# Aleutian Islands Golden King Crab 

May 2021

M.S.M. Siddeek ${ }^{1,}{ }^{\text {J. }}$ Zheng $^{1}$, C. Siddon ${ }^{1}$, B. Daly ${ }^{2}$, M.J. Westphal ${ }^{3}$, and L. Hulbert ${ }^{1}$

${ }^{1}$ Alaska Dept Fish and Game, Juneau
${ }^{2}$ ADF\&G, Kodiak
${ }^{3}$ ADF\&G, Dutch Harbor

## A. Executive Summary

## 1. Stock

Golden king crab, Lithodes aequispinus, Aleutian Islands, east of $174^{\circ} \mathrm{W}$ longitude (EAG) and west of $174^{\circ} \mathrm{W}$ longitude (WAG).

## 2. Catches

The Aleutian Islands golden king crab (AIGKC) commercial fishery has been prosecuted every year since $1981 / 82$. Retained catch peaked in $1986 / 87$ at $2,686 \mathrm{t}(5,922,425 \mathrm{lb})$ and $3,999 \mathrm{t}$ $(8,816,319 \mathrm{lb})$, respectively, for EAG and WAG, but the retained catch dropped sharply from 1989/90 to 1990/91. The fishery has been managed separately east (EAG) and west (WAG) of $174^{\circ}$ W longitude since 1996/97, and Guideline Harvest Levels (GHLs) of 1,452 $t(3,200,000$ $\mathrm{lb})$ for EAG and $1,225 \mathrm{t}(2,700,000 \mathrm{lb})$ for WAG were introduced into management. The GHL was subsequently reduced to $1,361 \mathrm{t}(3,000,000 \mathrm{lb})$ beginning in $1998 / 99$ for EAG. The reduced harvest levels remained at $1,361 \mathrm{t}(3,000,000 \mathrm{lb})$ for EAG and $1,225 \mathrm{t}(2,700,000 \mathrm{lb})$ for WAG through 2007/08, but were increased to $1,429 \mathrm{t}(3,150,000 \mathrm{lb})$ for EAG and $1,294 \mathrm{t}$ (2,835,000 lb) for WAG beginning with the 2008/09 fishing season following an Alaska Board of Fisheries (BOF) decision. The management specification changed from GHL to TAC (Total Allowable Catch) with adoption of the Crab Rationalization Program in 2005/06 (NPFMC 2007b). The TACs were increased by another BOF decision to $1,501 \mathrm{t}(3,310,000 \mathrm{lb})$ for EAG and $1,352 \mathrm{t}(2,980,000 \mathrm{lb})$ for WAG beginning with the $2012 / 13$ fishing season. The below par fishery performance in WAG in 2014/15 and 2015/16 lead to reduction in TAC to $1,014 \mathrm{t}$ $(2,235,000 \mathrm{lb})$, which reflected a $25 \%$ reduction in the TAC for WAG, while the TAC for EAG was kept at the same level, $1,501 \mathrm{t}(3,310,000 \mathrm{lb})$ for the $2016 / 17$ through $2017 / 18$ fishing seasons. With the improved fishery performance and stock status in 2017/18, the TACs were further increased to $1,134 \mathrm{t}(2,500,000 \mathrm{lb})$ for WAG and $1,749 \mathrm{t}(3,856,000 \mathrm{lb})$ for EAG beginning with the 2018/19 fishing season. With the implementation of a revised state harvest strategy in 2019, the TACs were further increased to $1,302 \mathrm{t}(2,870,000 \mathrm{lb})$ for WAG and 1,955 $t(4,310,000 \mathrm{lb})$ for EAG. Based on the model estimated abundances, the 2020/21 fishing season TACs were adjusted to $1,343 \mathrm{t}(2,960,000 \mathrm{lb})$ for WAG and $1,656 \mathrm{t}(3,650,000 \mathrm{lb})$ for EAG.

Catches have been regularly under the GHL/TAC and the fishery has harvested close to allowable levels since 1996/97. These TAC levels were set below the ABCs determined under Tier 5 criteria (considering 1991-1995 mean catch for the whole Aleutian Islands region, 3,145 $t(6,933,822 \mathrm{lb})$, as the catch limit) under the most recent crab management plan. A new harvest strategy based on model estimated mature male abundance was accepted by the BOF in March

2019, specifying a $15 \%$ maximum harvest rate for EAG and $20 \%$ maximum harvest rate for WAG, and implemented during the 2019/20 fishery. In addition to the retained catch allotted as TAC, there was retained catch in a cost-recovery fishery towards a \$300,000 goal in 2013/14 and 2014/15 to fund an onboard observer program, and towards a $\$ 500,000$ goal in 2015/16 to 2020/21 to fund an onboard observer program and stock survey.

Total mortality of Aleutian Islands golden king crab includes retained catch in the directed and the cost recovery fisheries, mortality of discarded catch, and bycatch in fixed-gear and trawl groundfish fisheries, though bycatch in other fisheries is low compared to mortality in the directed fishery. Total retained catch in the post-rationalized fishery (2005/06-2020/21) has ranged from $2,387 \mathrm{t}(5,262,000 \mathrm{lb})$ to $3,319 \mathrm{t}(7,316,853 \mathrm{lb})$. Total mortality ranged from 2,506 $\mathrm{t}(5,525,000 \mathrm{lb})$ to $3,733 \mathrm{t}(8,230,000 \mathrm{lb})$ for the same period. Total retained catch in 2020/21 was $2,770 \mathrm{t}(6,106,746 \mathrm{lb}): 1,733 \mathrm{t}(3,821,118 \mathrm{lb})$ from the EAG fishery (which included costrecovery catch), and $1,037 \mathrm{t}(2,285,628 \mathrm{lb})$ from the WAG fishery. Discarded (non-retained) catch occurs mainly during the directed fishery. Although low levels of discarded catch can occur during other crab fisheries, there have been no such fisheries prosecuted locally since 2004/05, except as surveys for red king crab conducted under an Alaska Department of Fish and Game (ADF\&G) Commissioner's Permit (and no golden king crab were caught during the cooperative red king crab survey performed by industry and ADF\&G in the Adak area in September 2015; Hilsinger et al. 2016). Estimates of the bycatch mortality during crab fisheries decreased during 1995/96-2005/06, both in absolute value and relative to the retained catch weight, and stabilized during 2005/06-2014/15. Total estimated bycatch mortality during crab fisheries in 2020/21 was $241 \mathrm{t}(531,000 \mathrm{lb})$ for EAG and $81 \mathrm{t}(179,000 \mathrm{lb})$ for WAG (WAG fishery on-going). Discarded catch also occurs during fixed-gear and trawl groundfish fisheries but is small relative to the directed fishery. Groundfish fisheries are a minor contributor to total fishery discard mortality, $40 \mathrm{t}(88,000 \mathrm{lb})$ for EAG and $17 \mathrm{t}(37,000 \mathrm{lb})$ for WAG in 2020/21.

Catch per unit effort (CPUE, i.e., catch per pot lift) of retained legal males decreased from the 1980s into the mid-1990s, but increased after 1994/95, particularly with the initiation of the Crab Rationalization Program in 2005/06. Although CPUE for the two areas showed similar trends through 2010/11, during 2011/12-2014/15 CPUE trends have diverged (increasing for EAG and decreasing for WAG).

A cooperative golden king crab survey was performed by the Aleutian Islands King Crab Foundation (an industry group) and ADF\&G in the EAG and WAG (beginning in August 2018) fisheries, by vessels that were commercial fishing (i.e., each vessel fishing an allotted share of total allowable catch). The cooperative survey was not conducted in 2020/21 due to COVID-19. For catch accounting, it was assumed that bycatch mortality that occurred during any survey was accounted for by reported discards for the season's fishery.

## 3. Stock biomass

Estimated mature male biomass (MMB) for EAG under all scenarios decreased from the 1980s to the 1990s, then increased during the 2000s and systematically increased since 2014. Estimated MMB for WAG decreased during the late 1980s and 1990s, increased during the 2000s, decreased for several years since 2009 and has increased since 2014. The low levels of MMB for EAG were observed in 1995-1997 and in 1990s for WAG. Stock trends reflected the fishery standardized CPUE trends in both regions.

## 4. Recruitment

The numbers of recruits to the model size groups under all scenarios have fluctuated in both EAG and WAG. For EAG, model recruitment was high during 2017-2019, highest in 2018; and lowest in 1986. The model recruitment for WAG was high during 1984 to 1986, highest in 1985, and lowest in 2011. A slightly increasing trend in recruitment was observed since 2011 in WAG.

## 5. Management performance

The size-based assessment model was accepted at the September 2016 CPT and October 2016 SSC meetings for OFL determination for the 2017/18 fishery cycle. In addition, the CPT in January 2017 and SSC in February 2017 recommended using the Tier 3 method to compute OFL and ABC. The assessment model was first used for setting OFL and ABC for the 2017/18 fishing season. The CPT in May 2017 and SSC in June 2017 accepted the authors' recommendation of using scenario 9 (i.e., model using the knife-edge maturity to determine MMB) for OFL and ABC calculation. During the May 2017 meeting, the CPT noted that a single OFL and ABC are defined for Aleutian Islands golden king crab (AIGKC), however; separate models are available by area. Hence, following previous assessments, OFLs and ABCs by area were summed to calculate OFL and ABC for the entire stock.

All models for EAG and WAG used the previous season's fishery information (i.e., 2020/21 fishery, concluded in EAG and 77\% of TAC achieved in WAG). We recommend any one from three models for EAG and WAG: model 21.1a (selection of a fixed period, 1987-2017, for mean number of recruit calculation for reference points estimation; and standardization of observer and fishery CPUE by the negative binomial generalized linear model); model 21.1b (same as 21.1a but consideration of three total selectivity periods to reduce retrospective pattern in EAG); and 21.1c (same as 21.1a but consideration of year and area interaction factor for observer CPUE standardization).

We also proposed variants of the above three models: 21.1a1, 21.1b1, and 21.1c1 [knife-edge maturity size increased by one size bin to 116 mm carapace length for mature male biomass (MMB) estimation]; and 21.1a2, 21.1b2, and 21.1c2 (logistic maturity curve for MMB estimation) for evaluation.

Model 19.1 is the base model (accepted model 19.1 in 2019) with the knife-edge male maturity at 111 mm CL , an $M$ of $0.21 \mathrm{yr}^{-1}$, selection of a fixed period, 1987-2012, for mean number of recruit calculation for reference points estimation; and addition of up to 2020/21 data. Models 21.1a, 21.1b, and 21.1c are modifications from the base model. Although we list the references points of model 19.1, we do not recommend selecting this model because different fixed period was considered for the mean recruit calculation.

The data for the EAG are complete through the 2020/21 season. The fishery in the WAG was still operating when the assessment was conducted, with $77 \%$ of the TAC taken (as of the 26 March 2021 reported summary). In a hypothetical scenario when the WAG fishery were to achieve the whole TAC by end of the 2020/21 fishing season, based on current retained to total fishing mortality ratio in WAG, it would result in a total catch of $3,482 \mathrm{t}$, which would be less than 2020/21 ABC.

Status and catch specifications (1000 t) of Aleutian Islands golden king crab

| Year | MSST | Biomass <br> (MMB) | TAC | Retained <br> Catch | Total <br> Catch $^{\mathbf{a}}$ | OFL | ABC $^{\mathbf{b}}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $2017 / 18$ | 6.044 | 14.205 | 2.515 | 2.585 | 2.942 | 6.048 | 4.536 |
| $2018 / 19$ | 5.880 | 17.848 | 2.883 | 2.965 | 3.355 | 5.514 | 4.136 |
| $2019 / 20$ | 5.915 | 16.386 | 3.257 | 3.319 | 3.735 | 5.249 | 3.937 |
| $2020 / 21^{\mathrm{d}}$ | $5.998^{\mathrm{d}, \mathrm{c}}$ | $16.215^{\mathrm{d}, \mathrm{c}}$ | 2.999 | $2.770^{\mathrm{c}}$ | $3.148^{\mathrm{c}}$ | 4.798 | 3.599 |
| $2020 / 21^{\mathrm{e}}$ | 6.026 | 16.207 |  |  |  |  |  |
| $2020 / 21^{\mathrm{f}}$ | 5.964 | 14.871 |  |  |  |  |  |
| $2020 / 21^{\mathrm{g}}$ | 5.991 | 15.124 |  |  |  |  |  |
| $2021 / 22^{\mathrm{d}}$ |  | 14.821 |  |  |  | 4.796 | 3.615 |
| $2021 / 22^{\mathrm{e}}$ |  | 14.816 |  |  |  | 4.817 | 3.613 |
| $2021 / 22^{\mathrm{f}}$ |  | 13.832 |  |  |  | 4.407 | 3.305 |
| $2021 / 22^{\mathrm{g}}$ |  | 14.048 |  |  |  | 4.471 | 3.353 |

Status and catch specifications (million lb) of Aleutian Islands golden king crab

| Year | MSST | Biomass <br> (MMB) | TAC | Retained <br> Catch $^{2}$ | Total <br> Catch $^{\mathbf{a}}$ | OFL | ABC $^{\mathbf{b}}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $2017 / 18$ | 13.325 | 31.315 | 5.545 | 5.699 | 6.487 | 13.333 | 10.000 |
| $2018 / 19$ | 12.964 | 39.348 | 6.356 | 6.536 | 7.396 | 12.157 | 9.118 |
| $2019 / 20$ | 13.041 | 36.124 | 7.180 | 7.317 | 8.234 | 11.572 | 8.679 |
| $2020 / 21^{\mathrm{d}}$ | $13.223^{\mathrm{d}, \mathrm{c}}$ | $35.748^{\mathrm{d}, \mathrm{c}}$ | 6.610 | $6.107^{\mathrm{c}}$ | $6.940^{\mathrm{c}}$ | 10.579 | 7.934 |
| $2020 / 21^{\mathrm{e}}$ | 13.284 | 35.730 |  |  |  |  |  |
| $2020 / 21^{\mathrm{f}}$ | 13.148 | 32.784 |  |  |  |  |  |
| $2020 / 21^{\mathrm{g}}$ | 13.207 | 33.341 |  |  |  |  |  |
| $2021 / 22^{\mathrm{d}}$ |  | 32.675 |  |  |  | 10.573 | 7.969 |
| $2021 / 22^{\mathrm{e}}$ |  | 32.662 |  |  |  | 10.620 | 7.965 |
| $2021 / 22^{\mathrm{f}}$ |  | 30.494 |  |  |  | 9.715 | 7.287 |
| $2021 / 22^{\mathrm{g}}$ |  | 30.971 |  |  |  | 9.857 | 7.393 |

a. Total retained catch plus estimated bycatch mortality of discarded bycatch during crab fisheries and groundfish fisheries.
b. $25 \%$ buffer was applied to total catch OFL to determine ABC.
c. WAG fishery was still being prosecuted when the assessment was conducted.
d. Model 19.1(base model)
e. Model 21.1a.
f. Model 21.1b.
g. Model 21.1c.

## 6. Basis for the OFL

The length-based model developed for the Tier 3 analysis estimated mature male biomass (MMB) on February 15 each year for the period 1961 through 2021. The terminal year mature male biomass was projected by an additional year to determine OFL and ABC for the 2021/22 season. The Tier 3 approach uses a constant annual natural mortality ( $M$ ), knifeedge maturity size/maturity curve, and the mean number of recruits for different time periods for OFL and ABC calculation. Previously derived $M$ of $0.21 \mathrm{yr}^{-1}$ from the combined data from the EAG and WAG data was used (Siddeek et al. 2018).

We provided OFL and ABC estimates for EAG and WAG separately and combined (i.e., for the entire Aleutian Islands; AI) from ten models, 19.1, 21.1a, 21.1b, and 21.1c (CPT/SSC suggested models); and 21.1a1, 21.1b1, 21.1c1, 21.1a2, 21.1b2, and 21.1c2 for EAG, WAG, and AI in the following six tables. The stock statuses of all models were above $\mathrm{MMB}_{35 \%}$ (i.e., they were in Tier 3a) except two WAG models, 21.1c1 and 21.1c2. The mean recruitment estimation period for those WAG models was reduced to 1993-1997 to bring up the current stock statuses above $\mathrm{MMB}_{35 \%}$ for combining the two regions' OFL and ABC (following Siddeek, et al. 2017). The modified forms of WAG models were labeled as 21.1c1Ver2 and 21.1 c 2 Ver 2 and reference points only for those models are listed in the tables.

## EAG (Tier 3):

Biomass, total OFL, and ABC for the next fishing season in millions of pounds. Current MMB $=\mathrm{MMB}$ on 15 Feb .2022.

| Model | Tier | $M M B_{35 \%}$ | Current |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| MMB | MMB/ <br> $M M B_{35 \%}$ | $F_{\text {OFL }}$ | Recruitment Years <br> to define $M M B_{35 \%}$ | $F_{35 \%}$ |  | OFL | ABC <br> $\left(\mathrm{P}^{*}=0.49\right)$ | ABC <br> $(0.75 * O F L)$ |  |  |
| EAG19.1 | 3a | 14.762 | 19.229 | 1.30 | 0.61 | $1987-2012$ | 0.61 | 6.470 | 6.432 | 4.853 |
| EAG21.1a | 3a | 14.907 | 19.219 | 1.29 | 0.61 | $1987-2017$ | 0.61 | 6.466 | 6.429 | 4.850 |
| EAG21.1b | 3a | 14.658 | 16.960 | 1.16 | 0.56 | $1987-2017$ | 0.56 | 5.545 | 5.516 | 4.159 |
| EAG21.1c | 3a | 14.893 | 19.111 | 1.28 | 0.61 | $1987-2017$ | 0.61 | 6.433 | 6.401 | 4.825 |
| EAG21.1a1 | 3a | 14.353 | 18.510 | 1.29 | 0.55 | $1987-2017$ | 0.55 | 5.956 | 5.921 | 4.467 |
| EAG21.1a2 | 3a | 13.068 | 16.047 | 1.23 | 0.48 | $1987-2017$ | 0.48 | 5.330 | 5.299 | 3.998 |
| EAG21.1b1 | 3a | 14.202 | 16.245 | 1.14 | 0.50 | $1987-2017$ | 0.50 | 5.063 | 5.037 | 3.798 |
| EAG21.1b2 | 3a | 12.928 | 14.124 | 1.09 | 0.44 | $1987-2017$ | 0.44 | 4.559 | 4.535 | 3.419 |
| EAG21.1c1 | 3a | 14.468 | 18.480 | 1.28 | 0.54 | $1987-2017$ | 0.54 | 5.838 | 5.809 | 4.379 |
| EAG21.1c2 | 3a | 13.055 | 15.955 | 1.22 | 0.48 | $1987-2017$ | 0.48 | 5.303 | 5.276 | 3.977 |

Biomass, total OFL, and ABC for the next fishing season in 1000 t .

| Model | Tier | MMB $35 \%$ | Current <br> MMB | $\begin{gathered} \mathrm{MMB} / \\ M M B_{35 \%} \end{gathered}$ | $F_{\text {OFL }}$ | Recruitment Years to Define $M M B_{35 \%}$ | $F_{35 \%}$ | OFL | $\begin{gathered} \mathrm{ABC} \\ \left(\mathrm{P}^{*}=0.49\right) \end{gathered}$ | $\begin{gathered} \mathrm{ABC} \\ (0.75 * \mathrm{OFL}) \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| EAG19.1 | 3a | 6.69585 | 8.72217 | 1.30 | 0.61 | 1987-2012 | 0.61 | 2.935086 | 2.918135 | 2.201314 |
| EAG21.1a | 3 a | 6.76171 | 8.71756 | 1.29 | 0.61 | 1987-2017 | 0.61 | 2.932993 | 2.916042 | 2.199745 |
| EAG21.1b | 3a | 6.64877 | 7.69293 | 1.16 | 0.56 | 1987-2017 | 0.56 | 2.515111 | 2.502109 | 1.886333 |
| EAG21.1c | 3a | 6.75540 | 8.66857 | 1.28 | 0.61 | 1987-2017 | 0.61 | 2.918103 | 2.903490 | 2.188577 |
| EAG21.1a1 | 3a | 6.51066 | 8.39605 | 1.29 | 0.55 | 1987-2017 | 0.55 | 2.701414 | 2.685591 | 2.026060 |
| EAG21.1a2 | 3a | 5.92744 | 7.27904 | 1.23 | 0.48 | 1987-2017 | 0.48 | 2.417822 | 2.403729 | 1.813367 |
| EAG21.1b1 | 3 a | 6.44179 | 7.36860 | 1.14 | 0.50 | 1987-2017 | 0.50 | 2.296729 | 2.284992 | 1.722546 |
| EAG21.1b2 | 3 a | 5.86180 | 6.40658 | 1.09 | 0.44 | 1987-2017 | 0.44 | 2.067846 | 2.057141 | 1.550884 |
| EAG21.1c1 | 3 a | 6.56279 | 8.38248 | 1.28 | 0.54 | 1987-2017 | 0.54 | 2.648313 | 2.635166 | 1.986235 |
| EAG21.1c2 | 3a | 5.92153 | 7.23736 | 1.22 | 0.48 | 1987-2017 | 0.48 | 2.405558 | 2.393397 | 1.804168 |

## WAG (Tier 3):

Biomass, total OFL, and ABC for the next fishing season in millions of pounds. Current MMB $=\mathrm{MMB}$ on 15 Feb .2022.

| Model | Tier | MMB ${ }_{35 \%}$ | Current <br> MMB | $\begin{gathered} \hline \mathrm{MMB} / \\ M M B_{35 \%} \end{gathered}$ | $F_{\text {OFL }}$ | Recruitment Years to Define $M M B_{35 \%}$ | $F_{35 \%}$ | OFL | $\begin{gathered} \mathrm{ABC} \\ \left(\mathrm{P}^{*}=0.49\right) \end{gathered}$ | $\begin{gathered} \hline \mathrm{ABC} \\ (0.75 * \mathrm{OFL}) \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| WAG19.1 | 3a | 11.684 | 13.446 | 1.15 | 0.57 | 1987-2012 | 0.57 | 4.155 | 4.140 | 3.116 |
| WAG21.1a | 3a | 11.662 | 13.444 | 1.15 | 0.57 | 1987-2017 | 0.57 | 4.154 | 4.138 | 3.115 |
| WAG21.1b | 3 a | 11.638 | 13.534 | 1.16 | 0.57 | 1987-2017 | 0.57 | 4.171 | 4.155 | 3.128 |
| WAG21.1c | 3 a | 11.521 | 11.860 | 1.03 | 0.57 | 1987-2017 | 0.57 | 3.424 | 3.410 | 2.568 |
| WAG21.1a1 | 3 a | 11.346 | 12.750 | 1.12 | 0.50 | 1987-2017 | 0.50 | 3.741 | 3.728 | 2.806 |
| WAG21.1a2 | 3 a | 10.297 | 11.108 | 1.08 | 0.44 | 1987-2017 | 0.44 | 3.308 | 3.291 | 2.481 |
| WAG21.1b1 | 3 a | 11.319 | 12.837 | 1.13 | 0.50 | 1987-2017 | 0.50 | 3.757 | 3.742 | 2.818 |
| WAG21.1b2 | 3 a | 10.270 | 11.165 | 1.09 | 0.44 | 1987-2017 | 0.44 | 3.384 | 3.371 | 2.538 |
| WAG21.1c1Ver2 | 3 a | 10.097 | 11.082 | 1.10 | 0.51 | 1993-1997 | 0.51 | 3.135 | 3.122 | 2.351 |
| WAG21.1c2Ver2 | 3 a | 9.145 | 9.655 | 1.06 | 0.45 | 1993-1997 | 0.45 | 2.830 | 2.818 | 2.122 |

Biomass, total OFL and ABC for the next fishing season in 1000 t .

