Jewett et al., 1985 |
|  |  | $\begin{aligned} & 92(2.4) \\ & 107(4.6) \end{aligned}$ | Breakpoint analysis on $\log$ (chela height) vs. | St. Matthew Is. District | Somerton and Otto, 1986 |
|  |  | 130 (4.0) | $\log$ (carapace length) | Pribilof Is District Eastern Aleutian Is |  |
|  |  | $\begin{aligned} & 117.9 \text { to } \\ & 158.0 \end{aligned}$ | Breakpoint analysis on $\log$ (chela height) vs. $\log$ (carapace length) | Various water inlets in southeast Alaska | Olson, 2016 |
|  |  | 108.6 (2.6) | Breakpoint analysis on | Bowers Ridge | Otto and |
|  |  | 120.8 (2.9) | $\log$ (chela height) vs. $\log$ (carapace length) | Seguam Pass | Cummiskey, 1985 |
|  |  | $107.8 \text { (5.2) }$ | Breakpoint analysis on | Bowers Ridge | Current analysis |
|  |  | $107.0 \text { (6.2) }$ | log (chela height/carapace length) vs. carapace length; median estimates | Seguam Pass |  |
|  |  | 110 | Minimum size of successful mating (lab observation) | Prince William Sound | Paul and Paul, 2001 |
|  | Female | 105.5 (0.7) | Size at $50 \%$ ovigerity logistic regression | British Columbia, Canada | Jewett et al., 1985 |


| Species | Sex | Size-atMaturity (mm CL) | Method | Area | Sources |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{aligned} & 97.7(0.5) \\ & 99.9(0.2) \\ & 110.7(0.8) \end{aligned}$ | Size at $50 \%$ ovigerity logistic regression | St. Matthew Is. District, Pribilof Is District, Eastern Aleutian Is | Somerton and Otto, 1986 |
|  |  | $\begin{aligned} & 106.4(0.5) \\ & 113.2(0.3) \\ & 102.2(0.3) \end{aligned}$ | Size at $50 \%$ ovigerity logistic regression | Bowers Ridge Seguam Pass Petrel Bank | Otto and Cummiskey, 1985 |
| Paralithodes platypus | Male | $\begin{aligned} & 77(9.8) \\ & 108(12.8) \\ & 87(7.2) \\ & 93(13.9) \end{aligned}$ | Breakpoint analysis on $\log$ (chela height) vs. $\log$ (carapace length) | St. Matthew Is. <br> Pribilof Is. <br> Olga Bay <br> Prince William Sound | Somerton and MacIntosh, 1983 |
|  |  | $\sim 100$ | Lab study: Asymptote of the spermatophore diameter vs. carapace length | St. Matthew Is. | Paul et al., 1991 |
|  | Female | 80.6 (0.6) <br> 96.3 (0.3) <br> 93.7 (0.4) <br> 87.4 (0.5) | Size at $50 \%$ ovigerity logistic regression | St. Matthew Is. Pribilof Is. Olga Bay Prince William Sound | Somerton and MacIntosh, 1983 |
| Paralithodes camtschatica | Male | 102.8 | Breakpoint analysis on $\log$ (chela height) vs. $\log$ (carapace length) | Eastern Bering Sea | $\begin{aligned} & \text { Somerton, D.A., } \\ & 1980 \end{aligned}$ |
|  |  | 120 | Smallest male grasping female (in situ observation on mating pairs) | Kodiak | Powell et al., 2002 |
|  |  | 104.3 | Breakpoint analysis on $\log$ (chela height) vs. $\log$ (carapace length) | Barents Sea, Norway | Rafter et al., 1996 |
|  |  | 105 | Lab study: Asymptote of the spermatophore diameter vs. carapace length | Bristol Bay | Paul et al., 1991 |
|  | Female | 101.9 | Breakpoint analysis on $\log$ (chela height) vs. $\log$ (carapace length) | Eastern Bering Sea | $\begin{aligned} & \text { Somerton, D.A., } \\ & 1980 \end{aligned}$ |
|  |  | 88.8 (0.5) | Size at $50 \%$ ovigerity logistic regression | Bristol Bay | Otto et al., 1990 |
|  |  | 89 (1.3) | Size at $50 \%$ ovigerity logistic regression | Adak Island | Blau et al., 1990 |

## References:

Blau, S.F. 1990. Size at maturity of female red king crabs (Paralithodes camtschatica) in the Adak Management Area, Alaska. In proceedings of the international symposium on king and

Tanner crabs. Alaska Sea Grant College Program Report No. 90-04. University of Alaska Fairbanks, Fairbanks, AK, pp. 105-116.

Jewett, S.C., N.A. Sloan, and D.A. Somerton, 1985. Size at sexual maturity and fecundity of the fjord-dwelling golden king crab Lithodes aequispina Benedict from northern British Columbia. Journal of Crustacean Biology, 5(3): 377-385.

Olson, A.P. 2016. Spatial variability in size at maturity and reproductive timing of golden king crab (Lithodes aequispina) in Southeast Alaska. MSc thesis, University of Alaska Fairbanks, Fairbanks, Alaska.

Otto, R.S. and P.A. Cummiskey, 1985. Observations on the reproductive biology of the golden king crab (Lithodes aequispina) in the Bering Sea and Aleutian Islands. In Proceedings of the International king crab symposium, Edited by B. R. Melteff. University of Alaska Sea Grant, Anchorage, Alaska. pp. 123-136.

Otto, R.S., R.A. Macintosh, and P.A. Cummiskey. 1990. Fecundity and other reproductive parameters of female red king crab (Paralithodes camtschatica) in Bristol Bay and Norton Sound, Alaska. In Proceedings of the International Symposium on King and Tanner Crabs. Alaska Sea Grant College Program Report No. 90-04, University of Alaska Fairbanks, Fairbanks, AK, pp. 65-90.

Paul, A.J. and J.M. Paul, 2001. Size of maturity in male golden king crab, Lithodes aequispinus (Anomura: Lithodidae). Journal of Crustacean Biology, 21(2): 384-387.

Somerton, D.A. 1980. A computer technique for estimating the size of sexual maturity in crabs. Can.J.Fish.Aquat.Sci., 37:1488-1494.

Paul, J.M, A.J. Paul, R.S. Otto, and R. Macintosh. 1991. Spermatophore presence in relation to carapace length for eastern Bering Sea blue king crab (Paralithodes platypus, Brandt, 1850) and red king crab ( $P$. camtschaticus, Tilesius, 1815) . Journal of Shellfish Research 10(1): 157-163.

Powell, G.C., D. Pengilly, and S.F Blau. 2002. Mating pairs of red king crab (Paralithodes camtschaticus) in the Kodiak Archipelago, Alaska, 1960-1984. In A.J. Paul, E.G. Dawe, R. Elner, G.S. Jameison, G.H. Kruse, R.S. Otto, B. Sainte-Marie, T.C. Shirley, and D. Woodby (eds.), Crab in Cold Water Regions: Biology, Management, and Economics. University of Alaska Sea Grant College Program AK-SG-02-01, University of Alaska Fairbanks, Fairbanks, AK, pp 225-245.

Rafter E.E., M.Nilssen, and J.H. Sundet. 1996. Stomach content, life history, maturation, and morphometric parameters of red king crab, Paralithodes camtschaticus, from Varangerfjord area, North Norway, ICES CM 1996/K:10.

Somerton, D.A. and R.S. Otto, 1986. Distribution and reproductive biology of the golden king crab, Lithodes aequispina, in the Eastern Bering Sea. Fishery Bulletin, 84(3): 571-584.

Somerton, D.A. and MacIntosh, R.A. 1983. The size at sexual maturity of blue king crab, Paralithodes platypus, in Alaska. Fishery Bulletin, 81:621-628.

Webb, J., 2014. Reproductive ecology of commercially important lithodid crabs. In King crabs of the world: biology and fisheries management, Edited by B. G. Stevens. CRC Press, Boca Raton, FL. pp. 285-314.

## Appendix D: Francis and McAllister and Ianelli re-weighting methods

## Stage-1 effective sample size:

We considered number of vessel-days as the initial input annual effective sample sizes (i.e., Stage-1) for retained and total size compositions and number of trips for groundfish discard catch size composition without enforcing any upper limit. The number of vessel-days was calculated using,
Vessel - days $_{t}=$ mean trip $^{\text {day }} \times$ number of trips made by all vessels ${ }_{t}$
The groundfish bycatch of golden king crab comes from bottom trawlers, fish pot, and longlines. Vessel-days are difficult to calculate for the groundfish bycatch and hence we used annual number of trips as the Stage-1 effective sample size. Please note that we did not use the groundfish discard size compositions in any of the scenario's optimization although the predicted effective sample sizes were produced as a byproduct. We refer to the Stage-1 effective samples sizes for the size-composition of the retained catch, total catch, and the groundfish crab bycatch for year t as $\tau_{1, \mathrm{t}}^{r} \tau_{1, \mathrm{t}}^{T}$, and $\tau_{1, \mathrm{t}}^{\tau_{r}}$ respectively.

We estimated the Stage-2 effective sample sizes iteratively from Stage-1 input effective sample sizes. The reiterated effective sample sizes' subscripts replace 1 by 2.

## Francis' method:

The Francis' (2011) mean length based method [i.e., Francis TA1.8 method, Punt (2017)] uses the following formulas:

Observed mean length for year $t$,
$\overline{l_{t}}=\sum_{i=1}^{n} l_{t, i} \times P_{t, i}$
Predicted mean length for year $t$,
$\hat{\bar{l}_{t}}=\sum_{i=1}^{n} l_{t, i} \times \hat{P}_{t, i}$
Variance of the predicted mean length in year $t$,

$$
\begin{equation*}
\operatorname{var}\left(\hat{\bar{l}}_{t}\right)=\frac{\sum_{i=1}^{n} \hat{P}_{t, i}\left(l_{t, i}-\hat{l}_{t}\right)^{2}}{S_{t}} \tag{D.4}
\end{equation*}
$$

Francis' re-weighting parameter $W$,

$$
\begin{equation*}
W=\frac{1}{\operatorname{var}\left\{\frac{\bar{l}_{t}-\hat{l}_{t}}{\sqrt{\operatorname{var}\left(\hat{l}_{t}\right)}}\right\}} \tag{D.5}
\end{equation*}
$$

where $\widehat{\mathrm{P}}_{\mathrm{t}, \mathrm{i}}$ and $\mathrm{P}_{\mathrm{t}, \mathrm{i}}$ are the estimated and observed proportions of the catch during year t in lengthclass $i, l_{t, i}$ is the mid length of the length-class i during year $t, S_{t}$ is the effective sample size in
year $t, \hat{\bar{I}}_{t}$ and $\overline{l_{t}}$ are predicted and observed mean lengths of the catch during year $t, n$ is the number of length bins, and W is the re-weighting multiplier of Stage- 1 sample sizes.

Francis (2017) suggested that a good stopping criterion for the iteration process is when there are no appreciable changes in the key outputs. Hence, we considered a stopping criterion of no appreciable change ( $<0.01 \%$ ) in W and terminal year MMB (Equation A.27).
$S_{t}$ is related to the initial input (Stage-1) effective sample size according to:

$$
\begin{equation*}
S_{t, i}=W_{i} \tau_{1, t} \tag{D.6}
\end{equation*}
$$

where $S_{t, i}$ is the effective sample size for year $t$ in iteration $i$ and $W_{i}$ is the Francis weight calculated using Equation D. 5 during iteration i.

We did the re-weighting for combined data (for $M$ estimation), individual scenarios, and MMB profiles. For brevity, we provide the iteration process for Francis Stage-2 weight calculation for individual scenarios for EAG and WAG respectively in Table D.1.

## McAllister's and Ianelli's method:

Based on the assumption that the size-composition data are a multinomial sample, McAllister and Ianelli (1997) provided an estimator for the Stage-2 effective sample sizes based on the ratio of the theoretical variance of expected proportions to the actual variance of proportions,

$$
\begin{equation*}
\tau_{2, t}=\frac{\sum_{l} \hat{P}_{t, l}\left(1-\hat{P}_{t, l}\right)}{\sum_{l}\left(P_{t, l}-\hat{P}_{t, l}\right)^{2}} \tag{D.7}
\end{equation*}
$$

McAllister and Ianelli (1997) set the effective sample size for each size-composition data set for eastern Bering Sea yellowfin sole (Limanda aspera) as the arithmetic mean of $\tau_{2, t}$ over years $t$ (i.e., a year-invariant effective sample size) and iterated the model fitting, updating the effective sample sizes, until convergence occurred. Equation D. 7 ignores correlation among the residuals for the catch proportions so likely overestimates effective sample sizes (Francis, 2011). Punt (2015) suggests using the harmonic mean of $\tau_{2, t}$ if the McAllister and Ianelli formula is used. A harmonic mean (constant) multiplier was consequently used to update the effective sample sizes at each iteration of model fitting until convergence occurred; i.e.

$$
\begin{equation*}
\tau_{2, t, i}=\left\{\frac{1}{n_{t}} \sum_{t}\left[\frac{\tau_{2, t, i-1}}{\tau_{2, t, i-1}}\right]^{-1}\right\}^{-1} \tau_{2, t, i-1} \tag{D.8}
\end{equation*}
$$

where $\tau_{2, t, i}$ is the Stage-2 effective sample size for year $t$ in iteration $i\left(\tau_{2, t, 0}=\tau_{1, \mathrm{t}}\right)$ and $\dot{\tau}_{2, t, i}$ is the result of applying Equation D.7. Convergence of the process of setting the Stage-2 effective sample sizes using Equation D. 8 was assessed similar to Francis' procedure, but the weight ( $W$ ) at the final iteration was allowed to reach 1 . We considered this re-weighting process for scenarios EAG17_0e and WAG17_0e (Table D.1).

Table D.1. Iteration process for Stage-2 effective sample size re-weighting multiplier, $W$, by Francis' (scenarios 17b0, 17_0, 17_0a, 17_0b, 17_0c, 17_0d, and 17_0f) and McAllister and Ianelli (scenario 17_0e) methods for retained, total, and groundfish discard catch size compositions of golden king crab for EAG and WAG. Sc. =scenario. Note: For certain scenarios we have done over six iterations, but we provide only the last three iteration results.

| Area | Sc. | Iteration No. | Retained <br> Catch Size <br> Comp <br> Effective <br> Sample Size <br> Multiplier <br> (W) | Total Catch <br> Size Comp <br> Effective <br> Sample <br> Size <br> Multiplier <br> (W) | Groundfish <br> Discard Catch <br> Size Comp <br> Effective <br> Sample Size <br> Multiplier <br> (W) | Terminal <br> MMB (t) | $M \mathrm{yr}^{-1}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| EAGpart | 17b0 | 1 | 0.8384 | 0.5053 | 0.4469 | 14,342 | 0.2274 |
|  |  | 2 | 0.8384 | 0.5066 | 0.4458 | 14,142 | 0.2254 |
|  |  | 3 | 0.8384 | 0.5053 | 0.4469 | 14,141 | 0.2254 |
| WAGpart | 17b0 | 1 | 0.5176 | 0.4685 | 0.7542 | 6,646 | 0.2274 |
|  |  | 2 | 0.5175 | 0.4684 | 0.7584 | 6,603 | 0.2254 |
|  |  | 3 | 0.5176 | 0.4685 | 0.7542 | 6,603 | 0.2254 |
| EAG | 17_0 | 1 | 0.8343 | 0.5084 | 0.4476 | 13,455 | 0.21 |
|  |  | 2 | 0.8339 | 0.5086 | 0.4476 | 13,455 |  |
|  |  | 3 | 0.8338 | 0.5086 | 0.4476 | 13,455 |  |
| WAG | 17_0 | 1 | 0.5171 | 0.4698 | 0.7596 | 6,269 | 0.21 |
|  |  | 2 | 0.5172 | 0.4697 | 0.7598 | 6,269 |  |
|  |  | 3 | 0.5173 | 0.4697 | 0.7597 | 6,269 |  |
| EAG | 17_0a | 1 | 0.8343 | 0.5349 | 0.4488 | 13,579 | 0.21 |
|  |  | 2 | 0.8340 | 0.5349 | 0.4487 | 13,579 |  |
|  |  | 3 | 0.8340 | 0.5349 | 0.4488 | 13,579 |  |
| WAG | 17_0a | 1 | 0.5180 | 0.4707 | 0.7625 | 6,280 | 0.21 |
|  |  | 2 | 0.5183 | 0.4706 | 0.7627 | 6,280 |  |
|  |  | 3 | 0.5183 | 0.4705 | 0.7627 | 6,280 |  |
| EAG | 17_0b | 1 | 0.8351 | 0.5066 | 0.4498 | 11,842 | 0.21 |
|  |  | 2 | 0.8353 | 0.5066 | 0.4497 | 11,842 |  |
|  |  | 3 | 0.8354 | 0.5065 | 0.4497 | 11,842 |  |
| WAG | 17_0b | 1 | 0.5084 | 0.4831 | 0.7643 | 5,884 | 0.21 |
|  |  | 2 | 0.5083 | 0.4831 | 0.7643 | 5,884 |  |
|  |  | 3 | 0.5082 | 0.4832 | 0.7642 | 5,884 |  |
| EAG | 17_0c | 1 | 0.8182 | 0.5387 | 0.4468 | 13,766 | 0.21 |
|  |  | 2 | 0.8181 | 0.5388 | 0.4467 | 13,767 |  |
|  |  | 3 | 0.8181 | 0.5388 | 0.4467 | 13,767 |  |
| WAG | 17_0c | 1 | 0.4979 | 0.4774 | 0.7581 | 6,154 | 0.21 |
|  |  | 2 | 0.4979 | 0.4774 | 0.7579 | 6,154 |  |
|  |  | 3 | 0.4979 | 0.4774 | 0.7581 | 6,154 |  |
| EAG | 17_0d | 1 | 0.8450 | 0.5311 | 0.4495 | 8,833 | 0.21 |
|  |  | 2 | 0.8452 | 0.5310 | 0.4495 | 8,833 |  |
|  |  | 3 | 0.8452 | 0.5310 | 0.4495 | 8,833 |  |


| WAG | 17_0d | 1 | 0.5582 | 0.4830 | 0.7604 | 6,136 | 0.21 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  |  | 2 | 0.5578 | 0.4826 | 0.7611 | 6,136 |  |
|  |  | 3 | 0.5577 | 0.4826 | 0.7610 | 6,136 |  |
| EAG | 17_0e | 1 | 1.4025 | 0.7873 | 1.6475 | 13,453 | 0.21 |
|  |  | 2 | 1.0640 | 0.9582 | 1.0022 | 13,444 |  |
|  |  | 3 | 1.0100 | 0.9908 | 1.0001 | 13,440 |  |
| WAG | 17_0e | 1 | 1.1639 | 0.9202 | 0.9948 | 6,348 | 0.21 |
|  |  | 2 | 1.0384 | 0.9526 | 0.9989 | 6,353 |  |
|  |  | 3 | 1.0097 | 0.9817 | 0.9997 | 6,355 |  |
| EAG | $17 \_0 f$ | 1 | 0.8396 | 0.5065 | 0.4487 | 12,484 | 0.21 |
|  |  | 2 | 0.8410 | 0.5060 | 0.4488 | 12,485 |  |
|  |  | 3 | 0.8411 | 0.5060 | 0.4488 | 12,485 |  |

## Appendix E: Jittering

Jittering of scenario 17_0 parameter estimates:
We followed the Stock Synthesis approach to do 100 jitter runs of scenarios EAG17_0 and WAG17_0 parameter estimates to use as initial parameter values (as .PIN file in ADMB) to assess model stability and to determine whether a global as opposed to local minima has been found by the search algorithm:

The Jitter factor of 0.3 was multiplied by a random normal deviation $r d e v=N(0,1)$, to a transformed parameter value based upon the predefined parameter:

$$
\begin{equation*}
\text { temp }=0.5 * \text { rdev } * \text { Jitterfactor } * \ln \left(\frac{P_{\max }-P_{\min }+0.0000002}{P_{v a l}-P_{\min }+0.0000001}-1\right), \tag{E.1}
\end{equation*}
$$

with the final jittered initial parameter value back transformed as:

$$
\begin{equation*}
P_{\text {new }}=P_{\min }+\frac{P_{\max }-P_{\min }}{1.0+\exp (-2.0 \text { temp })}, \tag{E.2}
\end{equation*}
$$

where $P_{\text {max }}$ and $P_{\text {min }}$ are upper and lower bounds of parameter search space and $P_{\text {val }}$ is the estimated parameter value before the jittering.

The jitter results are summarized for scenario 17_0 in Tables E. 1 and E. 2 for EAG and WAG, respectively. Almost all runs converged to the highest log likelihood values for EAG. On the other hand, some jitter runs for WAG produced smaller objective function values compared to the base estimate (run 0 ). However, those fits predicted extremely large groundfish bycatches in certain years, consequently we ignored those runs. Thus we selected scenario 17_0 as the base scenario for EAG and WAG.