| Model | Tier | MMB ${ }_{35 \%}$ | Current <br> MMB | $\begin{gathered} \hline \mathrm{MMB} / \\ M M B_{35 \%} \end{gathered}$ | $F_{\text {OFL }}$ | Recruitment Years to Define $M M B_{35 \%}$ | $F_{35 \%}$ | OFL | $\begin{gathered} \mathrm{ABC} \\ \left(\mathrm{P}^{*}=0.49\right) \end{gathered}$ | $\begin{gathered} \hline \mathrm{ABC} \\ (0.75 * \mathrm{OFL}) \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| WAG19.1 | 3a | 5.29998 | 6.09902 | 1.15 | 0.57 | 1987-2012 | 0.57 | 1.884520 | 1.877668 | 1.413390 |
| WAG21.1a | 3a | 5.28969 | 6.09798 | 1.15 | 0.57 | 1987-2017 | 0.57 | 1.884023 | 1.877172 | 1.413017 |
| WAG21.1b | 3a | 5.27901 | 6.13891 | 1.16 | 0.57 | 1987-2017 | 0.57 | 1.891769 | 1.884480 | 1.418827 |
| WAG21.1c | 3a | 5.22569 | 5.37973 | 1.03 | 0.57 | 1987-2017 | 0.57 | 1.553061 | 1.546571 | 1.164796 |
| WAG21.1a1 | 3a | 5.14652 | 5.78336 | 1.12 | 0.50 | 1987-2017 | 0.50 | 1.697025 | 1.690875 | 1.272768 |
| WAG21.1a2 | 3a | 4.67050 | 5.03845 | 1.08 | 0.44 | 1987-2017 | 0.44 | 1.528387 | 1.522805 | 1.146290 |
| WAG21.1b1 | 3a | 5.13443 | 5.82280 | 1.13 | 0.50 | 1987-2017 | 0.50 | 1.704275 | 1.697558 | 1.278206 |
| WAG21.1b2 | 3a | 4.65828 | 5.06458 | 1.09 | 0.44 | 1987-2017 | 0.44 | 1.534941 | 1.529012 | 1.151206 |
| WAG21.1c1Ver2 | 3a | 4.57997 | 5.02675 | 1.10 | 0.51 | 1993-1997 | 0.51 | 1.422027 | 1.416168 | 1.066520 |
| WAG21.1c2Ver2 | 3a | 4.14813 | 4.37926 | 1.06 | 0.45 | 1993-1997 | 0.45 | 1.283679 | 1.278425 | 0.962759 |

Aleutian Islands (AI)
Total OFL and ABC for the next fishing season in millions of pounds.

| Model | OFL | ABC <br> $\left(\mathrm{P}^{*}=0.49\right)$ | ABC <br> $(0.75 * \mathrm{OFL})$ |
| :---: | :---: | :---: | :---: |
| 19.1 | 10.625 | 10.573 | 7.969 |
| 21.1a | 10.620 | 10.567 | 7.965 |
| 21.1b | 9.715 | 9.671 | 7.287 |
| 21.1c | 9.857 | 9.811 | 7.393 |
| 21.1a1 | 9.697 | 9.648 | 7.273 |
| 21.1a2 | 8.700 | 8.656 | 6.525 |
| 21.1b1 | 8.821 | 8.780 | 6.615 |
| 21.1b2 | 7.943 | 7.906 | 5.957 |
| 21.1c1 (with Ver 2 model for WAG) | 8.973 | 8.932 | 6.730 |
| 21.1c2(with Ver 2 model for WAG) | 8.133 | 8.095 | 6.100 |

Aleutian Islands (AI)
Total OFL and ABC for the next fishing season in 1000 t .

| Model | OFL | ABC <br> $\left(\mathrm{P}^{*}=0.49\right)$ | ABC <br> $(0.75 * \mathrm{OFL})$ |
| :---: | :---: | :---: | :---: |
| 19.1 | 4.81961 | 4.79580 | 3.61470 |
| 21.1a | 4.81702 | 4.79321 | 3.61276 |
| 21.1 b | 4.40688 | 4.38659 | 3.30516 |
| 21.1 c | 4.47116 | 4.45006 | 3.35337 |
| 21.1a1 | 4.39844 | 4.37647 | 3.29883 |
| 21.1a2 | 3.94621 | 3.92653 | 2.95966 |
| 21.1b1 | 4.00100 | 3.98255 | 3.00075 |
| 21.1b2 | 3.60279 | 3.58615 | 2.70209 |
| 21.1c1 (with Ver 2 model for WAG) | 4.07034 | 4.05133 | 3.05276 |
| 21.1c2(with Ver 2 model for WAG) | 3.68924 | 3.67182 | 2.76693 |

7. Probability density functions of the OFL

Assuming a lognormal distribution of total OFL, we determined the cumulative distributions of OFL and selected the median as the OFL.
8. Basis for the $A B C$ recommendation

A $x$ proportion buffer on the OFL ; i.e., $\mathrm{ABC}=(1.0-\mathrm{x}) * \mathrm{OFL}$. The CPT recommended $\mathrm{x}=0.25$.

Please see also the section G on ABC .

## 9. A summary of the results of any rebuilding analysis:

Not applicable.

## 10. Summary of Major Changes

1. Changes (if any) to management of the fishery

- None.

2. Changes to input data

- Commercial fisheries data were updated with values from the most recent observer and fish ticket data for 2020/21: retained catch for the directed fishery and discarded catch estimates for the directed fishery, non-directed crab fisheries, and groundfish fisheries. Thus, the time series of data used in the model are retained catch (1981/82-2020/21), total catch (1990/91-2020/21), and groundfish bycatch (1989/90-2020/21) biomass and size compositions.
- We detected some errors in preparing observer and fish ticket size composition data for 2016-2019 and rectified them in the current assessment.
- Fish ticket retained CPUE were standardized by the generalized linear model (GLM) with the negative binomial link function for the 1985/86-1998/98 period.
- Observer pot sample legal size crab CPUE data were standardized by the GLM with the negative binomial link function with variable selection by CAIC (modified AIC) followed by R square criterion, separately for 1995/96-2004/05 and 2005/06-2020/21 periods. A Year and Area interaction factor was considered in one model (21.1c) to estimate a set of CPUE indices. The habitat areas were determined from observer historical pot locations as fishing footprints (Appendix B).

3. Changes to assessment methodology

None.
4. Changes to assessment results

As expected, the addition of the 2020/21 data changed the OFL and ABC estimates, but changes in parameter or abundance estimates were not dramatic.

## B. Responses to SSC and CPT Comments

## 1. Response to January 2021 CPT comments

Comment\#1: The current GMACS model has some unexpected behavior (e.g., an inability to fit the catch data and unrealistically good fits to the CPUE data) so is not viable for adoption. However, progress on a GMACS-based assessment should be included as an appendix to the assessment report.

Response: We continue to work with the GMACS group to tailor the base model to implement EAG 19.1 model. We try to mimic the status quo model (relabeled as 19.1A) results with GMACS output. Once this is satisfactorily matched, we will be ready to apply GMACS on EAG and WAG assessment.

Comparison of CPUE trends by EAG19.1A with GMACS run\#10AUpdate (most parameters were fixed to modified EAG19.1A estimated parameter values and GMACS employing the same formula
as that of the status quo model for CPUE calculation) is depicted in the following figure (Figure CPT1).

The GMACS ctl, dat, and prj files for EAG 19.1 are included in Appendix $F$.


Figure CPT1. Comparison of input CPUE indices (open circles with $+/-2$ SE for model 19.1A) with predicted CPUE indices (colored solid lines) under original model 19.1A, modified 19.1A to satisfy GMACS coding procedure, and GMACS Run\#10AUpdate (most parameters were fixed to modified 19.1A model output) for EAG golden king crab data, 1985/86-2020/21. Model estimated additional standard error was added to each input standard error.

Comment\# 2: During the January 2021 CPT meeting, authors presented the following model scenarios that include cooperative survey data: 19.1 d : model 21.1a, but with the EAG cooperative survey CPUE indices for 2015-2019; 19.1e, as for 19.1d, but using the fish ticket CPUE data for the entire period. Model 19.1d had cooperative survey and CPUE data for same years (which could lead to double use of some data) and model 19.1e made no use of the observer CPUE data, even though these data are preferable to the fish ticket CPUE data. The CPT recommends that the results of an exploratory model in which model 21.1a is modified to use the EAG cooperative survey data, and the observer CPUE data from 2015 are ignored, be provided in an appendix to the assessment report. This model could be considered for use in the 2022 assessment.

Response: We provided results of an exploratory model (21.1d) in Appendix $C$ in which model 21.1a was modified to use the EAG cooperative survey data for 2015-19, and the observer CPUE data from 2015-2019 were ignored.

Comment\#3: The CPT was unclear why a predicted CPUE from the CPUE standardization with Year:Area interaction could be negative. Also, the formulae used to compute the variances for
years*areas with no data should be provided and a bias correction factor should be applied to Equation A. 15.
$\widehat{B_{l, j}}=e^{A_{i}+C_{j}}$
Response: We provided the clarification and formula to compute the standard deviations (hence variances) for years*areas with no data in Appendix B.

Comment\#4: The analysts should consider a range of alternative standardization models for the EAG cooperative survey data (including that suggested by the SSC) and provide a rationale (including model fit, whether the analysis converged, etc.) for the selected model.

Response: We provided a few alternative standardization models for EAG cooperative survey data and provided the rationale for the selected model in Appendix C. The limitations on time series of data and factor levels (for example, only a maximum of three vessels operate in EAG) prevent us exploring variants of the simple random slope model.

Comment\# 5: The assessment should more clearly provide the rationale for separate EAG and WAG assessments, including consideration of differences in trends in CPUE and agecompositions between the EAG and the WAG.

Response: We provided a few reasons in the January 2021 CPT document. We added more reasons to justify separate stock assessments in the two regions in our response to February 2021 SSC comment\#4.

Comment\# 6: Model results should be presented only for the best fit (lowest objective function). If a best fit model exhibits unusual features (e.g., outlying $F$ estimates or parameters on bounds), this can often be rectified by implementing bounds or smoothing penalties on some of the parameters.

Response: We detected some errors in input observer and fish ticket size measurement data for 2016-2019 and rectified them. This surprisingly eliminated multiple minima issues found earlier when jittering WAG parameter estimates.

Comment\# 7: The runs used to create the retrospective plots should be checked as one run (1994 for WAG) appears to be a case where the minimizer has failed to converge.

Response: The minimizer did not fail in the retrospective run. The 1994 MMB jump was the result of an increase in total catch removal in the model (Figure CPT2):


Figure CPT2. Estimated total catch (in 10s of crab) to the model size range (101 to $185+\mathrm{mm}$ CL ) and effort (pot lifts) by year in WAG.

Comment \# 8: Additional suggestions related to presentation:
a. Add the historical TACs to the plot of catches and CPUEs.

Response: done (Figure 6 in the main text).
b. When plotting MMB, the $x$-axis value should be the second of the two years (e.g., the MMB for 2019/20 should be plotted against 2020, not 2019).

Response: done. Please see an example plot in Figure CPT3. We used the same procedure in constructing other plots.


Figure CPT3. Trends in golden king crab mature male biomass for models 19.1, 21.1a, 21.1b, and 21.1c fits to EAG (left) and WAG (right) data, 2007-2021. Year 2021 refers to 2020/21 fishing season.
c. Provide a figure that shows the size composition data for the fleet with time-varying selectivity in a vertical manner (e.g., Figure 12 in the snow crab assessment) to see changes in the data over time. This can be done using the R package ggridges.

Response: done (Figures 9 to 11 for EAG and 27 to 29 for WAG in the main text).
d. Additionally, a plot that shows the fits to the aggregate size-composition data would be useful.

Response: done (Figure CPT4). Cumulative size compositions of retained catch fitted well for all models but total catch fits slightly differ.


Figure CPT4. Cumulative size compositions vs carapace length for retained catch (left panels) and total catch (right panels) for EAG (top) and WAG (bottom). Observed proportions are marked by solid circles and predicted proportions by different models' (19.1, 21.1a, 21.1b, and 21.1c) are shown by different colored lines.
e. Include plots of the estimates of the smoothing functions (soak for the observer CPUE; depth and soak for the cooperative survey data) when conducting the CPUE and survey standardizations to check that these behave sensibly.

Response: We have addressed the soak time and depth effects on CPUE in the next section (f). We have not addressed the smoothing functions output because of time limitation. We will address this in the next run.
f. Also plot the data on CPUE vs. depth and CPUE vs. soak time to allow the underlying patterns to be identified.

Response: done. We plotted annual observer nominal legal male CPUE against annual mean soak time and annual mean depth (Figure CPT5). Soak time appears to positively influence CPUE (adjusted $R$ square for EAG and WAG: 0.81); but not depth (adjusted $R$ square for EAG: -0.02, WAG: 0.25). Please note that their effects were already modelled in CPUE standardization.


Figure CPT5. Annual observer nominal legal male CPUE vs annual mean soak time and annual mean depth in EAG (top) and WAG (bottom).

## 2. Response to one of June 2020 SSC comments:

Comment 4: The SSC recommends that, if this approach [i.e., 0.7 Sigma approach] were used, the CV would be a better choice. This is because the standard deviations of high recruitments with the same CV might be higher than a cutoff and would result in reference points that are biased lower. The SSC also recommends exploring choosing a reasonable lag from the current year that would include most crabs that have recruited into the fishery. For example, if most crabs are observed by age 6, the 2020 assessment would use recruitments from 1987-2014, and the 2021 assessment would use 1987-2015, and so on.

Response: We did not adequately address the above comment earlier. We try to address the points below.

First point: Considering CV instead of standard deviation of rec_dev provided mixed results for choosing an appropriate year range for different levels of CV (e.g., CV 100\%: 1985-2011; CV 125\%: 1984-2013; and CV 150\%: 1984-2016). Furthermore, years in the selected ranges were non-contiguous. We provide the case for $150 \%$ CV cutoff level below (Figure SSC1). Although CV standardizes different magnitudes of variability for a comparison, they appeared to be not useful here.

Second point: The fixed period 1987-2012 for mean number of recruit calculation used in the base model 19.1 considers an 8-year time lag from the terminal year, which appears to be the recruitment age for golden king crab. However, it is necessary to locate cut off points on both
ends of the time range to choose an appropriate fixed period for mean recruit calculation. The method used in our selection process takes care of this.


Figure SSC1. Coefficient of variation (CV) of recruit_dev vs. year for EAG (top) and WAG (bottom) for model 21.1a. The 1981-2020 marked years have input information for model fitting.

## 3. Response to February 2021 SSC comments

Comment\#1: With respect to Model 21.1b (Same structure and data inputs as M21.1a, but with three-time blocks for selectivity (1960-2004, 2005-2015, and 2016+), the SSC agrees that justification will need to be provided as to why allowing time-varying selectivity based on these time blocks is appropriate, relative to other time-varying parameterizations.

Response: The first block, 1960-2004, pertains to the pre-rationalization period, which was the feature of all models. In the new model scenario (21.1b), the post-rationalization period was divided into two, 2005-2015 and 2016+, to create two selectivity blocks based on observed different fishing patterns (total size distribution shifted to small size groups, please see reduced proportions in large size groups in Figure SSC2) between the two post-rationalization periods in EAG. The visualization of shift is clear in Figures 10a, b, and c in the main text. This pattern was not seen in WAG though. A three-selectivity-block design was created specifically to reduce the retrospective pattern in MMB in EAG (please see the reduction in Mohn rho value for model $21.1 b$ in EAG in Figure 23). Although several other factors separately or in combination may contribute to retrospective patterns (e.g., change in growth, change in catchability, etc.), change in selectivity was the most easily explained by data available.


Figure SSC2. Observer total size composition for 2008-2018 in the EAG.
Comment\#2: The SSC supports continued efforts to recreate existing operational model structures and explore novel structures, within the standardized GMACS platform.

Response: We continue effort to implement golden king crab stock assessment models in GMACS. Appendix F provides preliminary comparison of our model results with GMACS results for EAG19.1. Please see our response to CPT comment\#1 as well.

Comment\#3: An exploratory model in which M21.1a ignores 2015- observer CPUE data but incorporates estimates from the cooperative survey for EAG. The SSC encourages exploration of whether overlapping these indices may help in estimating catchability.

Response: We explored this issue but did not observe much difference in the post-rationalization catchability estimates.

Comment\#4: The SSC would also like to reiterate its support for several previous suggestions:
a. Exploration of a single-area model, or possibly a two-area model with larval connectivity, for the AIGKC crab stock.

Response: We expanded the list of justifications that were presented at the January 2021 CPT and February 2021 SSC meetings, for assessing EAG and WAG separately below:

We modelled EAG and WAG stocks separately for several reasons:
(a) Fishery catch data (e.g., CPUE magnitude and CPUE temporal trends) suggest that the productivity is different between the two areas.
(b) WAG has wider area of stock distribution compared to limited area distribution in EAG.
(c) The fishing areas are spatially separated with an area gap between EAG and WAG (Figure 8 in the main text). Regions of low fishery catch suggest that availability of suitable habitat may vary longitudinally.
(d) Tagging studies have shown little mixing between the two areas (Watson and Gish 2002).
(e) Currents are known to be strong around the Aleutian Islands, thus larval mixing between the two regions may occur. Yet needed data to confirm larval drift trajectories or horizontal displacement are lacking. Unlike other king crabs, golden king crab females carry large, yolk-rich, eggs, which hatch into lecithotrophic (non-feeding) larvae that do not require a pelagic distribution for encountering food items. Depth at larval release, the lecithotrophic nature of larvae, and swimming inactivity in lab studies implies benthic distributions, which may limit larval drift between areas if horizontal current velocities are reduced at depth.
(f) Integrating contrasting data in one single model may provide parameter estimates in between the two extremes which would not be applicable to either (Richards 1991; Schnute and Hilborn 1993).
(g) Area specific assessment is superior to a holistic approach for this stock because of patchy nature of golden king crab distribution.
h) Alaska Board of Fisheries decided to manage the two areas with separate total allowable catches.
i) Genetic analysis shows no significant differentiation between areas within the Aleutian Island population (Grant and Siddon 2018), thus there is no genetic support for subdividing this population; however, above listed factors support separate stock assessments in the two regions.
b. Evaluation of whether catches of AIGKC caught in the NMFS Aleutian Island trawl survey could be utilized as an additional index of abundance.

Response: We have not looked at the NMFS trawl data for AIGKC catches. We will explore any feasibility of including AIGKC caught in the NMFS Aleutian Island trawl survey as an additional index of abundance in the future.
c. Continued exploration of the Year:Area effect in the CPUE standardization, specifically by fitting two area models and combining the results and comparing to the Year:Area model. Diagnostic plots of the data and model predictions of time trends by area (holding all other predictors at their median or mean value) might shed light on the nature of the interaction and aid interpretation.

Response: We addressed the Year:Area effect in the CPUE standardization including some of SSC's suggestions in Appendix B. We have not made SSC suggested diagnostic plots in this report because of time limitation. We will address this in the next run.

## C. Introduction

## 1. Scientific name:

Golden king crab, Lithodes aequispinus J.E. Benedict, 1895.

## 2. Distribution:

General distribution of golden king crab is summarized by NMFS (2004). Golden king crab, also called brown king crab, occur from the Sea of Japan to the northern Bering Sea (ca. $61^{\circ} \mathrm{N}$ latitude), around the Aleutian Islands, generally in high-relief habitat such as inter-island passes, on various sea mounts, and as far south as northern British Columbia (Alice Arm) (Jewett et al. 1985). They are typically found on the continental slope at depths of $300-1,000 \mathrm{~m}$ on extremely rough bottom. They are frequently found on coral bottom.

The Aleutian Islands king crab stock boundary is defined by the boundaries of the Aleutian Islands king crab Registration Area O (Figure 1). In this chapter, "Aleutian Islands Area" means the area described by the current definition of Aleutian Islands king crab Registration Area O. Nichols et al. (2021) define the boundaries of Aleutian Islands king crab Registration Area O:

The Aleutian Islands king crab Registration Area $O$ eastern boundary is the longitude of Scotch Cap Light ( $164^{\circ} 44.72^{\prime} W$ long); the northern boundary is a line from Cape Sarichef ( $54^{\circ} 36^{\prime} \mathrm{N}$ lat) to $171^{\circ} \mathrm{W}$ long, north to $55^{\circ} 30^{\prime} \mathrm{N}$ lat; and the western boundary the United States-Russia Maritime Boundary Line of 1990.

During 1984/85-1995/96, the Aleutian Islands king crab populations had been managed using the Adak and Dutch Harbor Registration Areas, which were divided at $171^{\circ} \mathrm{W}$ longitude (Figure 2), but from the 1996/97 season to present the fishery has been managed using a division at $174^{\circ} \mathrm{W}$ longitude (Figure 1). In March 1996, the Alaska Board of Fisheries (BOF) replaced the Adak and Dutch Harbor areas with the newly created Aleutian Islands Registration Area O and directed ADF\&G to manage the golden king crab fishery in the areas east and west of $174^{\circ} \mathrm{W}$ longitude as two distinct stocks. That re-designation of management areas was intended to reflect golden king crab stock distribution, congruent with the longitudinal pattern in fishery production prior to 1996/97 (Figure 3). The longitudinal pattern in fishery production relative to $174^{\circ} \mathrm{W}$ longitude since 1996/97 is like that observed prior to the change in management area definition, although there have been some changes in the longitudinal pattern in fishery production within the areas east and west of $174^{\circ} \mathrm{W}$ longitude (Figure 4).

Commercial fishing for golden king crab in the Aleutian Islands Area typically occurs at depths of 100-275 fathoms (183-503 m). Pots sampled by at-sea fishery observers in 2013/14 were fished at an average depth of 176 fathoms ( $322 \mathrm{~m} ; \mathrm{N}=499$ ) in the area east of $174^{\circ} \mathrm{W}$ longitude and 158 fathoms ( $289 \mathrm{~m} ; \mathrm{N}=1,223$ ) for the area west of $174^{\circ} \mathrm{W}$ longitude (Gaeuman 2014).

## 3. Evidence of stock structure:

Given the expansiveness of the Aleutian Islands Area and the existence of deep ( $>1,000$ m ) canyons between some islands, at least some weak structuring of the stock within the area would be expected. Data for making inferences on stock structure of golden king crab within the Aleutian Islands are largely limited to the geographic distribution of commercial fishery catch and effort. Catch data by statistical area from fish tickets and catch data by location from pots sampled by observers suggest that habitat for legal-sized males may be continuous throughout the waters adjacent to the islands in the Aleutian chain. However, regions of low fishery catch suggest that availability of suitable habitat, in which golden king crab are present at only low densities, may vary longitudinally. Catch has been low in the fishery in the area between $174^{\circ} \mathrm{W}$ longitude and $176^{\circ} \mathrm{W}$ longitude (the Adak Island area, Figures 3 and 4) in comparison to adjacent areas, a pattern that is consistent with low CPUE for golden king crab between $174^{\circ} \mathrm{W}$ longitude and $176^{\circ} \mathrm{W}$ longitude (Figure 5) during the 2002, 2004, 2006, 2010, and 2012 NMFS Aleutian Islands bottom trawl surveys (von Szalay et al. 2011). In addition to longitudinal variation in density, there is also a gap in fishery catch and effort between the Petrel Bank-Petrel Spur area and the Bowers Bank area; both of those areas, which are separated by Bowers Canyon, have reported effort and catch. Recoveries during commercial fisheries of golden king crab tagged during ADF\&G surveys (Blau and Pengilly 1994; Blau et al. 1998; Watson and Gish 2002; Watson 2004, 2007) provided no evidence of substantial movements by crab in the size classes that were tagged (males and females $\geq 90-\mathrm{mm}$ carapace length [CL]). Maximum straight-line distance between release and recovery location of 90 golden king crab released prior to the 1991/92 fishery and recovered through the 1992/93 fishery was 61.2 km (Blau and Pengilly 1994). Of the 4,567 recoveries reported through April 12, 2016 for the male and female golden king crab tagged and released between $170.5^{\circ} \mathrm{W}$ longitude and $171.5^{\circ} \mathrm{W}$ longitude during the 1991, 1997, 2000, 2003, and 2006 ADF\&G Aleutian Island golden king pot surveys, none of the 3,807 with recovery locations specified by latitude and longitude were recovered west of $173^{\circ} \mathrm{W}$ longitude and only fifteen were recovered west of $172^{\circ} \mathrm{W}$ longitude (V. Vanek, ADF\&G, Kodiak, pers. comm.). Similarly, of 139 recoveries in which only the statistical area of recovery was reported, none were recovered in statistical areas west of $173^{\circ} \mathrm{W}$ longitude and only one was in a statistical area west of $172^{\circ} \mathrm{W}$ longitude.

## 4. Life history characteristics relevant to management:

There is a paucity of information on golden king crab life history characteristics due in part to the deep depth distribution ( $\sim 200-1000 \mathrm{~m}$ ) and the asynchronous nature of life history events (Otto and Cummiskey 1985; Somerton and Otto 1986). The reproductive cycle is thought to last approximately 24 months and at any time of year ovigerous females can be found carrying egg clutches in highly disparate developmental states (Otto and Cummiskey 1985). Females carry large, yolk-rich, eggs, which hatch into lecithotrophic (i.e., the larvae can develop successfully to juvenile crab without eating; Shirley and Zhou 1997) larvae that are negatively phototactic (Adams and Paul 1999). Molting and mating are also
asynchronous and protracted (Otto and Cummiskey 1985; Shirley and Zhou 1997) with some indications of seasonality (Hiramoto 1985). Molt increment for large males (adults) in Southeast Alaska is 16.3 mm CL per molt (Koeneman and Buchanan 1985) and was estimated at 14.4 mm CL for legal males in the EAG (Watson et al. 2002). Annual molting probability of males decreases with increasing size, which results in a protracted inter-molt period and creates difficulty in determining annual molt probability (Watson et al. 2002). Male size-at-maturity varies among stocks (Webb 2014) and declines with increasing latitude from about 130 mm CL in the Aleutian Islands to 90 mm CL in Saint Matthew Island section (Somerton and Otto 1986). Along with a lack of annual survey data, limited stock-specific life history stock information prevents development of the standard lengthbased assessment model.

## 5. Summary of management history:

A complete summary of the management history through 2015/16 is provided in Leon et al. (2017). The first commercial landing of golden king crab in the Aleutian Islands was in 1975/76 but directed fishing did not occur until 1981/82.

The Aleutian Islands golden king crab fishery was restructured beginning in 1996/97 to replace the Adak and Dutch Harbor areas with the newly created Aleutian Islands Registration Area O and golden king crab in the areas east and west of $174^{\circ} \mathrm{W}$ longitude were managed separately as two stocks (ADF\&G 2002). Hereafter, the east of $174^{\circ} \mathrm{W}$ longitude stock segment is referred to as EAG and the west of $174^{\circ} \mathrm{W}$ longitude stock segment is referred to as WAG. Table 1 provides the historical summary of number of vessels, GHL/TAC, harvest, effort, CPUE and average weight in the Aleutian Islands golden king crab fishery.

The fisheries in 1996/97-1997/98 were managed with GHLs of $1,452 \mathrm{t}(3,200,000 \mathrm{lb})$ in EAG and $1,225 \mathrm{t}(2,700,000 \mathrm{lb})$ in WAG (Table 1). During 1998/99-2004/05 the fisheries were managed with GHLs of $1,361 \mathrm{t}(3,000,000 \mathrm{lb})$ for EAG and $1,225 \mathrm{t}(2,700,000 \mathrm{lb})$ for WAG. During 2005/06-2007/08 the fisheries were managed with a total allowable catch (TAC) of $1,361 \mathrm{t}(3,000,000 \mathrm{lb})$ for EAG and a TAC of $1,225 \mathrm{t}(2,700,000 \mathrm{lb})$ for WAG. By state regulation ( 5 AAC 34.612), TAC for the Aleutian Islands golden king crab fishery during 2008/09-2011/12 was $1,429 \mathrm{t}(3,150,000 \mathrm{lb})$ for EAG and $1,286 \mathrm{t}(2,835,000 \mathrm{lb})$ for WAG. In March 2012, the BOF changed 5 AAC 34.612 so that the TAC beginning in 2012/13 would be $1,501 \mathrm{t}(3,310,000 \mathrm{lb})$ for the EAG and $1,352 \mathrm{t}(2,980,000 \mathrm{lb})$ for WAG. Additionally, the BOF added a provision to 5 AAC 34.612 that allows ADF\&G to lower the TAC below the specified level if conservation concerns arise. The TAC for 2016/17 (and 2017/18) was reduced by $25 \%$ for WAG to $1,014 \mathrm{t}(2,235,000 \mathrm{lb})$ while keeping the TAC for EAG at the same level as the previous season.

During 1996/97-2020/21 the annual retained catch during commercial fishing (including cost-recovery fishing that occurred during 2013/14-2020/21) has averaged $2 \%$ below the annual GHL/TACs. During 1996/97-2020/21, the retained catch has been as much as $13 \%$ below (1998/99) and as much as $6 \%$ above (2000/01) the GHL/TAC.

A summary of other relevant State of Alaska fishery regulations and management actions pertaining to the Aleutian Islands golden king crab fishery is provided below:

Beginning in 2005/06 the Aleutian Islands golden king crab fishery has been prosecuted under the Crab Rationalization Program. Accompanying the adoption of crab rationalization program was implementation of a community development quota (CDQ) fishery for golden king crab in the eastern Aleutians (i.e., EAG) and the Adak Community Allocation (ACA) fishery for golden king crab in the western Aleutians (i.e., WAG; Hartill 2012; Nichols et al. 2021). The CDQ fishery in the eastern Aleutians is allocated $10 \%$ of the golden king crab TAC for the area east of $174^{\circ} \mathrm{W}$ longitude and the ACA fishery in the western Aleutians is allocated $10 \%$ of the golden king crab TAC for the area west of $174^{\circ} \mathrm{W}$ longitude. The CDQ fishery and the ACA fishery are managed by ADF\&G and prosecuted concurrently with the individual fishing quota (IFQ) fishery.

Golden king crab may be commercially fished only with king crab pots (defined in state regulation 5 AAC 34.050). Pots used to fish for golden king crab in the Aleutian Islands Area must be longlined and, since 1996, each pot must have at least four escape rings of five and one-half inches minimum inside diameter installed on the vertical plane or at least one-third of one vertical surface of the pot composed of not less than nine-inch stretched mesh webbing to permit escapement of undersized golden king crab ( 5 AAC 34.625 (b)). Prior to the regulation requiring an escape mechanism on pots, some participants in the Aleutian Islands golden king crab fishery voluntarily sewed escape rings (typically 139 mm [5.5 inches]) into their gear or, more rarely, included panels with escape mesh (Beers 1992). Regarding the gear used since the establishment of 5 AAC 34.625 (b) in 1996, Linda Kozak, a representative of the industry, reported in a 19 September 2008 email to the Crab Plan Team, "... the golden king crab fleet has modified their gear to allow for small crab sorting," and provided a written statement from Lance Nylander, of Dungeness Gear Works in Seattle, who "believes he makes all the gear for the golden king crab harvesting fleet," saying that, "Since 1999, DGW has installed 9 [-inch] escape web on the door of over $95 \%$ of Golden Crab pot orders we manufactured." A study to estimate the contactselection curve for male golden king crab was conducted aboard one vessel commercial fishing for golden king crab during the 2012/13 season and found gear and fishing practices used by that vessel were highly effective in reducing bycatch of sublegal-sized males and females (Vanek et al. 2013). In March 2011 (effective for 2011/12), the BOF amended 5 $A A C 34.625$ (b) to relax the "biotwine" specification for pots used in the Aleutian Islands golden king crab fishery relative to the requirement in 5 AAC 39.145 that "(1) a sidewall ...of all shellfish and bottom fish pots must contain an opening equal to or exceeding 18 inches in length... The opening must be laced, sewn, or secured together by a single length of untreated, 100 percent cotton twine, no larger than 30 thread." Regulation 5 AAC 34.625 (b)(1) allows the opening described in $5 A A C 39.145$ (1) to be "laced, sewn, or secured together by a single length of untreated, 100 percent cotton twine, no larger than 60 [rather than 30] thread."

Regulation (5 AAC 34.610 (b)) sets the commercial fishing season for golden king crab in the Aleutian Islands Area as 1 August through 30 April. That regulatory fishing season
became effective in 2015/16 (the commercial fishing season was set in regulation as 15 August through 15 May during 2005/06-2014/15).

Current regulations (5 AAC 39.645 (d)(4)(A)) stipulate that onboard observers are required on catcher vessels during the time that at least $50 \%$ of the retained catch is captured in each of the three trimesters of the 9 -month fishing season. Onboard observers are required for $100 \%$ of fishing activity on catcher-processor vessels during the crab fishing season.

In addition, the commercial golden king crab fishery in the Aleutian Islands Area may only retain males at least 6.0 -inches ( 152.4 mm ) carapace width (CW), including spines ( 5 AAC 34.620 (b)), which is at least one annual molt increment larger than the $50 \%$ maturity length of 120.8 mm CL for males estimated by Otto and Cummiskey (1985). A carapace length (CL) $\geq 136 \mathrm{~mm}$ is used to identify legal-size males when CW measurements are not available (Table 3-5 in NPFMC 2007b). Note that the size limit for golden king crab has been 6-inches ( 152.4 mm ) CW for the entire Aleutian Islands Area since the 1985/86 season. Prior to the $1985 / 86$ season, the legal-size limit was 6.5 -inches ( 165.1 mm ) CW for at least one of the now-defunct Adak or Dutch Harbor Registration Areas.

We re-evaluated the male maturity size using 1991 pot survey measurements of carapace length and chela height in EAG and 1984 NMFS measurements in WAG (Siddeek et al. 2018). Bootstrap analysis of chela height and carapace length data provided the median $50 \%$ male maturity length estimates of 107.02 mm CL in EAG and 107.85 mm CL in WAG. We used a knife-edge $50 \%$ maturity length of 111.0 mm CL, which is the lower limit of the next upper size bin, for mature male biomass (MMB) estimation. We analyzed the recently collected (2018 to 2020) chela height and carapace length data and proposed a higher $50 \%$ maturity length of 116.0 mm CL and a maturity curve for MMB estimation (Appendix D).

Daily catch and catch-per-unit effort (CPUE) are determined in-season to monitor fishery performance and progress towards the respective TACs. Figures 6 to 8 provide the 1985/86-2020/21 time series of catches, CPUE, and the geographic distribution of catch during the 2020/21 fishing season. Increases in CPUE were observed during the late 1990s through the early 2000s, and with the implementation of crab rationalization in 2005. This is likely due to changes in gear configurations in the late 1990s (crab harvesters, personal communication, 1 July 2008) and, after rationalization due to increased soak time (Siddeek et al. 2015), and decreased competition owing to the reduced number of vessels fishing. Decreased competition could allow crab vessels to target only the most productive fishing areas. Trends in fishery nominal CPUE within the areas EAG and WAG generally paralleled each other during 1985/86-2010/11 but diverged thereafter (EAG CPUE exceeded one and half times of that in WAG). A moderate decreasing trend in CPUE was observed since 2014 in EAG. A sharp drop in CPUE was detected in 2020/21 for WAG (Figures 6 and 7).
6. Brief description of the annual ADF\&G harvest strategy:

In March 2019, the BOF adopted a revised harvest strategy (Daly et al. 2019). The annual TAC is set by state regulation, 5 AAC 34.612 (Harvest Levels for Golden King Crab in Registration Area O), per:
(a) In that portion of the Registration Area O east of $174^{\circ} \mathrm{W}$. long., the total allowable catch level shall be established as follows:
(1) if MMA $_{E}$ is less than 25 percent of MMA $_{\mathrm{E},(1985-2017) \text {, the fishery will not open; }}$
(2) if MMA ${ }_{E}$ is at least 25 percent but not greater than 100 percent of MMAE.(19852017), the number of legal male golden king crab available for harvest will be computed as $(0.15) x\left(\mathrm{MMA}_{\mathrm{E}} / \mathrm{MMA}_{\mathrm{E} .(1985-2017)}\right) \mathrm{x}\left(\mathrm{MMA}_{\mathrm{E}}\right)$ or 25 percent of LMA $_{E}$, whichever is less; and
(3) if MMA ${ }_{\mathrm{E}}$ is greater than 100 percent of $\mathrm{MMA}_{\mathrm{E},(1985-2017) \text {, the number of legal }}$ male golden king crab available for harvest will be computed as $(0.15) \times\left(\right.$ MMA $\left._{E}\right)$ or 25 percent of LMA $E_{2}$ whichever is less.
(b) In that portion of the Registration Area O west of $174^{\circ} \mathrm{W}$. long., the total allowable catch level shall be established as follows:
(1) if MMAW is less than 25 percent of MMAw,(1985-2017), the fishery will not open
(2) if MMA ${ }_{w}$ is at least 25 percent but not greater than 100 percent of MMAW,(19852017), the number of legal male golden king crab available for harvest will be computed as $(0.20) x\left(\right.$ MMA $_{W} /$ MMA $_{W},(1985-2017)$ x(MMAW) or 25 percent of LMAw, whichever is less; and $^{2}$
(3) if MMAw is greater than 100 percent of MMAw.(1985-2017), the number of legal male golden king crab available for harvest will be computed as ( 0.20 ) x(MMAw) or 25 percent of LMAw, whichever is less.
(c) In implementing this harvest strategy, the department shall consider the reliability of estimates of golden king crab, the manageability of the fishery, and other factors the department determines necessary to be consistent with sustained yield principles and to use the best scientific information available and consider all sources of uncertainty as necessary to avoid overfishing.
(d) In this section,
(1) $\mathrm{MMA}_{\mathrm{E}}$ means the abundance of male golden king crab in the portion of the Aleutian Islands Management Area O east of $174^{\circ} \mathrm{W}$. long that are greater than or equal to 111 millimeters in carapace length estimated by the stock assessment model for the time prior to the start of the fishery;
(2) $\mathrm{MMA}_{\mathrm{E},(1985-2017)}$ means the mean value of the abundance of male golden king crab in the portion of the Aleutian Islands Management Area O east of $174^{\circ} \mathrm{W}$. long that are greater than or equal to 111 millimeters in carapace length estimated by the stock assessment model for the time prior to the start of the fishery for the period 1985-2017;
(3) LMAE means the abundance of male golden king crab in the portion of the $^{\text {the }}$ Aleutian Islands Management Area O east of $174^{\circ} \mathrm{W}$. long that are greater than or equal to 136 millimeters in carapace length estimated by the stock assessment model for the time prior to the start of the fishery;
(4) $\mathrm{MMA}_{\mathrm{W}}$ means the abundance of male golden king crab in the portion of the Aleutian Islands Management Area O west of $174^{\circ} \mathrm{W}$. long that are greater than
or equal to 111 millimeters in carapace length estimated by the stock assessment model for the time prior to the start of the fishery;
(5) MMAw,(1985-2017) means the mean value of the abundance of male golden king crab in the portion of the Aleutian Islands Management Area O west of $174^{\circ}$ W. long that are greater than or equal to 111 millimeters in carapace length estimated by the stock assessment model for the time prior to the start of the fishery for the period 1985-2017; and
(6) LMAw means the abundance of male golden king crab in the portion of the Aleutian Islands Management Area O west of $174^{\circ} \mathrm{W}$. long that are greater than or equal to 136 millimeters in carapace length estimated by the stock assessment model for the time prior to the start of the fishery.

In addition to the retained catch that is limited by the TAC established by ADF\&G under 5 AAC 34.612, ADF\&G has authority to annually receive receipts up to $\$ 500,000$ through cost-recovery fishing on Aleutian Islands golden king crab. The retained catch from that cost-recovery fishing is not counted against attainment of the annually established TAC.
7. Summary of the history of the basis and estimates of MMBMSY or proxy MMBMSY:

We estimated the proxy $M M B_{M S Y}$ as $M M B_{35 \%}$ using the Tier 3 estimation procedure, which is explained in a subsequent section.

## D. Data

1. Summary of new information:
(a) Commercial fishery retained catch by size, estimated total catch by size, groundfish male discard catch by size, observer CPUE index, and commercial fishery CPUE index were updated to include 2020/21 information. Available data by year are shown below.


| Tag release |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  | $\nabla$ |  |  |  |  |  |  | $\Delta$ |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Tag Recovery |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |

## 2. Data presented as time series:

## a. Total Catch:

Fish ticket data on retained catch weight, catch numbers, effort (pot lifts), CPUE, and average weight of retained catch for 1981/82-2020/21 (Table 1). Estimated total catch weight for 1990/91-2020/21 (Table 2a).

## b. Bycatch and discards:

Retained catch, bycatch mortality (male and female) separated by the crab fishery and groundfish fishery, and total fishery mortality for 1981/82-2020/21 (Table 2). Crab fishery discards are available after observer sampling was established in 1988/89. Observer data for the 1988/89-1989/90 seasons are not considered reliable. Table 2 provides crab fishery discards and groundfish fishery bycatch for 1991/92-2020/21 seasons.

## c. Catch-per-unit-effort:

- Pot fishery and observer nominal retained and total CPUE, pot fishery effort, observer sample size, and estimated observer CPUE index delineated by EAG and WAG for 1985/86-2020/21 (Table 3).
- Estimated commercial fishery CPUE index with coefficient of variation (Table 4 for EAG and Table 14 for WAG). The estimation methods, and CPUE fits are described in Appendix B.
d. Catch-at-length:

Information on length compositions is provided (Figures 9a, b, c to 11a, b, c for EAG; and $27 \mathrm{a}, \mathrm{b}, \mathrm{c}$ to $29 \mathrm{a}, \mathrm{b}$, c for WAG for models 21.1a, 21.1b, and 21.1c, respectively.
e. Survey biomass estimates:

Estimates are not available for the area because no systematic surveys, covering the entire fishing area, have occurred.
f. Survey catch-at-length:

Not available.
g. Other time series data: None.
3. Data which may be aggregated over time:

- Molt and size transition matrix: Tag release - recapture -time at liberty records from 1991, 1997, 2000, 2003, and 2006 male tag crab releases were aggregated by year at liberty to determine the molt increment and size transition matrix by the integrated model.
- Weight-at-length: Male length-weight relationship: $\mathrm{W}=\mathrm{aL}^{\mathrm{b}}$ where $a=$ $1.445 * 10^{-4}, b=3.28113[\sigma=0.00737$ (bias correction for $a$ was not required because of the very small value of $\sigma$ ); updated estimates from WAG data].
- Natural mortality: A previous model estimated fixed natural mortality value of $0.21 \mathrm{yr}^{-1}$, was used in the assessment.

4. Information on any data sources that were available, but were excluded from the assessment:

Data from triennial ADF\&G pot surveys for Aleutian Islands golden king crab in a limited area in EAG (between $170^{\circ} 21^{\prime}$ and $171^{\circ} 33^{\prime} \mathrm{W}$ longitude) that were performed during 1997 (Blau et al. 1998), 2000 (Watson and Gish 2002), 2003 (Watson 2004), and 2006 (Watson 2007) are available, but were not used in this assessment. However, the tag release and recapture data from these surveys were used.

Data from the cooperative pot surveys conducted during 2015 to 2019 are available but is limited in the time series. The EAG survey covers the full time series but WAG survey started only in 2018. We incorporated the EAG data in a model scenario (21.1d) as a test run in this assessment (Appendix C).

## E. Analytic Approach

1. History of modeling approaches for this stock:

A size structured assessment model based on only fisheries data was under development for several years for the EAG and WAG golden king crab stocks and accepted in 2016 for OFL and ABC setting for the 2017/18 season. The CPT in January 2017 and SSC in February 2017 recommended using the Tier 3 procedure to set the OFL and ABC. They also suggested using the maturity data to estimate the male mature biomass (MMB). We followed these suggestions in this report to estimate the model based OFL and ABC.

## 2. Model Description:

## a. Description of overall modeling approach:

The underlying population dynamics model is male-only and length-based (Appendix A). This model combines commercial retained catch, total catch, groundfish fishery discarded catch, standardized observer legal size catch-per-unit-effort (CPUE) indices, fishery retained catch size composition, total catch size composition, and tag recaptures by release-recapture length to estimate stock assessment parameters. The tagging data were used to calculate the size transition matrix. To estimate the MMB, we used the
knife-edge $50 \%$ maturity based on the chela height and carapace length data analysis. To include a long time series of CPUE indices for stock abundance contrast, we also considered the 1985/86-1998/99 legal size standardized fishery CPUE indices as a separate likelihood component in all scenarios (Table T1).

There were significant changes in fishing practice associated with changes in management regulations (e.g., constant TAC since 1996/97 and crab rationalization since 2005/06), pot configuration (escape web on the pot door increased to 9 -inch since 1999), and improved observer recording in Aleutian Islands golden king crab fisheries since 1998. These changes prompted us to consider two sets of catchability and total selectivity parameters with only one set of retention parameters for the periods 1985/86-2004/05 and 2005/06-2020/21. We also considered a model (21.1b) with three total selectivity curves to reduce the retrospective pattern of EAG MMB.

We fitted the observer and commercial fishery CPUE indices with standard errors (estimated by GLM) and an additional assessment model estimated constant variance. The assessment model predicted total and retained CPUEs. However, we compared only the predicted retained CPUE with the observer legal size crab CPUE indices in the likelihood function because observer recordings of legal-size crabs are reliable.

The data series ranges used for the WAG are the same as those for EAG.
b. Software:

AD Model Builder (Fournier et al. 2012).
c. -f . Details are given in Appendix A.
g. Critical assumptions and consequences of assumption failures:

Because of the lack of an annual stock survey, we relied heavily on standardized CPUE indices (Appendix B) and catch and size composition information to determine the stock abundance trends in both regions. We assumed that the observer and fish ticket CPUE indices are linearly related to exploitable abundance. We kept $M$ constant at 0.21 $\mathrm{yr}^{-1}$ and knife-edge maturity size at 111 mm CL (Siddeek et al. 2018). We also considered a higher knife-edge maturity size of 116 mm CL and maturity curves for MMB estimation in different model scenarios. We assumed directed pot fishery discard mortality at $0.20 \mathrm{yr}^{-1}$, overall groundfish fishery mortality at $0.65 \mathrm{yr}^{-1}$ (mean of groundfish pot fishery mortality $\left[0.5 \mathrm{yr}^{-1}\right]$ and groundfish trawl fishery mortality [ 0.8 $\left.\mathrm{yr}^{-1}\right]$ ), groundfish fishery selectivity at full selection for all length classes (selectivity $=$ 1.0). Any discard of legal-size males in the directed pot fishery was not considered in this analysis. These fixed values invariably reduced the number of model parameters to be estimated and helped in convergence. We assumed different $q$ 's (scaling parameter for standardized CPUE in the model, Equation A.13) and logistic selectivity patterns (Equation A.9) for different periods for the pot fishery.
h. Changes to any of the above since the previous assessment:

None.
i. Model code has been checked and validated.

The codes have been checked at various times by independent reviewers and the current codes are available from the first author.

## 3. Model Selection and Evaluation

## a. Description of alternative model configurations:

We considered ten models for EAG and WAG (Table T1). We presented OFL and ABC results for all models separately for EAG, WAG, and the entire AI in the executive summary tables. We considered model 19.1 as the base model. It considers:
i) Initial abundance by the equilibrium condition considering the mean number of recruits for 1987-2012: The equilibrium abundance was determined for 1960 (Equations A. 4 and A.5), projected forward with only $M$ and annual recruits until 1980, then retained catches removed during 1981-1984 and projected to obtain the initial abundance in 1985.
ii) Observer CPUE indices for 1995/96-2020/21.
iii) Fishery CPUE indices for 1985/86-1998/99.
iv) Initial (Stage-1) weighting of effective sample sizes: number of vessel-days for retained and total catch size compositions, and number of fishing trips for groundfish discard size composition (the groundfish size composition was not used in model fitting); and (Stage-2) iterative re-weighting of effective sample sizes by the Francis method.
v) Two catchabilities and two sets of logistic total selectivities for the periods 1985/86-2004/05 and 2005/06-2020/21, and a single set of logistic retention curve parameters.
vi) Full selectivity (selectivity $=1.0$ ) for groundfish fishery bycatch.
vii) Knife-edge $50 \%$ maturity size of 111 mm CL.
viii) Stock dynamics $M=0.21 \mathrm{yr}^{-1}$, pot fishery handling mortality $=0.2 \mathrm{yr}^{-1}$, and mean groundfish bycatch handling mortality $=0.65 \mathrm{yr}^{-1}$.
ix) Size transition matrix using tagging data estimated by the normal probability function with the logistic molt probability sub-model. The tag-recaptures were treated as Bernoulli trials (i.e., Stage-1 weighting).
x) The period, 1987-2012, was used to determine the mean number of recruits for $M M B_{35 \%}$ (a proxy for $M M B_{M S Y}$ ) estimation under Tier 3.

The salient features and variations from the base scenario of all other scenarios are listed in Table T1. The list of fixed and estimable parameters is provided in Table A1 and detail weights with coefficient of variations (CVs) assigned to each type of data are listed in Table A2.

Best estimates of parameter values for the base model 19.1 were jittered to confirm model global convergence. The results indicated that global convergence was achieved for most runs (Appendix E).