Table E.1. Results from 100 jitter runs for scenario 17_0 for EAG. Jitter run 0 corresponds to the original optimized estimates. Note: $\mathrm{B}_{\text {MSY }}$ reference points were based on average recruitment for 1986-2016.

| Jitter <br> Run | Objective Function | Maximum Gradient | $\mathrm{B}_{35 \%}$ (t) | OFL (t) | Current MMB <br> (t) |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 0 | 285.91650 | 0.0000222934 | 6,954.48 | 3,917.74 | 11,554.70 |
| 1 | 285.91650 | 0.0000174504 | 6,954.48 | 3,917.74 | 11,554.70 |
| 2 | 285.91650 | 0.0001631845 | 6,954.48 | 3,917.74 | 11,554.70 |
| 3 | 285.91650 | 0.0000062988 | 6,954.48 | 3,917.74 | 11,554.70 |
| 4 | 285.91650 | 0.0002805318 | 6,954.48 | 3,917.74 | 11,554.70 |
| 5 | 285.91650 | 0.0001137684 | 6,954.48 | 3,917.74 | 11,554.70 |
| 6 | 285.91650 | 0.0001572297 | 6,954.48 | 3,917.74 | 11,554.70 |
| 7 | 285.91650 | 0.0001488496 | 6,954.48 | 3,917.74 | 11,554.70 |
| 8 | 285.91650 | 0.0003391617 | 6,954.48 | 3,917.74 | 11,554.70 |
| 9 | 285.91650 | 0.0001285458 | 6,954.48 | 3,917.74 | 11,554.70 |


| 10 | 285.91650 | 0.0000977588 | 6,954.48 | 3,917.74 | 11,554.70 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 11 | 285.91650 | 0.0001231468 | 6,954.48 | 3,917.74 | 11,554.70 |
| 12 | 285.91650 | 0.0000890800 | 6,954.48 | 3,917.74 | 11,554.70 |
| 13 | 285.91650 | 0.0000399059 | 6,954.48 | 3,917.74 | 11,554.70 |
| 14 | 285.91650 | 0.0002567647 | 6,954.48 | 3,917.74 | 11,554.70 |
| 15 | 285.91650 | 0.0000064600 | 6,954.48 | 3,917.74 | 11,554.70 |
| 16 | 285.91650 | 0.0002346045 | 6,954.48 | 3,917.74 | 11,554.70 |
| 17 | 285.91650 | 0.0002820026 | 6,954.48 | 3,917.74 | 11,554.70 |
| 18 | 285.91650 | 0.0000241932 | 6,954.48 | 3,917.74 | 11,554.70 |
| 19 | 285.91650 | 0.0000365975 | 6,954.48 | 3,917.74 | 11,554.70 |
| 20 | 285.91650 | 0.0003771734 | 6,954.48 | 3,917.74 | 11,554.70 |
| 21 | 285.91650 | 0.0001375338 | 6,954.48 | 3,917.74 | 11,554.70 |
| 22 | 285.91650 | 0.0001120951 | 6,954.48 | 3,917.74 | 11,554.70 |
| 23 | 285.91650 | 0.0000285661 | 6,954.48 | 3,917.74 | 11,554.70 |
| 24 | 285.91650 | 0.0006714663 | 6,954.48 | 3,917.74 | 11,554.70 |
| 25 | 285.91650 | 0.0001187696 | 6,954.48 | 3,917.74 | 11,554.70 |
| 26 | 285.91650 | 0.0000138714 | 6,954.48 | 3,917.74 | 11,554.70 |
| 27 | 285.91650 | 0.0000495531 | 6,954.48 | 3,917.74 | 11,554.70 |
| 28 | 285.91650 | 0.0005756958 | 6,954.48 | 3,917.74 | 11,554.70 |
| 29 | 285.91650 | 0.0000373670 | 6,954.48 | 3,917.74 | 11,554.70 |
| 30 | 285.91650 | 0.0001517096 | 6,954.48 | 3,917.74 | 11,554.70 |
| 31 | 285.91650 | 0.0003618456 | 6,954.48 | 3,917.74 | 11,554.70 |
| 32 | 285.91650 | 0.0013670960 | 6,954.48 | 3,917.74 | 11,554.70 |
| 33 | 285.91650 | 0.0000539773 | 6,954.48 | 3,917.74 | 11,554.70 |
| 34 | 285.91650 | 0.0000154992 | 6,954.48 | 3,917.74 | 11,554.70 |
| 35 | 285.91650 | 0.0000760394 | 6,954.48 | 3,917.74 | 11,554.70 |
| 36 | 285.91650 | 0.0000046526 | 6,954.48 | 3,917.74 | 11,554.70 |
| 37 | 285.91650 | 0.0002455134 | 6,954.48 | 3,917.74 | 11,554.70 |
| 38 | 285.91650 | 0.0001081487 | 6,954.48 | 3,917.74 | 11,554.70 |
| 39 | 285.91650 | 0.0001221035 | 6,954.48 | 3,917.74 | 11,554.70 |
| 40 | 285.91650 | 0.0001775793 | 6,954.48 | 3,917.74 | 11,554.70 |
| 41 | 285.91650 | 0.0000850537 | 6,954.48 | 3,917.74 | 11,554.70 |
| 42 | 285.91650 | 0.0000655746 | 6,954.48 | 3,917.74 | 11,554.70 |
| 43 | 285.91650 | 0.0001097075 | 6,954.48 | 3,917.74 | 11,554.70 |
| 44 | 285.91650 | 0.0005359162 | 6,954.48 | 3,917.74 | 11,554.70 |
| 45 | 285.91650 | 0.0000582206 | 6,954.48 | 3,917.74 | 11,554.70 |
| 46 | 285.91650 | 0.0001263718 | 6,954.48 | 3,917.74 | 11,554.70 |
| 47 | 285.91650 | 0.0001669157 | 6,954.48 | 3,917.74 | 11,554.70 |
| 48 | 285.91650 | 0.0001184376 | 6,954.48 | 3,917.74 | 11,554.70 |
| 49 | 285.91650 | 0.0001850153 | 6,954.48 | 3,917.74 | 11,554.70 |
| 50 | 285.91650 | 0.0001171299 | 6,954.48 | 3,917.74 | 11,554.70 |
| 51 | 285.91650 | 0.0000927041 | 6,954.48 | 3,917.74 | 11,554.70 |
| 52 | 285.91650 | 0.0001977530 | 6,954.48 | 3,917.74 | 11,554.70 |


| 53 | 285.91650 | 0.0000502208 | 6,954.48 | 3,917.74 | 11,554.70 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 54 | 285.91650 | 0.0002810899 | 6,954.48 | 3,917.74 | 11,554.70 |
| 55 | 285.91650 | 0.0002931756 | 6,954.48 | 3,917.74 | 11,554.70 |
| 56 | 285.91650 | 0.0001466994 | 6,954.48 | 3,917.74 | 11,554.70 |
| 57 | 285.91650 | 0.0001492200 | 6,954.48 | 3,917.74 | 11,554.70 |
| 58 | 285.91650 | 0.0000375202 | 6,954.48 | 3,917.74 | 11,554.70 |
| 59 | 285.91650 | 0.0004659215 | 6,954.48 | 3,917.74 | 11,554.70 |
| 60 | 285.91650 | 0.0000479571 | 6,954.48 | 3,917.74 | 11,554.70 |
| 61 | 285.91650 | 0.0000159505 | 6,954.48 | 3,917.74 | 11,554.70 |
| 62 | 285.91650 | 0.0000466713 | 6,954.48 | 3,917.74 | 11,554.70 |
| 63 | 285.91650 | 0.0001467107 | 6,954.48 | 3,917.74 | 11,554.70 |
| 64 | 285.91650 | 0.0003362615 | 6,954.48 | 3,917.74 | 11,554.70 |
| 65 | 285.91650 | 0.0003528916 | 6,954.48 | 3,917.74 | 11,554.70 |
| 66 | 285.91650 | 0.0001518528 | 6,954.48 | 3,917.74 | 11,554.70 |
| 67 | 285.91650 | 0.0000965183 | 6,954.48 | 3,917.74 | 11,554.70 |
| 68 | 285.91650 | 0.0001700814 | 6,954.48 | 3,917.74 | 11,554.70 |
| 69 | 285.91650 | 0.0001150075 | 6,954.48 | 3,917.74 | 11,554.70 |
| 70 | 285.91650 | 0.0001708935 | 6,954.48 | 3,917.74 | 11,554.70 |
| 71 | 285.91650 | 0.0000843366 | 6,954.48 | 3,917.74 | 11,554.70 |
| 72 | 285.91650 | 0.0000147518 | 6,954.48 | 3,917.74 | 11,554.70 |
| 73 | 285.91650 | 0.0000711309 | 6,954.48 | 3,917.74 | 11,554.70 |
| 74 | 285.91650 | 0.0000831972 | 6,954.48 | 3,917.74 | 11,554.70 |
| 75 | 285.91650 | 0.0001249322 | 6,954.48 | 3,917.74 | 11,554.70 |
| 76 | 285.91650 | 0.0000950038 | 6,954.48 | 3,917.74 | 11,554.70 |
| 77 | 285.91650 | 0.0000930142 | 6,954.48 | 3,917.74 | 11,554.70 |
| 78 | 285.91650 | 0.0005069687 | 6,954.48 | 3,917.74 | 11,554.70 |
| 79 | 285.91650 | 0.0001041060 | 6,954.48 | 3,917.74 | 11,554.70 |
| 80 | 285.91650 | 0.0000268403 | 6,954.48 | 3,917.74 | 11,554.70 |
| 81 | 285.91650 | 0.0001235642 | 6,954.48 | 3,917.74 | 11,554.70 |
| 82 | 285.91650 | 0.0001945769 | 6,954.48 | 3,917.74 | 11,554.70 |
| 83 | 285.91650 | 0.0004412037 | 6,954.48 | 3,917.74 | 11,554.70 |
| 84 | 285.91650 | 0.0000976698 | 6,954.48 | 3,917.74 | 11,554.70 |
| 85 | 285.91650 | 0.0000551057 | 6,954.48 | 3,917.74 | 11,554.70 |
| 86 | 285.91650 | 0.0000495026 | 6,954.48 | 3,917.74 | 11,554.70 |
| 87 | 285.91650 | 0.0005078082 | 6,954.48 | 3,917.74 | 11,554.70 |
| 88 | 285.91650 | 0.0001855834 | 6,954.48 | 3,917.74 | 11,554.70 |
| 89 | 285.91650 | 0.0001687559 | 6,954.48 | 3,917.74 | 11,554.70 |
| 90 | 285.91650 | 0.0000065286 | 6,954.48 | 3,917.74 | 11,554.70 |
| 91 | 285.91650 | 0.0000599673 | 6,954.48 | 3,917.74 | 11,554.70 |
| 92 | 285.91650 | 0.0003389603 | 6,954.48 | 3,917.74 | 11,554.70 |
| 93 | 285.91650 | 0.0000402791 | 6,954.48 | 3,917.74 | 11,554.70 |
| 94 | 285.91650 | 0.0002217916 | 6,954.48 | 3,917.74 | 11,554.70 |
| 95 | 285.91650 | 0.0000923698 | 6,954.48 | 3,917.74 | 11,554.70 |


| 96 | 285.91650 | 0.0000245177 | $6,954.48$ | $3,917.74$ | $11,554.70$ |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 97 | 285.91650 | 0.0001364416 | $6,954.48$ | $3,917.74$ | $11,554.70$ |
| 98 | 285.91650 | 0.0001427303 | $6,954.48$ | $3,917.74$ | $11,554.70$ |
| 99 | 285.91650 | 0.0000980820 | $6,954.48$ | $3,917.74$ | $11,554.70$ |
| 100 | 285.91650 | 0.0000929987 | $6,954.48$ | $3,917.74$ | $11,554.70$ |

Table E.2. Results from 100 jitter runs for scenario 17_0 for WAG. Jitter run 0 corresponds to the original optimized estimates. Since there were differences in the objective function estimates, we sorted out the jitter results from lowest to the highest objective function values. Note: $\mathrm{B}_{\mathrm{MSY}}$ reference points were based on average recruitment for 1986-2017.

| $\begin{array}{l}\text { Jitter } \\ \text { Run }\end{array}$ | $\begin{array}{l}\text { Objective } \\ \text { Function }\end{array}$ |  | $\begin{array}{l}\text { Maximum } \\ \text { Gradient }\end{array}$ | B $_{35 \%}(\mathrm{t})$ | OFL $(\mathrm{t})$ |
| ---: | ---: | :--- | ---: | ---: | ---: | \(\left.\begin{array}{l}Current <br>

MMB(\mathrm{t})\end{array}\right]\)

| 22 | 194.09020 | 0.0001078193 | 5,137.94 | 1,596.46 | 6,397.24 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 23 | 194.09020 | 0.0002701639 | 5,137.94 | 1,596.46 | 6,397.24 |
| 24 | 194.09020 | 0.0004094629 | 5,137.94 | 1,596.46 | 6,397.24 |
| 25 | 194.09020 | 0.0001398647 | 5,137.94 | 1,596.46 | 6,397.24 |
| 26 | 194.09020 | 0.0001581441 | 5,137.94 | 1,596.46 | 6,397.24 |
| 27 | 194.09020 | 0.0000172173 | 5,137.94 | 1,596.46 | 6,397.24 |
| 28 | 194.09020 | 0.0002431567 | 5,137.94 | 1,596.46 | 6,397.24 |
| 29 | 194.09020 | 0.0001333304 | 5,137.94 | 1,596.46 | 6,397.24 |
| 30 | 194.09020 | 0.0001117535 | 5,137.94 | 1,596.46 | 6,397.24 |
| 31 | 194.09020 | 0.0001606068 | 5,137.94 | 1,596.46 | 6,397.24 |
| 33 | 194.09020 | 0.0004427428 | 5,137.94 | 1,596.46 | 6,397.24 |
| 34 | 194.09020 | 0.0001611413 | 5,137.94 | 1,596.46 | 6,397.24 |
| 35 | 194.09020 | 0.0000631701 | 5,137.94 | 1,596.46 | 6,397.24 |
| 36 | 194.09020 | 0.0000459606 | 5,137.94 | 1,596.46 | 6,397.24 |
| 37 | 194.09020 | 0.0001064168 | 5,137.94 | 1,596.46 | 6,397.24 |
| 38 | 194.09020 | 0.0000172059 | 5,137.94 | 1,596.46 | 6,397.24 |
| 40 | 194.09020 | 0.0000038408 | 5,137.94 | 1,596.46 | 6,397.24 |
| 41 | 194.09020 | 0.0000859666 | 5,137.94 | 1,596.46 | 6,397.24 |
| 42 | 194.09020 | 0.0000537521 | 5,137.94 | 1,596.46 | 6,397.24 |
| 43 | 194.09020 | 0.0001620099 | 5,137.94 | 1,596.46 | 6,397.24 |
| 44 | 194.09020 | 0.0000315661 | 5,137.94 | 1,596.46 | 6,397.24 |
| 45 | 194.09020 | 0.0000738932 | 5,137.94 | 1,596.46 | 6,397.24 |
| 46 | 194.09020 | 0.0001887252 | 5,137.94 | 1,596.46 | 6,397.24 |
| 47 | 194.09020 | 0.0000429643 | 5,137.94 | 1,596.46 | 6,397.24 |
| 48 | 194.09020 | 0.0000776832 | 5,137.94 | 1,596.46 | 6,397.24 |
| 49 | 194.09020 | 0.0003267544 | 5,137.94 | 1,596.46 | 6,397.24 |
| 50 | 194.09020 | 0.0003924007 | 5,137.94 | 1,596.46 | 6,397.24 |
| 51 | 194.09020 | 0.0001833688 | 5,137.94 | 1,596.46 | 6,397.24 |
| 52 | 194.09020 | 0.0002360240 | 5,137.94 | 1,596.46 | 6,397.24 |
| 53 | 194.09020 | 0.0000717775 | 5,137.94 | 1,596.46 | 6,397.24 |
| 54 | 194.09020 | 0.0001178624 | 5,137.94 | 1,596.46 | 6,397.24 |
| 55 | 194.09020 | 0.0002562605 | 5,137.94 | 1,596.46 | 6,397.24 |
| 56 | 194.09020 | 0.0001003891 | 5,137.94 | 1,596.46 | 6,397.24 |
| 57 | 194.09020 | 0.0002306516 | 5,137.94 | 1,596.46 | 6,397.24 |
| 58 | 194.09020 | 0.0001687052 | 5,137.94 | 1,596.46 | 6,397.24 |
| 59 | 194.09020 | 0.0001481354 | 5,137.94 | 1,596.46 | 6,397.24 |
| 60 | 194.09020 | 0.0000907526 | 5,137.94 | 1,596.46 | 6,397.24 |
| 61 | 194.09020 | 0.0002972557 | 5,137.94 | 1,596.46 | 6,397.24 |
| 63 | 194.09020 | 0.0001718722 | 5,137.94 | 1,596.46 | 6,397.24 |
| 64 | 194.09020 | 0.0000443092 | 5,137.94 | 1,596.46 | 6,397.24 |
| 65 | 194.09020 | 0.0004282920 | 5,137.94 | 1,596.46 | 6,397.24 |
| 66 | 194.09020 | 0.0000609887 | 5,137.94 | 1,596.46 | 6,397.24 |
| 69 | 194.09020 | 0.0000496104 | 5,137.94 | 1,596.46 | 6,397.24 |


| 70 | 194.09020 | 0.0001474220 | $5,137.94$ | $1,596.46$ | $6,397.24$ |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 71 | 194.09020 | 0.0000817530 | $5,137.94$ | $1,596.46$ | $6,397.24$ |
| 72 | 194.09020 | 0.0002925135 | $5,137.94$ | $1,596.46$ | $6,397.24$ |
| 74 | 194.09020 | 0.0000172826 | $5,137.94$ | $1,596.46$ | $6,397.24$ |
| 75 | 194.09020 | 0.0001158849 | $5,137.94$ | $1,596.46$ | $6,397.24$ |
| 77 | 194.09020 | 0.0000685658 | $5,137.94$ | $1,596.46$ | $6,397.24$ |
| 78 | 194.09020 | 0.0000642759 | $5,137.94$ | $1,596.46$ | $6,397.24$ |
| 79 | 194.09020 | 0.0002103009 | $5,137.94$ | $1,596.46$ | $6,397.24$ |
| 80 | 194.09020 | 0.0000927951 | $5,137.94$ | $1,596.46$ | $6,397.24$ |
| 82 | 194.09020 | 0.0000092932 | $5,137.94$ | $1,596.46$ | $6,397.24$ |
| 83 | 194.09020 | 0.0002106457 | $5,137.94$ | $1,596.46$ | $6,397.24$ |
| 85 | 194.09020 | 0.0002154777 | $5,137.94$ | $1,596.46$ | $6,397.24$ |
| 86 | 194.09020 | 0.0002772188 | $5,137.94$ | $1,596.46$ | $6,397.24$ |
| 87 | 194.09020 | 0.0000738715 | $5,137.94$ | $1,596.46$ | $6,397.24$ |
| 88 | 194.09020 | 0.0000222923 | $5,137.94$ | $1,596.46$ | $6,397.24$ |
| 89 | 194.09020 | 0.0000501345 | $5,137.94$ | $1,596.46$ | $6,397.24$ |
| 91 | 194.09020 | 0.0004448138 | $5,137.94$ | $1,596.46$ | $6,397.24$ |
| 92 | 194.09020 | 0.0000542747 | $5,137.94$ | $1,596.46$ | $6,397.24$ |
| 93 | 194.09020 | 0.0002043152 | $5,137.94$ | $1,596.46$ | $6,397.24$ |
| 94 | 194.09020 | 0.0000163931 | $5,137.94$ | $1,596.46$ | $6,397.24$ |
| 95 | 194.09020 | 0.0001567686 | $5,137.94$ | $1,596.46$ | $6,397.24$ |
| 97 | 194.09020 | 0.0000887919 | $5,137.94$ | $1,596.46$ | $6,397.24$ |
| 99 | 194.09020 | 0.0001385326 | $5,137.94$ | $1,596.46$ | $6,397.24$ |
| 100 | 194.09020 | 0.0004455103 | $5,137.94$ | $1,596.46$ | $6,397.24$ |
| 68 | 194.56190 | 0.0002346658 | $5,667.26$ | $1,710.85$ | $6,808.70$ |
| 96 | 1575.75000 | 8813.9550000000 | $11,642.10$ | $5,338.79$ | $18,099.20$ |
| 67 | 1755.93500 | 7572.1610000000 | $112,920.00$ | $185,340.00$ | $446,363.00$ |
| 73 | 1783.22200 | 2679.4760000000 | $6,571.68$ | $1,390.86$ | $6,177.96$ |
| 81 | 2018.62300 | 5380.9700000000 | $11,434.00$ | $7,879.87$ | $29,661.80$ |
|  |  |  |  |  |  |
| 9 |  |  |  |  |  |

## Appendix F: Projection

Simulations on future projection and outlook of Aleutian Islands golden king crab under Tier 3 harvest control rule

## Simulation Method

We simulated the future male abundances from the 2018 model scenarios 17_0 and 17_0d estimated abundances by length-class and recruitment. We projected the abundances for 30 years with 100 random replicates and estimated various management parameters: legal male biomass (LMB), mature male biomass (MMB), OFL (total) catch, retained catch, CPUE indices, and probability of overfishing under federal overfishing control rule. Future population projections primarily depend on future recruitment, but crab recruitment is difficult to predict. Therefore, annual recruitment for the projections was selected by a random selection from estimated recruitments during 1987-2012 (CPT and SSC agreed time period, Siddeek et al., 2017). Besides recruitment, the other major uncertainty for the projections is estimated abundance in 2016 (terminal year). The estimated recruitments were randomly selected using a uniform random distribution whereas the 2016 abundance was randomized by a lognormal random error.
The simulation steps are as follows:

1) Run the assessment model scenario 17_0 and 17_0d from the start year to the terminal year of the data (1981/82-2016/17 fishing seasons). Model equations are provided in Appendix A.
2) After estimating the abundances and parameters, run the forecast function (at the standard deviation phase of the ADMB optimization).
a) Randomize the recruitment:

Random selection of model estimated recruits for 1987 to 2012 was done as follows:

$$
\begin{equation*}
R_{i}=e^{\left.\left[\operatorname{logMeanRec}+r e c_{\operatorname{dev}(1987+u n i f o r m} \text { random error selected year incrment }\right)\right]} \tag{F.1}
\end{equation*}
$$

where $\mathrm{i}=2$ to 30 years
b) Randomize the abundance:

The lognormal random error to abundance is added in the following steps:
We first scaled the standard error based on the terminal year abundance (number of crabs) on its standard error (i.e., CV= $\frac{\text { Std.Error of terminal year mature male abundance }}{\text { terminal year mature male abundance }}$ ). Then we added the lognormal random error to abundance as follows:
$N_{i}=N_{i} e^{\varepsilon_{i}-\frac{\sigma_{\varepsilon}^{2}}{2}}$
where $\sigma_{\varepsilon}=\frac{\text { Std.Error of terminal year mature male abundance }}{\text { terminal year mature male abundance }}$
$\mathrm{N}=$ abundance, and $i=$ projection year.
The scaled standard error estimates (CV) are:
Scenario 17_0:
WAG: $\sigma_{\varepsilon}=0.18108$
EAG: $\quad \sigma_{\varepsilon}=0.18726$
Scenario 17_0d:
WAG: $\sigma_{\varepsilon}=0.23771$
EAG: $\quad \sigma_{\varepsilon}=0.23674$
3. Projection.