Table T1. Features of all model scenarios: Initial condition was estimated in year 1960 by the equilibrium condition; two catchability and two sets of logistic total selectivity curves were used for the pre- and post-rationalization periods; a single retention curve was used for the whole period; and a common $M$ of $0.21 \mathrm{yr}^{-1}$ was used. The effective sample sizes for size compositions were estimated in two stages: Stage-1: as the number of vessel days/trips and Stage-2: as the Francis re-iteration method. Changes in base model specifications are highlighted by the shaded text.

| Model | CPUE Data Type and Maturity Option | Period for Mean Number of Recruit Calculation for (a) Initial Equilibrium Abundance and (b) Reference Points Estimations and Remarks |
| :---: | :---: | :---: |
| 19.1 (accepted model in May 2019, implemented with up to 2020/21 data) | Observer data from 1995/96-2020/21 Fish ticket data from 1985/86-1998/99; Observer and fish ticket CPUE standardization by negative binomial model; a knife-edge minimum maturity size of 111 mm CL. | 1987-2012; CPT/SSC suggested base model. |
| 21.1a | 19.1+ | 1987-2017; CPT/SSC suggested model. |
| 21.1b | ```21.1a+ three total selectivity periods (1960-2004; 2005-2015; 2016+).``` | CPT/SSC suggested model. |
| 21.1c | 21.1a+ the observer CPUE data standardized including Year:Area interactions. | CPT/SSC suggested model. |
| 21.1al | $21.1 \mathrm{a}+\mathrm{a}$ knife-edge minimum maturity size of 116 mm CL . | Authors proposed additional model. |
| 21.1a2 | $21.1 \mathrm{a}+$ maturity curve. | Authors proposed additional model. |
| 21.1 b 1 | $21.1 \mathrm{~b}+$ a knife-edge minimum maturity size of 116 mm CL . | Authors proposed additional model. |
| 21.1 b 2 | $21.1 \mathrm{~b}+$ maturity curve. | Authors proposed additional model. |
| 21.1 c 1 | $21.1 \mathrm{c}+$ a knife-edge minimum maturity size of 116 mm CL . | Authors proposed additional model. |
| 21.1c2 | $21.1 \mathrm{c}+$ maturity curve. | Authors proposed additional model. |

## b. Progression of results:

The OFL and ABC estimates are like those estimates made in 2019.
c. Label the approved model from the previous year as model:

We used the notation 19.1 for the base model which came from the last year accepted assessment model, 19.1.
d. Evidence of search for balance between realistic and simpler models:

Unlike annually surveyed stocks, Aleutian Islands golden king crab stock biomass is difficult to track, and several biological parameters are assumed based on knowledge from red king crab (e.g., handling mortality rate of $0.2 \mathrm{yr}-1$ ) due to a lack of species/stock specific information. We fixed several model parameters after initially running the model with free parameters to reduce the number of parameters to be estimated (e.g., groundfish bycatch selectivity parameters were fixed). In CPUE standardization, instead of using the traditional AIC we used the Consistent Akaike Information Criteria (Bozdogan 1987) that considers number of parameters and data points used for fitting models when selecting the final model. The models also considered different configuration of parameters to select parsimonious models. The detailed results of all models are provided in tables and figures.
e. Convergence status and criteria:

ADMB default convergence criteria were used.
f. Table of the sample sizes assumed for the size compositional data:

We estimated the initial input effective sample sizes (i.e., Stage-1) either as number of vessel-days for retained and total catch compositions or number of fishing trips for groundfish size composition (note: we did not use the groundfish size composition in model fitting) for all model scenarios. Then we estimated the Stage- 2 effective sample sizes iteratively from Stage-1 input effective sample sizes using the Francis' (2011, 2017) mean length-based method.

We provide the initial input sample sizes (Stage-1) and Stage-2 effective sample sizes for models 21.1a, 21.1b, and 21.1c in Tables 5 to 7 for EAG and Tables 15 to 17 for WAG.
g. Provide the basis for data weighting, including whether the input effective sample sizes are tuned, and the survey CV adjusted:
Described previously (f).
h. Do parameter estimates make sense and are they credible?

The estimated parameter values are within the bounds and various plots suggest that the parameter values are reasonable for a fixed $M$ value for the golden king crab stocks.
i. Model selection criteria:

We used several diagnostic criteria to select appropriate models for our recommendation: CPUE fits, observed vs. predicted tag recapture numbers by time at
large and release size, retained and total catch, and groundfish bycatch fits. Figures are provided for all model scenarios in the Results section.

## j. Residual analysis:

We illustrated residual fits by bubble plots for retained and total catch size composition predictions in various figures in the Results section.

## k. Model evaluation:

Only one base model with several model variations is presented and the evaluations are presented in the Results section below.

## 4. Results

## 1. List of effective sample sizes and weighting factors:

The Stage-1 and Stage-2 effective sample sizes are listed for various models in Tables 5 to 7 for EAG and Tables 15 to 17 for WAG. The weights, with the corresponding coefficient of variations specifications, for different data sets are provided in Table A2 for various models for both EAG and WAG. These weights (with the corresponding coefficient of variations) adequately fitted the length compositions, and no further changes were examined.

We used weighting factors for catch biomass, recruitment deviation, pot fishery F , and groundfish fishery F. We set the retained catch biomass weight to an arbitrarily large value (500.0) because retained catches are more reliable than any other data sets. We scaled the total catch biomass weight in accordance with the observer annual sample sizes (number of pots) with a maximum of 250.0 . The total catches were derived from observer nominal total CPUE and effort. In some years, observer sample sizes were low (Tables 3). We chose a small groundfish bycatch weight (0.2) based on the September 2015 CPT suggestion for a lower its weight. We used the best fit criteria to choose the lower weight for the groundfish bycatch. Groundfish bycatch of Aleutian Islands golden king crab is very low (Table 2). We set the CPUE weights to 1.0 for all models. We included a constant (model estimated) variance in addition to input CPUE variance for the CPUE fit. We used the Burnham et al. (1987) suggested formula for $\ln ($ CPUE ) (and $\ln (M M B)$ ) variance estimation (Equation A.14). However, the estimated additional variance values were small for both observer and fish ticket CPUE indices for the two regions. Nevertheless, the CPUE index variances estimated from the negative binomial GLM were adequate to fit the model, as confirmed by the fit diagnostics (Fox and Weisberg 2011). Parameter estimates are provided in Tables 8 for EAG and 18 for WAG for models 19.1, 21.1a, 21.1b, and 21.1c. The numbers of estimable parameters are listed in Table A1.

## 2. Include tables showing differences in likelihood:

Tables 13 and 23 list the total and component negative log likelihood values for EAG and WAG, respectively.

## 3. Tables of estimates:

a. The parameter estimates with coefficient of variation for models 19.1, 21.1a, 21.1b, and 21.1c are summarized in Tables 8 and 18 for EAG and WAG, respectively. We have also provided the boundaries for parameter searches in those tables. All parameter estimates were within the bounds.
b. All models considered molt probability parameters in addition to the linear growth increment and normally distributed growth variability parameters to determine the size transition matrix.
c. The mature male and legal male abundance time series for selected models (19.1, 21.1a, 21.1b, and 21.1c) are summarized in Tables 9 to 12 for EAG and Tables 19 to 22 for WAG.
d. The recruitment estimates for those models are summarized in Tables 9 to 12 for EAG and Tables 19 to 22 for WAG.
e. The negative log-likelihood component values and total negative log-likelihood values for models 19.1, 21.1a, 21.1b, and 21.1c are summarized in Table 13 for EAG and Table 23 for WAG. Although loglikelihood values of different models are not comparable because of data weighting (i.e., different magnitude of effective sample sizes), nevertheless, model 21.1c has the minimum total negative log likelihood for EAG and WAG. However, the total negative log likelihood values for the four models were not widely different. We may conclude that the input observer CPUE indices with Year and Area interaction appears to have positively influenced the overall fit.

## 4. Graphs of estimates:

a. Selectivity:

Total selectivity and retention curves of the pre- and post-rationalization periods for selected models are illustrated in Figures 12a and 12b for EAG and Figures 30a and 30 b for WAG. Figures 12 b and 30 b correspond to second part (2016-2020) of the post-rationalization period in the three total selectivity model. Total selectivity for the pre-rationalization period was used in the tagging model. The groundfish bycatch selectivity appeared flat in the preliminary analysis, indicating that all size groups were vulnerable to the gear. This is also shown in the size compositions of groundfish bycatch (Figures 11 and 29 for model 21.1a for EAG and WAG, respectively). Thus, we set the groundfish bycatch selectivity to 1.0 for all lengthclasses in the subsequent analysis.

## b. Mature male biomass:

The mature male biomass time series for models 19.1, 21.1a, 21.1b, and 21.1c are depicted in Figure 26 for EAG and WAG. Mature male biomass tracked the CPUE trends well for selected models for EAG and WAG. The biomass variance was estimated using the Burnham et al. (1987) suggested formula (Equation A.14). We
determined the mature male biomass values on 15 February each year and considered a fixed period time series of recruits (Table T1) for estimating mean number of recruits for the $M M B_{35 \%}$ calculation under a Tier 3 approach.

## c. Fishing mortality:

The full selection pot fishery $F$ values over time for models 19.1, 21.1a, 21.1b, and 21.1c are shown in Figure 25 for EAG and WAG. The $F$ peaked in late 1980s and early to mid-1990s and systematically declined in the EAG. Slight increases in $F$ were observed from 2014 to 2016, followed by a decline in the EAG. On the other hand, the $F$ in the WAG peaked in late 1980s, 1990s, and early 2000s, declined in late 2000s, and slightly increased in 2013-2014 before declining.

## d. F vs. MMB:

We provide these plots for models 21.1a, 21.1b, and 21.1c for EAG and WAG in Figure 42. The 2020/21 $F$ was below the overfishing levels in both regions.
e. Stock-Recruitment relationship: None.

## f. Recruitment:

Temporal changes in total number of recruits to the modeled (19.1, 21.1a, 21.1b, and 21.1c) population are illustrated in Figure 14 for EAG and in Figure 32 for WAG. The recruitment distribution to the model size group ( $101-185 \mathrm{~mm} \mathrm{CL}$ ) is shown in Figures 15 and 33 for EAG and WAG, for the respective models.

## 5. Evaluation of the fit to the data:

## g. Fits to catches:

The fishery retained and total catch, and groundfish bycatch (observed vs. estimated) plots are illustrated in Figure 17 for EAG and in Figure 35 for WAG for models 19.1, 21.1a, 21.1b, and 21.1c. The 1981/82-1984//85 retained catch plots for respective models are depicted in Figures 18 and 36 for EAG and WAG, respectively. All predicted fits were very close to observed values, especially for retained catch and groundfish bycatch mortality. However, pre-1995 total catch data did not fit well.

## h. Survey data plot:

We provide some cooperative pot survey data plots in Appendix C.
i. CPUE index data:

The comparison of predicted CPUE with input indices (open circles with $95 \%$ confidence intervals) for models 19.1, 21.1a, 21.1b, and 21.1c are shown in Figure 24 for EAG and Figure 41 for WAG. The CPUE variance was estimated using the Burnham et al. (1987) suggested formula (Equation A.14). These figures illustrate varying matches of CPUE predictions with input values by different models.

## j. Tagging data:

The predicted vs. observed tag recaptures by length-class for years 1 to 6 post tagging are depicted in Figure 13 for EAG and Figure 31 for WAG. The predictions appear reasonable. Note that we used the EAG tagging information for a fixed size transition matrix estimation for both stocks (EAG and WAG). The size transition matrices estimated using EAG tagging data in the EAG and WAG models were similar.

## k. Molt probability:

The predicted molt probabilities vs. CL are depicted for models 19.1, 21.1a, 21.1b, and 21.1c in Figures 16 for EAG and in Figure 34 for WAG. The fitted curves appear to be satisfactory.

## 1. Fit to catch size compositions:

Retained and total length compositions are shown in Figures 9a, b, c and 10a, b, c for EAG and 27a, b, c and 28a, b, c for WAG for models 21.1a, 21.1b, and 21.1c, respectively. The groundfish discard length compositions for model 21.1a are shown in Figures 11 for EAG and 29 for WAG. The retained and total catch size composition fits appear satisfactory. But the fits to groundfish bycatch size compositions are bad. Note that we did not use the groundfish size compositions in any of the model fits.

We illustrate the standardized residual plots as bubble plots of size composition over time for retained catch (Figures 19 and 20 for EAG, and 37 and 38 for WAG) and for total catch (Figures 21 and 22 for EAG, and 39 and 40 for WAG) for two models (21.1a and 21.1c). The retained catch bubble plots do not appear to exhibit major pronounced patterns among residuals for the selected models.
m . Marginal distributions for the fits to the composition data:
We did not provide this plot in this report.
n. Plots of implied versus input effective sample sizes and time series of implied effective sample sizes:
We did not provide the plots or table values of implied vs. input effective sample sizes in this report. However, we provide the Stage-1 and the optimized re-weighted Stage-2 effective sample sizes in Tables 5 to 7 for EAG and in Tables 15 to 17 for WAG, respectively for models 21.1a, 21.1b, and 21.1c.
o. Tables of RMSEs for the indices:

We did not provide this table in this report.
p. Quantile-quantile (Q-Q) plots:

We did not provide these plots for model fits in this report. However, we provide a Q-Q plot for cooperative survey CPUE fit in Appendix C.

## 6. Retrospective and historical analysis:

The retrospective fits for scenarios 21.1a, 21.1b, and 21.1c are shown in Figure 23 for EAG and WAG. The retrospective fits for the whole time series, 1961 to 2020, did not show severe departure when nine terminal years' data were sequentially removed, especially for WAG, and hence the current formulation of the model appears stable. The modified Mohn rho (1999) values are also given in the figure.

The Mohn rho ( $\rho$ ) formula, modified by Deroba (2014), is:

$$
\text { Mohn } \rho=\frac{\sum_{n=1}^{x} \frac{\left[\widehat{M M B}_{y=T-n, T-n}-\widehat{M M B}_{y=T-n, T}\right]}{\widehat{M B B}_{y=T-n, T}}}{x}
$$

where, $\widehat{M M B}_{y=T-n, T-n}$ is the MMB estimated for terminal year T-n (left subscript) using data up to $\mathrm{T}-\mathrm{n}$ years (right subscript), T is the terminal year of the entire data, x is the total number of peels, most recent year's data is "peeled off" recursively $n$ times, where $\mathrm{n}=1,2,3 \ldots \mathrm{x}$. We used nine peels $(\mathrm{x}=9)$ and our $\mathrm{T}=2020$.

The low values (rule of thumb: closer to zero / between -0.2 to 0.2 ) of Mohn rho indicate no severe model misspecification. The Mohn rho values show no severe model misspecification for WAG. The model 21.1b for EAG shows some reduction in the Mohn rho value compared to that of either model 21.1a or model 21.1c.

A severe drop in modeled biomass from initial MMB occurred when the fishery time series started in 1981 in both regions.

## 7. Uncertainty and sensitivity analysis:

The main task was to determine a plausible size transition matrix to project the population over time. In a previous study, we investigated the sensitivity of the model to determining the size transition matrix by using or not using a molt probability function (Siddeek et al. 2016a). The model fit improved when a molt probability model was included. Therefore, we included a molt probability sub-model for size transition matrix calculation in all model scenarios.

## 8. Conduct 'jitter analysis':

We conducted jitter analysis on the base model 19.1 (Appendix E). The results indicated that global convergence was achieved for most runs.

## F. Calculation of the OFL

## 1. Specification of the Tier level:

In the following section, we provide the Tier 3 method to determine OFL and ABC.
2. List of parameter and stock size estimates (or best available proxies thereof) required by limit and target control rules specified in the fishery management plan:

The critical assumptions for $M M B_{M S Y}$ reference point estimation of Aleutian Islands golden king crab are:
a. Natural mortality is constant.
b. A fixed growth transition matrix is adequately estimated from tagging data and a molt probability sub-model.
c. Total fishery selectivity and retention curves are length-dependent and the 2005/062020/21 period selectivity estimates are applicable.
d. Groundfish bycatch fishery selectivity is kept constant at 1.0 for all length groups.
e. Model estimated recruits (in millions of crab) are valid for different periods considered in selected models.
f. Model estimated groundfish bycatch mortality values are appropriately averaged for the period 2011/12-2020/21 (10 years).
g. The knife-edge $50 \%$ maturity size used for MMB estimation is correct.

## Method:

We simulated the population abundance starting from the model estimated terminal year stock size by length, model estimated parameter values, a fishing mortality value $(F)$, and a constant number of annual recruits. Once stock dynamics were stabilized (we used the $99^{\text {th }}$ year estimates) for an $F$, we calculated the $M M B / R$ for that $F$. We computed the relative $M M B / R$ in percentage, $\left(\frac{M M B}{R}\right)_{x \%}$ (where $\mathrm{x} \%=\frac{\frac{M M B_{F}}{R}}{\frac{M M B_{0}}{R}} \times 100$ and $M M B_{0} / R$ is the virgin $M M B / R$ ) for different $F$ values.
$F_{35 \%}$ is the $F$ value producing an $\mathrm{MMB} / \mathrm{R}$ value equal to $35 \%$ of $M M B_{0} / R$.
$M M B_{35 \%}$ is estimated using the following formula:
$M M B_{35 \%}=\left(\frac{M M B}{R}\right)_{35 \%} \times \bar{R}$,
where $\bar{R}$ is the mean number of model estimated recruits for a selected period.

## 3. Specification of the OFL:

a. Provide the equations (from Amendment 24) on which the OFL is to be based:
$F_{O F L}$ uses Equation A.28. The OFL is estimated by an iterative procedure accounting for intervening total removals (Appendix A).
b. Basis for projecting MMB to the time of mating:

We followed the NPFMC (2007a) guideline.
c. Specification of Fofl, OFL, and other applicable measures (if any) relevant to determining whether the stock is overfished or if overfishing is occurring:

The 2020/21 fishery data indicated that overfishing did not occur (Total Catch < OFL) and the stock did not reach an overfished status (MMB > MSST). Please see Management Performance table below. The OFL and ABC values for $2021 / 22$ in the table below are the authorsrecommended values.

Status and catch specifications (1000 t) of Aleutian Islands golden king crab

| Year | MSST | Biomass <br> (MMB) | TAC | Retained <br> Catch | Total <br> Catch $^{\mathbf{a}}$ | OFL | ABC $^{\text {b }}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $2017 / 18$ | 6.044 | 14.205 | 2.515 | 2.585 | 2.942 | 6.048 | 4.536 |
| $2018 / 19$ | 5.880 | 17.848 | 2.883 | 2.965 | 3.355 | 5.514 | 4.136 |
| $2019 / 20$ | 5.915 | 16.386 | 3.257 | 3.319 | 3.735 | 5.249 | 3.937 |
| $2020 / 21^{\mathrm{d}}$ | $5.998^{\mathrm{d}, \mathrm{c}}$ | $16.215^{\text {d,c }}$ | 2.999 | $2.770^{\mathrm{c}}$ | $3.148^{\mathrm{c}}$ | 4.798 | 3.599 |
| $2020 / 21^{\mathrm{e}}$ | 6.026 | 16.207 |  |  |  |  |  |
| $2020 / 21^{\mathrm{f}}$ | 5.964 | 14.871 |  |  |  |  |  |
| $2020 / 21^{\mathrm{g}}$ | 5.991 | 15.124 |  |  |  |  |  |
| $2021 / 22^{\mathrm{d}}$ |  | 14.821 |  |  |  | 4.796 | 3.615 |
| $2021 / 22^{\mathrm{e}}$ |  | 14.816 |  |  |  | 4.817 | 3.613 |
| $2021 / 22^{\mathrm{f}}$ |  | 13.832 |  |  |  | 4.407 | 3.305 |
| $2021 / 22^{\mathrm{g}}$ |  | 14.048 |  |  |  | 4.471 | 3.353 |

Status and catch specifications (million lb) of Aleutian Islands golden king crab

| Year | MSST | Biomass <br> (MMB) | TAC | Retained <br> Catch | Total <br> Catch $^{\mathbf{a}}$ | OFL | ABC $^{\mathbf{b}}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $2017 / 18$ | 13.325 | 31.315 | 5.545 | 5.699 | 6.487 | 13.333 | 10.000 |
| $2018 / 19$ | 12.964 | 39.348 | 6.356 | 6.536 | 7.396 | 12.157 | 9.118 |
| $2019 / 20$ | 13.041 | 36.124 | 7.180 | 7.317 | 8.234 | 11.572 | 8.679 |
| $2020 / 21^{\mathrm{d}}$ | $13.223^{\mathrm{d}, \mathrm{c}}$ | $35.748^{\mathrm{d}, \mathrm{c}}$ | 6.610 | $6.107^{\mathrm{c}}$ | $6.940^{\mathrm{c}}$ | 10.579 | 7.934 |
| $2020 / 21^{\mathrm{e}}$ | 13.284 | 35.730 |  |  |  |  |  |
| $2020 / 21^{\mathrm{f}}$ | 13.148 | 32.784 |  |  |  |  |  |
| $2020 / 21^{\mathrm{g}}$ | 13.207 | 33.341 |  |  |  |  |  |
| $2021 / 22^{\mathrm{d}}$ |  | 32.675 |  |  |  | 10.573 | 7.969 |
| $2021 / 22^{\mathrm{e}}$ |  | 32.662 |  |  |  | 10.620 | 7.965 |
| $2021 / 22^{\mathrm{f}}$ |  | 30.494 |  |  |  | 9.715 | 7.287 |
| $2021 / 22^{\mathrm{g}}$ |  | 30.971 |  |  |  | 9.857 | 7.393 |

a. Total retained catch plus estimated bycatch mortality of discarded bycatch during crab fisheries and groundfish fisheries.
b. $25 \%$ buffer was applied to total catch OFL to determine ABC.
c. WAG fishery was still operating when the assessment was conducted.
d. Model 19.1(base model)
e. Model 21.1a.
f. Model 21.1b.
g. Model 21.1c.

## 4. Specification of the retained portion of the total catch OFL:

The retained catch portions of the total-catch OFL for EAG, WAG, and the entire Aleutian Islands $(A I=E A G+W A G)$ stock were calculated for the three models (21.1a, 21.1b, and 21.1c):

Model 21.1a:
EAG: $2,801 \mathrm{t}$ ( 6.176 million lb)
WAG: $1,782 \mathrm{t}$ ( 3.928 million lb )
AI: $\quad 4,583 \mathrm{t}$ ( 10.104 million lb).
Model 21.1b:
EAG: 2,386 t (5.260 million lb)
WAG: $1,790 \mathrm{t}$ ( 3.947 million lb)
AI: $\quad 4,176 \mathrm{t}$ ( 9.207 million lb).

Model 21.1c:
EAG: $2,787 \mathrm{t}$ ( 6.143 million lb)
WAG: $1,461 \mathrm{t}$ ( 3.222 million lb)
AI: $\quad 4,248 \mathrm{t}$ ( 9.365 million lb).

## G. Calculation of ABC

We estimated the cumulative probability distribution of OFL assuming a $\log$ normal distribution of OFL. We calculated the OFL at the 0.5 probability and the maximum ABC at the 0.49 probability and considered an additional buffer by setting $\mathrm{ABC}=0.75^{*}$ OFL.

We provide the ABC estimates with the $25 \%$ buffer for EAG, WAG, and AI considering models 21.1a, 21.1b, and 21.1c:

> Model 21.1a:
> EAG: $\mathrm{ABC}=2,200 \mathrm{t}(4.850$ million lb)
> $\mathrm{WAG}: \mathrm{ABC}=1,413 \mathrm{t}(3.115$ million lb$)$
> $\mathrm{AI}: \mathrm{ABC}=3,613 \mathrm{t}(7.965$ million lb$)$.

Model 21.1b:
EAG: $\mathrm{ABC}=1,886 \mathrm{t}$ ( 4.159 million lb)
WAG: $\mathrm{ABC}=1,419 \mathrm{t}(3.128$ million lb$)$
$\mathrm{AI}: \mathrm{ABC}=3,305 \mathrm{t}(7.287$ million lb).
Model 21.1c:
EAG: $\mathrm{ABC}=2,188 \mathrm{t}(4.825$ million lb)
WAG: $\mathrm{ABC}=1,165 \mathrm{t}(2.568$ million lb$)$
$\mathrm{AI}: \mathrm{ABC}=3.353 \mathrm{t}(7.393$ million lb$)$.

## 1. List of variables related to scientific uncertainty:

- Models rely largely on fisheries data.
- Observer and fisheries CPUE indices played a major role in the assessment model.
- Natural mortality, $0.21 \mathrm{yr}^{-1}$, was estimated in the previous model and not independently estimated here.
- The period to compute the average number of recruits relative to the assumption that this represents "a period determined to be representative of the production potential of the stock."
- Fixed bycatch mortality rates were used in each fishery (crab fishery and the groundfish fishery) that discarded golden king crab.
- Discarded catch and bycatch mortality for each fishery in which bycatch occurred during 1981/82-1989/90 were not available.

2. List of additional uncertainties for alternative sigma-b.

We recommend a buffer of $25 \%$ to account for additional uncertainties.

## 3. Author recommended ABC :

Authors recommend three ABC options based on $25 \%$ buffer on the OFL under models 21.1a, 21.1b, and 21.1c.

## H. Rebuilding Analysis

Not applicable. This stock has not been declared overfished.

## I. Data Gaps and Research Priorities

1. Recruit abundances were tied to commercial catch sampling data. The implicit assumption in the analysis was that the estimated recruits come solely from the same exploited stock through growth and mortality. The current analysis did not consider that additional recruitment may occur through immigration from neighboring areas and possibly separate sub-stocks. The analysis also did not consider emigration from the study area, which would result in an assumption of increased $M$ or a reduced estimate of recruits. Extensive tagging experiments or resource surveys are needed to investigate stock distributions.
2. We estimated $M$ in the model. However, an independent estimate of $M$ is needed for comparison, which could be achieved with tagging experiments.
3. An extensive tagging study may provide independent estimates of molting probability and growth. We used historical tagging data to determine the size transition matrix.
4. An arbitrary $20 \%$ handling mortality rate on discarded males was used, which was obtained from the red king crab literature (Kruse et al. 2000; Siddeek 2002). An experimentally based independent estimate of handling mortality is needed for Aleutian Islands golden king crab.
5. The Aleutian King Crab Research Foundation recently initiated crab survey programs in the Aleutian Islands. This program needs to be strengthened and
continued for golden king crab research to address some of the data gaps and establish a fishery independent data source.
6. It is unclear how the recent changes in environmental conditions in the Bering Sea will affect golden king crab growth and survival. Limited length-weight data from the cooperative survey and independent biological sampling in 2018 and 2020 from WAG were used in the current assessment; however, more measurements are needed from both regions to increase the sample size to refine the length-weight model.
7. We used male maturity information to determine MMB. The ADF\&G observer sampling, dock side sampling, and cooperative survey programs collected male maturity data during 2018/19 through 2020/21. Preliminary analysis on these data is presented in this assessment. The CPT previously recommended to collect additional data on small size crab (sublegal) to improve maturity fit. The maturity data collection needs to be continued to accumulate more measurements on small crab.
8. Morphometric measurements provide size at maturity. Ideally, an experimental study under natural environment conditions is needed to collect male size at functional maturity data to determine functional maturity size.

## J. Acknowledgments

We thank ADF\&G personnel in Kodiak, Dutch Harbor, and Juneau for preparing/providing various fisheries and biological data and plots for this assessment. Apart from co-authors, William Bechtol and Katie Palof reviewed an earlier draft of this manuscript. We appreciate their editorial suggestions. We also appreciate technical and editorial input from CPT and SSC members, and industry personnel.

## K. Literature Cited

Adams, C.F., and A.J. Paul. 1999. Phototaxis and geotaxis of light-adapted zoeae of the golden king crab Lithodes aequispinus (Anomura: Lithodidae) in the laboratory. Journal of Crustacean Biology. 19(1): 106-110.
ADF\&G (Alaska Department of Fish and Game). 2002. Annual management report for the shellfish fisheries of the Westward Region, 2001. Alaska Department of Fish and Game, Division of Commercial Fisheries, Regional Information Report 4K02-54, Kodiak, Alaska.

Barnard, D.R., and R. Burt. 2004. Summary of the 2002 mandatory shellfish observer program database for the general and CDQ fisheries. Alaska Department of Fish and Game, Division of Commercial Fisheries, Regional Information Report 4K04-27, Kodiak, Alaska.

Barnard, D.R., R. Burt, and H. Moore. 2001. Summary of the 2000 mandatory shellfish observer program database for the open access fisheries. Alaska Department of Fish and Game, Division of Commercial Fisheries, Regional Information Report 4K01-39, Kodiak, Alaska.

Beers, D.E. 1992. Annual biological summary of the Westward Region shellfish observer database, 1991. Alaska Department of Fish and game, Division of Commercial Fisheries, Regional Information Report 4K92-33, Kodiak.
Blau, S.F., and D. Pengilly. 1994. Findings from the 1991 Aleutian Islands golden king crab survey in the Dutch Harbor and Adak management areas including analysis of recovered tagged crabs. Alaska Department of Fish and Game, Commercial Fisheries Management and Development Division, Regional Information Report 4K94-35, Kodiak.
Blau, S.F., L.J. Watson, and I. Vining. 1998. The 1997 Aleutian Islands golden king crab survey. Alaska Department of Fish and Game, Commercial Fisheries Management and Development Division, Regional Information Report 4K98-30, Kodiak.
Bowers, F.R., M. Schwenzfeier, S. Coleman, B.J. Failor-Rounds, K. Milani, K. Herring, M. Salmon, and M. Albert. 2008. Annual management report for the commercial and subsistence shellfish fisheries of the Aleutian Islands, Bering Sea, and the Westward Region's shellfish observer program, 2006/07. Alaska Department of Fish and Game, Divisions of Sport Fish and Commercial Fisheries, Fishery Management Report No. 08-02, Anchorage, Alaska.

Bowers, F.R., M. Schwenzfeier, K. Herring, M. Salmon, J. Shaishnikoff, H. Fitch, J. Alas, and B. Baechler. 2011. Annual management report for the commercial and subsistence shellfish fisheries of the Aleutian Islands, Bering Sea, and the Westward Region's shellfish observer program, 2009/10. Alaska Department of Fish and Game, Divisions of Sport Fish and Commercial Fisheries, Fishery Management Report No. 11-05, Anchorage, Alaska.
Bozdogan, H. 1987. Model selection and Akaike's Information Criterion (AIC): The general theory and its analytical extensions. Psychometrika, 52, 345-370.

Burnham, K.P., D.R. Anderson, G.C. White, C. Brownie, and K.H. Pollock. 1987. Design and analysis methods for fish survival experiments based on release-recapture. American Fisheries Society, Monograph 5, 437p.

Burnham, K.P. and D.R. Anderson. 2002. Model Selection and Multimodal Inference, A practical Information- Theoretic Approach. 2nd edition. Springer-Verlag, NY, 488p.
Campbell, R.A. 2004. CPUE standardization and the construction of indices of stock abundance in a spatially varying fishery using general linear models. Fish. Res., 70:209-227.
Daly, B., M.S.M. Siddeek, M. Stichert, S. Martell, and J. Zheng. 2019. Recommended harvest strategy for Aleutian Islands golden king crab. Alaska Department of Fish and Game, Fishery Manuscript Series No. 19-03, Anchorage.

Deroba, J.J. 2014. Evaluating the consequences of adjusting fish stock assessment estimates of biomass for retrospective patterns using Mohn's rho. North American Journal of Fisheries Management 34:380-390.
Feenstra, J., R. McGarvey, A. Linnane, M. Haddon, J. Matthews, and A.E. Punt. 2019. Impacts on CPUE from vessel fleet composition changes in an Australian lobster (Jasus edwardsii) fishery, New Zealand Journal of Marine and Freshwater Research, 53: 292-302.

Fournier, D.A., H.J. Skaug, J. Ancheta, J. Ianelli, A. Magnusson, M.N. Maunder, A. Nielsen, and J. Sibert. 2012. AD Model Builder: using automatic differentiation for statistical inference of highly parameterized complex nonlinear models. Optim. Methods Softw. 27:233-249.

Fox, J., and S. Weisberg. 2011. An R Companion to Applied Regression. Second edition. Sage Publications, Inc. 449 p.
Francis, R.I.C.C. 2011. Data weighting in statistical fisheries stock assessment models. Canadian Journal of. Fisheries and Aquatic Sciences 68: 1124-1138.
Francis, R.I.C.C. 2017. Revisiting data weighting in fisheries stock assessment models. Fisheries Research 192: 5-15.
Gaeuman, W.B. 2011. Summary of the 2009/10 mandatory crab observer program database for the BSAI commercial crab fisheries. Fishery Data Series No. 11-04. Alaska Department of Fish and Game, Kodiak.
Gaeuman, W.B. 2014. Summary of the 2013/2014 Mandatory Crab Observer Program Database for the Bering Sea/Aleutian Islands commercial crab fisheries. Alaska Department of Fish and Game, Fishery Data Series No. 14-49, Anchorage.

Gelman A, and J. Hill. 2007. Data Analysis Using Regression and Hierarchical/Multilevel Models. New York: Cambridge University Press; 2007.
Grant W., Siddon C. 2018. Phylogeography and management of golden king crab populations in Alaska. NPRB Project 1526 Final Report. 42 pp.
Hartill, T. 2012. Annual management report for the community development quota and Adak Community Allocation crab fisheries in the Bering Sea and Aleutian Islands, 2010/11. Pages 177-194 in Fitch, H., M. Schwenzfeier, B. Baechler, T. Hartill, M. Salmon, M. Deiman, E. Evans, E. Henry, L. Wald, J. Shaishnikoff, and K. Herring. Annual management report for the commercial and subsistence shellfish fisheries of the Aleutian Islands, Bering Sea, and the Westward Region's Shellfish Observer Program, 2010/11. Alaska Department of Fish and Game, Fishery Management Report No. 12-22, Anchorage.

Hausman, J. 1978. Specification Tests in Econometrics. Econometrica 46, 1251-1271.
Hilsinger, J., C. Siddon, and L. Hulbert. 2016. Cooperative red king crab survey in the Adak area, 2015. Anchorage., Alaska Department of Fish and Game, Fishery Data Series No. 16-18.

Hiramoto, K. 1985. Overview of the golden king crab, Lithodes aequispina, fishery and its fishery biology in the Pacific waters of Central Japan. Pages 297-315, In: Proceedings of the International King Crab Symposium. Alaska Sea Grant College Program, AK-SG-85-12, Fairbanks, Alaska.
Jewett, S.C., N.A. Sloan, and D.A. Somerton. 1985. Size at sexual maturity and fecundity of the fjord-dwelling golden king crab Lithodes aequispina Benedict from northern British Columbia. Journal of Crustacean Biology 5: 377-385.
Koeneman, T.M., and D.V. Buchanan. 1985. Growth of the golden king crab, Lithodes aequispina, in Southeast Alaskan waters. Pages 281-297, in Proceedings of the International King Crab Symposium. Alaska Sea Grant College Program, AK-SG-85-12, Fairbanks, Alaska.

Kruse, G.H., L.C. Byrne, F.C. Funk, S.C. Matulich, and J. Zheng. 2000. Analysis of minimum size limit for the red king crab fishery in Bristol Bay, Alaska. N. Am. J. Fish. Manage. 20:307-319.

Leon, J. M., J. Shaishnikoff, E. Nichols, and M. Westphal. 2017. Annual management report for shellfish fisheries of the Bering Sea-Aleutian Islands management area, 2015/16. Alaska Department of Fish and Game, Fishery Management Report No. 17-10, Anchorage.

Maunder, M.N., and A.E. Punt. 2004. Standardizing catch and effort data: a review of recent approaches. Fisheries Research 70: 141-159.

Mohn, R. 1999. The retrospective problem in sequential population analysis: An investigation using cod fishery and simulated data. ICES Journal of Marine Science 56:473-488.

Moore, H., L.C. Byrne, and M.C. Schwenzfeier. 2000. Summary of the 1999 mandatory shellfish observer program database for the open access fisheries. Alaska Department of Fish and Game, Division of Commercial Fisheries, Regional Information Report 4K00-50, Kodiak, Alaska.
Morrison, R., R.K. Gish, and M. Ruccio. 1998. Annual management report for the shellfish fisheries of the Aleutian Islands. Pages 82-139 in ADF\&G. Annual management report for the shellfish fisheries of the Westward Region. Alaska Department of Fish and Game, Division of Commercial Fisheries, Regional Information Report 4K98-39, Kodiak.
NMFS (National Marine Fisheries Service). 2004. Bering Sea Aleutian Islands Crab Fisheries Final Environmental Impact Statement. National Marine Fisheries Service, Alaska Region, Juneau.

Nichols, E., J. Shaishnikoff, and M. Westphal. 2021. Annual management report for shellfish fisheries of the Bering Sea/Aleutian Islands Management Area, 2019/20. Alaska Department of Fish and Game, Fishery Management Report No. 21-06, Anchorage.

NPFC (North Pacific Fishery Management Council). 2007a. Initial Review Draft: Environmental Assessment for proposed Amendment 24 to the Fishery Management Plan for Bering Sea and Aleutian Islands King and Tanner Crabs to Revise Overfishing Definitions. 17 January 2007. North Pacific Fishery Management Council, Anchorage.
North Pacific Fishery Management Council (NPFMC). 2007b. Public Review Draft: Environmental Assessment for proposed Amendment 24 to the Fishery Management Plan for Bering Sea and Aleutian Islands King and Tanner Crabs to Revise Overfishing Definitions. 14 November 2007. North Pacific Fishery Management Council, Anchorage.

Olson, A.P. 2016. Spatial variability in size at maturity and reproductive timing of golden king crab (Lithodes aequispina) in Southeast Alaska. M.S. thesis, University of Alaska Fairbanks, Fairbanks, Alaska.

Otto, R.S., and P.A. Cummiskey. 1985. Observations on the reproductive biology of golden king crab (Lithodes aequispina) in the Bering Sea and Aleutian Islands. Pages 123-136 in Proceedings of the International King Crab Symposium. Alaska Sea Grant College Program, AK-SG-85-12, Fairbanks, Alaska.

Paul, A.J. and J.M. Paul, 2001. Size of maturity in male golden king crab, Lithodes aequispinus (Anomura: Lithodidae). Journal of Crustacean Biology, 21(2): 384-387.

R Core Team. 2020. R: A language and environment for statistical computing. R Foundation for Statistical Computing, Vienna, Austria. URL http://www.R-project.org/.

Richards, L.J. 1991. Use of contradictory data sources in stock assessments. Fisheries Research 11(3-4): 225-238.

Schnute, J.T. and R. Hilborn. 1993. Analysis of contradictory data sources in fish stock assessment. Canadian Journal of Fisheries and Aquatic Sciences 50(9):1916-1923.
Shirley, T.C., and S. Zhou. 1997. Lecithotrophic development of the golden king crab Lithodes aequispinus (Anomura: Lithodidae). J. Crust. Biol. 17(2):207-216.

Siddeek, M.S.M. 2002. Review of biological reference points used in Bering Sea and Aleutian Islands (king and Tanner) crab management. Alaska Department of Fish and Game, Division of Commercial Fisheries, Regional Information Report 5J02-06, Juneau, Alaska.

Siddeek, M.S.M., D.R. Barnard, L.J. Watson, and R.K. Gish. 2005. A modified catch-length analysis model for golden king crab (Lithodes aequispinus) stock assessment in the eastern Aleutian Islands. Pages 783-805 in Fisheries assessment and management in data limited situations, Alaska Sea Grant College Program, AK-SG-05-02, Fairbanks, Alaska.

Siddeek, M.S.M., J. Zheng, and D. Pengilly. 2015. Aleutian Islands Golden King Crab (Lithodes aequispinus) Model-Based Stock Assessment in Fall 2015. Draft report submitted for the September 2015 Crab Plan Team Meeting. North Pacific Fishery Management Council, Anchorage, Alaska.

Siddeek, M.S.M., J. Zheng, A.E. Punt, and V. Vanek. 2016a. Estimation of size-transition matrices with and without moult probability for Alaska golden king crab using tag-recapture data. Fisheries Research 180:161-168.

Siddeek, M.S.M., J. Zheng, and D. Pengilly 2016b. Standardizing CPUE from the Aleutian Islands golden king crab observer data. Pages 97-116 in T.J. Quinn II, J.L. Armstrong, M.R. Baker, J. Heifetz, and D. Witherell (eds.), Assessing and Managing Data-Limited Fish Stocks. Alaska Sea Grant, University of Alaska Fairbanks, Alaska.
Siddeek, M.S.M., J. Zheng, C. Siddon, B. Daly, and D. Pengilly. 2017. Aleutian Islands golden king crab model-based stock assessment in Spring 2017. CRAB2017SAFE chapter. North Pacific Fishery Management Council, Anchorage, Alaska.

Siddeek, M.S.M., J. Zheng, C. Siddon, B. Daly, J. Runnebaum, and M.J. Westphal. 2018. Aleutian Islands Golden king crab model-based stock assessment in Spring 2018. CRAB2018SAFE chapter. North Pacific Fishery Management Council, Anchorage, Alaska.

Somerton, D.A., and R.S. Otto. 1986. Distribution and reproductive biology of the golden king crab, Lithodes aequispina, in the Eastern Bering Sea. Fishery Bulletin 81(3): 571-584.

Starr, P.J. 2012. Standardized CPUE analysis exploration: using the rock lobster voluntary logbook and observer catch sampling programmes. New Zealand Fisheries Assessment Report 2012/34, 75 p.

Vanek, V., D. Pengilly, and M.S.M. Siddeek. 2013. A study of commercial fishing gear selectivity during the 2012/13 Aleutian Islands golden king crab fishery east of $174^{\circ} \mathrm{W}$ longitude. Alaska Department of Fish and Game, Division of Sport Fish and Commercial Fisheries, Fishery Data Series No. 13-41.

Von Szalay, P.G., C.N. Roper, N.W. Raring, and M.H. Martin. 2011. Data report: 2010 Aleutian Islands bottom trawl survey. U.S. Dep. Commerce., NOAA Technical Memorandum NMFS-AFSC-215.

Watson, L.J. 2004. The 2003 triennial Aleutian Islands golden king crab survey and comparisons to the 1997 and 2000 surveys (revised October 17, 2005). Alaska Department of Fish and Game, Division of Commercial Fisheries, Regional Information Report 4K04-42, Kodiak. [Revised 10/17/2005].
Watson, L.J. 2007. The 2006 triennial Aleutian Islands golden king crab survey. Alaska Department of Fish and Game, Fishery Management Report No. 07-07, Anchorage.
Watson, L.J., and R.K. Gish. 2002. The 2000 Aleutian Islands golden king crab survey and recoveries of tagged crabs in the 1997-1999 and 2000-2002 fishing seasons. Alaska Department of Fish and Game, Division of Commercial Fisheries, Regional Information Report 4K02-6, Kodiak.
Watson, L.J., D. Pengilly, and S.F. Blau. 2002. Growth and molting of golden king crabs (Lithodes aequispinus) in the eastern Aleutian Islands, Alaska. Pages 169-187 in Crabs in cold water regions: biology, management, and economics, Alaska Sea Grant College Program, AK-SG-02-01, Fairbanks, Alaska.
Webb. J. 2014. Reproductive ecology of commercially important Lithodid crabs. Pages 285-314 in B.G. Stevens (ed.). King Crabs of the World: Biology and Fisheries Management. CRC Press, Taylor \& Francis Group, New York.

## Tables

Page intentionally left blank

Table 1. Commercial fishery history for the Aleutian Islands golden king crab fishery 1981/82-2020/21: number of vessels, guideline harvest level (GHL; established in lb, converted to t ) for 1996/97-2004/05, total allowable catch (TAC; established in lb, converted to $t$ ) for 2005/06-2020/21, weight of retained catch (harvest; t), number of retained crab, pot lifts, fishery catch-per-uniteffort (CPUE; retained crab per pot lift), and average weight ( kg ) of landed crab. The values are separated by EAG and WAG beginning in 1996/97.

| Crab Fishing Season | Vessels | GHL/TAC | Harvest ${ }^{\text {a }}$ | Crab | Pot Lifts | CPUE ${ }^{\text {b }}$ | Average Weight ${ }^{\text {c }}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1981/82 | 14-20 | - | 599 | 240,458 | 27,533 | 9 | $2.5{ }^{\text {d }}$ |
| 1982/83 | 99-148 | - | 4,169 | 1,737,109 | 179,472 | 10 | $2.4{ }^{\text {d }}$ |
| 1983/84 | 157-204 | - | 4,508 | 1,773,262 | 256,393 | 7 | $2.5{ }^{\text {d }}$ |
| 1984/85 | 38-51 | - | 2,132 | 971,274 | 88,821 | 11 | $2.2{ }^{\text {e }}$ |
| 1985/86 | 53 | - | 5,776 | 2,816,313 | 236,601 | 12 | $2.1{ }^{\text {f }}$ |
| 1986/87 | 64 | - | 6,685 | 3,345,680 | 433,870 | 8 | $2.0^{\text {f }}$ |
| 1987/88 | 66 | - | 4,199 | 2,177,229 | 307,130 | 7 | $1.9{ }^{\text {f }}$ |
| 1988/89 | 76 | - | 4,820 | 2,488,433 | 321,927 | 8 | $1.9{ }^{\text {f }}$ |
| 1989/90 | 68 | - | 5,453 | 2,902,913 | 357,803 | 8 | $1.9{ }^{\text {f }}$ |
| 1990/91 | 24 | - | 3,153 | 1,707,618 | 215,840 | 8 | $1.9{ }^{\text {f }}$ |
| 1991/92 | 20 | - | 3,494 | 1,847,398 | 234,857 | 8 | $1.9{ }^{\text {f }}$ |
| 1992/93 | 22 | - | 2,854 | 1,528,328 | 203,221 | 8 | $1.9{ }^{\text {f }}$ |
| 1993/94 | 21 | - | 2,518 | 1,397,530 | 234,654 | 6 | $1.8{ }^{\text {f }}$ |
| 1994/95 | 35 | - | 3,687 | 1,924,271 | 386,593 | 5 | $1.9{ }^{\text {f }}$ |


| Crab <br> Fishing <br> Season | Vessels |  | GHL/TAC |  | Harvest ${ }^{\text {a }}$ |  | Crab |  | Pot Lifts |  | CPUE ${ }^{\text {b }}$ |  | Average Weight ${ }^{\text {c }}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1995/96 | 28 |  | - |  | 3,157 |  | 1,582,333 |  | 293,021 |  | 5 |  | $2.0{ }^{\text {f }}$ |  |
|  | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG |
| 1996/97 | 14 | 13 | 1,452 | 1,225 | 1,493 | 1,145 | 731,909 | 602,968 | 113,460 | 99,267 | 7 | 6 | $2.04{ }^{\text {f }}$ | $1.91{ }^{\text {f }}$ |
| 1997/98 | 13 | 9 | 1,452 | 1,225 | 1,588 | 1,109 | 780,610 | 569,550 | 106,403 | 86,811 | 7 | 7 | $2.04{ }^{\text {f }}$ | $1.95{ }^{\text {f }}$ |
| 1998/99 | 14 | 3 | 1,361 | 1,225 | 1,473 | 768 | 740,011 | 410,018 | 83,378 | 35,975 | 9 | 11 | $2.00^{\text {f }}$ | $1.86{ }^{\text {f }}$ |
| 1999/00 | 15 | 15 | 1,361 | 1,225 | 1,392 | 1,256 | 709,332 | 676,558 | 79,129 | 107,040 | 9 | 6 | $1.95{ }^{\text {f }}$ | $1.86{ }^{\text {f }}$ |
| 2000/01 | 15 | 12 | 1,361 | 1,225 | 1,422 | 1,308 | 704,702 | 705,613 | 71,551 | 101,239 | 10 | 7 | $2.00^{\text {f }}$ | $1.86{ }^{\text {f }}$ |
| 2001/02 | 19 | 9 | 1,361 | 1,225 | 1,442 | 1,243 | 730,030 | 686,738 | 62,639 | 105,512 | 12 | 7 | $2.00^{\text {f }}$ | $1.81{ }^{\text {f }}$ |
| 2002/03 | 19 | 6 | 1,361 | 1,225 | 1,280 | 1,198 | 643,886 | 664,823 | 52,042 | 78,979 | 12 | 8 | $2.00^{\text {f }}$ | $1.81{ }^{\text {f }}$ |
| 2003/04 | 18 | 6 | 1,361 | 1,225 | 1,350 | 1,220 | 643,074 | 676,633 | 58,883 | 66,236 | 11 | 10 | $2.09^{\text {f }}$ | $1.81{ }^{\text {f }}$ |
| 2004/05 | 19 | 6 | 1,361 | 1,225 | 1,309 | 1,219 | 637,536 | 685,465 | 34,848 | 56,846 | 18 | 12 | $2.04{ }^{\text {f }}$ | $1.77{ }^{\text {f }}$ |
| 2005/06 | 7 | 3 | 1,361 | 1,225 | 1,300 | 1,204 | 623,971 | 639,368 | 24,569 | 30,116 | 25 | 21 | $2.09^{\text {f }}$ | $1.91{ }^{\text {f }}$ |
| 2006/07 | 6 | 4 | 1,361 | 1,225 | 1,357 | 1,030 | 650,587 | 527,734 | 26,195 | 26,870 | 25 | 20 | $2.09^{\text {f }}$ | $1.95{ }^{\text {f }}$ |
| 2007/08 | 4 | 3 | 1,361 | 1,225 | 1,356 | 1,142 | 633,253 | 600,595 | 22,653 | 29,950 | 28 | 20 | $2.13{ }^{\text {f }}$ | $1.91{ }^{\text {f }}$ |
| 2008/09 | 3 | 3 | 1,361 | 1,286 | 1,426 | 1,150 | 666,946 | 587,661 | 24,466 | 26,200 | 27 | 22 | $2.13{ }^{\text {f }}$ | $1.95{ }^{\text {f }}$ |
| 2009/10 | 3 | 3 | 1,429 | 1,286 | 1,429 | 1,253 | 679,886 | 628,332 | 29,298 | 26,489 | 26 | 24 | $2.09^{\text {f }}$ | $2.00^{\mathrm{f}}$ |
| 2010/11 | 3 | 3 | 1,429 | 1,286 | 1,428 | 1,279 | 670,983 | 626,246 | 25,851 | 29,994 | 26 | 21 | $2.13{ }^{\text {f }}$ | $2.04{ }^{\text {f }}$ |