Two scenarios of fishing mortality for the directed pot fishery were used in the projections under Tier 3 control rule (i.e., Federal overfishing control rule):
i) No directed fishery. This was used as a base projection.
ii) $F_{35 \%}$. This is the maximum fishing mortality allowed under the current Tier 3 overfishing definitions.

The groundfish bycatch mortality was kept constant at the last 10-year mean fishing mortality level.

Each scenario was replicated 100 times and projections made over 30 years beginning in 2016
At each time step in the future:
a) Calculated legal male biomass (LMB) and mature male biomass (MMB).
b) Calculated the overfishing level total catch (OFL), acceptable biological catch (ABC), retained catch (RETC), and catch-per-unit effort (CPUE) indices using the Tier 3 OFL control rule.
c) Implemented the fishery under Tier 3 OFL control rule and removed the OFL catch from the simulated population.
d) Drew new recruitment numbers from historical distribution.
e) Updated the number-at-length.
4) Repeated step- 3 for 30 years into the future.
5) Repeated steps 3 and 4 for 100s of Monte Carlo trials, randomizing recruitment and abundance.
6) Used the annual distribution of simulated OFL catch, ABC catch, RETC, CPUE, LMB, and MMB to calculate performance statistics:
a) Median and mean annual MMB, LMB, OFL, ABC, RETC, and CPUE with standard errors and $95 \%$ confidence limits (by Efron's and Tibshirani's (1986) method: $2.5 \%$ and $97.5 \%$ percentile points).
b) Probability that the median MMB remains above the threshold reference points ( $0.25 \mathrm{MMB}_{35 \%}$, 2016), median ABC and median OFL exceeding $\mathrm{ABC}_{2016}$ and $\mathrm{OFL}_{2016}$ respectively during the $30-\mathrm{yr}$ projection period. The subscript 2016 refers to estimates by the respective assessment model scenarios.

The state harvest control rule simulation procedures are under development; therefore, we are not presenting any results of the state harvest strategy in this report.

## Results

The simulations compared the projection outputs for 17_0 and 17_0d scenarios and also investigated the probability of the stock being overfished (median MMB $<0.25 \mathrm{MMB}_{35 \%, 2016}$ ) and overfishing occurred [i.e., median OFL catch (i.e., median total catch under FOFL) exceeded $\mathrm{OFL}_{2016}$ or $\mathrm{ABC}_{2016}$ estimates] during the 30 yr projection time horizon. The standard deviation of the total catch (OFL), retained catch (RETC), and CPUE are provided to assess the variability of the harvest under Tier 3 control rule.

We provide the results in the subsequent tables for the Tier 3 control rule for both scenarios. Tables F. 1 and F. 2 compare the $30-\mathrm{yr}$ projected OFL catches with that of the model estimated OFL and ABC and provide the probability of overfishing and overfished under Tier 3 control rule for 17_0 and 17_0d scenarios for EAG and WAG, respectively. Subsequent tables (Tables F. 3 to F.14) provide the mean, median, standard deviation, and $95 \%$ confidence intervals for projected OFL, ABC, RETC, CPUE, LMB, and MMB during time horizon. We can make the following general conclusion from the simulation results:

If the Tier 3 control rule were directly applied as the harvest strategy, the probability of median MMB declining below the threshold (overfished) would be zero for both scenarios for EAG and WAG. However, probability of median OFL (total) catch exceeding ABC would be 0.067 for scenario 17_0, but zero for scenario 17_0d for EAG. On the other hand, probability of median OFL exceeding ABC would be 1.0 for both scenarios (17_0 and 17_0d) for WAG (Tables F. 1 and F.2).

## Reference

Efron, B., and Tibshirani, R. 1986. Bootstrap methods for standard errors, confidence intervals, and other measures of statistical accuracy. Statistical Science, 1(1): 54-75.
M.S.M. Siddeek, J. Zheng, C. Siddon, B. Daly. 2017. Aleutian Islands golden king crab (Lithodes aequispinus) model-based stock assessment in Spring 2017. Draft report for the May 2017 CPT meeting, Juneau.

Table F.1. Comparison of projected median OFL (i.e., total catch under Tire 3 FofL) with OFL 2016 and $\mathrm{ABC}_{2016}$ in metric tons ( t ) for scenario (Sc) 17_0 and 17_0d with $\mathrm{F}=\mathrm{F}_{35 \%}\left(0.64 \mathrm{yr}^{-1}\right)$ and $\mathrm{F}=0$ for EAG. Probability of projected median OFL exceeding OFL 2016 and $\mathrm{ABC}_{2016}$ and projected median MMB ( t ) depleting below the threshold $\mathrm{MMB}_{2016}$ are also listed. Thresh ${ }_{2016}=$ threshold MMB in 2016.

| Projection Year | Sc17_0 |  |  | Sc 17_0, F=0 |  |  | Sc 17_0, F=F $\mathrm{F}_{35}$ |  |  | Sc17_0d |  |  | Sc 17_0d, F=0 |  |  | Sc 17_0d, F=F ${ }_{35 \%}$ |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\mathrm{OFL}_{2016}$ | $\mathrm{ABC}_{2016}$ | Thresh 2016 | OFL | ABC | MMB | OFL | ABC | MMB | OFL 2016 | $\mathrm{ABC}_{2016}$ | Thresh 2016 | OFL | ABC | MMB | OFL | ABC | MMB |
| 2016 | 3,918 | 2,938 | 1738.615 | 1.974 | 1.480 | 14,832 | 3,359 | 2,519 | 11,445 | 2,481 | 1,860 | 1672.03 | 1.672 | 1.254 | 10,201 | 2,262 | 1,696 | 7,922 |
| 2017 |  |  |  | 2.157 | 1.618 | 16,281 | 3,194 | 2,395 | 10,017 |  |  |  | 1.911 | 1.434 | 11,708 | 2,170 | 1,628 | 7,470 |
| 2018 |  |  |  | 2.252 | 1.689 | 17,228 | 2,909 | 2,182 | 8,790 |  |  |  | 2.090 | 1.567 | 13,021 | 2,074 | 1,555 | 7,191 |
| 2019 |  |  |  | 2.361 | 1.771 | 17,851 | 2,559 | 1,919 | 7,889 |  |  |  | 2.272 | 1.704 | 13,957 | 1,975 | 1,481 | 6,957 |
| 2020 |  |  |  | 2.407 | 1.805 | 18,401 | 2,256 | 1,692 | 7,416 |  |  |  | 2.376 | 1.782 | 15,000 | 1,907 | 1,430 | 6,804 |
| 2021 |  |  |  | 2.460 | 1.845 | 18,660 | 2,073 | 1,555 | 7,092 |  |  |  | 2.506 | 1.879 | 15,573 | 1,884 | 1,413 | 6,716 |
| 2022 |  |  |  | 2.471 | 1.854 | 18,959 | 1,951 | 1,463 | 7,064 |  |  |  | 2.598 | 1.949 | 16,324 | 1,832 | 1,374 | 6,736 |
| 2023 |  |  |  | 2.507 | 1.880 | 19,082 | 1,937 | 1,453 | 7,014 |  |  |  | 2.674 | 2.005 | 16,848 | 1,831 | 1,373 | 6,747 |
| 2024 |  |  |  | 2.548 | 1.911 | 19,380 | 1,899 | 1,424 | 7,019 |  |  |  | 2.763 | 2.072 | 17,381 | 1,819 | 1,365 | 6,705 |
| 2025 |  |  |  | 2.547 | 1.910 | 19,483 | 1,904 | 1,428 | 6,877 |  |  |  | 2.805 | 2.104 | 17,715 | 1,835 | 1,376 | 6,667 |
| 2026 |  |  |  | 2.544 | 1.908 | 19,434 | 1,893 | 1,420 | 6,884 |  |  |  | 2.844 | 2.133 | 17,969 | 1,831 | 1,373 | 6,653 |
| 2027 |  |  |  | 2.560 | 1.920 | 19,527 | 1,877 | 1,407 | 6,881 |  |  |  | 2.877 | 2.158 | 18,051 | 1,814 | 1,361 | 6,675 |
| 2028 |  |  |  | 2.587 | 1.940 | 19,632 | 1,871 | 1,403 | 6,828 |  |  |  | 2.923 | 2.192 | 18,398 | 1,813 | 1,360 | 6,723 |
| 2029 |  |  |  | 2.563 | 1.923 | 19,701 | 1,872 | 1,404 | 6,882 |  |  |  | 2.930 | 2.198 | 18,507 | 1,817 | 1,363 | 6,657 |
| 2030 |  |  |  | 2.571 | 1.928 | 19,630 | 1,883 | 1,413 | 6,930 |  |  |  | 2.947 | 2.210 | 18,649 | 1,818 | 1,363 | 6,743 |
| 2031 |  |  |  | 2.576 | 1.932 | 19,694 | 1,883 | 1,412 | 6,881 |  |  |  | 2.962 | 2.221 | 18,770 | 1,814 | 1,360 | 6,604 |
| 2032 |  |  |  | 2.564 | 1.923 | 19,639 | 1,889 | 1,417 | 6,845 |  |  |  | 2.968 | 2.226 | 18,733 | 1,836 | 1,377 | 6,656 |
| 2033 |  |  |  | 2.577 | 1.933 | 19,657 | 1,882 | 1,411 | 6,913 |  |  |  | 2.974 | 2.231 | 18,869 | 1,814 | 1,360 | 6,682 |
| 2034 |  |  |  | 2.571 | 1.928 | 19,676 | 1,875 | 1,406 | 6,945 |  |  |  | 2.978 | 2.233 | 18,846 | 1,818 | 1,364 | 6,674 |
| 2035 |  |  |  | 2.585 | 1.939 | 19,709 | 1,885 | 1,414 | 6,921 |  |  |  | 2.991 | 2.243 | 18,909 | 1,816 | 1,362 | 6,664 |
| 2036 |  |  |  | 2.582 | 1.936 | 19,710 | 1,881 | 1,411 | 6,920 |  |  |  | 2.985 | 2.239 | 19,008 | 1,806 | 1,355 | 6,690 |
| 2037 |  |  |  | 2.595 | 1.947 | 19,715 | 1,885 | 1,413 | 6,895 |  |  |  | 3.002 | 2.252 | 18,942 | 1,838 | 1,378 | 6,644 |
| 2038 |  |  |  | 2.596 | 1.947 | 19,804 | 1,880 | 1,410 | 6,983 |  |  |  | 3.017 | 2.262 | 19,059 | 1,812 | 1,359 | 6,768 |
| 2039 |  |  |  | 2.615 | 1.962 | 19,894 | 1,896 | 1,422 | 7,008 |  |  |  | 3.048 | 2.286 | 19,206 | 1,823 | 1,367 | 6,823 |
| 2040 |  |  |  | 2.617 | 1.962 | 19,979 | 1,905 | 1,429 | 7,025 |  |  |  | 3.038 | 2.278 | 19,305 | 1,831 | 1,374 | 6,870 |
| 2041 |  |  |  | 2.629 | 1.972 | 20,013 | 1,924 | 1,443 | 7,098 |  |  |  | 3.065 | 2.299 | 19,363 | 1,856 | 1,392 | 6,856 |
| 2042 |  |  |  | 2.623 | 1.967 | 20,115 | 1,915 | 1,436 | 6,976 |  |  |  | 3.045 | 2.284 | 19,400 | 1,857 | 1,393 | 6,716 |
| 2043 |  |  |  | 2.613 | 1.959 | 20,005 | 1,925 | 1,444 | 6,861 |  |  |  | 3.060 | 2.295 | 19,346 | 1,851 | 1,389 | 6,685 |
| 2044 |  |  |  | 2.605 | 1.953 | 19,920 | 1,898 | 1,423 | 6,871 |  |  |  | 3.060 | 2.295 | 19,359 | 1,842 | 1,381 | 6,695 |
| 2045 |  |  |  | 2.606 | 1.955 | 19,945 | 1,871 | 1,404 | 6,876 |  |  |  | 3.045 | 2.284 | 19,300 | 1,810 | 1,358 | 6,653 |
| Prob OFL> |  |  |  | 0 |  |  | 0.067 |  |  |  |  |  | 0 |  |  | 0 |  |  |
| $\mathrm{ABC}_{2016}$ |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Prob OFL> |  |  |  | 0 |  |  | 0 |  |  |  |  |  | 0 |  |  | 0 |  |  |
| $\mathrm{OFL}_{2016}$ |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Prob MMB< |  |  |  |  |  | 0 |  |  | 0 |  |  |  |  |  | 0 |  |  | 0 |
| Thresh 2016 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |

Table F.2. Comparison of projected median OFL (i.e., total catch under Tire 3 FofL) with OFL 2016 and $\mathrm{ABC}_{2016}$ in metric tons ( t ) for scenario (Sc) 17_0 and 17_0d with $\mathrm{F}=\mathrm{F}_{35} \%\left(0.6 \mathrm{yr}^{-1}\right.$ and $0.68 \mathrm{yr}^{-1}$ for $\mathrm{Sc} 17 \_0$ and $\mathrm{Sc} 17 \_0 \mathrm{~d}$, respectively) and $\mathrm{F}=0$ for WAG.

Probability of projected median OFL exceeding $\mathrm{OFL}_{2016}$ and $\mathrm{ABC}_{2016}$ and projected median MMB ( t ) depleting below the threshold $\mathrm{MMB}_{2016}$ are also listed. Thresh ${ }_{2016}=$ threshold MMB in 2016.

| Projection Year | Sc17_0 |  |  | Sc 17_0, F=0 |  |  | Sc 17_0, $\mathrm{F}=\mathrm{F}_{35 \%}$ |  |  | Sc17_0d |  |  | Sc 17_0d, F=0 |  |  | Sc 17_0d, $\mathrm{F}=\mathrm{F}_{35}{ }^{\text {\% }}$ |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\mathrm{OFL}_{2016}$ | $\mathrm{ABC}_{2016}$ | Thresh 2016 | OFL | ABC | MMB | OFL | ABC | MMB | $\mathrm{OFL}_{2016}$ | $\mathrm{ABC}_{2016}$ | Thresh $_{2016}$ | OFL | ABC | MMB | OFL | ABC | MMB |
| 2016 | 1,597 | 1,197 | 1,284.485 | 3.367 | 2.526 | 7,243 | 1,297 | 973 | 5,934 | 1,482 | 1,112 | 1,276.905 | 3.434 | 2.576 | 7,015 | 1,119 | 840 | 5,886 |
| 2017 |  |  |  | 3.955 | 2.966 | 8,745 | 1,489 | 1,116 | 6,057 |  |  |  | 4.065 | 3.048 | 8,565 | 1,416 | 1,062 | 6,132 |
| 2018 |  |  |  | 4.335 | 3.251 | 9,814 | 1,614 | 1,211 | 5,778 |  |  |  | 4.491 | 3.368 | 9,717 | 1,625 | 1,219 | 5,811 |
| 2019 |  |  |  | 4.765 | 3.574 | 10,647 | 1,605 | 1,204 | 5,550 |  |  |  | 4.945 | 3.709 | 10,565 | 1,636 | 1,227 | 5,584 |
| 2020 |  |  |  | 5.091 | 3.818 | 11,457 | 1,525 | 1,143 | 5,425 |  |  |  | 5.298 | 3.973 | 11,415 | 1,548 | 1,161 | 5,475 |
| 2021 |  |  |  | 5.369 | 4.027 | 12,153 | 1,475 | 1,106 | 5,229 |  |  |  | 5.581 | 4.186 | 12,089 | 1,483 | 1,113 | 5,243 |
| 2022 |  |  |  | 5.566 | 4.174 | 12,666 | 1,443 | 1,082 | 5,195 |  |  |  | 5.811 | 4.358 | 12,637 | 1,449 | 1,087 | 5,138 |
| 2023 |  |  |  | 5.786 | 4.339 | 13,127 | 1,407 | 1,055 | 5,175 |  |  |  | 6.032 | 4.524 | 13,059 | 1,406 | 1,055 | 5,131 |
| 2024 |  |  |  | 5.883 | 4.413 | 13,532 | 1,404 | 1,053 | 5,277 |  |  |  | 6.121 | 4.591 | 13,437 | 1,394 | 1,045 | 5,238 |
| 2025 |  |  |  | 5.999 | 4.499 | 13,725 | 1,410 | 1,057 | 5,272 |  |  |  | 6.266 | 4.699 | 13,633 | 1,398 | 1,048 | 5,214 |
| 2026 |  |  |  | 6.074 | 4.555 | 13,941 | 1,410 | 1,057 | 5,192 |  |  |  | 6.326 | 4.745 | 13,888 | 1,409 | 1,057 | 5,152 |
| 2027 |  |  |  | 6.136 | 4.602 | 14,080 | 1,415 | 1,061 | 5,145 |  |  |  | 6.400 | 4.800 | 14,001 | 1,411 | 1,058 | 5,122 |
| 2028 |  |  |  | 6.179 | 4.634 | 14,154 | 1,395 | 1,047 | 5,121 |  |  |  | 6.447 | 4.835 | 14,104 | 1,391 | 1,043 | 5,111 |
| 2029 |  |  |  | 6.216 | 4.662 | 14,253 | 1,381 | 1,036 | 5,136 |  |  |  | 6.470 | 4.853 | 14,194 | 1,384 | 1,038 | 5,117 |
| 2030 |  |  |  | 6.234 | 4.676 | 14,348 | 1,384 | 1,038 | 5,117 |  |  |  | 6.500 | 4.875 | 14,230 | 1,386 | 1,039 | 5,106 |
| 2031 |  |  |  | 6.318 | 4.738 | 14,406 | 1,391 | 1,043 | 5,148 |  |  |  | 6.608 | 4.956 | 14,359 | 1,381 | 1,036 | 5,102 |
| 2032 |  |  |  | 6.333 | 4.749 | 14,574 | 1,389 | 1,041 | 5,244 |  |  |  | 6.602 | 4.952 | 14,512 | 1,384 | 1,038 | 5,207 |
| 2033 |  |  |  | 6.330 | 4.747 | 14,570 | 1,402 | 1,051 | 5,189 |  |  |  | 6.611 | 4.958 | 14,461 | 1,393 | 1,045 | 5,189 |
| 2034 |  |  |  | 6.389 | 4.791 | 14,631 | 1,406 | 1,054 | 5,204 |  |  |  | 6.650 | 4.987 | 14,508 | 1,409 | 1,057 | 5,234 |
| 2035 |  |  |  | 6.412 | 4.809 | 14,695 | 1,406 | 1,054 | 5,213 |  |  |  | 6.676 | 5.007 | 14,590 | 1,401 | 1,051 | 5,201 |
| 2036 |  |  |  | 6.410 | 4.808 | 14,774 | 1,401 | 1,051 | 5,273 |  |  |  | 6.709 | 5.032 | 14,728 | 1,408 | 1,056 | 5,256 |
| 2037 |  |  |  | 6.452 | 4.839 | 14,750 | 1,411 | 1,058 | 5,174 |  |  |  | 6.724 | 5.043 | 14,711 | 1,413 | 1,059 | 5,146 |
| 2038 |  |  |  | 6.468 | 4.851 | 14,872 | 1,403 | 1,052 | 5,177 |  |  |  | 6.714 | 5.035 | 14,802 | 1,406 | 1,055 | 5,135 |
| 2039 |  |  |  | 6.424 | 4.818 | 14,825 | 1,407 | 1,055 | 5,188 |  |  |  | 6.674 | 5.005 | 14,730 | 1,395 | 1,046 | 5,110 |
| 2040 |  |  |  | 6.431 | 4.823 | 14,754 | 1,399 | 1,049 | 5,160 |  |  |  | 6.696 | 5.022 | 14,647 | 1,392 | 1,044 | 5,151 |
| 2041 |  |  |  | 6.428 | 4.821 | 14,756 | 1,390 | 1,043 | 5,180 |  |  |  | 6.698 | 5.023 | 14,659 | 1,392 | 1,044 | 5,121 |
| 2042 |  |  |  | 6.430 | 4.823 | 14,814 | 1,393 | 1,045 | 5,150 |  |  |  | 6.691 | 5.018 | 14,745 | 1,394 | 1,046 | 5,164 |
| 2043 |  |  |  | 6.458 | 4.843 | 14,766 | 1,392 | 1,044 | 5,212 |  |  |  | 6.716 | 5.037 | 14,703 | 1,390 | 1,043 | 5,223 |
| 2044 |  |  |  | 6.467 | 4.850 | 14,828 | 1,391 | 1,043 | 5,266 |  |  |  | 6.770 | 5.078 | 14,726 | 1,394 | 1,045 | 5,223 |
| 2045 |  |  |  | 6.507 | 4.880 | 14,951 | 1,407 | 1,056 | 5,274 |  |  |  | 6.795 | 5.096 | 14,871 | 1,408 | 1,056 | 5,229 |
| Prob OFL> |  |  |  | 0 |  |  | 1.000 |  |  |  |  |  | 0 |  |  | 1.000 |  |  |
| $\mathrm{ABC}_{2016}$ |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Prob OFL> |  |  |  | 0 |  |  | 0.067 |  |  |  |  |  | 0 |  |  | 0.133 |  |  |
| OFL 2016 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Prob MMB< |  |  |  |  |  | 0 |  |  | 0 |  |  |  |  |  | 0 |  |  | 0 |

Table F.3. Projected mean, median, and standard deviation (SD) of legal male (LMB) and mature male (MMB) biomass ( t ) with $95 \%$ confidence limits under no directed fishery (top) and under $\mathrm{F}_{35 \%}$ (bottom) harvest control rule for scenario 17_0 for EAG, 2016-2045. The top table provides a base projection scenario with no directed fishery.

| No Directed Fishery |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | Mean <br> LMB | Median LMB | $\begin{gathered} \text { SD } \\ \text { LMB } \end{gathered}$ | $95 \%$ <br> Lower <br> Limit | 95\% <br> Upper <br> Limit | Mean <br> MMB | Median MMB | $\begin{gathered} \text { SD } \\ \text { MMB } \end{gathered}$ | $95 \%$ <br> Lower <br> Limit | $\begin{aligned} & \hline 95 \% \\ & \text { Upper } \\ & \text { Limit } \end{aligned}$ |
| 2016 | 9,433 | 9,400 | 1,737 | 6,595 | 13,218 | 14,883 | 14,832 | 2,740 | 10,405 | 20,855 |
| 2017 | 12,587 | 12,543 | 2,317 | 8,800 | 17,637 | 16,313 | 16,281 | 2,879 | 11,472 | 22,419 |
| 2018 | 14,564 | 14,507 | 2,665 | 10,191 | 20,352 | 17,318 | 17,228 | 2,765 | 12,462 | 22,966 |
| 2019 | 15,658 | 15,592 | 2,660 | 11,087 | 21,174 | 17,972 | 17,851 | 2,547 | 13,564 | 23,105 |
| 2020 | 16,274 | 16,134 | 2,483 | 11,898 | 21,322 | 18,445 | 18,401 | 2,267 | 14,667 | 22,953 |
| 2021 | 16,686 | 16,579 | 2,208 | 12,937 | 21,095 | 18,764 | 18,660 | 2,036 | 15,385 | 22,643 |
| 2022 | 16,951 | 16,865 | 1,936 | 13,656 | 20,784 | 19,051 | 18,959 | 1,885 | 16,157 | 22,485 |
| 2023 | 17,171 | 17,116 | 1,753 | 14,400 | 20,388 | 19,259 | 19,082 | 1,742 | 16,473 | 22,652 |
| 2024 | 17,366 | 17,243 | 1,618 | 14,781 | 20,347 | 19,375 | 19,380 | 1,619 | 16,511 | 22,420 |
| 2025 | 17,493 | 17,327 | 1,513 | 14,933 | 20,420 | 19,432 | 19,483 | 1,497 | 16,608 | 22,483 |
| 2026 | 17,553 | 17,605 | 1,405 | 14,934 | 20,344 | 19,470 | 19,434 | 1,389 | 16,841 | 22,141 |
| 2027 | 17,573 | 17,560 | 1,309 | 15,118 | 20,291 | 19,532 | 19,527 | 1,273 | 17,123 | 21,747 |
| 2028 | 17,600 | 17,568 | 1,213 | 15,298 | 19,926 | 19,598 | 19,632 | 1,168 | 17,154 | 21,722 |
| 2029 | 17,654 | 17,645 | 1,103 | 15,492 | 19,590 | 19,630 | 19,701 | 1,132 | 17,275 | 21,915 |
| 2030 | 17,699 | 17,785 | 1,037 | 15,537 | 19,733 | 19,643 | 19,630 | 1,088 | 17,308 | 21,718 |
| 2031 | 17,718 | 17,723 | 1,013 | 15,618 | 19,809 | 19,658 | 19,694 | 1,027 | 17,590 | 21,531 |
| 2032 | 17,724 | 17,743 | 963 | 15,725 | 19,467 | 19,674 | 19,639 | 1,006 | 17,599 | 21,354 |
| 2033 | 17,738 | 17,788 | 922 | 15,861 | 19,497 | 19,713 | 19,657 | 998 | 17,616 | 21,727 |
| 2034 | 17,746 | 17,646 | 914 | 15,906 | 19,499 | 19,778 | 19,676 | 1,015 | 17,712 | 21,734 |
| 2035 | 17,810 | 17,732 | 923 | 15,948 | 19,565 | 19,789 | 19,709 | 1,037 | 17,929 | 21,948 |
| 2036 | 17,838 | 17,778 | 945 | 16,047 | 19,809 | 19,818 | 19,710 | 1,088 | 17,836 | 21,972 |
| 2037 | 17,864 | 17,785 | 978 | 16,167 | 19,867 | 19,833 | 19,715 | 1,110 | 17,931 | 22,094 |
| 2038 | 17,875 | 17,752 | 1,027 | 15,999 | 19,970 | 19,870 | 19,804 | 1,100 | 17,973 | 22,222 |
| 2039 | 17,901 | 17,835 | 1,021 | 16,166 | 19,971 | 19,898 | 19,894 | 1,152 | 17,873 | 22,330 |
| 2040 | 17,928 | 17,902 | 1,033 | 16,176 | 20,238 | 19,951 | 19,979 | 1,214 | 18,010 | 22,393 |
| 2041 | 17,961 | 17,984 | 1,092 | 16,044 | 20,048 | 20,007 | 20,013 | 1,207 | 18,006 | 22,526 |
| 2042 | 18,018 | 18,011 | 1,128 | 16,214 | 20,399 | 20,012 | 20,115 | 1,153 | 17,988 | 22,093 |
| 2043 | 18,053 | 18,078 | 1,099 | 16,200 | 20,182 | 19,987 | 20,005 | 1,115 | 17,941 | 22,148 |
| 2044 | 18,042 | 18,095 | 1,051 | 16,167 | 20,096 | 19,952 | 19,920 | 1,100 | 18,168 | 21,959 |
| 2045 | 18,012 | 18,009 | 1,020 | 16,199 | 19,887 | 19,941 | 19,945 | 1,114 | 18,181 | 22,370 |


| $\mathrm{F}_{35 \%}\left(0.64 \mathrm{yr}^{-1}\right)$ |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | $\begin{aligned} & \text { Mean } \\ & \text { LMB } \\ & \hline \end{aligned}$ | Median LMB | $\begin{gathered} \mathrm{SD} \\ \mathrm{LMB} \\ \hline \end{gathered}$ | 95\% Lower <br> Limit | 95\% <br> Upper <br> Limit | $\begin{aligned} & \text { Mean } \\ & \text { MMB } \\ & \hline \end{aligned}$ | Median MMB | $\begin{gathered} \text { SD } \\ \text { MMB } \\ \hline \end{gathered}$ | 95\% <br> Lower <br> Limit | 95\% <br> Upper <br> Limit |
| 2016 | 9,433 | 9,400 | 1,737 | 6,595 | 13,218 | 11,485 | 11,445 | 2,114 | 8,029 | 16,093 |
| 2017 | 9,210 | 9,178 | 1,696 | 6,439 | 12,905 | 10,054 | 10,017 | 1,731 | 7,096 | 13,655 |
| 2018 | 8,346 | 8,309 | 1,518 | 5,842 | 11,637 | 8,841 | 8,790 | 1,261 | 6,536 | 11,281 |
| 2019 | 7,275 | 7,264 | 1,135 | 5,304 | 9,522 | 7,968 | 7,889 | 885 | 6,497 | 9,837 |
| 2020 | 6,435 | 6,374 | 795 | 5,123 | 8,085 | 7,466 | 7,416 | 677 | 6,471 | 8,855 |
| 2021 | 5,943 | 5,923 | 561 | 5,148 | 7,087 | 7,180 | 7,092 | 592 | 6,206 | 8,342 |
| 2022 | 5,665 | 5,570 | 445 | 5,095 | 6,591 | 7,095 | 7,064 | 643 | 6,092 | 8,546 |
| 2023 | 5,563 | 5,509 | 438 | 4,915 | 6,516 | 7,063 | 7,014 | 658 | 6,042 | 8,493 |
| 2024 | 5,548 | 5,415 | 473 | 4,884 | 6,721 | 7,005 | 7,019 | 649 | 5,976 | 8,352 |
| 2025 | 5,526 | 5,456 | 459 | 4,870 | 6,508 | 6,939 | 6,877 | 635 | 5,932 | 8,313 |
| 2026 | 5,481 | 5,417 | 453 | 4,832 | 6,552 | 6,901 | 6,884 | 594 | 5,933 | 8,260 |
| 2027 | 5,434 | 5,383 | 431 | 4,826 | 6,452 | 6,919 | 6,881 | 562 | 6,129 | 8,227 |
| 2028 | 5,421 | 5,350 | 416 | 4,823 | 6,491 | 6,955 | 6,828 | 543 | 6,123 | 7,865 |
| 2029 | 5,449 | 5,348 | 388 | 4,909 | 6,390 | 6,957 | 6,882 | 555 | 6,067 | 8,016 |
| 2030 | 5,466 | 5,402 | 388 | 4,929 | 6,245 | 6,939 | 6,930 | 511 | 6,098 | 7,974 |
| 2031 | 5,455 | 5,383 | 385 | 4,874 | 6,272 | 6,931 | 6,881 | 471 | 6,104 | 7,756 |
| 2032 | 5,439 | 5,407 | 329 | 4,925 | 6,037 | 6,933 | 6,845 | 539 | 6,057 | 8,148 |
| 2033 | 5,446 | 5,388 | 348 | 4,928 | 6,245 | 6,964 | 6,913 | 582 | 6,030 | 8,224 |
| 2034 | 5,447 | 5,365 | 406 | 4,875 | 6,442 | 7,020 | 6,945 | 611 | 6,111 | 8,336 |
| 2035 | 5,500 | 5,377 | 428 | 4,924 | 6,437 | 7,007 | 6,921 | 614 | 6,049 | 8,323 |
| 2036 | 5,503 | 5,378 | 446 | 4,933 | 6,486 | 7,011 | 6,920 | 626 | 6,025 | 8,452 |
| 2037 | 5,507 | 5,386 | 441 | 4,901 | 6,459 | 7,005 | 6,895 | 634 | 6,041 | 8,437 |
| 2038 | 5,498 | 5,376 | 456 | 4,859 | 6,683 | 7,026 | 6,983 | 618 | 6,101 | 8,368 |
| 2039 | 5,507 | 5,430 | 450 | 4,875 | 6,508 | 7,038 | 7,008 | 630 | 6,076 | 8,276 |
| 2040 | 5,520 | 5,453 | 432 | 4,927 | 6,452 | 7,074 | 7,025 | 679 | 6,090 | 8,398 |
| 2041 | 5,536 | 5,493 | 474 | 4,887 | 6,436 | 7,109 | 7,098 | 665 | 6,100 | 8,294 |
| 2042 | 5,570 | 5,488 | 492 | 4,881 | 6,509 | 7,083 | 6,976 | 628 | 6,086 | 8,478 |
| 2043 | 5,571 | 5,505 | 465 | 4,884 | 6,562 | 7,027 | 6,861 | 622 | 6,085 | 8,636 |
| 2044 | 5,534 | 5,426 | 452 | 4,899 | 6,669 | 6,982 | 6,871 | 612 | 6,077 | 8,428 |
| 2045 | 5,498 | 5,358 | 448 | 4,919 | 6,586 | 6,978 | 6,876 | 605 | 6,036 | 8,197 |

Table F.4. Projected mean, median, and standard deviation (SD) of total catch (OFL) and retained catch (RETC) in biomass ( t ) with $95 \%$ confidence limits under no directed fishery (top) and under $\mathrm{F}_{35 \%}$ (bottom) harvest control rule for scenario 17_0 for EAG, 2016-2045.

| No Directed Fishery |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | $\begin{gathered} \text { Mean } \\ \text { OFL } \end{gathered}$ | Median OFL | $\begin{gathered} \text { SD } \\ \text { OFL } \end{gathered}$ | $\begin{gathered} 95 \% \\ \text { Lower } \\ \text { Limit } \\ \hline \end{gathered}$ | 95\% <br> Upper <br> Limit | Mean <br> RETC | Median RETC | $\begin{gathered} \text { SD } \\ \text { RETC } \end{gathered}$ | $\begin{gathered} \hline 95 \% \\ \text { Lower } \\ \text { Limit } \\ \hline \end{gathered}$ | 95\% <br> Upper <br> Limit |
| 2016 | 1.981 | 1.974 | 0.365 | 1.385 | 2.775 | 0 | 0 | 0 | 0 | 0 |
| 2017 | 2.163 | 2.157 | 0.363 | 1.533 | 2.907 | 0 | 0 | 0 | 0 | 0 |
| 2018 | 2.279 | 2.252 | 0.349 | 1.667 | 2.995 | 0 | 0 | 0 | 0 | 0 |
| 2019 | 2.367 | 2.361 | 0.316 | 1.841 | 2.993 | 0 | 0 | 0 | 0 | 0 |
| 2020 | 2.419 | 2.407 | 0.283 | 1.932 | 2.973 | 0 | 0 | 0 | 0 | 0 |
| 2021 | 2.466 | 2.460 | 0.262 | 2.048 | 2.943 | 0 | 0 | 0 | 0 | 0 |
| 2022 | 2.501 | 2.471 | 0.239 | 2.127 | 2.955 | 0 | 0 | 0 | 0 | 0 |
| 2023 | 2.523 | 2.507 | 0.223 | 2.157 | 2.959 | 0 | 0 | 0 | 0 | 0 |
| 2024 | 2.536 | 2.548 | 0.205 | 2.158 | 2.930 | 0 | 0 | 0 | 0 | 0 |
| 2025 | 2.542 | 2.547 | 0.190 | 2.177 | 2.951 | 0 | 0 | 0 | 0 | 0 |
| 2026 | 2.550 | 2.544 | 0.176 | 2.218 | 2.878 | 0 | 0 | 0 | 0 | 0 |
| 2027 | 2.560 | 2.560 | 0.159 | 2.240 | 2.837 | 0 | 0 | 0 | 0 | 0 |
| 2028 | 2.566 | 2.587 | 0.151 | 2.246 | 2.852 | 0 | 0 | 0 | 0 | 0 |
| 2029 | 2.569 | 2.563 | 0.147 | 2.263 | 2.862 | 0 | 0 | 0 | 0 | 0 |
| 2030 | 2.570 | 2.571 | 0.138 | 2.284 | 2.818 | 0 | 0 | 0 | 0 | 0 |
| 2031 | 2.574 | 2.576 | 0.134 | 2.296 | 2.829 | 0 | 0 | 0 | 0 | 0 |
| 2032 | 2.574 | 2.564 | 0.130 | 2.295 | 2.833 | 0 | 0 | 0 | 0 | 0 |
| 2033 | 2.587 | 2.577 | 0.131 | 2.319 | 2.832 | 0 | 0 | 0 | 0 | 0 |
| 2034 | 2.586 | 2.571 | 0.132 | 2.325 | 2.870 | 0 | 0 | 0 | 0 | 0 |
| 2035 | 2.592 | 2.585 | 0.137 | 2.352 | 2.859 | 0 | 0 | 0 | 0 | 0 |
| 2036 | 2.593 | 2.582 | 0.145 | 2.334 | 2.881 | 0 | 0 | 0 | 0 | 0 |
| 2037 | 2.598 | 2.595 | 0.140 | 2.344 | 2.879 | 0 | 0 | 0 | 0 | 0 |
| 2038 | 2.601 | 2.596 | 0.145 | 2.347 | 2.915 | 0 | 0 | 0 | 0 | 0 |
| 2039 | 2.606 | 2.615 | 0.155 | 2.331 | 2.901 | 0 | 0 | 0 | 0 | 0 |
| 2040 | 2.615 | 2.617 | 0.158 | 2.367 | 2.952 | 0 | 0 | 0 | 0 | 0 |
| 2041 | 2.619 | 2.629 | 0.152 | 2.356 | 2.922 | 0 | 0 | 0 | 0 | 0 |
| 2042 | 2.616 | 2.623 | 0.147 | 2.357 | 2.909 | 0 | 0 | 0 | 0 | 0 |
| 2043 | 2.614 | 2.613 | 0.143 | 2.358 | 2.862 | 0 | 0 | 0 | 0 | 0 |
| 2044 | 2.609 | 2.605 | 0.144 | 2.385 | 2.930 | 0 | 0 | 0 | 0 | 0 |
| 2045 | 2.614 | 2.606 | 0.147 | 2.383 | 2.909 | 0 | 0 | 0 | 0 | 0 |