| Crab Fishing | Vessels |  | GHL/TAC |  | Harvest ${ }^{\text {a }}$ |  | Crab |  | Pot Lifts |  | CPUE ${ }^{\text {b }}$ |  | Average Weight ${ }^{\text {c }}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG |
| 2011/12 | 3 | 3 | 1,429 | 1,286 | 1,429 | 1,276 | 668,828 | 616,118 | 17,915 | 26,326 | 37 | 23 | $2.13{ }^{\text {f }}$ | $2.09^{\text {f }}$ |
| 2012/13 | 3 | 3 | 1,501 | 1,352 | 1,504 | 1,339 | 687,666 | 672,916 | 20,827 | 32,716 | 33 | 21 | $2.18^{\text {f }}$ | $2.00^{\text {f }}$ |
| 2013/14 | 3 | 3 | 1,501 | 1,352 | 1,546 | 1,347 | 720,220 | 686,883 | 21,388 | 41,835 | 34 | 16 | $2.13{ }^{\text {f }}$ | $1.95{ }^{\text {f }}$ |
| 2014/15 | 3 | 2 | 1,501 | 1,352 | 1,554 | 1,217 | 719,064 | 635,312 | 17,002 | 41,548 | 42 | 15 | $2.18^{\text {f }}$ | $1.91{ }^{\text {f }}$ |
| 2015/16 | 3 | 2 | 1,501 | 1,352 | 1,590 | 1,139 | 763,604 | 615,355 | 19,376 | 41,108 | 39 | 15 | $2.09^{\text {f }}$ | $1.85{ }^{\text {f }}$ |
| 2016/17 | 3 | 3 | 1,501 | 1,014 | 1,578 | 1,015 | 793,983 | 543,796 | 24,470 | 38,118 | 32 | 14 | $1.99{ }^{\text {f }}$ | $1.87{ }^{\text {f }}$ |
| 2017/18 | 3 | 3 | 1,501 | 1,014 | 1,571 | 1,014 | 802,610 | 519,051 | 25,516 | 30,885 | 31 | 17 | $1.96{ }^{\text {f }}$ | $1.95{ }^{\text {f }}$ |
| 2018/19 | 3 | 3 | 1,749 | 1,134 | 1,830 | 1,135 | 940,336 | 578,221 | 25,553 | 29,156 | 37 | 20 | $1.95{ }^{\text {f }}$ | $1.96{ }^{\text {f }}$ |
| 2019/20 | 3 | 3 | 1,955 | 1,302 | 2,031 | 1,288 | 1,057,464 | 649,832 | 30,998 | 42,924 | 34 | 15 | $1.92{ }^{\text {f }}$ | $1.98{ }^{\text {f }}$ |
| 2020/21 | 3 | 3 | 1,656 | 1,343 | 1,733 | 1,037* | 902,122 | 560,519* | 30,072 | 37,478* | 30 | $15^{*}$ | $1.92{ }^{\text {f }}$ | $1.85{ }^{\text {f* }}$ |

Note:
a. Includes deadloss.
b. Number of crab per pot lift.
c. Average weight of landed crab, including dead loss.
d. Managed with $6.5^{\prime \prime}$ carapace width (CW) minimum size limit.
e. Managed with $6.5^{\prime \prime} \mathrm{CW}$ minimum size limit west of $171^{\circ} \mathrm{W}$ longitude and $6.0^{\prime \prime}$ minimum size limit east of $171^{\circ} \mathrm{W}$ longitude.
. Managed with $6.0^{\prime \prime}$ minimum size limit.
*As of March 26, 2021, WAG fishery is ongoing.
Catch and effort data include cost recovery fishery.

Table 2. Annual weight of total fishery mortality to Aleutian Islands golden king crab, 1981/82 2020/21, partitioned by source of mortality: retained catch, bycatch mortality during crab fisheries, and bycatch mortality during groundfish fisheries. For bycatch in the federal groundfish fisheries, historical data (1991-2008) are not available for areas east and west of 174 W , and are listed for federal groundfish reporting areas 541,542 , and 543 combined. The 2009- present data are available by separate EAG and WAG fisheries and are listed as such. A mortality rate of $20 \%$ was applied for crab fisheries bycatch, and a mortality rate of $50 \%$ for groundfish pot fisheries and $80 \%$ for the trawl fisheries were applied.

| Season |  |  | Bycatch Mortality by Fishery Type (t) |  |  |  | Total Fishery Mortality (t) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Retained Catch ( t ) |  | Crab |  | Groundfish |  |  |  |  |
|  | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | Entire AI |
| 1981/82 | 490 | 95 |  |  |  |  |  |  | 585 |
| 1982/83 | 1,260 | 2,655 |  |  |  |  |  |  | 3,914 |
| 1983/84 | 1,554 | 2,991 |  |  |  |  |  |  | 4,545 |
| 1984/85 | 1,839 | 424 |  |  |  |  |  |  | 2,263 |
| 1985/86 | 2,677 | 1,996 |  |  |  |  |  |  | 4,673 |
| 1986/87 | 2,798 | 4,200 |  |  |  |  |  |  | 6,998 |
| 1987/88 | 1,882 | 2,496 |  |  |  |  |  |  | 4,379 |
| 1988/89 | 2,382 | 2,441 |  |  |  |  |  |  | 4,823 |
| 1989/90 | 2,738 | 3,028 |  |  |  |  |  |  | 5,766 |
| 1990/91 | 1,623 | 1,621 |  |  |  |  |  |  | 3,244 |
| 1991/92 | 2,035 | 1,397 | 515 | 344 |  | 0 |  |  | 4,291 |
| 1992/93 | 2,112 | 1,025 | 1,206 | 373 |  | 0 |  |  | 4,716 |
| 1993/94 | 1,439 | 686 | 383 | 258 |  | 4 |  |  | 2,770 |
| 1994/95 | 2,044 | 1,540 | 687 | 823 |  | 1 |  |  | 5,095 |
| 1995/96 | 2,259 | 1,203 | 725 | 530 |  | 2 |  |  | 4,719 |
| 1996/97 | 1,738 | 1,259 | 485 | 439 |  | 5 |  |  | 3,926 |
| 1997/98 | 1,588 | 1,083 | 441 | 343 |  | 1 |  |  | 3,455 |
| 1998/99 | 1,473 | 955 | 434 | 285 |  | 1 |  |  | 3,149 |
| 1999/00 | 1,392 | 1,222 | 313 | 385 |  | 3 |  |  | 3,316 |
| 2000/01 | 1,422 | 1,342 | 82 | 437 |  | 2 |  |  | 3,285 |
| 2001/02 | 1,442 | 1,243 | 74 | 387 |  | 0 |  |  | 3,146 |
| 2002/03 | 1,280 | 1,198 | 52 | 303 |  | 8 |  |  | 2,850 |
| 2003/04 | 1,350 | 1,220 | 53 | 148 |  | 0 |  |  | 2,792 |
| 2004/05 | 1,309 | 1,219 | 41 | 143 |  | 1 |  |  | 2,715 |
| 2005/06 | 1,300 | 1,204 | 22 | 73 |  | 2 |  |  | 2,601 |
| 2006/07 | 1,357 | 1,022 | 28 | 81 |  | 8 |  |  | 2,506 |
| 2007/08 | 1,356 | 1,142 | 24 | 114 |  | 9 |  |  | 2,695 |
| 2008/09 | 1,426 | 1,150 | 61 | 102 |  | 3 |  |  | 2,772 |
| 2009/10 | 1,429 | 1,253 | 111 | 108 | 18 | 5 | 1,558 | 1,366 | 2,923 |
| 2010/11 | 1,428 | 1,279 | 123 | 124 | 49 | 3 | 1,600 | 1,407 | 3,006 |
| 2011/12 | 1,429 | 1,276 | 106 | 117 | 25 | 4 | 1,560 | 1,398 | 2,957 |
| 2012/13 | 1,504 | 1,339 | 118 | 145 | 9 | 6 | 1,631 | 1,491 | 3,122 |
| 2013/14 | 1,546 | 1,347 | 113 | 174 | 5 | 7 | 1,665 | 1,528 | 3,192 |


| $2014 / 15$ | 1,554 | 1,217 | 127 | 175 | 9 | 5 | 1,691 | 1,397 | 3,088 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $2015 / 16$ | 1,590 | 1,139 | 165 | 157 | 23 | 2 | 1,778 | 1,298 | 3,076 |
| $2016 / 17$ | 1,578 | 1,015 | 203 | 145 | 101 | 4 | 1,882 | 1,164 | 3,046 |
| $2017 / 18$ | 1,571 | 1,014 | 219 | 126 | 47 | 2 | 1,837 | 1,142 | 2,979 |
| $2018 / 19$ | 1,830 | 1,135 | 240 | 140 | 24 | 3 | 2,094 | 1,278 | 3,372 |
| $2019 / 20$ | 2,031 | 1,288 | 275 | 112 | 18 | 6 | 2,327 | 1,406 | 3,733 |
| $2020 / 21$ | 1,733 | $1,037^{*}$ | 241 | 81 | 40 | 17 | 2,014 | 1,134 | 3,148 |

*As of March 26, 2021, WAG fishery is ongoing.

Table 2a. Time series of estimated total male catch (weight of crabs on the deck without applying any handling mortality) for the EAG and WAG golden king crab stocks (1990/91-2020/21). The crab weights are for the size range $\geq 101 \mathrm{~mm}$ CL and a length-weight formula was used to predict weight at the mid-point of each size bin. NA: no observer sampling to compute catch.

| Year | Total Catch Biomass (t) <br> EAG | Total Catch Biomass (t) <br> WAG |
| :---: | :---: | :---: |
| $1990 / 91$ | 1,405 | 3,657 |
| $1991 / 92$ | 5,861 | 2,555 |
| $1992 / 93$ | 5,532 | 1,508 |
| $1993 / 94$ | NA | 2,804 |
| $1994 / 95$ | 1,971 | 4,911 |
| $1995 / 96$ | 3,711 | 2,115 |
| $1996 / 97$ | 2,052 | 1,754 |
| $1997 / 98$ | 2,540 | 1,789 |
| $1998 / 99$ | 2,783 | 1,079 |
| $1999 / 00$ | 2,275 | 2,079 |
| $2000 / 01$ | 2,554 | 2,215 |
| $2001 / 02$ | 2,099 | 2,123 |
| $2002 / 03$ | 1,806 | 1,880 |
| $2003 / 04$ | 1,825 | 1,856 |
| $2004 / 05$ | 1,629 | 1,871 |
| $2005 / 06$ | 1,737 | 1,793 |
| $2006 / 07$ | 1,636 | 1,553 |
| $2007 / 08$ | 1,818 | 1,610 |
| $2008 / 09$ | 1,815 | 1,733 |
| $2009 / 10$ | 1,771 | 1,687 |
| $2010 / 11$ | 1,755 | 1,597 |
| $2011 / 12$ | 1,770 | 1,523 |
| $2012 / 13$ | 1,948 | 1,831 |
| $2013 / 14$ | 1,839 | 1,916 |
| $2014 / 15$ | 1,966 | 1,593 |
| $2015 / 16$ | 2,125 | 1,558 |
| $2016 / 17$ | 2,234 | 1,568 |
| $2017 / 18$ | 2,376 | 1,435 |
|  |  |  |


| $2018 / 19$ | 2,724 | 1,632 |
| :--- | :---: | :---: |
| $2019 / 20$ | 3,026 | 1,709 |
| $2020 / 21$ | 2,584 | $1,561^{*}$ |

*As of March 26, 2021, WAG fishery is ongoing.
Table 3. Time series of nominal annual pot fishery retained, observer retained, and observer total catch-per-unit-effort (CPUE, number of crabs per pot lift), total pot fishing effort (number of pot lifts), observer sample size (number of sampled pots), and GLM estimated observer CPUE Index (for non-interaction model) for the EAG and WAG golden king crab stocks, 1985/86-2020/21. Observer retained CPUE includes retained and non-retained legal-size crabs.

| Year | Pot Fishery Nominal Retained CPUE |  | Obs. Nominal Retained CPUE |  | Obs. Nominal Total CPUE |  | Pot Fishery Effort (no.pot lifts) |  | Obs. Sample <br> Size (no.pot lifts) |  | Obs. CPUE Index |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG |
| 1985/86 | 11.90 | 11.90 |  |  |  |  | 117,718 | 118,563 |  |  |  |  |
| 1986/87 | 8.42 | 7.32 |  |  |  |  | 155,240 | 277,780 |  |  |  |  |
| 1987/88 | 7.03 | 7.15 |  |  |  |  | 146,501 | 160,229 |  |  |  |  |
| 1988/89 | 7.52 | 7.93 |  |  |  |  | 155,518 | 166,409 |  |  |  |  |
| 1989/90 | 8.49 | 7.83 |  |  |  |  | 155,262 | 202,541 |  |  |  |  |
| 1990/91 | 8.90 | 7.00 | 6.84 | 8.34 | 13.00 | 26.67 | 106,281 | 108,533 | 138 | 340 |  |  |
| 1991/92 | 8.20 | 7.40 | 9.84 | 6.14 | 36.91 | 19.17 | 133,428 | 101,429 | 377 | 857 |  |  |
| 1992/93 | 8.40 | 5.90 | 10.44 | 4.26 | 38.52 | 16.83 | 133,778 | 69,443 | 199 | 690 |  |  |
| 1993/94 | 7.80 | 4.40 | 5.91 | 12.75 | 20.81 | 17.23 | 106,890 | 127,764 | 31 | 174 |  |  |
| 1994/95 | 5.90 | 4.10 | 4.66 | 6.62 | 12.91 | 19.23 | 191,455 | 195,138 | 127 | 1,270 |  |  |
| 1995/96 | 5.90 | 4.70 | 6.03 | 6.03 | 16.98 | 14.28 | 177,773 | 115,248 | 6,388 | 5,598 | 1.00 | 1.16 |
| 1996/97 | 6.50 | 6.10 | 6.02 | 5.90 | 13.81 | 13.54 | 113,460 | 99,267 | 8,360 | 7,194 | 0.94 | 0.98 |
| 1997/98 | 7.30 | 6.60 | 7.99 | 6.72 | 18.25 | 15.03 | 106,403 | 86,811 | 4,670 | 3,985 | 0.87 | 0.98 |
| 1998/99 | 8.90 | 11.40 | 9.82 | 9.43 | 25.77 | 23.09 | 83,378 | 35,975 | 3,616 | 1,876 | 1.00 | 1.09 |
| 1999/00 | 9.00 | 6.30 | 10.28 | 6.09 | 20.77 | 14.49 | 79,129 | 107,040 | 3,851 | 4,523 | 0.92 | 0.91 |
| 2000/01 | 9.90 | 7.00 | 10.40 | 6.46 | 25.39 | 16.64 | 71,551 | 101,239 | 5,043 | 4,740 | 0.82 | 0.84 |
| 2001/02 | 11.70 | 6.50 | 11.73 | 6.04 | 22.48 | 14.66 | 62,639 | 105,512 | 4,626 | 4,454 | 1.04 | 0.82 |
| 2002/03 | 12.40 | 8.40 | 12.70 | 7.47 | 22.59 | 17.37 | 52,042 | 78,979 | 3,980 | 2,509 | 1.10 | 0.91 |
| 2003/04 | 10.90 | 10.20 | 11.34 | 9.33 | 19.43 | 18.17 | 58,883 | 66,236 | 3,960 | 3,334 | 0.97 | 1.16 |
| 2004/05 | 18.30 | 12.10 | 18.34 | 11.14 | 28.48 | 22.45 | 34,848 | 56,846 | 2,206 | 2,619 | 1.44 | 1.25 |
| 2005/06 | 25.40 | 21.20 | 29.52 | 23.89 | 38.55 | 36.23 | 24,569 | 30,116 | 1,193 | 1,365 | 0.98 | 1.14 |
| 2006/07 | 24.80 | 19.60 | 25.13 | 23.93 | 33.39 | 33.47 | 26,195 | 26,870 | 1,098 | 1,183 | 0.80 | 1.07 |
| 2007/08 | 28.00 | 20.00 | 31.10 | 21.01 | 40.38 | 32.46 | 22,653 | 29,950 | 998 | 1,082 | 0.89 | 0.99 |
| 2008/09 | 27.30 | 22.40 | 29.97 | 24.50 | 38.23 | 38.16 | 24,466 | 26,200 | 613 | 979 | 0.88 | 1.11 |
| 2009/10 | 25.90 | 23.70 | 26.60 | 26.54 | 35.88 | 34.08 | 26,298 | 26,489 | 408 | 892 | 0.72 | 1.18 |
| 2010/11 | 26.00 | 20.90 | 26.40 | 22.43 | 37.10 | 29.05 | 25,851 | 29,994 | 436 | 867 | 0.75 | 1.04 |
| 2011/12 | 37.30 | 23.40 | 39.48 | 23.63 | 52.04 | 31.13 | 17,915 | 26,326 | 361 | 837 | 1.08 | 1.08 |
| 2012/13 | 33.02 | 20.57 | 37.82 | 22.88 | 47.57 | 30.76 | 20,827 | 32,716 | 438 | 1,109 | 1.03 | 1.10 |
| 2013/14 | 33.67 | 16.42 | 35.94 | 16.89 | 46.16 | 25.01 | 21,388 | 41,835 | 499 | 1,223 | 1.01 | 0.81 |
| 2014/15 | 42.29 | 15.29 | 47.01 | 15.25 | 60.00 | 22.67 | 17,002 | 41,548 | 376 | 1,137 | 1.33 | 0.77 |
| 2015/16 | 39.41 | 14.97 | 43.27 | 15.81 | 58.68 | 22.14 | 19,376 | 41,108 | 478 | 1,296 | 1.26 | 0.77 |
| 2016/17 | 32.45 | 14.29 | 36.89 | 16.65 | 52.82 | 24.41 | 24,470 | 38,118 | 617 | 1,060 | 1.06 | 0.88 |
| 2017/18 | 31.46 | 16.81 | 35.18 | 19.30 | 54.62 | 25.54 | 25,516 | 30,885 | 585 | 760 | 1.01 | 1.01 |
| 2018/19 | 36.80 | 19.83 | 41.57 | 22.90 | 62.97 | 30.69 | 25,553 | 29,156 | 475 | 688 | 1.24 | 1.23 |
| 2019/20 | 34.11 | 15.10 | 40.88 | 16.30 | 57.46 | 22.73 | 30,998 | 42,963 | 540 | 967 | 1.15 | 0.97 |
| 2020/21 | 30.00 | 14.96* | 36.15 | 16.44* | 57.21 | 24.12* | 30,072 | 37,478* | 552 | 672* | 1.04 | 1.02* |

[^0]Table 4. Time series of negative binomial GLM estimated CPUE indices and coefficient of variation (CV) for the fish ticket based retained catch-per-pot lift for the EAG golden king crab stock. The GLM was fitted to the 1985/86 to 1998/99 time series of data.

| Year | CPUE <br> Index | CV |
| :--- | :---: | :---: |
| $1985 / 86$ | 1.63 | 0.03 |
| $1986 / 87$ | 1.23 | 0.04 |
| $1987 / 88$ | 0.96 | 0.05 |
| $1988 / 89$ | 1.04 | 0.04 |
| $1989 / 90$ | 1.08 | 0.03 |
| $1990 / 91$ | 0.99 | 0.05 |
| $1991 / 92$ | 0.90 | 0.05 |
| $1992 / 93$ | 0.92 | 0.05 |
| $1993 / 94$ | 0.91 | 0.05 |
| $1994 / 95$ | 0.81 | 0.05 |
| $1995 / 96$ | 0.78 | 0.06 |
| $1996 / 97$ | 0.78 | 0.06 |
| $1997 / 98$ | 1.05 | 0.04 |
| $1998 / 99$ | 1.21 | 0.04 |

Table 5. The initial input number of vessel-days/trips and Stage-2 effective sample sizes iteratively estimated by the Francis method for retained, total, and groundfish discard catch size compositions of golden king crab for model 21.1a fit to EAG data. NA: not available.

| Year | Initial <br> Input <br> Retained <br> Vessel- <br> Days <br> Sample <br> Size (no) | Stage-2 <br> Retained <br> Effective <br> Sample <br> Size (no) | Initial <br> Input <br> Total <br> Vessel- <br> Days <br> Sample <br> Size <br> (no) | Stage-2 <br> Total Effective Sample Size (no) | Initial Input Groundfish Trip Sample Size (no) | Stage-2 <br> Groundfish Effective Sample Size (no) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1985/86 | 57 | 49 |  |  |  |  |
| 1986/87 | 11 | 9 |  |  |  |  |
| 1987/88 | 61 | 53 |  |  |  |  |
| 1988/89 | 352 | 304 |  |  |  |  |
| 1989/90 | 792 | 683 |  |  | 9 | 4 |
| 1990/91 | 163 | 141 | 22 | 13 | 13 | 6 |
| 1991/92 | 140 | 121 | 48 | 29 | NA | NA |
| 1992/93 | 49 | 42 | 41 | 24 | 2 | 1 |
| 1993/94 | 340 | 293 | NA | NA | 2 | 1 |
| 1994/95 | 319 | 275 | 34 | 20 | 4 | 2 |
| 1995/96 | 879 | 758 | 1,117 | 665 | 5 | 2 |
| 1996/97 | 547 | 472 | 509 | 303 | 4 | 2 |
| 1997/98 | 538 | 464 | 711 | 423 | 8 | 4 |
| 1998/99 | 541 | 467 | 574 | 342 | 15 | 7 |
| 1999/00 | 463 | 399 | 607 | 361 | 14 | 7 |
| 2000/01 | 436 | 376 | 495 | 295 | 16 | 8 |
| 2001/02 | 488 | 421 | 510 | 303 | 13 | 6 |
| 2002/03 | 406 | 350 | 438 | 261 | 15 | 7 |
| 2003/04 | 405 | 349 | 416 | 248 | 17 | 8 |
| 2004/05 | 280 | 242 | 299 | 178 | 10 | 5 |
| 2005/06 | 266 | 230 | 232 | 138 | 12 | 6 |
| 2006/07 | 234 | 202 | 143 | 85 | 14 | 7 |
| 2007/08 | 199 | 172 | 134 | 80 | 17 | 8 |
| 2008/09 | 197 | 170 | 113 | 67 | 15 | 7 |
| 2009/10 | 170 | 147 | 95 | 57 | 16 | 8 |
| 2010/11 | 183 | 158 | 108 | 64 | 26 | 13 |
| 2011/12 | 160 | 138 | 107 | 64 | 13 | 6 |
| 2012/13 | 187 | 161 | 99 | 59 | 18 | 9 |
| 2013/14 | 193 | 167 | 122 | 73 | 17 | 8 |
| 2014/15 | 168 | 145 | 99 | 59 | 16 | 8 |
| 2015/16 | 190 | 164 | 125 | 74 | 10 | 5 |
| 2016/17 | 223 | 192 | 155 | 92 | 12 | 6 |
| 2017/18 | 213 | 184 | 133 | 79 | 12 | 6 |
| 2018/19 | 218 | 188 | 234 | 139 | 9 | 4 |
| 2019/20 | 214 | 185 | 148 | 88 | 8 | 4 |
| 2020/21 | 227 | 196 | 155 | 92 | 6 | 3 |