| $\mathrm{F}_{35 \%}\left(0.64 \mathrm{yr}^{-1}\right)$ |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | $\begin{gathered} \text { Mean } \\ \text { OFL } \end{gathered}$ | Median OFL | $\begin{gathered} \text { SD } \\ \text { OFL } \\ \hline \end{gathered}$ | 95\% <br> Lower <br> Limit | 95\% <br> Upper <br> Limit | $\begin{aligned} & \text { Mean } \\ & \text { RETC } \\ & \hline \end{aligned}$ | $\begin{gathered} \text { Median } \\ \text { RETC } \\ \hline \end{gathered}$ | $\begin{gathered} \text { SD } \\ \text { RETC } \\ \hline \end{gathered}$ | 95\% <br> Lower <br> Limit | 95\% <br> Upper <br> Limit |
| 2016 | 3,371 | 3,359 | 621 | 2,357 | 4,723 | 3,198 | 3,187 | 589 | 2,236 | 4,481 |
| 2017 | 3,205 | 3,194 | 589 | 2,241 | 4,488 | 3,051 | 3,042 | 564 | 2,134 | 4,278 |
| 2018 | 2,925 | 2,909 | 526 | 2,052 | 4,058 | 2,796 | 2,788 | 517 | 1,878 | 3,906 |
| 2019 | 2,566 | 2,559 | 401 | 1,877 | 3,372 | 2,447 | 2,449 | 405 | 1,714 | 3,247 |
| 2020 | 2,274 | 2,256 | 281 | 1,801 | 2,856 | 2,154 | 2,143 | 292 | 1,657 | 2,734 |
| 2021 | 2,092 | 2,073 | 196 | 1,813 | 2,500 | 1,964 | 1,964 | 213 | 1,641 | 2,390 |
| 2022 | 1,991 | 1,951 | 153 | 1,802 | 2,312 | 1,853 | 1,823 | 177 | 1,571 | 2,201 |
| 2023 | 1,949 | 1,937 | 145 | 1,740 | 2,266 | 1,806 | 1,791 | 170 | 1,504 | 2,137 |
| 2024 | 1,939 | 1,899 | 155 | 1,720 | 2,321 | 1,792 | 1,776 | 185 | 1,494 | 2,202 |
| 2025 | 1,931 | 1,904 | 152 | 1,718 | 2,268 | 1,781 | 1,781 | 184 | 1,469 | 2,142 |
| 2026 | 1,917 | 1,893 | 149 | 1,704 | 2,257 | 1,769 | 1,741 | 179 | 1,455 | 2,138 |
| 2027 | 1,902 | 1,877 | 142 | 1,700 | 2,231 | 1,758 | 1,753 | 165 | 1,493 | 2,108 |
| 2028 | 1,896 | 1,871 | 135 | 1,708 | 2,250 | 1,753 | 1,738 | 156 | 1,529 | 2,119 |
| 2029 | 1,904 | 1,872 | 127 | 1,732 | 2,219 | 1,760 | 1,734 | 151 | 1,504 | 2,097 |
| 2030 | 1,909 | 1,883 | 126 | 1,733 | 2,158 | 1,768 | 1,743 | 152 | 1,508 | 2,036 |
| 2031 | 1,906 | 1,883 | 125 | 1,720 | 2,172 | 1,766 | 1,756 | 149 | 1,510 | 2,060 |
| 2032 | 1,901 | 1,889 | 108 | 1,734 | 2,123 | 1,757 | 1,747 | 131 | 1,504 | 2,006 |
| 2033 | 1,903 | 1,882 | 112 | 1,737 | 2,156 | 1,759 | 1,741 | 136 | 1,529 | 2,031 |
| 2034 | 1,905 | 1,875 | 130 | 1,723 | 2,224 | 1,763 | 1,733 | 153 | 1,509 | 2,110 |
| 2035 | 1,919 | 1,885 | 140 | 1,730 | 2,230 | 1,777 | 1,752 | 163 | 1,496 | 2,108 |
| 2036 | 1,923 | 1,881 | 146 | 1,732 | 2,254 | 1,779 | 1,749 | 171 | 1,489 | 2,135 |
| 2037 | 1,924 | 1,885 | 146 | 1,728 | 2,243 | 1,780 | 1,773 | 170 | 1,507 | 2,128 |
| 2038 | 1,922 | 1,880 | 149 | 1,716 | 2,314 | 1,780 | 1,749 | 173 | 1,503 | 2,192 |
| 2039 | 1,924 | 1,896 | 148 | 1,720 | 2,253 | 1,779 | 1,763 | 174 | 1,503 | 2,133 |
| 2040 | 1,929 | 1,905 | 144 | 1,732 | 2,247 | 1,785 | 1,782 | 172 | 1,518 | 2,138 |
| 2041 | 1,935 | 1,924 | 154 | 1,729 | 2,230 | 1,793 | 1,808 | 178 | 1,522 | 2,105 |
| 2042 | 1,944 | 1,915 | 161 | 1,728 | 2,258 | 1,805 | 1,784 | 183 | 1,511 | 2,141 |
| 2043 | 1,945 | 1,925 | 154 | 1,726 | 2,263 | 1,803 | 1,782 | 177 | 1,515 | 2,135 |
| 2044 | 1,934 | 1,898 | 149 | 1,733 | 2,300 | 1,790 | 1,761 | 171 | 1,516 | 2,175 |
| 2045 | 1,922 | 1,871 | 147 | 1,731 | 2,289 | 1,777 | 1,749 | 170 | 1,518 | 2,174 |

Table F.5. Projected mean, median, and standard deviation (SD) of retained CPUE indices with $95 \%$ confidence limits under $\mathrm{F}_{35 \%}$ harvest control rule for scenario 17_0 for EAG, 2016-2045.

|  |  |  | $\mathrm{F}_{35 \%}\left(0.64 \mathrm{yr}^{-1}\right)$ |  |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | $95 \%$ <br> Mean |  |  |  | Median | SD | Lower <br> Upper |
| CPUE | CPUE | CPUE | Limit | Limit |  |  |  |  |  |
| 2016 | 0.997 | 0.993 | 0.183 | 0.697 | 1.396 |  |  |  |  |
| 2017 | 0.977 | 0.974 | 0.180 | 0.683 | 1.369 |  |  |  |  |
| 2018 | 0.894 | 0.890 | 0.162 | 0.631 | 1.245 |  |  |  |  |
| 2019 | 0.781 | 0.779 | 0.121 | 0.573 | 1.023 |  |  |  |  |
| 2020 | 0.691 | 0.684 | 0.084 | 0.553 | 0.866 |  |  |  |  |
| 2021 | 0.638 | 0.636 | 0.059 | 0.555 | 0.759 |  |  |  |  |
| 2022 | 0.609 | 0.597 | 0.046 | 0.549 | 0.705 |  |  |  |  |
| 2023 | 0.598 | 0.592 | 0.045 | 0.533 | 0.697 |  |  |  |  |
| 2024 | 0.597 | 0.586 | 0.049 | 0.533 | 0.719 |  |  |  |  |
| 2025 | 0.595 | 0.588 | 0.047 | 0.528 | 0.699 |  |  |  |  |
| 2026 | 0.591 | 0.585 | 0.047 | 0.524 | 0.701 |  |  |  |  |
| 2027 | 0.585 | 0.579 | 0.045 | 0.524 | 0.692 |  |  |  |  |
| 2028 | 0.584 | 0.578 | 0.043 | 0.524 | 0.697 |  |  |  |  |
| 2029 | 0.587 | 0.578 | 0.040 | 0.533 | 0.685 |  |  |  |  |
| 2030 | 0.589 | 0.580 | 0.040 | 0.536 | 0.669 |  |  |  |  |
| 2031 | 0.587 | 0.579 | 0.040 | 0.529 | 0.673 |  |  |  |  |
| 2032 | 0.586 | 0.583 | 0.034 | 0.533 | 0.649 |  |  |  |  |
| 2033 | 0.587 | 0.580 | 0.036 | 0.534 | 0.667 |  |  |  |  |
| 2034 | 0.587 | 0.578 | 0.042 | 0.528 | 0.691 |  |  |  |  |
| 2035 | 0.592 | 0.580 | 0.044 | 0.533 | 0.690 |  |  |  |  |
| 2036 | 0.593 | 0.579 | 0.046 | 0.537 | 0.697 |  |  |  |  |
| 2037 | 0.593 | 0.580 | 0.046 | 0.533 | 0.693 |  |  |  |  |
| 2038 | 0.592 | 0.577 | 0.047 | 0.527 | 0.716 |  |  |  |  |
| 2039 | 0.593 | 0.584 | 0.047 | 0.529 | 0.697 |  |  |  |  |
| 2040 | 0.595 | 0.587 | 0.045 | 0.535 | 0.694 |  |  |  |  |
| 2041 | 0.596 | 0.589 | 0.049 | 0.528 | 0.690 |  |  |  |  |
| 2042 | 0.599 | 0.590 | 0.051 | 0.527 | 0.699 |  |  |  |  |
| 2043 | 0.600 | 0.593 | 0.048 | 0.531 | 0.702 |  |  |  |  |
| 2044 | 0.596 | 0.586 | 0.047 | 0.529 | 0.713 |  |  |  |  |
| 2045 | 0.592 | 0.578 | 0.047 | 0.533 | 0.707 |  |  |  |  |
|  |  |  |  |  |  |  |  |  |  |

Table F.6. Projected mean, median, and standard deviation (SD) of legal male (LMB) and mature male (MMB) biomass ( t ) with $95 \%$ confidence limits under no directed fishery (top) and under $\mathrm{F}_{35 \%}$ (bottom) harvest control rule for scenario 17_0d for EAG, 2016-2045. The top table provides a base projection scenario with no directed fishery.

| No Directed Fishery |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | Mean <br> LMB | Median LMB | $\begin{gathered} \text { SD } \\ \text { LMB } \end{gathered}$ |  |  | Mean <br> MMB | Median <br> MMB | $\begin{gathered} \text { SD } \\ \text { MMB } \\ \hline \end{gathered}$ |  | $\begin{aligned} & \hline 95 \% \\ & \text { Upper } \\ & \text { Limit } \end{aligned}$ |
| 2016 | 6,380 | 6,317 | 1,492 | 4,036 | 9,720 | 10,303 | 10,201 | 2,410 | 6,517 | 15,696 |
| 2017 | 8,560 | 8,476 | 2,002 | 5,415 | 13,042 | 11,798 | 11,708 | 2,604 | 7,590 | 17,467 |
| 2018 | 10,234 | 10,117 | 2,373 | 6,490 | 15,526 | 13,110 | 13,021 | 2,533 | 8,872 | 18,403 |
| 2019 | 11,540 | 11,427 | 2,419 | 7,551 | 16,666 | 14,156 | 13,957 | 2,346 | 10,271 | 19,084 |
| 2020 | 12,560 | 12,405 | 2,290 | 8,730 | 17,353 | 15,033 | 15,000 | 2,094 | 11,533 | 19,428 |
| 2021 | 13,400 | 13,345 | 2,041 | 10,045 | 17,723 | 15,743 | 15,573 | 1,905 | 12,803 | 19,627 |
| 2022 | 14,066 | 14,004 | 1,804 | 11,079 | 17,776 | 16,385 | 16,324 | 1,787 | 13,789 | 19,940 |
| 2023 | 14,643 | 14,507 | 1,655 | 12,140 | 17,967 | 16,917 | 16,848 | 1,679 | 14,335 | 20,394 |
| 2024 | 15,150 | 15,076 | 1,552 | 12,825 | 18,258 | 17,331 | 17,381 | 1,587 | 14,705 | 20,452 |
| 2025 | 15,560 | 15,517 | 1,476 | 13,189 | 18,526 | 17,647 | 17,715 | 1,498 | 15,057 | 20,612 |
| 2026 | 15,873 | 15,903 | 1,395 | 13,550 | 18,569 | 17,894 | 17,969 | 1,422 | 15,326 | 20,455 |
| 2027 | 16,102 | 16,185 | 1,329 | 13,784 | 18,564 | 18,129 | 18,051 | 1,321 | 15,632 | 20,465 |
| 2028 | 16,296 | 16,311 | 1,255 | 13,967 | 18,635 | 18,336 | 18,398 | 1,228 | 15,841 | 20,412 |
| 2029 | 16,487 | 16,473 | 1,160 | 14,221 | 18,403 | 18,480 | 18,507 | 1,194 | 16,107 | 20,815 |
| 2030 | 16,641 | 16,671 | 1,097 | 14,440 | 18,678 | 18,578 | 18,649 | 1,160 | 16,190 | 20,835 |
| 2031 | 16,743 | 16,802 | 1,078 | 14,591 | 18,980 | 18,666 | 18,770 | 1,112 | 16,481 | 20,835 |
| 2032 | 16,812 | 16,905 | 1,036 | 14,719 | 18,805 | 18,749 | 18,733 | 1,090 | 16,680 | 20,769 |
| 2033 | 16,886 | 16,976 | 1,003 | 14,950 | 18,922 | 18,836 | 18,869 | 1,067 | 16,746 | 20,946 |
| 2034 | 16,944 | 16,978 | 987 | 15,039 | 18,916 | 18,942 | 18,846 | 1,063 | 16,902 | 20,851 |
| 2035 | 17,043 | 17,010 | 977 | 15,185 | 18,859 | 18,999 | 18,909 | 1,073 | 16,939 | 20,976 |
| 2036 | 17,109 | 17,018 | 981 | 15,220 | 19,041 | 19,056 | 19,008 | 1,139 | 16,994 | 21,238 |
| 2037 | 17,171 | 17,079 | 1,016 | 15,281 | 19,027 | 19,088 | 18,942 | 1,177 | 17,131 | 21,443 |
| 2038 | 17,196 | 17,107 | 1,080 | 15,291 | 19,465 | 19,154 | 19,059 | 1,177 | 17,201 | 21,705 |
| 2039 | 17,242 | 17,141 | 1,091 | 15,448 | 19,480 | 19,194 | 19,206 | 1,225 | 17,118 | 21,872 |
| 2040 | 17,289 | 17,194 | 1,105 | 15,403 | 19,809 | 19,250 | 19,305 | 1,296 | 17,135 | 21,839 |
| 2041 | 17,325 | 17,357 | 1,160 | 15,458 | 19,682 | 19,324 | 19,363 | 1,305 | 17,086 | 22,242 |
| 2042 | 17,391 | 17,377 | 1,210 | 15,319 | 20,080 | 19,337 | 19,400 | 1,249 | 17,209 | 21,814 |
| 2043 | 17,437 | 17,530 | 1,190 | 15,452 | 19,940 | 19,323 | 19,346 | 1,189 | 17,243 | 21,651 |
| 2044 | 17,433 | 17,426 | 1,134 | 15,472 | 19,561 | 19,314 | 19,359 | 1,145 | 17,356 | 21,601 |
| 2045 | 17,418 | 17,423 | 1,073 | 15,548 | 19,448 | 19,332 | 19,300 | 1,148 | 17,505 | 21,690 |


| $\mathrm{F}_{35 \%}\left(0.64 \mathrm{yr}^{-1}\right)$ |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | $\begin{aligned} & \text { Mean } \\ & \text { LMB } \end{aligned}$ | Median LMB | $\begin{gathered} \text { SD } \\ \text { LMB } \\ \hline \end{gathered}$ | 95\% Lower <br> Limit | 95\% <br> Upper <br> Limit | $\begin{aligned} & \text { Mean } \\ & \text { MMB } \\ & \hline \end{aligned}$ | Median MMB | $\begin{gathered} \text { SD } \\ \text { MMB } \\ \hline \end{gathered}$ | 95\% <br> Lower <br> Limit | 95\% <br> Upper <br> Limit |
| 2016 | 6,380 | 6,317 | 1,492 | 4,036 | 9,720 | 8,001 | 7,922 | 1,871 | 5,061 | 12,189 |
| 2017 | 6,314 | 6,215 | 1,418 | 4,237 | 9,563 | 7,556 | 7,470 | 1,585 | 5,035 | 10,977 |
| 2018 | 6,068 | 5,930 | 1,305 | 4,234 | 9,084 | 7,254 | 7,191 | 1,184 | 5,386 | 9,669 |
| 2019 | 5,783 | 5,651 | 1,010 | 4,299 | 7,907 | 7,004 | 6,957 | 846 | 5,772 | 8,950 |
| 2020 | 5,550 | 5,422 | 732 | 4,538 | 7,199 | 6,865 | 6,804 | 656 | 5,887 | 8,205 |
| 2021 | 5,416 | 5,374 | 517 | 4,682 | 6,620 | 6,773 | 6,716 | 594 | 5,751 | 7,937 |
| 2022 | 5,323 | 5,245 | 425 | 4,730 | 6,200 | 6,783 | 6,736 | 666 | 5,800 | 8,264 |
| 2023 | 5,306 | 5,228 | 446 | 4,636 | 6,330 | 6,802 | 6,747 | 684 | 5,760 | 8,404 |
| 2024 | 5,331 | 5,189 | 497 | 4,684 | 6,583 | 6,787 | 6,705 | 669 | 5,685 | 8,226 |
| 2025 | 5,336 | 5,258 | 482 | 4,651 | 6,458 | 6,746 | 6,667 | 665 | 5,727 | 8,092 |
| 2026 | 5,318 | 5,239 | 473 | 4,619 | 6,401 | 6,709 | 6,653 | 649 | 5,673 | 8,157 |
| 2027 | 5,281 | 5,158 | 464 | 4,650 | 6,293 | 6,723 | 6,675 | 618 | 5,769 | 8,189 |
| 2028 | 5,264 | 5,184 | 458 | 4,611 | 6,434 | 6,751 | 6,723 | 599 | 5,782 | 7,855 |
| 2029 | 5,287 | 5,208 | 433 | 4,702 | 6,286 | 6,742 | 6,657 | 606 | 5,749 | 7,852 |
| 2030 | 5,298 | 5,204 | 428 | 4,684 | 6,199 | 6,711 | 6,743 | 561 | 5,785 | 7,775 |
| 2031 | 5,279 | 5,185 | 422 | 4,662 | 6,178 | 6,698 | 6,604 | 523 | 5,814 | 7,689 |
| 2032 | 5,254 | 5,244 | 369 | 4,705 | 5,924 | 6,707 | 6,656 | 571 | 5,745 | 8,044 |
| 2033 | 5,261 | 5,190 | 384 | 4,667 | 6,165 | 6,734 | 6,682 | 607 | 5,869 | 8,116 |
| 2034 | 5,264 | 5,189 | 427 | 4,652 | 6,382 | 6,789 | 6,674 | 646 | 5,798 | 8,073 |
| 2035 | 5,315 | 5,201 | 452 | 4,724 | 6,296 | 6,789 | 6,664 | 657 | 5,794 | 8,084 |
| 2036 | 5,324 | 5,182 | 477 | 4,695 | 6,315 | 6,791 | 6,690 | 685 | 5,817 | 8,413 |
| 2037 | 5,336 | 5,261 | 481 | 4,697 | 6,379 | 6,776 | 6,644 | 695 | 5,809 | 8,323 |
| 2038 | 5,318 | 5,170 | 502 | 4,691 | 6,679 | 6,808 | 6,768 | 672 | 5,760 | 8,295 |
| 2039 | 5,328 | 5,229 | 498 | 4,696 | 6,487 | 6,815 | 6,823 | 669 | 5,771 | 8,214 |
| 2040 | 5,345 | 5,271 | 470 | 4,662 | 6,455 | 6,839 | 6,870 | 720 | 5,662 | 8,357 |
| 2041 | 5,352 | 5,314 | 498 | 4,647 | 6,415 | 6,883 | 6,856 | 719 | 5,749 | 8,302 |
| 2042 | 5,386 | 5,327 | 530 | 4,601 | 6,494 | 6,858 | 6,716 | 669 | 5,847 | 8,387 |
| 2043 | 5,391 | 5,273 | 507 | 4,664 | 6,545 | 6,806 | 6,685 | 646 | 5,932 | 8,428 |
| 2044 | 5,352 | 5,257 | 483 | 4,721 | 6,534 | 6,778 | 6,695 | 636 | 5,862 | 8,344 |
| 2045 | 5,322 | 5,185 | 469 | 4,755 | 6,561 | 6,794 | 6,653 | 634 | 5,820 | 8,047 |

Table F.7. Projected mean, median, and standard deviation (SD) of total catch (OFL) and retained catch (RETC) in biomass ( t ) with $95 \%$ confidence limits under no directed fishery (top) and under $\mathrm{F}_{35 \%}$ (bottom) harvest control rule for scenario 17_0d for EAG, 2016-2045.

| No Directed Fishery |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | $\begin{gathered} \text { Mean } \\ \text { OFL } \end{gathered}$ | Median OFL | $\begin{gathered} \text { SD } \\ \text { OFL } \end{gathered}$ | $\begin{gathered} 95 \% \\ \text { Lower } \\ \text { Limit } \\ \hline \end{gathered}$ | 95\% <br> Upper <br> Limit | Mean <br> RETC | Median RETC | $\begin{gathered} \text { SD } \\ \text { RETC } \end{gathered}$ | $\begin{gathered} \hline 95 \% \\ \text { Lower } \\ \text { Limit } \\ \hline \end{gathered}$ | 95\% <br> Upper <br> Limit |
| 2016 | 1.689 | 1.672 | 0.395 | 1.068 | 2.573 | 0 | 0 | 0 | 0 | 0 |
| 2017 | 1.929 | 1.911 | 0.397 | 1.267 | 2.758 | 0 | 0 | 0 | 0 | 0 |
| 2018 | 2.115 | 2.090 | 0.386 | 1.473 | 2.933 | 0 | 0 | 0 | 0 | 0 |
| 2019 | 2.278 | 2.272 | 0.350 | 1.707 | 3.016 | 0 | 0 | 0 | 0 | 0 |
| 2020 | 2.403 | 2.376 | 0.317 | 1.882 | 3.051 | 0 | 0 | 0 | 0 | 0 |
| 2021 | 2.517 | 2.506 | 0.297 | 2.057 | 3.099 | 0 | 0 | 0 | 0 | 0 |
| 2022 | 2.613 | 2.598 | 0.275 | 2.206 | 3.175 | 0 | 0 | 0 | 0 | 0 |
| 2023 | 2.690 | 2.674 | 0.261 | 2.277 | 3.222 | 0 | 0 | 0 | 0 | 0 |
| 2024 | 2.750 | 2.763 | 0.244 | 2.346 | 3.218 | 0 | 0 | 0 | 0 | 0 |
| 2025 | 2.795 | 2.805 | 0.232 | 2.402 | 3.244 | 0 | 0 | 0 | 0 | 0 |
| 2026 | 2.836 | 2.844 | 0.218 | 2.429 | 3.250 | 0 | 0 | 0 | 0 | 0 |
| 2027 | 2.874 | 2.877 | 0.200 | 2.472 | 3.203 | 0 | 0 | 0 | 0 | 0 |
| 2028 | 2.902 | 2.923 | 0.191 | 2.527 | 3.248 | 0 | 0 | 0 | 0 | 0 |
| 2029 | 2.921 | 2.930 | 0.187 | 2.544 | 3.296 | 0 | 0 | 0 | 0 | 0 |
| 2030 | 2.936 | 2.947 | 0.179 | 2.569 | 3.262 | 0 | 0 | 0 | 0 | 0 |
| 2031 | 2.953 | 2.962 | 0.175 | 2.614 | 3.300 | 0 | 0 | 0 | 0 | 0 |
| 2032 | 2.962 | 2.968 | 0.170 | 2.623 | 3.298 | 0 | 0 | 0 | 0 | 0 |
| 2033 | 2.984 | 2.974 | 0.167 | 2.664 | 3.285 | 0 | 0 | 0 | 0 | 0 |
| 2034 | 2.991 | 2.978 | 0.165 | 2.668 | 3.318 | 0 | 0 | 0 | 0 | 0 |
| 2035 | 3.005 | 2.991 | 0.172 | 2.686 | 3.309 | 0 | 0 | 0 | 0 | 0 |
| 2036 | 3.007 | 2.985 | 0.184 | 2.676 | 3.382 | 0 | 0 | 0 | 0 | 0 |
| 2037 | 3.019 | 3.002 | 0.181 | 2.726 | 3.379 | 0 | 0 | 0 | 0 | 0 |
| 2038 | 3.026 | 3.017 | 0.186 | 2.708 | 3.450 | 0 | 0 | 0 | 0 | 0 |
| 2039 | 3.032 | 3.048 | 0.199 | 2.714 | 3.429 | 0 | 0 | 0 | 0 | 0 |
| 2040 | 3.046 | 3.038 | 0.205 | 2.696 | 3.501 | 0 | 0 | 0 | 0 | 0 |
| 2041 | 3.052 | 3.065 | 0.199 | 2.701 | 3.484 | 0 | 0 | 0 | 0 | 0 |
| 2042 | 3.050 | 3.045 | 0.192 | 2.719 | 3.433 | 0 | 0 | 0 | 0 | 0 |
| 2043 | 3.050 | 3.060 | 0.182 | 2.726 | 3.389 | 0 | 0 | 0 | 0 | 0 |
| 2044 | 3.049 | 3.060 | 0.181 | 2.744 | 3.421 | 0 | 0 | 0 | 0 | 0 |
| 2045 | 3.059 | 3.045 | 0.182 | 2.782 | 3.474 | 0 | 0 | 0 | 0 | 0 |


| $\mathrm{F}_{35 \%}\left(0.64 \mathrm{yr}^{-1}\right)$ |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | $\begin{gathered} \text { Mean } \\ \text { OFL } \end{gathered}$ | Median OFL | $\begin{gathered} \text { SD } \\ \text { OFL } \\ \hline \end{gathered}$ | 95\% Lower <br> Limit | 95\% <br> Upper <br> Limit | $\begin{aligned} & \text { Mean } \\ & \text { RETC } \\ & \hline \end{aligned}$ | $\begin{gathered} \text { Median } \\ \text { RETC } \\ \hline \end{gathered}$ | $\begin{gathered} \text { SD } \\ \text { RETC } \\ \hline \end{gathered}$ | 95\% <br> Lower <br> Limit | 95\% <br> Upper <br> Limit |
| 2016 | 2,284 | 2,262 | 534 | 1,445 | 3,480 | 2,124 | 2,137 | 556 | 1,116 | 3,289 |
| 2017 | 2,206 | 2,170 | 493 | 1,487 | 3,336 | 2,046 | 2,052 | 519 | 1,148 | 3,157 |
| 2018 | 2,126 | 2,074 | 446 | 1,500 | 3,154 | 1,981 | 1,966 | 467 | 1,209 | 3,012 |
| 2019 | 2,027 | 1,975 | 351 | 1,521 | 2,782 | 1,887 | 1,855 | 373 | 1,299 | 2,657 |
| 2020 | 1,947 | 1,907 | 254 | 1,595 | 2,523 | 1,809 | 1,778 | 276 | 1,406 | 2,397 |
| 2021 | 1,897 | 1,884 | 178 | 1,652 | 2,316 | 1,756 | 1,752 | 201 | 1,459 | 2,206 |
| 2022 | 1,865 | 1,832 | 144 | 1,675 | 2,154 | 1,721 | 1,709 | 169 | 1,441 | 2,043 |
| 2023 | 1,857 | 1,831 | 147 | 1,641 | 2,195 | 1,712 | 1,691 | 170 | 1,424 | 2,068 |
| 2024 | 1,863 | 1,819 | 162 | 1,647 | 2,268 | 1,717 | 1,686 | 187 | 1,429 | 2,134 |
| 2025 | 1,864 | 1,835 | 159 | 1,645 | 2,244 | 1,715 | 1,703 | 188 | 1,401 | 2,105 |
| 2026 | 1,859 | 1,831 | 156 | 1,634 | 2,227 | 1,710 | 1,688 | 185 | 1,400 | 2,106 |
| 2027 | 1,848 | 1,814 | 153 | 1,641 | 2,165 | 1,704 | 1,676 | 177 | 1,407 | 2,034 |
| 2028 | 1,842 | 1,813 | 149 | 1,636 | 2,227 | 1,699 | 1,689 | 171 | 1,418 | 2,102 |
| 2029 | 1,847 | 1,817 | 142 | 1,652 | 2,193 | 1,702 | 1,686 | 166 | 1,425 | 2,073 |
| 2030 | 1,851 | 1,818 | 140 | 1,650 | 2,151 | 1,708 | 1,695 | 167 | 1,406 | 2,021 |
| 2031 | 1,846 | 1,814 | 137 | 1,641 | 2,136 | 1,703 | 1,691 | 161 | 1,409 | 2,018 |
| 2032 | 1,838 | 1,836 | 121 | 1,653 | 2,076 | 1,695 | 1,683 | 143 | 1,434 | 1,952 |
| 2033 | 1,839 | 1,814 | 124 | 1,658 | 2,128 | 1,694 | 1,678 | 145 | 1,429 | 2,008 |
| 2034 | 1,842 | 1,818 | 137 | 1,645 | 2,199 | 1,698 | 1,677 | 159 | 1,454 | 2,081 |
| 2035 | 1,856 | 1,816 | 147 | 1,667 | 2,176 | 1,712 | 1,687 | 169 | 1,439 | 2,049 |
| 2036 | 1,861 | 1,806 | 157 | 1,652 | 2,191 | 1,716 | 1,673 | 182 | 1,420 | 2,066 |
| 2037 | 1,864 | 1,838 | 160 | 1,655 | 2,217 | 1,717 | 1,706 | 183 | 1,431 | 2,081 |
| 2038 | 1,861 | 1,812 | 166 | 1,653 | 2,313 | 1,718 | 1,690 | 187 | 1,424 | 2,182 |
| 2039 | 1,862 | 1,823 | 164 | 1,649 | 2,244 | 1,717 | 1,691 | 188 | 1,403 | 2,118 |
| 2040 | 1,868 | 1,831 | 156 | 1,640 | 2,238 | 1,721 | 1,706 | 185 | 1,422 | 2,115 |
| 2041 | 1,872 | 1,856 | 163 | 1,645 | 2,213 | 1,728 | 1,741 | 187 | 1,405 | 2,086 |
| 2042 | 1,881 | 1,857 | 173 | 1,624 | 2,255 | 1,741 | 1,730 | 192 | 1,422 | 2,125 |
| 2043 | 1,882 | 1,851 | 168 | 1,644 | 2,254 | 1,742 | 1,721 | 187 | 1,458 | 2,126 |
| 2044 | 1,871 | 1,842 | 159 | 1,666 | 2,258 | 1,729 | 1,702 | 177 | 1,478 | 2,139 |
| 2045 | 1,861 | 1,810 | 155 | 1,678 | 2,254 | 1,719 | 1,685 | 173 | 1,464 | 2,134 |

Table F.8. Projected mean, median, and standard deviation (SD) of retained CPUE indices with $95 \%$ confidence limits under $\mathrm{F}_{35 \%}$ harvest control rule for scenario 17_0d for EAG, 2016-2045.

|  |  |  |  | $\mathrm{F}_{35 \%}\left(0.64 \mathrm{yr}^{-1}\right)$ |  |  |  |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | $95 \%$ <br> Mean |  |  |  | Median <br> Year | CPUE | CPUE | CPUE | Lower <br> Limit | Uper <br> Limit |
| 2016 | 0.661 | 0.652 | 0.151 | 0.432 | 1.004 |  |  |  |  |  |  |  |
| 2017 | 0.657 | 0.644 | 0.144 | 0.455 | 0.991 |  |  |  |  |  |  |  |
| 2018 | 0.636 | 0.619 | 0.134 | 0.455 | 0.947 |  |  |  |  |  |  |  |
| 2019 | 0.606 | 0.591 | 0.104 | 0.457 | 0.829 |  |  |  |  |  |  |  |
| 2020 | 0.583 | 0.571 | 0.075 | 0.477 | 0.753 |  |  |  |  |  |  |  |
| 2021 | 0.568 | 0.565 | 0.053 | 0.497 | 0.692 |  |  |  |  |  |  |  |
| 2022 | 0.559 | 0.551 | 0.043 | 0.498 | 0.649 |  |  |  |  |  |  |  |
| 2023 | 0.557 | 0.549 | 0.045 | 0.491 | 0.662 |  |  |  |  |  |  |  |
| 2024 | 0.560 | 0.545 | 0.050 | 0.495 | 0.687 |  |  |  |  |  |  |  |
| 2025 | 0.561 | 0.553 | 0.049 | 0.492 | 0.676 |  |  |  |  |  |  |  |
| 2026 | 0.559 | 0.551 | 0.048 | 0.490 | 0.670 |  |  |  |  |  |  |  |
| 2027 | 0.555 | 0.544 | 0.047 | 0.493 | 0.657 |  |  |  |  |  |  |  |
| 2028 | 0.553 | 0.546 | 0.047 | 0.487 | 0.673 |  |  |  |  |  |  |  |
| 2029 | 0.556 | 0.548 | 0.044 | 0.494 | 0.660 |  |  |  |  |  |  |  |
| 2030 | 0.557 | 0.546 | 0.043 | 0.498 | 0.647 |  |  |  |  |  |  |  |
| 2031 | 0.555 | 0.545 | 0.043 | 0.494 | 0.646 |  |  |  |  |  |  |  |
| 2032 | 0.552 | 0.551 | 0.037 | 0.497 | 0.622 |  |  |  |  |  |  |  |
| 2033 | 0.553 | 0.546 | 0.039 | 0.493 | 0.643 |  |  |  |  |  |  |  |
| 2034 | 0.553 | 0.545 | 0.043 | 0.494 | 0.668 |  |  |  |  |  |  |  |
| 2035 | 0.558 | 0.546 | 0.046 | 0.500 | 0.658 |  |  |  |  |  |  |  |
| 2036 | 0.560 | 0.544 | 0.048 | 0.498 | 0.662 |  |  |  |  |  |  |  |
| 2037 | 0.561 | 0.552 | 0.049 | 0.495 | 0.667 |  |  |  |  |  |  |  |
| 2038 | 0.559 | 0.544 | 0.051 | 0.494 | 0.699 |  |  |  |  |  |  |  |
| 2039 | 0.560 | 0.549 | 0.050 | 0.496 | 0.678 |  |  |  |  |  |  |  |
| 2040 | 0.562 | 0.555 | 0.047 | 0.497 | 0.676 |  |  |  |  |  |  |  |
| 2041 | 0.562 | 0.559 | 0.050 | 0.491 | 0.671 |  |  |  |  |  |  |  |
| 2042 | 0.565 | 0.558 | 0.054 | 0.488 | 0.680 |  |  |  |  |  |  |  |
| 2043 | 0.566 | 0.555 | 0.052 | 0.492 | 0.684 |  |  |  |  |  |  |  |
| 2044 | 0.563 | 0.552 | 0.049 | 0.499 | 0.684 |  |  |  |  |  |  |  |
| 2045 | 0.559 | 0.544 | 0.048 | 0.503 | 0.685 |  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |  |  |  |  |  |

Table F.9. Projected mean, median, and standard deviation (SD) of legal male (LMB) and mature male (MMB) biomass ( t ) with $95 \%$ confidence limits under no directed fishery (top) and under $\mathrm{F}_{35 \%}$ (bottom) harvest control rule for scenario 17_0 for WAG, 2016-2045. The top table provides a base projection scenario with no directed fishery.

| No Directed Fishery |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | Mean LMB | Median <br> LMB | $\begin{gathered} \text { SD } \\ \text { LMB } \end{gathered}$ | $\begin{gathered} 95 \% \\ \text { Lower } \\ \text { Limit } \\ \hline \end{gathered}$ | 95\% <br> Upper <br> Limit | Mean <br> MMB | Median <br> MMB | $\begin{gathered} \text { SD } \\ \text { MMB } \\ \hline \end{gathered}$ | 95\% <br> Lower <br> Limit | 95\% <br> Upper <br> Limit |
| 2016 | 3,581 | 3,571 | 637 | 2,534 | 4,964 | 7,263 | 7,243 | 1,292 | 5,141 | 10,070 |
| 2017 | 5,526 | 5,511 | 983 | 3,911 | 7,662 | 8,749 | 8,745 | 1,452 | 6,428 | 11,840 |
| 2018 | 7,292 | 7,253 | 1,284 | 5,191 | 10,074 | 9,898 | 9,814 | 1,403 | 7,709 | 12,874 |
| 2019 | 8,572 | 8,561 | 1,349 | 6,452 | 11,405 | 10,815 | 10,647 | 1,312 | 8,797 | 13,603 |
| 2020 | 9,502 | 9,311 | 1,277 | 7,554 | 12,218 | 11,541 | 11,457 | 1,222 | 9,518 | 13,945 |
| 2021 | 10,211 | 10,095 | 1,167 | 8,321 | 12,665 | 12,115 | 12,153 | 1,130 | 10,264 | 14,185 |
| 2022 | 10,752 | 10,731 | 1,074 | 8,956 | 12,782 | 12,602 | 12,666 | 1,035 | 10,879 | 14,553 |
| 2023 | 11,195 | 11,267 | 979 | 9,618 | 13,002 | 13,025 | 13,127 | 967 | 11,284 | 14,629 |
| 2024 | 11,567 | 11,645 | 900 | 9,982 | 13,232 | 13,398 | 13,532 | 893 | 11,684 | 14,808 |
| 2025 | 11,909 | 12,028 | 843 | 10,346 | 13,211 | 13,681 | 13,725 | 799 | 12,083 | 15,031 |
| 2026 | 12,192 | 12,286 | 768 | 10,682 | 13,403 | 13,902 | 13,941 | 730 | 12,413 | 15,167 |
| 2027 | 12,408 | 12,481 | 691 | 10,977 | 13,601 | 14,076 | 14,080 | 703 | 12,709 | 15,225 |
| 2028 | 12,574 | 12,550 | 639 | 11,283 | 13,678 | 14,210 | 14,154 | 717 | 12,847 | 15,469 |
| 2029 | 12,702 | 12,659 | 636 | 11,495 | 13,806 | 14,306 | 14,253 | 706 | 13,035 | 15,594 |
| 2030 | 12,796 | 12,752 | 658 | 11,545 | 13,971 | 14,382 | 14,348 | 683 | 13,082 | 15,585 |
| 2031 | 12,863 | 12,834 | 635 | 11,773 | 14,045 | 14,471 | 14,406 | 722 | 12,990 | 15,833 |
| 2032 | 12,923 | 12,862 | 644 | 11,715 | 14,092 | 14,561 | 14,574 | 733 | 13,329 | 15,953 |
| 2033 | 13,004 | 12,980 | 669 | 11,654 | 14,311 | 14,612 | 14,570 | 704 | 13,303 | 16,015 |
| 2034 | 13,066 | 13,072 | 670 | 11,871 | 14,304 | 14,655 | 14,631 | 672 | 13,403 | 16,083 |
| 2035 | 13,107 | 13,078 | 630 | 11,950 | 14,389 | 14,705 | 14,695 | 656 | 13,459 | 16,069 |
| 2036 | 13,143 | 13,138 | 610 | 11,963 | 14,475 | 14,745 | 14,774 | 647 | 13,540 | 15,999 |
| 2037 | 13,187 | 13,209 | 593 | 12,086 | 14,331 | 14,761 | 14,750 | 656 | 13,465 | 15,747 |
| 2038 | 13,208 | 13,203 | 597 | 12,045 | 14,247 | 14,780 | 14,872 | 667 | 13,502 | 16,029 |
| 2039 | 13,225 | 13,255 | 607 | 12,032 | 14,221 | 14,778 | 14,825 | 709 | 13,322 | 16,210 |
| 2040 | 13,232 | 13,300 | 626 | 11,973 | 14,443 | 14,775 | 14,754 | 765 | 13,298 | 15,976 |
| 2041 | 13,225 | 13,223 | 675 | 11,890 | 14,471 | 14,773 | 14,756 | 780 | 13,281 | 16,113 |
| 2042 | 13,223 | 13,198 | 711 | 11,855 | 14,335 | 14,764 | 14,814 | 784 | 13,091 | 16,148 |
| 2043 | 13,214 | 13,227 | 720 | 11,735 | 14,536 | 14,784 | 14,766 | 785 | 13,351 | 16,114 |
| 2044 | 13,214 | 13,261 | 724 | 11,863 | 14,509 | 14,828 | 14,828 | 776 | 13,538 | 16,103 |
| 2045 | 13,240 | 13,274 | 716 | 12,022 | 14,462 | 14,880 | 14,951 | 782 | 13,420 | 16,134 |


| $\mathrm{F}_{35 \%}\left(0.6 \mathrm{yr}^{-1}\right)$ |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | $\begin{aligned} & \text { Mean } \\ & \text { LMB } \end{aligned}$ | Median LMB | $\begin{gathered} \text { SD } \\ \text { LMB } \\ \hline \end{gathered}$ | 95\% Lower <br> Limit | 95\% <br> Upper <br> Limit | $\begin{aligned} & \text { Mean } \\ & \text { MMB } \\ & \hline \end{aligned}$ | Median MMB | $\begin{gathered} \text { SD } \\ \text { MMB } \\ \hline \end{gathered}$ | 95\% <br> Lower <br> Limit | 95\% <br> Upper <br> Limit |
| 2016 | 3,581 | 3,571 | 637 | 2,534 | 4,964 | 5,951 | 5,934 | 1,059 | 4,212 | 8,251 |
| 2017 | 4,237 | 4,208 | 727 | 3,117 | 5,851 | 6,061 | 6,057 | 961 | 4,582 | 8,101 |
| 2018 | 4,619 | 4,575 | 777 | 3,445 | 6,341 | 5,849 | 5,778 | 713 | 4,731 | 7,379 |
| 2019 | 4,546 | 4,531 | 623 | 3,646 | 5,907 | 5,597 | 5,550 | 543 | 4,544 | 6,679 |
| 2020 | 4,334 | 4,262 | 451 | 3,635 | 5,246 | 5,395 | 5,425 | 473 | 4,472 | 6,244 |
| 2021 | 4,160 | 4,135 | 346 | 3,554 | 4,858 | 5,258 | 5,229 | 420 | 4,438 | 6,005 |
| 2022 | 4,031 | 4,030 | 306 | 3,503 | 4,611 | 5,205 | 5,195 | 396 | 4,405 | 5,872 |
| 2023 | 3,967 | 3,937 | 268 | 3,506 | 4,459 | 5,209 | 5,175 | 415 | 4,496 | 5,911 |
| 2024 | 3,949 | 3,907 | 271 | 3,476 | 4,438 | 5,246 | 5,277 | 393 | 4,567 | 5,888 |
| 2025 | 3,973 | 3,951 | 283 | 3,499 | 4,516 | 5,243 | 5,272 | 345 | 4,581 | 5,859 |
| 2026 | 3,983 | 3,962 | 253 | 3,595 | 4,442 | 5,222 | 5,192 | 335 | 4,663 | 5,945 |
| 2027 | 3,971 | 3,962 | 230 | 3,554 | 4,500 | 5,200 | 5,145 | 389 | 4,550 | 6,031 |
| 2028 | 3,960 | 3,924 | 239 | 3,621 | 4,446 | 5,183 | 5,121 | 445 | 4,389 | 6,116 |
| 2029 | 3,954 | 3,890 | 283 | 3,524 | 4,622 | 5,164 | 5,136 | 427 | 4,373 | 5,867 |
| 2030 | 3,943 | 3,891 | 300 | 3,456 | 4,542 | 5,152 | 5,117 | 414 | 4,408 | 5,930 |
| 2031 | 3,927 | 3,895 | 271 | 3,468 | 4,454 | 5,174 | 5,148 | 440 | 4,240 | 5,997 |
| 2032 | 3,925 | 3,888 | 288 | 3,405 | 4,485 | 5,210 | 5,244 | 415 | 4,376 | 5,993 |
| 2033 | 3,952 | 3,945 | 291 | 3,470 | 4,554 | 5,204 | 5,189 | 373 | 4,586 | 5,960 |
| 2034 | 3,959 | 3,927 | 272 | 3,472 | 4,525 | 5,197 | 5,204 | 369 | 4,446 | 5,865 |
| 2035 | 3,952 | 3,927 | 244 | 3,529 | 4,457 | 5,207 | 5,213 | 371 | 4,465 | 5,907 |
| 2036 | 3,949 | 3,930 | 247 | 3,510 | 4,430 | 5,214 | 5,273 | 397 | 4,454 | 5,873 |
| 2037 | 3,965 | 3,955 | 254 | 3,490 | 4,399 | 5,200 | 5,174 | 422 | 4,448 | 6,031 |
| 2038 | 3,960 | 3,946 | 278 | 3,523 | 4,475 | 5,198 | 5,177 | 420 | 4,457 | 6,055 |
| 2039 | 3,956 | 3,939 | 288 | 3,494 | 4,571 | 5,179 | 5,188 | 432 | 4,323 | 6,002 |
| 2040 | 3,950 | 3,926 | 281 | 3,486 | 4,617 | 5,167 | 5,160 | 452 | 4,271 | 6,004 |
| 2041 | 3,938 | 3,885 | 295 | 3,426 | 4,539 | 5,164 | 5,180 | 442 | 4,300 | 5,915 |
| 2042 | 3,938 | 3,907 | 299 | 3,402 | 4,543 | 5,157 | 5,150 | 438 | 4,283 | 5,912 |
| 2043 | 3,930 | 3,901 | 292 | 3,442 | 4,486 | 5,180 | 5,212 | 420 | 4,404 | 5,889 |
| 2044 | 3,932 | 3,892 | 290 | 3,381 | 4,477 | 5,226 | 5,266 | 403 | 4,436 | 5,911 |
| 2045 | 3,954 | 3,947 | 274 | 3,536 | 4,410 | 5,266 | 5,274 | 415 | 4,412 | 6,051 |