Table 6. The initial input number of vessel-days/trips and Stage-2 effective sample sizes iteratively estimated by the Francis method for retained, total, and groundfish discard catch size compositions of golden king crab for model 21.1b fit to EAG data. NA: not available.
$\left.\begin{array}{ccccccc}\hline \text { Year } & \begin{array}{c}\text { Initial } \\ \text { Input } \\ \text { Retained } \\ \text { Vessel- } \\ \text { Days }\end{array} & \begin{array}{c}\text { Stage-2 } \\ \text { Retained } \\ \text { Effective } \\ \text { Sample } \\ \text { Size (no) }\end{array} & \begin{array}{c}\text { Initial } \\ \text { Input } \\ \text { Total } \\ \text { Vessel- } \\ \text { Days }\end{array} & \begin{array}{c}\text { Stage-2 } \\ \text { Total } \\ \text { Effective } \\ \text { Sample } \\ \text { Size (no) }\end{array} & \begin{array}{c}\text { Initial } \\ \text { Input } \\ \text { Sroundfish } \\ \text { Trip } \\ \text { Sample }\end{array} & \begin{array}{c}\text { Stage-2 } \\ \text { Groundfish } \\ \text { Effective } \\ \text { Sample }\end{array} \\ \text { Size (nize (no) }\end{array}\right]$

Table 7. The initial input number of vessel-days/trips and Stage-2 effective sample sizes iteratively estimated by the Francis method for retained, total, and groundfish discard catch size compositions of golden king crab for model 21.1c fit to EAG data. NA: not available.

| Year | Initial <br> Input Retained VesselDays Sample Size (no) | Stage-2 <br> Retained <br> Effective <br> Sample <br> Size (no) | Initial Input Total VesselDays Sample Size (no) | Stage-2 <br> Total <br> Effective <br> Sample <br> Size (no) | Initial Input Groundfish Trip Sample Size (no) | Stage-2 Groundfish Effective Sample Size (no) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1985/86 | 57 | 48 |  |  |  |  |
| 1986/87 | 11 | 9 |  |  |  |  |
| 1987/88 | 61 | 52 |  |  |  |  |
| 1988/89 | 352 | 298 |  |  |  |  |
| 1989/90 | 792 | 670 |  |  | 9 | 4 |
| 1990/91 | 163 | 138 | 22 | 13 | 13 | 6 |
| 1991/92 | 140 | 119 | 48 | 29 | NA | NA |
| 1992/93 | 49 | 41 | 41 | 24 | 2 | 1 |
| 1993/94 | 340 | 288 | NA | NA | 2 | 1 |
| 1994/95 | 319 | 270 | 34 | 20 | 4 | 2 |
| 1995/96 | 879 | 744 | 1,117 | 665 | 5 | 2 |
| 1996/97 | 547 | 463 | 509 | 303 | 4 | 2 |
| 1997/98 | 538 | 455 | 711 | 423 | 8 | 4 |
| 1998/99 | 541 | 458 | 574 | 342 | 15 | 7 |
| 1999/00 | 463 | 392 | 607 | 361 | 14 | 7 |
| 2000/01 | 436 | 369 | 495 | 295 | 16 | 8 |
| 2001/02 | 488 | 413 | 510 | 304 | 13 | 6 |
| 2002/03 | 406 | 344 | 438 | 261 | 15 | 7 |
| 2003/04 | 405 | 343 | 416 | 248 | 17 | 8 |
| 2004/05 | 280 | 237 | 299 | 178 | 10 | 5 |
| 2005/06 | 266 | 225 | 232 | 138 | 12 | 6 |
| 2006/07 | 234 | 198 | 143 | 85 | 14 | 7 |
| 2007/08 | 199 | 168 | 134 | 80 | 17 | 8 |
| 2008/09 | 197 | 167 | 113 | 67 | 15 | 7 |
| 2009/10 | 170 | 144 | 95 | 57 | 16 | 8 |
| 2010/11 | 183 | 155 | 108 | 64 | 26 | 13 |
| 2011/12 | 160 | 135 | 107 | 64 | 13 | 6 |
| 2012/13 | 187 | 158 | 99 | 59 | 18 | 9 |
| 2013/14 | 193 | 163 | 122 | 73 | 17 | 8 |
| 2014/15 | 168 | 142 | 99 | 59 | 16 | 8 |
| 2015/16 | 190 | 161 | 125 | 74 | 10 | 5 |
| 2016/17 | 223 | 189 | 155 | 92 | 12 | 6 |
| 2017/18 | 213 | 180 | 133 | 79 | 12 | 6 |
| 2018/19 | 218 | 185 | 234 | 139 | 9 | 4 |
| 2019/20 | 214 | 181 | 148 | 88 | 8 | 4 |
| 2020/21 | 227 | 192 | 155 | 92 | 6 | 3 |

Table 8. Parameter estimates and coefficient of variations (CV) with the 2020 MMB (MMB estimated on 15 Feb 2021) for models 19.1, 21.1a, 21.1b, and 21.1c for the golden king crab data from the EAG, 1985/86-2020/21. Recruitment and fishing mortality deviations and initial size frequency determination parameters were omitted from this list.

| Parameter | Model 19.1 |  | Model 21.1a |  | Model 21.1b |  | Model 21.1c |  | Limits |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Estimate | CV | Estimate | CV | Estimate | CV | Estimate | CV |  |
| $\log _{\_} \omega_{1}$ ( growth incr. intercept) | 2.538 | 0.01 | 2.538 | 0.01 | 2.539 | 0.01 | 2.538 | 0.01 | 1.0, 4.5 |
| $\omega_{2}$ ( growth incr. slope) | -8.175 | 0.21 | -8.175 | 0.21 | -8.079 | 0.21 | -8.166 | 0.21 | -12.0-5.0 |
| log_a (molt prob. slope) | -2.508 | 0.02 | -2.508 | 0.02 | -2.520 | 0.02 | -2.509 | 0.02 | -4.61-1.39 |
| $\log _{\sim} \mathrm{b}$ (molt prob. L50) | 4.949 | 0.001 | 4.949 | 0.001 | 4.949 | 0.001 | 4.949 | 0.001 | 3.869,5.05 |
| $\sigma$ (growth variability std) | 3.676 | 0.03 | 3.676 | 0.03 | 3.679 | 0.03 | 3.676 | 0.03 | 0.1,12.0 |
| log_total sel delta $\theta$, 1985-04 | 3.384 | 0.02 | 3.384 | 0.02 | 3.397 | 0.02 | 3.390 | 0.02 | 0.,4.4 |
| log_total sel delta $\theta$, 2005-20 | 2.989 | 0.02 | 2.989 | 0.02 | 3.070 | 0.03 | 2.989 | 0.02 | 0.,4.4 |
| $\log _{\text {_ }}$ total sel delta $0,2016-20$ |  |  |  |  | 2.909 | 0.04 |  |  | 0.,4.4 |
| $\log _{-}$ret. sel delta0, 1985-20 | 1.859 | 0.02 | 1.859 | 0.02 | 1.857 | 0.02 | 1.860 | 0.02 | 0.,4.4 |
| log_tot sel $\theta_{50}, 1985-04$ | 4.835 | 0.002 | 4.835 | 0.002 | 4.838 | 0.003 | 4.836 | 0.002 | 4.0,5.0 |
| log_tot sel $\theta_{50}, 2005-20$ | 4.919 | 0.002 | 4.919 | 0.002 | 4.934 | 0.002 | 4.918 | 0.002 | 4.0,5.0 |
| log_tot sel $\theta_{50}, 2016-20$ |  |  |  |  | 4.899 | 0.002 |  |  | 4.0,5.0 |
| $\log _{-}$ret. sel $\theta_{50}, 1985-20$ | 4.915 | 0.0003 | 4.915 | 0.0003 | 4.915 | 0.0003 | 4.915 | 0.0003 | 4.0,5.0 |
| $\log \_\beta_{\mathrm{r}}$ (rec.distribution par.) | -1.080 | 0.17 | -1.080 | 0.17 | -1.076 | 0.17 | -1.079 | 0.17 | -12.0, 12.0 |
| logq2 (catchability 1995-04) | -0.547 | 0.13 | -0.547 | 0.13 | -0.530 | 0.13 | -0.557 | 0.12 | -9.0, 2.25 |
| $\operatorname{logq3}$ (catchability 2005-20) | -0.743 | 0.16 | -0.743 | 0.16 | -0.691 | 0.17 | -0.713 | 0.15 | -9.0, 2.25 |
| log_mean_rec (mean rec.) | 0.840 | 0.05 | 0.840 | 0.05 | 0.825 | 0.05 | 0.840 | 0.05 | 0.01, 5.0 |
| log_mean_Fpot (Pot fishery F) | -0.959 | 0.07 | -0.959 | 0.07 | -0.921 | 0.08 | -0.958 | 0.07 | -15.0, -0.01 |
| log_mean_Fground (GF byc. F) | -9.155 | 0.08 | -9.155 | 0.08 | -9.128 | 0.08 | -9.155 | 0.08 | -15.0, -1.6 |
| $\sigma_{e}^{2}$ (observer CPUE additional var) | 0.058 | 0.36 | 0.058 | 0.36 | 0.035 | 0.38 | 0.018 | 0.64 | 0.0, 0.15 |
| $\sigma_{e}^{2}$ (fishery CPUE additional var) | 0.033 | 0.44 | 0.033 | 0.44 | 0.034 | 0.44 | 0.034 | 0.44 | 0.0,1.0 |
| 2020 MMB | 9,778 | 0.22 | 9,771 | 0.22 | 8,402 | 0.20 | 9,711 | 0.19 |  |

Table 9. Annual abundance estimates of model recruits (millions of crab), legal male biomass ( t ) with coefficient of variations (CV), and mature male biomass ( t ) with CV for model 19.1 for golden king crab in the EAG. Legal male biomass was estimated on July 1 (start of fishing year) of fishing year $y$. Mature male biomass for fishing year y was estimated on February 15 of year $y+1$, after the year y fishery total catch removal. Recruits estimates for 1961 to 2021 are restricted to 19852021. Equilibrium $\mathrm{MMB}_{\mathrm{eq}}$ and $\mathrm{MMB}_{35 \%}$ are also listed.

| Year | Recruits to the Model ( $\geq$ 101 mm CL) | Mature Male Biomass $(\geq 111 \mathrm{~mm} \mathrm{CL})$ | CV | Legal Size Male <br> Biomass ( $\geq 136$ <br> mm CL) | CV |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{aligned} & \mathrm{MMB}_{\mathrm{eq}}=22,976 \\ & M M B_{35 \%}=6,696 \end{aligned}$ |  |  |  |
| 1985 | 1.71 | 9,543 | 0.04 | 9,694 | 0.05 |
| 1986 | 1.01 | 7,321 | 0.04 | 8,224 | 0.04 |
| 1987 | 4.31 | 6,730 | 0.05 | 6,447 | 0.04 |
| 1988 | 3.60 | 6,762 | 0.05 | 5,392 | 0.05 |
| 1989 | 2.02 | 5,910 | 0.06 | 4,854 | 0.07 |
| 1990 | 2.90 | 6,003 | 0.05 | 4,389 | 0.06 |
| 1991 | 3.50 | 6,067 | 0.04 | 4,673 | 0.06 |
| 1992 | 2.25 | 5,985 | 0.04 | 4,485 | 0.05 |
| 1993 | 2.15 | 6,130 | 0.03 | 4,499 | 0.05 |
| 1994 | 2.44 | 5,658 | 0.04 | 4,917 | 0.03 |
| 1995 | 2.322 | 5,076 | 0.04 | 4,470 | 0.04 |
| 1996 | 2.25 | 5,196 | 0.04 | 3,865 | 0.04 |
| 1997 | 3.04 | 5,465 | 0.05 | 4,006 | 0.04 |
| 1998 | 2.80 | 6,047 | 0.05 | 4,122 | 0.05 |
| 1999 | 2.91 | 6,731 | 0.06 | 4,565 | 0.05 |
| 2000 | 2.71 | 7,341 | 0.06 | 5,235 | 0.06 |
| 2001 | 2.04 | 7,689 | 0.06 | 5,864 | 0.06 |
| 2002 | 2.53 | 7,956 | 0.07 | 6,396 | 0.07 |
| 2003 | 2.18 | 8,191 | 0.07 | 6,727 | 0.07 |
| 2004 | 1.89 | 8,215 | 0.07 | 6,944 | 0.07 |
| 2005 | 2.81 | 8,233 | 0.07 | 7,087 | 0.08 |
| 2006 | 2.17 | 8,411 | 0.07 | 6,974 | 0.08 |
| 2007 | 2.10 | 8,389 | 0.07 | 7,068 | 0.08 |
| 2008 | 3.04 | 8,476 | 0.07 | 7,174 | 0.08 |
| 2009 | 1.98 | 8,676 | 0.06 | 7,104 | 0.08 |
| 2010 | 1.84 | 8,472 | 0.06 | 7,305 | 0.07 |
| 2011 | 2.15 | 8,188 | 0.06 | 7,353 | 0.06 |
| 2012 | 1.86 | 7,876 | 0.06 | 7,092 | 0.06 |
| 2013 | 1.60 | 7,360 | 0.06 | 6,775 | 0.06 |
| 2014 | 2.71 | 7,019 | 0.06 | 6,374 | 0.06 |
| 2015 | 2.73 | 7,074 | 0.07 | 5,870 | 0.06 |
| 2016 | 2.92 | 7,280 | 0.08 | 5,654 | 0.07 |
| 2017 | 3.46 | 7809 | 0.10 | 5,755 | 0.08 |
| 2018 | 4.27 | 8,588 | 0.14 | 6,114 | 0.10 |
| 2019 | 3.19 | 9,307 | 0.19 | 6,571 | 0.13 |
| 2020 | 2.56 | 9,778 | 0.22 | 7,291 | 0.19 |
| 2021 | 2.32 |  |  |  |  |

Table 10. Annual abundance estimates of model recruits (millions of crab), legal male biomass ( t ) with coefficient of variations (CV), and mature male biomass ( $t$ ) with CV for model 21.1a for golden king crab in the EAG. Legal male biomass was estimated on July 1 (start of fishing year) of fishing year y. Mature male biomass for fishing year y was estimated on February 15 of year $\mathrm{y}+1$, after the year y fishery total catch removal. Recruits estimates for 1961 to 2021 are restricted to 1985-2021. Equilibrium $\mathrm{MMB}_{\mathrm{eq}}$ and $\mathrm{MMB}_{35 \%}$ are also listed.

| Year | Recruits to the Model ( $\geq$ 101 mm CL) | Mature Male Biomass $(\geq 111 \mathrm{~mm} \mathrm{CL})$ | CV | $\begin{gathered} \text { Legal Size Male } \\ \text { Biomass }(\geq 136 \\ \text { mm CL) } \end{gathered}$ | CV |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{aligned} & \mathrm{MMB}_{\mathrm{eq}}=23,202 \\ & M M B_{35 \%}=6,762 \end{aligned}$ |  |  |  |
| 1985 | 1.71 | 9,543 | 0.04 | 9,695 | 0.05 |
| 1986 | 1.01 | 7,321 | 0.04 | 8,224 | 0.04 |
| 1987 | 4.31 | 6,730 | 0.05 | 6,447 | 0.04 |
| 1988 | 3.60 | 6,762 | 0.05 | 5,392 | 0.05 |
| 1989 | 2.02 | 5,911 | 0.06 | 4,854 | 0.07 |
| 1990 | 2.90 | 6,003 | 0.05 | 4,389 | 0.06 |
| 1991 | 3.50 | 6,067 | 0.04 | 4,673 | 0.06 |
| 1992 | 2.25 | 5,985 | 0.04 | 4,486 | 0.05 |
| 1993 | 2.15 | 6,130 | 0.03 | 4,499 | 0.05 |
| 1994 | 2.44 | 5,658 | 0.04 | 4,917 | 0.03 |
| 1995 | 2.32 | 5,076 | 0.04 | 4,470 | 0.04 |
| 1996 | 2.25 | 5,196 | 0.04 | 3,865 | 0.04 |
| 1997 | 3.04 | 5,465 | 0.05 | 4,006 | 0.04 |
| 1998 | 2.80 | 6,048 | 0.05 | 4,122 | 0.05 |
| 1999 | 2.91 | 6,731 | 0.06 | 4,565 | 0.05 |
| 2000 | 2.71 | 7,341 | 0.06 | 5,235 | 0.06 |
| 2001 | 2.04 | 7,689 | 0.06 | 5,864 | 0.06 |
| 2002 | 2.53 | 7,956 | 0.07 | 6,396 | 0.07 |
| 2003 | 2.18 | 8,191 | 0.07 | 6,727 | 0.07 |
| 2004 | 1.89 | 8,215 | 0.07 | 6,944 | 0.07 |
| 2005 | 2.81 | 8,233 | 0.07 | 7,088 | 0.08 |
| 2006 | 2.17 | 8,411 | 0.07 | 6,974 | 0.08 |
| 2007 | 2.10 | 8,389 | 0.07 | 7,068 | 0.08 |
| 2008 | 3.04 | 8,476 | 0.07 | 7,174 | 0.08 |
| 2009 | 1.98 | 8,676 | 0.06 | 7,104 | 0.08 |
| 2010 | 1.84 | 8,472 | 0.06 | 7,305 | 0.07 |
| 2011 | 2.15 | 8,188 | 0.06 | 7,353 | 0.06 |
| 2012 | 1.86 | 7,876 | 0.06 | 7,091 | 0.06 |
| 2013 | 1.59 | 7,360 | 0.06 | 6,775 | 0.06 |
| 2014 | 2.71 | 7,018 | 0.06 | 6,374 | 0.06 |
| 2015 | 2.73 | 7,073 | 0.07 | 5,870 | 0.06 |
| 2016 | 2.92 | 7,279 | 0.08 | 5,653 | 0.07 |
| 2017 | 3.46 | 7,807 | 0.10 | 5,754 | 0.08 |
| 2018 | 4.27 | 8,584 | 0.14 | 6,113 | 0.10 |
| 2019 | 3.19 | 9,301 | 0.19 | 6,568 | 0.13 |
| 2020 | 2.56 | 9,771 | 0.22 | 7,287 | 0.19 |
| 2021 | 2.32 |  |  |  |  |

Table 11. Annual abundance estimates of model recruits (millions of crab), legal male biomass ( t ) with coefficient of variations (CV), and mature male biomass ( t ) with CV for model 21.1 b for golden king crab in the EAG. Legal male biomass was estimated on July 1 (start of fishing year) of fishing year $y$. Mature male biomass for fishing year y was estimated on February 15 of year $\mathrm{y}+1$, after the year y fishery total catch removal. Recruits estimates for 1961 to 2021 are restricted to 1985-2021. Equilibrium MMBeq and $\mathrm{MMB}_{35 \%}$ are also listed.

| Year | Recruits to the Model ( $\geq$ 101 mm CL) | $\begin{gathered} \text { Mature Male } \\ \text { Biomass } \\ (\geq 111 \mathrm{~mm} \mathrm{CL}) \end{gathered}$ | CV | Legal Size Male <br> Biomass ( $\geq 136$ mm CL) | CV |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{aligned} & \mathrm{MMB}_{\mathrm{eq}}=22,941 \\ & M M B_{35 \%}=6,649 \end{aligned}$ |  |  |  |
| 1985 | 1.70 | 9,509 | 0.04 | 9,650 | 0.06 |
| 1986 | 1.01 | 7,290 | 0.04 | 8,186 | 0.04 |
| 1987 | 4.30 | 6,703 | 0.05 | 6,412 | 0.04 |
| 1988 | 3.62 | 6,737 | 0.05 | 5,363 | 0.05 |
| 1989 | 2.01 | 5,890 | 0.06 | 4,818 | 0.07 |
| 1990 | 2.92 | 5,985 | 0.05 | 4,355 | 0.07 |
| 1991 | 3.49 | 6,053 | 0.04 | 4,641 | 0.06 |
| 1992 | 2.25 | 5,971 | 0.04 | 4,460 | 0.05 |
| 1993 | 2.15 | 6,119 | 0.03 | 4,474 | 0.05 |
| 1994 | 2.43 | 5,645 | 0.03 | 4,895 | 0.03 |
| 1995 | 2.29 | 5,051 | 0.04 | 4,450 | 0.04 |
| 1996 | 2.21 | 5,145 | 0.04 | 3,841 | 0.04 |
| 1997 | 2.96 | 5,372 | 0.05 | 3,963 | 0.04 |
| 1998 | 2.69 | 5,887 | 0.05 | 4,050 | 0.05 |
| 1999 | 2.79 | 6,489 | 0.05 | 4,441 | 0.05 |
| 2000 | 2.56 | 7,005 | 0.06 | 5,039 | 0.06 |
| 2001 | 1.93 | 7,260 | 0.06 | 5,586 | 0.06 |
| 2002 | 2.41 | 7,452 | 0.07 | 6,026 | 0.06 |
| 2003 | 2.12 | 7,636 | 0.07 | 6,280 | 0.07 |
| 2004 | 1.85 | 7,654 | 0.07 | 6,433 | 0.07 |
| 2005 | 2.77 | 7,686 | 0.07 | 6,551 | 0.08 |
| 2006 | 2.12 | 7,885 | 0.07 | 6,448 | 0.08 |
| 2007 | 2.07 | 7,889 | 0.07 | 6,565 | 0.08 |
| 2008 | 3.11 | 8,033 | 0.07 | 6,696 | 0.08 |
| 2009 | 2.09 | 8,349 | 0.07 | 6,658 | 0.08 |
| 2010 | 1.93 | 8,283 | 0.06 | 6,936 | 0.07 |
| 2011 | 2.30 | 8,140 | 0.06 | 7,108 | 0.07 |
| 2012 | 2.04 | 7,991 | 0.06 | 6,981 | 0.06 |
| 2013 | 1.77 | 7,641 | 0.06 | 6,808 | 0.06 |
| 2014 | 2.51 | 7,355 | 0.06 | 6,563 | 0.06 |
| 2015 | 2.67 | 7,313 | 0.07 | 6,197 | 0.07 |
| 2016 | 2.79 | 7,414 | 0.07 | 5,947 | 0.07 |
| 2017 | 3.11 | 7,743 | 0.09 | 5,937 | 0.08 |
| 2018 | 3.43 | 8,097 | 0.12 | 6,153 | 0.09 |
| 2019 | 2.70 | 8,248 | 0.17 | 6,349 | 0.11 |
| 2020 | 2.51 | 8,402 | 0.20 | 6,561 | 0.16 |
| 2021 | 2.28 |  |  |  |  |

Table 12. Annual abundance estimates of model recruits (millions of crab), legal male biomass ( t ) with coefficient of variations (CV), and mature male biomass ( t ) with CV for model 21.1c for golden king crab in the EAG. Legal male biomass was estimated on July 1 (start of fishing year) of fishing year $y$. Mature male biomass for fishing year y was estimated on February 15 of year $\mathrm{y}+1$, after the year y fishery total catch removal. Recruits estimates for 1961 to 2021 are restricted to 1985-2021. Equilibrium MMBeq and $\mathrm{MMB}_{35 \%}$ are also listed.