Table F.10. Projected mean, median, and standard deviation (SD) of total catch (OFL) and retained catch (RETC) in biomass ( t ) with $95 \%$ confidence limits under no directed fishery (top) and under $\mathrm{F}_{35 \%}$ (bottom) harvest control rule for scenario 17_0 for WAG, 2016-2045.

| No Directed Fishery |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | $\begin{gathered} \text { Mean } \\ \text { OFL } \end{gathered}$ | Median OFL | $\begin{gathered} \text { SD } \\ \text { OFL } \end{gathered}$ | $\begin{gathered} 95 \% \\ \text { Lower } \\ \text { Limit } \\ \hline \end{gathered}$ | 95\% <br> Upper <br> Limit | Mean <br> RETC | Median RETC | $\begin{gathered} \text { SD } \\ \text { RETC } \end{gathered}$ | $\begin{gathered} \hline 95 \% \\ \text { Lower } \\ \text { Limit } \\ \hline \end{gathered}$ | 95\% <br> Upper <br> Limit |
| 2016 | 3.377 | 3.367 | 0.601 | 2.390 | 4.682 | 0 | 0 | 0 | 0 | 0 |
| 2017 | 3.949 | 3.955 | 0.603 | 2.998 | 5.223 | 0 | 0 | 0 | 0 | 0 |
| 2018 | 4.416 | 4.335 | 0.585 | 3.520 | 5.656 | 0 | 0 | 0 | 0 | 0 |
| 2019 | 4.796 | 4.765 | 0.549 | 3.902 | 5.943 | 0 | 0 | 0 | 0 | 0 |
| 2020 | 5.085 | 5.091 | 0.517 | 4.241 | 6.051 | 0 | 0 | 0 | 0 | 0 |
| 2021 | 5.332 | 5.369 | 0.472 | 4.555 | 6.220 | 0 | 0 | 0 | 0 | 0 |
| 2022 | 5.532 | 5.566 | 0.439 | 4.785 | 6.356 | 0 | 0 | 0 | 0 | 0 |
| 2023 | 5.721 | 5.786 | 0.411 | 4.977 | 6.388 | 0 | 0 | 0 | 0 | 0 |
| 2024 | 5.864 | 5.883 | 0.369 | 5.149 | 6.482 | 0 | 0 | 0 | 0 | 0 |
| 2025 | 5.978 | 5.999 | 0.335 | 5.307 | 6.557 | 0 | 0 | 0 | 0 | 0 |
| 2026 | 6.068 | 6.074 | 0.312 | 5.459 | 6.595 | 0 | 0 | 0 | 0 | 0 |
| 2027 | 6.139 | 6.136 | 0.308 | 5.537 | 6.676 | 0 | 0 | 0 | 0 | 0 |
| 2028 | 6.192 | 6.179 | 0.314 | 5.605 | 6.735 | 0 | 0 | 0 | 0 | 0 |
| 2029 | 6.230 | 6.216 | 0.289 | 5.684 | 6.787 | 0 | 0 | 0 | 0 | 0 |
| 2030 | 6.265 | 6.234 | 0.306 | 5.680 | 6.822 | 0 | 0 | 0 | 0 | 0 |
| 2031 | 6.314 | 6.318 | 0.314 | 5.726 | 6.919 | 0 | 0 | 0 | 0 | 0 |
| 2032 | 6.339 | 6.333 | 0.315 | 5.761 | 6.896 | 0 | 0 | 0 | 0 | 0 |
| 2033 | 6.361 | 6.330 | 0.293 | 5.847 | 6.993 | 0 | 0 | 0 | 0 | 0 |
| 2034 | 6.381 | 6.389 | 0.291 | 5.811 | 6.992 | 0 | 0 | 0 | 0 | 0 |
| 2035 | 6.407 | 6.412 | 0.280 | 5.893 | 6.942 | 0 | 0 | 0 | 0 | 0 |
| 2036 | 6.412 | 6.410 | 0.283 | 5.859 | 6.905 | 0 | 0 | 0 | 0 | 0 |
| 2037 | 6.426 | 6.452 | 0.282 | 5.864 | 6.882 | 0 | 0 | 0 | 0 | 0 |
| 2038 | 6.429 | 6.468 | 0.292 | 5.831 | 6.993 | 0 | 0 | 0 | 0 | 0 |
| 2039 | 6.426 | 6.424 | 0.321 | 5.789 | 7.014 | 0 | 0 | 0 | 0 | 0 |
| 2040 | 6.430 | 6.431 | 0.333 | 5.803 | 6.972 | 0 | 0 | 0 | 0 | 0 |
| 2041 | 6.422 | 6.428 | 0.335 | 5.748 | 7.030 | 0 | 0 | 0 | 0 | 0 |
| 2042 | 6.425 | 6.430 | 0.340 | 5.775 | 7.013 | 0 | 0 | 0 | 0 | 0 |
| 2043 | 6.440 | 6.458 | 0.336 | 5.825 | 7.002 | 0 | 0 | 0 | 0 | 0 |
| 2044 | 6.462 | 6.467 | 0.336 | 5.872 | 7.017 | 0 | 0 | 0 | 0 | 0 |
| 2045 | 6.480 | 6.507 | 0.341 | 5.858 | 7.031 | 0 | 0 | 0 | 0 | 0 |


| $\mathrm{F}_{35 \%}\left(0.6 \mathrm{yr}^{-1}\right)$ |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | $\begin{gathered} \text { Mean } \\ \text { OFL } \\ \hline \end{gathered}$ | Median OFL | $\begin{gathered} \text { SD } \\ \text { OFL } \\ \hline \end{gathered}$ | 95\% <br> Lower <br> Limit | 95\% <br> Upper <br> Limit | $\begin{aligned} & \text { Mean } \\ & \text { RETC } \\ & \hline \end{aligned}$ | Median RETC | $\begin{gathered} \text { SD } \\ \text { RETC } \\ \hline \end{gathered}$ | 95\% <br> Lower <br> Limit | 95\% <br> Upper <br> Limit |
| 2016 | 1,301 | 1,297 | 232 | 921 | 1,804 | 1,177 | 1,188 | 235 | 726 | 1,652 |
| 2017 | 1,500 | 1,489 | 256 | 1,106 | 2,069 | 1,372 | 1,371 | 251 | 939 | 1,906 |
| 2018 | 1,629 | 1,614 | 267 | 1,232 | 2,219 | 1,514 | 1,506 | 264 | 1,116 | 2,087 |
| 2019 | 1,614 | 1,605 | 218 | 1,303 | 2,091 | 1,505 | 1,500 | 223 | 1,165 | 1,977 |
| 2020 | 1,546 | 1,525 | 158 | 1,301 | 1,864 | 1,433 | 1,412 | 171 | 1,136 | 1,767 |
| 2021 | 1,485 | 1,475 | 121 | 1,271 | 1,725 | 1,372 | 1,374 | 137 | 1,112 | 1,623 |
| 2022 | 1,440 | 1,443 | 104 | 1,264 | 1,636 | 1,325 | 1,336 | 120 | 1,080 | 1,538 |
| 2023 | 1,417 | 1,407 | 91 | 1,261 | 1,591 | 1,300 | 1,298 | 107 | 1,079 | 1,487 |
| 2024 | 1,411 | 1,404 | 92 | 1,254 | 1,577 | 1,296 | 1,281 | 105 | 1,115 | 1,474 |
| 2025 | 1,417 | 1,410 | 95 | 1,265 | 1,597 | 1,306 | 1,309 | 106 | 1,118 | 1,497 |
| 2026 | 1,420 | 1,410 | 86 | 1,288 | 1,572 | 1,310 | 1,314 | 96 | 1,144 | 1,465 |
| 2027 | 1,416 | 1,415 | 78 | 1,282 | 1,592 | 1,300 | 1,297 | 89 | 1,153 | 1,485 |
| 2028 | 1,413 | 1,395 | 82 | 1,300 | 1,587 | 1,291 | 1,269 | 100 | 1,127 | 1,482 |
| 2029 | 1,410 | 1,381 | 96 | 1,266 | 1,641 | 1,289 | 1,271 | 115 | 1,076 | 1,535 |
| 2030 | 1,407 | 1,384 | 101 | 1,246 | 1,606 | 1,286 | 1,272 | 118 | 1,076 | 1,512 |
| 2031 | 1,402 | 1,391 | 93 | 1,249 | 1,580 | 1,282 | 1,276 | 112 | 1,070 | 1,471 |
| 2032 | 1,402 | 1,389 | 97 | 1,238 | 1,593 | 1,285 | 1,283 | 113 | 1,074 | 1,488 |
| 2033 | 1,409 | 1,402 | 98 | 1,248 | 1,623 | 1,295 | 1,291 | 110 | 1,083 | 1,514 |
| 2034 | 1,412 | 1,406 | 91 | 1,255 | 1,595 | 1,297 | 1,296 | 106 | 1,111 | 1,492 |
| 2035 | 1,410 | 1,406 | 82 | 1,265 | 1,582 | 1,296 | 1,294 | 98 | 1,099 | 1,481 |
| 2036 | 1,409 | 1,401 | 83 | 1,264 | 1,575 | 1,293 | 1,293 | 101 | 1,092 | 1,468 |
| 2037 | 1,413 | 1,411 | 86 | 1,256 | 1,565 | 1,295 | 1,302 | 104 | 1,088 | 1,464 |
| 2038 | 1,413 | 1,403 | 93 | 1,268 | 1,595 | 1,297 | 1,302 | 110 | 1,091 | 1,488 |
| 2039 | 1,411 | 1,407 | 97 | 1,263 | 1,615 | 1,292 | 1,296 | 116 | 1,084 | 1,511 |
| 2040 | 1,409 | 1,399 | 96 | 1,253 | 1,631 | 1,288 | 1,296 | 118 | 1,053 | 1,528 |
| 2041 | 1,405 | 1,390 | 99 | 1,233 | 1,606 | 1,284 | 1,277 | 119 | 1,035 | 1,496 |
| 2042 | 1,405 | 1,393 | 101 | 1,222 | 1,611 | 1,283 | 1,284 | 121 | 1,067 | 1,511 |
| 2043 | 1,403 | 1,392 | 99 | 1,239 | 1,593 | 1,283 | 1,279 | 117 | 1,041 | 1,488 |
| 2044 | 1,404 | 1,391 | 98 | 1,220 | 1,595 | 1,289 | 1,284 | 113 | 1,097 | 1,490 |
| 2045 | 1,411 | 1,407 | 93 | 1,270 | 1,570 | 1,298 | 1,307 | 108 | 1,083 | 1,469 |

Table F.11. Projected mean, median, and standard deviation (SD) of retained CPUE indices with $95 \%$ confidence limits under $\mathrm{F}_{35 \%}$ harvest control rule for scenario 17_0 for WAG, 2016-2045.

| $\mathrm{F}_{35 \%}\left(0.6 \mathrm{yr}^{-1}\right)$ |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Year | $\begin{gathered} \text { Mean } \\ \text { CPUE } \\ \hline \end{gathered}$ | Median CPUE | $\begin{gathered} \text { SD } \\ \text { CPUE } \\ \hline \end{gathered}$ | $\begin{gathered} \hline 95 \% \\ \text { Lower } \\ \text { Limit } \\ \hline \end{gathered}$ | 95\% <br> Upper <br> Limit |
| 2016 | 0.753 | 0.749 | 0.131 | 0.546 | 1.041 |
| 2017 | 0.887 | 0.880 | 0.151 | 0.661 | 1.223 |
| 2018 | 0.970 | 0.960 | 0.161 | 0.725 | 1.328 |
| 2019 | 0.954 | 0.950 | 0.129 | 0.766 | 1.237 |
| 2020 | 0.908 | 0.893 | 0.092 | 0.772 | 1.095 |
| 2021 | 0.871 | 0.864 | 0.070 | 0.753 | 1.015 |
| 2022 | 0.844 | 0.845 | 0.062 | 0.738 | 0.964 |
| 2023 | 0.831 | 0.823 | 0.055 | 0.739 | 0.935 |
| 2024 | 0.827 | 0.818 | 0.056 | 0.727 | 0.927 |
| 2025 | 0.832 | 0.825 | 0.058 | 0.733 | 0.944 |
| 2026 | 0.834 | 0.829 | 0.052 | 0.754 | 0.930 |
| 2027 | 0.833 | 0.831 | 0.048 | 0.744 | 0.942 |
| 2028 | 0.831 | 0.823 | 0.049 | 0.759 | 0.932 |
| 2029 | 0.830 | 0.818 | 0.058 | 0.742 | 0.968 |
| 2030 | 0.827 | 0.819 | 0.062 | 0.728 | 0.949 |
| 2031 | 0.824 | 0.819 | 0.055 | 0.728 | 0.932 |
| 2032 | 0.823 | 0.814 | 0.059 | 0.717 | 0.939 |
| 2033 | 0.828 | 0.825 | 0.060 | 0.723 | 0.953 |
| 2034 | 0.830 | 0.823 | 0.056 | 0.724 | 0.945 |
| 2035 | 0.828 | 0.822 | 0.050 | 0.746 | 0.933 |
| 2036 | 0.828 | 0.822 | 0.050 | 0.739 | 0.928 |
| 2037 | 0.831 | 0.827 | 0.052 | 0.739 | 0.920 |
| 2038 | 0.831 | 0.829 | 0.057 | 0.736 | 0.937 |
| 2039 | 0.830 | 0.824 | 0.059 | 0.739 | 0.957 |
| 2040 | 0.829 | 0.825 | 0.057 | 0.738 | 0.966 |
| 2041 | 0.826 | 0.819 | 0.060 | 0.723 | 0.951 |
| 2042 | 0.826 | 0.819 | 0.061 | 0.725 | 0.949 |
| 2043 | 0.824 | 0.820 | 0.060 | 0.719 | 0.939 |
| 2044 | 0.824 | 0.818 | 0.060 | 0.710 | 0.936 |
| 2045 | 0.829 | 0.827 | 0.056 | 0.740 | 0.924 |

Table F.12. Projected mean, median, and standard deviation (SD) of legal male (LMB) and mature male (MMB) biomass ( t ) with $95 \%$ confidence limits under no directed fishery (top) and under $\mathrm{F}_{35 \%}$ (bottom) harvest control rule for scenario 17_0d for WAG, 2016-2045. The top table provides a base projection scenario with no directed fishery.

| No Directed Fishery |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | Mean <br> LMB | Median <br> LMB | $\begin{gathered} \text { SD } \\ \text { LMB } \end{gathered}$ | 95\% <br> Lower <br> Limit | 95\% <br> Upper <br> Limit | Mean <br> MMB | Median <br> MMB | $\begin{gathered} \text { SD } \\ \text { MMB } \end{gathered}$ | 95\% <br> Lower <br> Limit | 95\% <br> Upper <br> Limit |
| 2016 | 3,114 | 3,082 | 731 | 1,966 | 4,751 | 7,086 | 7,015 | 1,664 | 4,473 | 10,813 |
| 2017 | 5,231 | 5,179 | 1,229 | 3,302 | 7,982 | 8,653 | 8,565 | 1,897 | 5,747 | 12,848 |
| 2018 | 7,156 | 7,068 | 1,664 | 4,553 | 10,877 | 9,840 | 9,717 | 1,837 | 7,101 | 13,857 |
| 2019 | 8,514 | 8,461 | 1,771 | 5,843 | 12,382 | 10,778 | 10,565 | 1,706 | 8,248 | 14,532 |
| 2020 | 9,472 | 9,275 | 1,675 | 7,021 | 13,147 | 11,512 | 11,415 | 1,567 | 9,062 | 14,823 |
| 2021 | 10,190 | 10,001 | 1,513 | 7,874 | 13,523 | 12,083 | 12,089 | 1,431 | 9,855 | 14,907 |
| 2022 | 10,729 | 10,690 | 1,369 | 8,596 | 13,491 | 12,560 | 12,637 | 1,291 | 10,420 | 15,125 |
| 2023 | 11,162 | 11,189 | 1,228 | 9,209 | 13,492 | 12,975 | 13,059 | 1,184 | 10,983 | 15,295 |
| 2024 | 11,522 | 11,572 | 1,108 | 9,640 | 13,807 | 13,346 | 13,437 | 1,076 | 11,528 | 15,238 |
| 2025 | 11,857 | 11,918 | 1,017 | 10,115 | 13,710 | 13,624 | 13,633 | 955 | 11,823 | 15,302 |
| 2026 | 12,137 | 12,211 | 917 | 10,487 | 13,766 | 13,836 | 13,888 | 857 | 12,253 | 15,430 |
| 2027 | 12,347 | 12,367 | 815 | 10,812 | 13,810 | 14,005 | 14,001 | 801 | 12,511 | 15,313 |
| 2028 | 12,506 | 12,515 | 738 | 11,145 | 13,852 | 14,133 | 14,104 | 790 | 12,770 | 15,557 |
| 2029 | 12,629 | 12,587 | 712 | 11,343 | 13,832 | 14,222 | 14,194 | 758 | 12,864 | 15,647 |
| 2030 | 12,717 | 12,676 | 713 | 11,442 | 14,008 | 14,292 | 14,230 | 735 | 12,931 | 15,650 |
| 2031 | 12,778 | 12,765 | 679 | 11,586 | 14,050 | 14,379 | 14,359 | 782 | 12,929 | 15,883 |
| 2032 | 12,834 | 12,767 | 694 | 11,589 | 14,120 | 14,468 | 14,512 | 796 | 13,067 | 15,984 |
| 2033 | 12,913 | 12,933 | 725 | 11,546 | 14,342 | 14,515 | 14,461 | 770 | 13,146 | 16,056 |
| 2034 | 12,974 | 12,993 | 729 | 11,661 | 14,334 | 14,556 | 14,508 | 736 | 13,248 | 16,009 |
| 2035 | 13,012 | 13,001 | 690 | 11,836 | 14,390 | 14,604 | 14,590 | 721 | 13,224 | 15,910 |
| 2036 | 13,046 | 13,041 | 669 | 11,811 | 14,325 | 14,641 | 14,728 | 711 | 13,185 | 15,965 |
| 2037 | 13,088 | 13,120 | 652 | 11,800 | 14,262 | 14,651 | 14,711 | 724 | 13,144 | 15,807 |
| 2038 | 13,105 | 13,213 | 656 | 11,753 | 14,196 | 14,669 | 14,802 | 736 | 13,219 | 16,065 |
| 2039 | 13,119 | 13,203 | 669 | 11,748 | 14,269 | 14,661 | 14,730 | 785 | 13,079 | 16,111 |
| 2040 | 13,123 | 13,196 | 691 | 11,762 | 14,479 | 14,651 | 14,647 | 852 | 12,886 | 15,926 |
| 2041 | 13,110 | 13,122 | 750 | 11,557 | 14,376 | 14,651 | 14,659 | 867 | 12,829 | 16,081 |
| 2042 | 13,106 | 13,115 | 792 | 11,440 | 14,339 | 14,645 | 14,745 | 867 | 12,909 | 16,137 |
| 2043 | 13,098 | 13,171 | 799 | 11,350 | 14,481 | 14,666 | 14,703 | 859 | 13,009 | 16,142 |
| 2044 | 13,100 | 13,192 | 796 | 11,517 | 14,491 | 14,717 | 14,726 | 847 | 13,374 | 16,156 |
| 2045 | 13,130 | 13,179 | 781 | 11,737 | 14,429 | 14,775 | 14,871 | 857 | 13,271 | 16,112 |


| $\mathrm{F}_{35 \%}\left(0.68 \mathrm{yr}^{-1}\right)$ |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | $\begin{aligned} & \text { Mean } \\ & \text { LMB } \\ & \hline \end{aligned}$ | Median LMB | $\begin{gathered} \text { SD } \\ \text { LMB } \\ \hline \end{gathered}$ | 95\% Lower <br> Limit | 95\% <br> Upper <br> Limit | $\begin{aligned} & \text { Mean } \\ & \text { MMB } \\ & \hline \end{aligned}$ | Median MMB | $\begin{gathered} \text { SD } \\ \text { MMB } \\ \hline \end{gathered}$ | 95\% <br> Lower <br> Limit | 95\% <br> Upper <br> Limit |
| 2016 | 3,114 | 3,082 | 731 | 1,966 | 4,751 | 5,946 | 5,886 | 1,396 | 3,754 | 9,073 |
| 2017 | 4,116 | 4,051 | 930 | 2,738 | 6,244 | 6,187 | 6,132 | 1,300 | 4,262 | 9,067 |
| 2018 | 4,701 | 4,611 | 1,042 | 3,189 | 7,086 | 5,966 | 5,811 | 947 | 4,567 | 8,085 |
| 2019 | 4,652 | 4,571 | 843 | 3,482 | 6,545 | 5,664 | 5,584 | 669 | 4,407 | 6,948 |
| 2020 | 4,397 | 4,288 | 579 | 3,525 | 5,678 | 5,419 | 5,475 | 536 | 4,391 | 6,376 |
| 2021 | 4,182 | 4,150 | 406 | 3,466 | 4,969 | 5,251 | 5,243 | 464 | 4,352 | 6,075 |
| 2022 | 4,024 | 4,011 | 341 | 3,435 | 4,678 | 5,177 | 5,138 | 435 | 4,300 | 5,925 |
| 2023 | 3,944 | 3,903 | 294 | 3,435 | 4,487 | 5,174 | 5,131 | 459 | 4,381 | 6,014 |
| 2024 | 3,918 | 3,848 | 299 | 3,404 | 4,495 | 5,218 | 5,238 | 439 | 4,455 | 5,938 |
| 2025 | 3,943 | 3,909 | 315 | 3,432 | 4,541 | 5,218 | 5,214 | 388 | 4,486 | 5,899 |
| 2026 | 3,957 | 3,928 | 285 | 3,513 | 4,462 | 5,193 | 5,152 | 379 | 4,499 | 5,995 |
| 2027 | 3,945 | 3,925 | 259 | 3,484 | 4,507 | 5,170 | 5,122 | 429 | 4,479 | 6,055 |
| 2028 | 3,931 | 3,889 | 269 | 3,524 | 4,472 | 5,151 | 5,111 | 480 | 4,259 | 6,153 |
| 2029 | 3,924 | 3,871 | 308 | 3,475 | 4,666 | 5,129 | 5,117 | 465 | 4,251 | 5,881 |
| 2030 | 3,911 | 3,871 | 323 | 3,370 | 4,562 | 5,116 | 5,106 | 463 | 4,242 | 5,997 |
| 2031 | 3,895 | 3,842 | 297 | 3,380 | 4,463 | 5,140 | 5,102 | 490 | 4,156 | 6,064 |
| 2032 | 3,893 | 3,829 | 323 | 3,311 | 4,529 | 5,178 | 5,207 | 460 | 4,287 | 6,095 |
| 2033 | 3,922 | 3,897 | 324 | 3,390 | 4,561 | 5,173 | 5,189 | 419 | 4,478 | 6,009 |
| 2034 | 3,931 | 3,906 | 303 | 3,404 | 4,535 | 5,166 | 5,234 | 408 | 4,309 | 5,858 |
| 2035 | 3,923 | 3,903 | 270 | 3,450 | 4,443 | 5,175 | 5,201 | 408 | 4,258 | 5,849 |
| 2036 | 3,919 | 3,919 | 271 | 3,415 | 4,394 | 5,179 | 5,256 | 438 | 4,300 | 5,880 |
| 2037 | 3,934 | 3,941 | 278 | 3,407 | 4,371 | 5,162 | 5,146 | 476 | 4,329 | 6,099 |
| 2038 | 3,928 | 3,924 | 308 | 3,423 | 4,483 | 5,161 | 5,135 | 476 | 4,299 | 6,101 |
| 2039 | 3,923 | 3,892 | 326 | 3,413 | 4,608 | 5,138 | 5,110 | 487 | 4,100 | 6,016 |
| 2040 | 3,917 | 3,876 | 317 | 3,374 | 4,628 | 5,122 | 5,151 | 508 | 4,073 | 6,076 |
| 2041 | 3,901 | 3,857 | 331 | 3,290 | 4,587 | 5,125 | 5,121 | 489 | 4,151 | 5,903 |
| 2042 | 3,900 | 3,881 | 331 | 3,273 | 4,568 | 5,122 | 5,164 | 481 | 4,130 | 5,994 |
| 2043 | 3,896 | 3,880 | 320 | 3,330 | 4,481 | 5,148 | 5,223 | 458 | 4,316 | 5,948 |
| 2044 | 3,900 | 3,877 | 316 | 3,294 | 4,511 | 5,199 | 5,223 | 446 | 4,346 | 5,975 |
| 2045 | 3,924 | 3,918 | 302 | 3,465 | 4,446 | 5,244 | 5,229 | 461 | 4,336 | 6,093 |

Table F.13. Projected mean, median, and standard deviation (SD) of total catch (OFL) and retained catch (RETC) in biomass ( t ) with $95 \%$ confidence limits under no directed fishery (top) and under $\mathrm{F}_{35 \%}$ (bottom) harvest control rule for scenario 17_0d for WAG, 2016-2045.

| No Directed Fishery |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | $\begin{gathered} \text { Mean } \\ \text { OFL } \end{gathered}$ | Median OFL | $\begin{gathered} \text { SD } \\ \text { OFL } \end{gathered}$ | $\begin{gathered} 95 \% \\ \text { Lower } \\ \text { Limit } \\ \hline \end{gathered}$ | 95\% <br> Upper <br> Limit | Mean <br> RETC | Median RETC | $\begin{gathered} \text { SD } \\ \text { RETC } \end{gathered}$ | $\begin{gathered} \hline 95 \% \\ \text { Lower } \\ \text { Limit } \\ \hline \end{gathered}$ | 95\% <br> Upper <br> Limit |
| 2016 | 3.469 | 3.434 | 0.815 | 2.190 | 5.293 | 0 | 0 | 0 | 0 | 0 |
| 2017 | 4.094 | 4.065 | 0.825 | 2.846 | 5.883 | 0 | 0 | 0 | 0 | 0 |
| 2018 | 4.599 | 4.491 | 0.798 | 3.423 | 6.348 | 0 | 0 | 0 | 0 | 0 |
| 2019 | 5.006 | 4.945 | 0.742 | 3.853 | 6.633 | 0 | 0 | 0 | 0 | 0 |
| 2020 | 5.311 | 5.298 | 0.692 | 4.228 | 6.737 | 0 | 0 | 0 | 0 | 0 |
| 2021 | 5.568 | 5.581 | 0.624 | 4.560 | 6.782 | 0 | 0 | 0 | 0 | 0 |
| 2022 | 5.773 | 5.811 | 0.571 | 4.831 | 6.990 | 0 | 0 | 0 | 0 | 0 |
| 2023 | 5.970 | 6.032 | 0.523 | 5.072 | 6.943 | 0 | 0 | 0 | 0 | 0 |
| 2024 | 6.119 | 6.121 | 0.467 | 5.242 | 6.949 | 0 | 0 | 0 | 0 | 0 |
| 2025 | 6.233 | 6.266 | 0.417 | 5.471 | 6.978 | 0 | 0 | 0 | 0 | 0 |
| 2026 | 6.325 | 6.326 | 0.379 | 5.633 | 7.010 | 0 | 0 | 0 | 0 | 0 |
| 2027 | 6.398 | 6.400 | 0.364 | 5.747 | 7.011 | 0 | 0 | 0 | 0 | 0 |
| 2028 | 6.450 | 6.447 | 0.358 | 5.810 | 7.071 | 0 | 0 | 0 | 0 | 0 |
| 2029 | 6.487 | 6.470 | 0.326 | 5.877 | 7.147 | 0 | 0 | 0 | 0 | 0 |
| 2030 | 6.522 | 6.500 | 0.347 | 5.872 | 7.152 | 0 | 0 | 0 | 0 | 0 |
| 2031 | 6.572 | 6.608 | 0.357 | 5.945 | 7.276 | 0 | 0 | 0 | 0 | 0 |
| 2032 | 6.597 | 6.602 | 0.360 | 5.931 | 7.239 | 0 | 0 | 0 | 0 | 0 |
| 2033 | 6.618 | 6.611 | 0.336 | 6.046 | 7.321 | 0 | 0 | 0 | 0 | 0 |
| 2034 | 6.638 | 6.650 | 0.334 | 6.005 | 7.264 | 0 | 0 | 0 | 0 | 0 |
| 2035 | 6.665 | 6.676 | 0.323 | 6.041 | 7.233 | 0 | 0 | 0 | 0 | 0 |
| 2036 | 6.667 | 6.709 | 0.327 | 6.019 | 7.245 | 0 | 0 | 0 | 0 | 0 |
| 2037 | 6.681 | 6.724 | 0.326 | 6.004 | 7.246 | 0 | 0 | 0 | 0 | 0 |
| 2038 | 6.683 | 6.714 | 0.337 | 6.003 | 7.328 | 0 | 0 | 0 | 0 | 0 |
| 2039 | 6.674 | 6.674 | 0.375 | 5.930 | 7.320 | 0 | 0 | 0 | 0 | 0 |
| 2040 | 6.680 | 6.696 | 0.388 | 5.880 | 7.305 | 0 | 0 | 0 | 0 | 0 |
| 2041 | 6.673 | 6.698 | 0.390 | 5.819 | 7.354 | 0 | 0 | 0 | 0 | 0 |
| 2042 | 6.676 | 6.691 | 0.391 | 5.871 | 7.341 | 0 | 0 | 0 | 0 | 0 |
| 2043 | 6.693 | 6.716 | 0.384 | 6.008 | 7.317 | 0 | 0 | 0 | 0 | 0 |
| 2044 | 6.719 | 6.770 | 0.384 | 6.074 | 7.328 | 0 | 0 | 0 | 0 | 0 |
| 2045 | 6.741 | 6.795 | 0.393 | 6.020 | 7.378 | 0 | 0 | 0 | 0 | 0 |


| $\mathrm{F}_{35 \%}\left(0.68 \mathrm{yr}^{-1}\right)$ |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | $\begin{aligned} & \text { Mean } \\ & \text { OFL } \end{aligned}$ | Median OFL | $\begin{gathered} \text { SD } \\ \mathrm{OFL} \\ \hline \end{gathered}$ | 95\% Lower <br> Limit | 95\% <br> Upper <br> Limit | $\begin{aligned} & \text { Mean } \\ & \text { RETC } \\ & \hline \end{aligned}$ | Median RETC | $\begin{gathered} \text { SD } \\ \text { RETC } \\ \hline \end{gathered}$ | 95\% <br> Lower <br> Limit | 95\% <br> Upper <br> Limit |
| 2016 | 1,131 | 1,119 | 266 | 714 | 1,725 | 1,017 | 1,029 | 275 | 516 | 1,586 |
| 2017 | 1,440 | 1,416 | 323 | 964 | 2,181 | 1,322 | 1,314 | 321 | 784 | 2,025 |
| 2018 | 1,656 | 1,625 | 360 | 1,143 | 2,480 | 1,553 | 1,532 | 358 | 1,028 | 2,356 |
| 2019 | 1,664 | 1,636 | 302 | 1,250 | 2,342 | 1,569 | 1,557 | 308 | 1,131 | 2,240 |
| 2020 | 1,584 | 1,548 | 211 | 1,281 | 2,066 | 1,487 | 1,457 | 225 | 1,129 | 1,978 |
| 2021 | 1,508 | 1,483 | 146 | 1,255 | 1,805 | 1,410 | 1,399 | 164 | 1,108 | 1,729 |
| 2022 | 1,452 | 1,449 | 118 | 1,249 | 1,684 | 1,350 | 1,351 | 135 | 1,062 | 1,601 |
| 2023 | 1,421 | 1,406 | 101 | 1,240 | 1,610 | 1,316 | 1,300 | 119 | 1,068 | 1,525 |
| 2024 | 1,410 | 1,394 | 100 | 1,238 | 1,607 | 1,309 | 1,299 | 116 | 1,092 | 1,517 |
| 2025 | 1,416 | 1,398 | 105 | 1,254 | 1,618 | 1,319 | 1,307 | 119 | 1,111 | 1,534 |
| 2026 | 1,420 | 1,409 | 97 | 1,269 | 1,588 | 1,323 | 1,322 | 109 | 1,123 | 1,502 |
| 2027 | 1,417 | 1,411 | 88 | 1,263 | 1,599 | 1,314 | 1,305 | 101 | 1,135 | 1,510 |
| 2028 | 1,413 | 1,391 | 91 | 1,278 | 1,603 | 1,305 | 1,276 | 111 | 1,119 | 1,509 |
| 2029 | 1,410 | 1,384 | 103 | 1,269 | 1,658 | 1,302 | 1,286 | 126 | 1,072 | 1,568 |
| 2030 | 1,406 | 1,386 | 108 | 1,231 | 1,628 | 1,298 | 1,284 | 130 | 1,070 | 1,551 |
| 2031 | 1,401 | 1,381 | 101 | 1,234 | 1,598 | 1,294 | 1,300 | 124 | 1,068 | 1,506 |
| 2032 | 1,400 | 1,384 | 107 | 1,220 | 1,620 | 1,296 | 1,293 | 125 | 1,071 | 1,531 |
| 2033 | 1,408 | 1,393 | 109 | 1,230 | 1,635 | 1,306 | 1,300 | 124 | 1,087 | 1,547 |
| 2034 | 1,412 | 1,409 | 102 | 1,238 | 1,624 | 1,310 | 1,314 | 120 | 1,114 | 1,539 |
| 2035 | 1,410 | 1,401 | 91 | 1,260 | 1,584 | 1,309 | 1,318 | 110 | 1,086 | 1,502 |
| 2036 | 1,409 | 1,408 | 90 | 1,244 | 1,568 | 1,305 | 1,302 | 113 | 1,072 | 1,484 |
| 2037 | 1,413 | 1,413 | 93 | 1,234 | 1,565 | 1,306 | 1,318 | 116 | 1,075 | 1,482 |
| 2038 | 1,412 | 1,406 | 102 | 1,240 | 1,604 | 1,307 | 1,310 | 123 | 1,061 | 1,508 |
| 2039 | 1,410 | 1,395 | 110 | 1,239 | 1,638 | 1,302 | 1,300 | 133 | 1,033 | 1,554 |
| 2040 | 1,408 | 1,392 | 108 | 1,224 | 1,644 | 1,298 | 1,303 | 135 | 1,011 | 1,559 |
| 2041 | 1,403 | 1,392 | 111 | 1,199 | 1,626 | 1,294 | 1,284 | 134 | 1,006 | 1,535 |
| 2042 | 1,402 | 1,394 | 111 | 1,195 | 1,634 | 1,293 | 1,292 | 135 | 1,048 | 1,548 |
| 2043 | 1,401 | 1,390 | 108 | 1,216 | 1,599 | 1,294 | 1,287 | 129 | 1,032 | 1,509 |
| 2044 | 1,402 | 1,394 | 106 | 1,204 | 1,622 | 1,301 | 1,304 | 124 | 1,099 | 1,536 |
| 2045 | 1,410 | 1,408 | 102 | 1,258 | 1,590 | 1,310 | 1,326 | 120 | 1,084 | 1,509 |

Table F.14. Projected mean, median, and standard deviation (SD) of retained CPUE indices with $95 \%$ confidence limits under $\mathrm{F}_{35 \%}$ harvest control rule for scenario 17_0d for WAG, 2016-2045.

| $\mathrm{F}_{35 \%}\left(0.68 \mathrm{yr}^{-1}\right)$ |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Year | $\begin{aligned} & \text { Mean } \\ & \text { CPUE } \\ & \hline \end{aligned}$ | Median CPUE | $\begin{gathered} \text { SD } \\ \text { CPUE } \\ \hline \end{gathered}$ | 95\% <br> Lower <br> Limit | 95\% <br> Upper <br> Limit |
| 2016 | 0.562 | 0.554 | 0.128 | 0.369 | 0.855 |
| 2017 | 0.733 | 0.719 | 0.163 | 0.499 | 1.109 |
| 2018 | 0.846 | 0.828 | 0.185 | 0.576 | 1.272 |
| 2019 | 0.841 | 0.825 | 0.152 | 0.628 | 1.184 |
| 2020 | 0.795 | 0.778 | 0.104 | 0.645 | 1.029 |
| 2021 | 0.755 | 0.746 | 0.071 | 0.638 | 0.894 |
| 2022 | 0.726 | 0.723 | 0.059 | 0.628 | 0.843 |
| 2023 | 0.711 | 0.706 | 0.051 | 0.622 | 0.808 |
| 2024 | 0.706 | 0.697 | 0.052 | 0.613 | 0.805 |
| 2025 | 0.710 | 0.703 | 0.055 | 0.621 | 0.815 |
| 2026 | 0.713 | 0.707 | 0.050 | 0.634 | 0.801 |
| 2027 | 0.712 | 0.709 | 0.045 | 0.635 | 0.808 |
| 2028 | 0.710 | 0.703 | 0.046 | 0.643 | 0.807 |
| 2029 | 0.708 | 0.697 | 0.052 | 0.635 | 0.839 |
| 2030 | 0.706 | 0.695 | 0.056 | 0.614 | 0.820 |
| 2031 | 0.703 | 0.693 | 0.051 | 0.615 | 0.801 |
| 2032 | 0.702 | 0.693 | 0.056 | 0.607 | 0.813 |
| 2033 | 0.707 | 0.703 | 0.057 | 0.608 | 0.820 |
| 2034 | 0.709 | 0.705 | 0.053 | 0.611 | 0.813 |
| 2035 | 0.707 | 0.702 | 0.047 | 0.629 | 0.797 |
| 2036 | 0.707 | 0.704 | 0.046 | 0.622 | 0.787 |
| 2037 | 0.710 | 0.706 | 0.047 | 0.616 | 0.786 |
| 2038 | 0.709 | 0.708 | 0.053 | 0.621 | 0.804 |
| 2039 | 0.708 | 0.701 | 0.056 | 0.619 | 0.826 |
| 2040 | 0.707 | 0.699 | 0.054 | 0.625 | 0.831 |
| 2041 | 0.705 | 0.698 | 0.057 | 0.602 | 0.822 |
| 2042 | 0.704 | 0.700 | 0.057 | 0.605 | 0.821 |
| 2043 | 0.703 | 0.698 | 0.055 | 0.599 | 0.805 |
| 2044 | 0.703 | 0.701 | 0.055 | 0.597 | 0.814 |
| 2045 | 0.707 | 0.707 | 0.052 | 0.626 | 0.799 |

## Appendix H. B0 Analysis

For proper B0 analysis, a stock-recruitment relationship and impacts of environmental factors on recruitment are needed. We did not establish a stock-recruitment relationship for Aleutian Islands golden king crab. Furthermore, the impacts of environmental factors on recruitment have not been studied in the Aleutian Islands areas. Therefore, we approached the B0 analysis in a simple way. We computed the time series of B0 values using the same recruitment time series estimated by the base assessment model 17_0 and setting all directed and bycatch fishing mortality to zero. Figure H. 1 compares the time series of estimated B0 and MMB with fishing and MMB ratio (MMB/B0) for scenario 17_0 separately for EAG and WAG. It is clear that the fishery has a great impact on the biomass dynamics with MMB dropping precipitously with the onset of significant fishery removals in 1981.


Figure H.1. Estimated B0 (t) (dark green curve) and MMB (t) with fishing (black curve with +/2SE) (top panel ); and MMB/B0 ratio (bottom panel) from 1960 to 2016 for scenario 17_0 for Aleutian Islands golden king crab in EAG (left) and WAG (right). (Note: 2016 MMB= MMB estimated on 15 February 2017).