| Year | Recruits to the Model ( $\geq$ 101 mm CL) | $\begin{gathered} \text { Mature Male } \\ \text { Biomass } \\ (\geq 111 \mathrm{~mm} \mathrm{CL}) \end{gathered}$ | CV | Legal Size Male <br> Biomass ( $\geq 136$ mm CL) | CV |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{aligned} & \mathrm{MMB}_{\mathrm{eq}}=23,196 \\ & M M B_{35 \%}=6,755 \end{aligned}$ |  |  |  |
| 1985 | 1.71 | 9,557 | 0.04 | 9,718 | 0.06 |
| 1986 | 1.02 | 7,332 | 0.04 | 8,241 | 0.04 |
| 1987 | 4.32 | 6,744 | 0.05 | 6,456 | 0.04 |
| 1988 | 3.58 | 6,776 | 0.05 | 5,401 | 0.05 |
| 1989 | 2.03 | 5,918 | 0.06 | 4,865 | 0.07 |
| 1990 | 2.89 | 6,008 | 0.05 | 4,396 | 0.06 |
| 1991 | 3.49 | 6,064 | 0.04 | 4,678 | 0.06 |
| 1992 | 2.22 | 5,973 | 0.04 | 4,485 | 0.05 |
| 1993 | 2.14 | 6,105 | 0.03 | 4,491 | 0.05 |
| 1994 | 2.44 | 5,628 | 0.03 | 4,898 | 0.03 |
| 1995 | 2.32 | 5,050 | 0.04 | 4,440 | 0.04 |
| 1996 | 2.25 | 5,173 | 0.04 | 3,837 | 0.04 |
| 1997 | 3.03 | 5,443 | 0.05 | 3,980 | 0.04 |
| 1998 | 2.80 | 6,024 | 0.05 | 4,100 | 0.05 |
| 1999 | 2.92 | 6,713 | 0.05 | 4,541 | 0.05 |
| 2000 | 2.70 | 7,327 | 0.06 | 5,212 | 0.06 |
| 2001 | 2.05 | 7,675 | 0.06 | 5,848 | 0.06 |
| 2002 | 2.55 | 7,950 | 0.06 | 6,380 | 0.06 |
| 2003 | 2.14 | 8,186 | 0.07 | 6,714 | 0.06 |
| 2004 | 1.86 | 8,193 | 0.07 | 6,940 | 0.07 |
| 2005 | 2.78 | 8,188 | 0.07 | 7,076 | 0.07 |
| 2006 | 2.17 | 8,348 | 0.07 | 6,944 | 0.07 |
| 2007 | 2.11 | 8,330 | 0.07 | 7,013 | 0.07 |
| 2008 | 3.03 | 8,424 | 0.07 | 7,111 | 0.07 |
| 2009 | 2.00 | 8,634 | 0.06 | 7,048 | 0.07 |
| 2010 | 1.87 | 8,454 | 0.06 | 7,257 | 0.07 |
| 2011 | 2.21 | 8,205 | 0.06 | 7,320 | 0.06 |
| 2012 | 1.90 | 7,937 | 0.05 | 7,085 | 0.06 |
| 2013 | 1.60 | 7,447 | 0.06 | 6,810 | 0.06 |
| 2014 | 2.70 | 7,107 | 0.06 | 6,445 | 0.06 |
| 2015 | 2.73 | 7,149 | 0.07 | 5,957 | 0.06 |
| 2016 | 2.94 | 7,347 | 0.07 | 5,732 | 0.07 |
| 2017 | 3.41 | 7,859 | 0.09 | 5,817 | 0.08 |
| 2018 | 4.20 | 8,587 | 0.12 | 6,170 | 0.09 |
| 2019 | 3.17 | 9,258 | 0.16 | 6,595 | 0.11 |
| 2020 | 2.56 | 9,711 | 0.19 | 7,263 | 0.16 |
| 2021 | 2.32 |  |  |  |  |

Table 13. Negative log-likelihood values of the fits for models 19.1 (last year's accepted model with additional 2020/21 data), 20.1a, 21.1b, and 21.1c for golden king crab in the EAG. Likelihood components with zero entry in the entire rows are omitted. RetdcatchB = retained catch biomass.

| Likelihood Component | Model 19.1 | Model 21.1a | Model 21.1b | Model 21.1c | 21.1a-19.1 | 21.1b-19.1 | 21.1c-19.1 |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  |  |  |
| Number of free parameters | 152 | 152 | 154 | 152 |  |  |  |
| Retlencomp | -1334.9600 | -1334.9600 | -1325.3100 | -1331.1500 | 0 | 9.65 | 3.81 |
| Totallencomp | -1486.6000 | -1486.6000 | -1481.4000 | -1487.0300 | 0 | 5.2 | -0.43 |
| Observer cpue | 0.1376 | 0.1441 | -6.5913 | -6.3306 | 0.0065 | -6.7289 | -6.4682 |
| RetdcatchB | 7.6797 | 7.6795 | 7.5856 | 7.8838 | -0.0002 | -0.0941 | 0.2041 |
| TotalcatchB | 23.7070 | 23.7071 | 23.5424 | 23.7389 | 0.00010 | -0.1646 | 0.0319 |
| GdiscdcatchB | 0.0003 | 0.0003 | 0.0003 | 0.0003 | 0 | 0 | 0 |
| Rec_dev | 6.8557 | 6.8548 | 5.7315 | 6.6954 | -0.0009 | -1.1242 | -0.1603 |
| Pot F_dev | 0.0128 | 0.0128 | 0.0120 | 0.0128 | 0 | -0.0008 | 0 |
| Gbyc_F_dev | 0.0267 | 0.0267 | 0.0261 | 0.0265 | 0 | -0.0006 | -0.0002 |
| Tag | 2692.6000 | 2692.6000 | 2692.9000 | 2692.5700 | 0 | 0.3 | -0.03 |
| Fishery cpue | -3.5625 | -3.5628 | -3.3106 | -3.2839 | -0.0003 | 0.2519 | 0.2786 |
| RetcatchN | 0.0062 | 0.0062 | 0.0058 | 0.0061 | 0 | -0.0004 | -0.00010 |
| Total | -94.0875 | -94.0851 | -86.8140 | -96.8547 | 0.0024 | 7.2735 | -2.7672 |

Table 14. Time series of negative binomial GLM estimated CPUE indices and coefficient of variations (CV) for the fish ticket based retained catch-per-pot lift for the WAG golden king crab stock. The GLM was fitted to the 1985/86 to 1998/99 time series of data. GLM predictor variables were selected by R square criteria.

| Year | CPUE <br> Index | CV |
| :---: | :---: | :---: |
| $1985 / 86$ | 2.07 | 0.02 |
| $1986 / 87$ | 1.59 | 0.03 |
| $1987 / 88$ | 1.22 | 0.04 |
| $1988 / 89$ | 1.41 | 0.02 |
| $1989 / 90$ | 1.15 | 0.02 |
| $1990 / 91$ | 0.87 | 0.04 |
| $1991 / 92$ | 0.76 | 0.05 |
| $1992 / 93$ | 0.61 | 0.07 |
| $1993 / 94$ | 0.76 | 0.07 |
| $1994 / 95$ | 0.83 | 0.04 |
| $1995 / 96$ | 0.90 | 0.04 |
| $1996 / 97$ | 0.84 | 0.04 |
| $1997 / 98$ | 0.76 | 0.04 |
| $1998 / 99$ | 1.06 | 0.03 |

Table 15. The initial input number of vessel-days/trips and Stage-2 effective sample sizes iteratively estimated by the Francis method for retained, total, and groundfish discard catch size compositions of golden king crab for model 21.1 a model fit to WAG data. NA: not available.

| Year | Initial <br> Input <br> Retained <br> Vessel- <br> Days <br> Sample <br> Size (no) | Stage-2 <br> Retained <br> Effective <br> Sample <br> Size (no) | Initial Input Total VesselDays Sample Size (no) | Stage-2 <br> Total Effective <br> Sample <br> Size (no) | Initial Input Groundfish Trip Sample Size (no) | Stage-2 <br> Groundfish Effective Sample Size (no) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1985/86 | 45 | 24 |  |  |  |  |
| 1986/87 | 23 | 12 |  |  |  |  |
| 1987/88 | 8 | 4 |  |  |  |  |
| 1988/89 | 286 | 152 |  |  |  |  |
| 1989/90 | 513 | 274 |  |  | 7 | 5 |
| 1990/91 | 205 | 109 | 190 | 93 | 6 | 4 |
| 1991/92 | 102 | 54 | 104 | 51 | 1 | 1 |
| 1992/93 | 76 | 41 | 94 | 46 | 3 | 2 |
| 1993/94 | 378 | 202 | 62 | 30 | NA | NA |
| 1994/95 | 367 | 196 | 119 | 58 | 2 | , |
| 1995/96 | 705 | 376 | 907 | 445 | 5 | 3 |
| 1996/97 | 817 | 436 | 1,061 | 520 | 8 | 5 |
| 1997/98 | 984 | 525 | 1,116 | 547 | 6 | 4 |
| 1998/99 | 613 | 327 | 638 | 313 | 14 | 9 |
| 1999/00 | 915 | 488 | 1,155 | 567 | 18 | 12 |
| 2000/01 | 1,029 | 549 | 1,205 | 591 | 11 | 7 |
| 2001/02 | 898 | 479 | 975 | 478 | 11 | 7 |
| 2002/03 | 628 | 335 | 675 | 331 | 16 | 10 |
| 2003/04 | 688 | 367 | 700 | 343 | 8 | 5 |
| 2004/05 | 449 | 239 | 488 | 239 | 9 | 6 |
| 2005/06 | 337 | 180 | 220 | 108 | 6 | 4 |
| 2006/07 | 337 | 180 | 321 | 157 | 14 | 9 |
| 2007/08 | 276 | 147 | 257 | 126 | 17 | 11 |
| 2008/09 | 318 | 170 | 258 | 127 | 19 | 12 |
| 2009/10 | 362 | 193 | 292 | 143 | 24 | 15 |
| 2010/11 | 328 | 175 | 222 | 109 | 13 | 8 |
| 2011/12 | 295 | 157 | 252 | 124 | 14 | 9 |
| 2012/13 | 288 | 154 | 241 | 118 | 18 | 12 |
| 2013/14 | 327 | 174 | 236 | 116 | 17 | 11 |
| 2014/15 | 305 | 163 | 219 | 107 | 18 | 12 |
| 2015/16 | 287 | 153 | 243 | 119 | 10 | 6 |
| 2016/17 | 392 | 209 | 253 | 124 | 12 | 8 |
| 2017/18 | 299 | 159 | 222 | 109 | 10 | 6 |
| 2018/19 | 328 | 175 | 318 | 156 | 5 | 3 |
| 2019/20 | 338 | 180 | 224 | 110 | 6 | 4 |
| 2020/21 | 372 | 198 | 176 | 86 | 7 | 5 |

Table 16. The initial input number of vessel-days/trips and Stage-2 effective sample sizes iteratively estimated by the Francis method for retained, total, and groundfish discard catch size compositions of golden king crab for model 21.1b model fit to WAG data. NA: not available.

| Year | Initial <br> Input Retained VesselDays Sample Size (no) | Stage-2 <br> Retained <br> Effective <br> Sample <br> Size (no) | Initial <br> Input <br> Total VesselDays Sample Size (no) | Stage-2 <br> Total Effective Sample Size (no) | Initial Input Groundfish Trip Sample Size (no) | Stage-2 <br> Groundfish Effective Sample Size (no) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1985/86 | 45 | 24 |  |  |  |  |
| 1986/87 | 23 | 12 |  |  |  |  |
| 1987/88 | 8 | 4 |  |  |  |  |
| 1988/89 | 286 | 153 |  |  |  |  |
| 1989/90 | 513 | 275 |  |  | 7 | 5 |
| 1990/91 | 205 | 110 | 190 | 94 | 6 | 4 |
| 1991/92 | 102 | 55 | 104 | 51 | 1 | 1 |
| 1992/93 | 76 | 41 | 94 | 46 | 3 | 2 |
| 1993/94 | 378 | 202 | 62 | 31 | NA | NA |
| 1994/95 | 367 | 197 | 119 | 59 | 2 | 1 |
| 1995/96 | 705 | 377 | 907 | 448 | 5 | 3 |
| 1996/97 | 817 | 437 | 1,061 | 524 | 8 | 5 |
| 1997/98 | 984 | 527 | 1,116 | 551 | 6 | 4 |
| 1998/99 | 613 | 328 | 638 | 315 | 14 | 9 |
| 1999/00 | 915 | 490 | 1,155 | 570 | 18 | 12 |
| 2000/01 | 1,029 | 551 | 1,205 | 595 | 11 | 7 |
| 2001/02 | 898 | 481 | 975 | 481 | 11 | 7 |
| 2002/03 | 628 | 336 | 675 | 333 | 16 | 10 |
| 2003/04 | 688 | 368 | 700 | 346 | 8 | 5 |
| 2004/05 | 449 | 240 | 488 | 241 | 9 | 6 |
| 2005/06 | 337 | 180 | 220 | 109 | 6 | 4 |
| 2006/07 | 337 | 180 | 321 | 158 | 14 | 9 |
| 2007/08 | 276 | 148 | 257 | 127 | 17 | 11 |
| 2008/09 | 318 | 170 | 258 | 127 | 19 | 12 |
| 2009/10 | 362 | 194 | 292 | 144 | 24 | 15 |
| 2010/11 | 328 | 176 | 222 | 110 | 13 | 8 |
| 2011/12 | 295 | 158 | 252 | 124 | 14 | 9 |
| 2012/13 | 288 | 154 | 241 | 119 | 18 | 12 |
| 2013/14 | 327 | 175 | 236 | 116 | 17 | 11 |
| 2014/15 | 305 | 163 | 219 | 108 | 18 | 12 |
| 2015/16 | 287 | 154 | 243 | 120 | 10 | 6 |
| 2016/17 | 392 | 210 | 253 | 125 | 12 | 8 |
| 2017/18 | 299 | 160 | 222 | 110 | 10 | 6 |
| 2018/19 | 328 | 176 | 318 | 157 | 5 | 3 |
| 2019/20 | 338 | 181 | 224 | 111 | 6 | 4 |
| 2020/21 | 372 | 199 | 176 | 87 | 7 | 5 |

Table 17. The initial input number of vessel-days/trips and Stage-2 effective sample sizes iteratively estimated by the Francis method for retained, total, and groundfish discard catch size compositions of golden king crab for model 21.1c model fit to WAG data. NA: not available.

| Year | Initial <br> Input Retained VesselDays Sample Size (no) | Stage-2 <br> Retained <br> Effective <br> Sample <br> Size (no) | Initial <br> Input <br> Total VesselDays Sample Size (no) | Stage-2 <br> Total Effective Sample Size (no) | Initial Input Groundfish Trip Sample Size (no) | Stage-2 <br> Groundfish Effective Sample Size (no) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1985/86 | 45 | 24 |  |  |  |  |
| 1986/87 | 23 | 12 |  |  |  |  |
| 1987/88 | 8 | 4 |  |  |  |  |
| 1988/89 | 286 | 154 |  |  |  |  |
| 1989/90 | 513 | 275 |  |  | 7 | 5 |
| 1990/91 | 205 | 110 | 190 | 96 | 6 | 4 |
| 1991/92 | 102 | 55 | 104 | 53 | 1 | 1 |
| 1992/93 | 76 | 41 | 94 | 47 | 3 | 2 |
| 1993/94 | 378 | 203 | 62 | 31 | NA | NA |
| 1994/95 | 367 | 197 | 119 | 60 | 2 | 1 |
| 1995/96 | 705 | 378 | 907 | 458 | 5 | 3 |
| 1996/97 | 817 | 439 | 1,061 | 536 | 8 | 5 |
| 1997/98 | 984 | 528 | 1,116 | 564 | 6 | 4 |
| 1998/99 | 613 | 329 | 638 | 322 | 14 | 9 |
| 1999/00 | 915 | 491 | 1,155 | 583 | 18 | 12 |
| 2000/01 | 1,029 | 552 | 1,205 | 609 | 11 | 7 |
| 2001/02 | 898 | 482 | 975 | 492 | 11 | 7 |
| 2002/03 | 628 | 337 | 675 | 341 | 16 | 10 |
| 2003/04 | 688 | 369 | 700 | 354 | 8 | 5 |
| 2004/05 | 449 | 241 | 488 | 246 | 9 | 6 |
| 2005/06 | 337 | 181 | 220 | 111 | 6 | 4 |
| 2006/07 | 337 | 181 | 321 | 162 | 14 | 9 |
| 2007/08 | 276 | 148 | 257 | 130 | 17 | 11 |
| 2008/09 | 318 | 171 | 258 | 130 | 19 | 12 |
| 2009/10 | 362 | 194 | 292 | 147 | 24 | 16 |
| 2010/11 | 328 | 176 | 222 | 112 | 13 | 8 |
| 2011/12 | 295 | 158 | 252 | 127 | 14 | 9 |
| 2012/13 | 288 | 155 | 241 | 122 | 18 | 12 |
| 2013/14 | 327 | 176 | 236 | 119 | 17 | 11 |
| 2014/15 | 305 | 164 | 219 | 111 | 18 | 12 |
| 2015/16 | 287 | 154 | 243 | 123 | 10 | 7 |
| 2016/17 | 392 | 210 | 253 | 128 | 12 | 8 |
| 2017/18 | 299 | 161 | 222 | 112 | 10 | 7 |
| 2018/19 | 328 | 176 | 318 | 161 | 5 | 3 |
| 2019/20 | 338 | 181 | 224 | 113 | 6 | 4 |
| 2020/21 | 372 | 200 | 176 | 89 | 7 | 5 |

Table 18. Parameter estimates and coefficient of variations (CV) with the 2020 MMB (MMB estimated on 15 Feb 2021 ) for models 19.1, 21.1a, 21.1b, and 21.1c for the golden king crab data from the WAG, 1985/86-2020/21. Recruitment and fishing mortality deviations and initial size frequency determination parameters were omitted from this list.

| Parameter | Model 19.1 |  | Model 21.1a |  | Model 21.1b |  | Model 21.1c |  | Limits |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Estimate | CV | Estimate | CV | Estimate | CV | Estimate | CV |  |
| $\log _{\_} \omega_{1}$ ( growth incr. intercept) | 2.538 | 0.01 | 2.538 | 0.01 | 2.538 | 0.01 | 2.539 | 0.01 | 1.0, 4.5 |
| $\omega_{2}$ ( growth incr. slope) | -7.479 | 0.23 | -7.479 | 0.23 | -7.469 | 0.23 | -7.441 | -0.23 | -12.0-5.0 |
| log_a (molt prob. slope) | -2.644 | 0.03 | -2.644 | 0.03 | -2.645 | 0.03 | -2.647 | -0.03 | -4.61-1.39 |
| $\log _{\sim} \mathrm{b}$ (molt prob. L50) | 4.948 | 0.001 | 4.948 | 0.001 | 4.948 | 0.001 | 4.948 | 0.001 | 3.869,5.05 |
| $\sigma$ (growth variability std) | 3.691 | 0.03 | 3.691 | 0.03 | 3.691 | 0.03 | 3.692 | 0.03 | 0.1,12.0 |
| log_total sel delta $\theta$, 1985-04 | 3.405 | 0.01 | 3.405 | 0.01 | 3.406 | 0.01 | 3.408 | 0.01 | 0.,4.4 |
| log_total sel delta $0,2005-20$ | 2.878 | 0.02 | 2.878 | 0.02 | 2.898 | 0.03 | 2.882 | 0.02 | 0.,4.4 |
| $\log _{-}$total sel delta $0,2016-20$ |  |  |  |  | 2.841 | 0.03 |  |  | 0.,4.4 |
| $\log _{-}$ret. sel delta $\theta$, 1985-20 | 1.783 | 0.02 | 1.783 | 0.02 | 1.783 | 0.02 | 1.784 | 0.02 | 0.,4.4 |
| log_tot sel $\theta_{50}, 1985-04$ | 4.869 | 0.002 | 4.869 | 0.002 | 4.870 | 0.002 | 4.870 | 0.002 | 4.0,5.0 |
| log_tot sel $\theta_{50}, 2005-20$ | 4.901 | 0.001 | 4.901 | 0.001 | 4.901 | 0.002 | 4.901 | 0.001 | 4.0,5.0 |
| log_tot sel $\theta_{50}, 2016-20$ |  |  |  |  | 4.901 | 0.002 |  |  | 4.0,5.0 |
| $\log _{-}$ret. sel $\theta_{50}, 1985-20$ | 4.915 | 0.0002 | 4.915 | 0.0002 | 4.915 | 0.0002 | 4.915 | 0.0002 | 4.0,5.0 |
| $\log \_\beta_{\mathrm{r}}$ (rec.distribution par.) | -1.029 | 0.16 | -1.029 | 0.16 | -1.030 | 0.16 | -1.028 | -0.16 | -12.0, 12.0 |
| logq2 (catchability 1995-04) | -0.032 | 2.12 | -0.032 | 2.12 | -0.031 | 2.23 | -0.038 | -1.80 | -9.0, 2.25 |
| logq3 (catchability 2005-20) | -0.430 | 0.20 | -0.430 | 0.20 | -0.426 | 0.20 | -0.392 | -0.22 | -9.0, 2.25 |
| log_mean_rec (mean rec.) | 0.722 | 0.05 | 0.722 | 0.05 | 0.722 | 0.06 | 0.706 | 0.06 | 0.01, 5.0 |
| log_mean_Fpot (Pot fishery F) | -0.726 | 0.08 | -0.726 | 0.08 | -0.724 | 0.08 | -0.709 | -0.09 | -15.0, -0.01 |
| log_mean_Fground (GF byc. F) | -8.383 | 0.09 | -8.383 | 0.09 | -8.382 | 0.09 | -8.362 | -0.09 | -15.0, -1.6 |
| $\sigma_{e}^{2}$ (observer CPUE additional var) | 0.018 | 0.34 | 0.018 | 0.34 | 0.018 | 0.34 | 0.022 | 0.36 | 0.0, 0.15 |
| $\sigma_{e}^{2}$ (fishery CPUE additional var) | 0.025 | 0.60 | 0.025 | 0.60 | 0.025 | 0.60 | 0.025 | 0.57 | 0.0,1.0 |
| 2020 MMB | 6,438 | 0.15 | 6,436 | 0.15 | 6,469 | 0.17 | 5,413 | 0.16 |  |

Table 19. Annual abundance estimates of model recruits (millions of crab), legal male biomass ( t ) with coefficient of variations (CV), and mature male biomass ( t ) with CV for model 19.1 for golden king crab in the WAG. Legal male biomass was estimated on July 1 (start of fishing year) of fishing year $y$. Mature male biomass for fishing year y was estimated on February 15 of year $\mathrm{y}+1$, after the year y fishery total catch removal. Recruits estimates for 1961 to 2021 are restricted to 19852021. Equilibrium $\mathrm{MMB}_{\text {eq }}$ and $\mathrm{MMB}_{35 \%}$ are also listed.

| Year | Recruits to the Model ( $\geq$ 101 mm CL) | Mature Male Biomass $(\geq 111 \mathrm{~mm} \mathrm{CL})$ | CV | Legal Size Male <br> Biomass ( $\geq 136$ mm CL) | CV |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{aligned} & \mathrm{MMB}_{\mathrm{eq}}=18,279 \\ & M M B_{35 \%}=6,696 \end{aligned}$ |  |  |  |
| 1985 | 4.00 | 10,590 | 0.05 | 9,037 | 0.10 |
| 1986 | 3.46 | 8,141 | 0.05 | 8,478 | 0.08 |
| 1987 | 2.65 | 7,472 | 0.04 | 5,995 | 0.06 |
| 1988 | 1.87 | 6,390 | 0.04 | 5,602 | 0.05 |
| 1989 | 2.53 | 4,397 | 0.04 | 4,943 | 0.04 |
| 1990 | 1.93 | 4,102 | 0.05 | 3,109 | 0.06 |
| 1991 | 1.62 | 3,871 | 0.05 | 2,846 | 0.05 |
| 1992 | 2.02 | 4,008 | 0.04 | 2,793 | 0.05 |
| 1993 | 1.59 | 4,578 | 0.03 | 2,947 | 0.05 |
| 1994 | 1.95 | 3,876 | 0.03 | 3,525 | 0.03 |
| 1995 | 1.89 | 3,863 | 0.04 | 2,843 | 0.03 |
| 1996 | 1.73 | 3,871 | 0.04 | 2,778 | 0.04 |
| 1997 | 1.85 | 3,940 | 0.04 | 2,815 | 0.04 |
| 1998 | 1.89 | 4,256 | 0.04 | 2,896 | 0.04 |
| 1999 | 2.24 | 4,282 | 0.04 | 3,177 | 0.03 |
| 2000 | 2.51 | 4,433 | 0.04 | 3,111 | 0.04 |
| 2001 | 2.56 | 4,871 | 0.05 | 3,112 | 0.04 |
| 2002 | 2.51 | 5,432 | 0.05 | 3,438 | 0.05 |
| 2003 | 1.76 | 5,769 | 0.05 | 3,962 | 0.05 |
| 2004 | 2.28 | 5,886 | 0.06 | 4,464 | 0.05 |
| 2005 | 2.34 | 6,183 | 0.06 | 4,666 | 0.06 |
| 2006 | 2.47 | 6,720 | 0.05 | 4,847 | 0.06 |
| 2007 | 1.74 | 6,916 | 0.05 | 5,299 | 0.06 |
| 2008 | 1.52 | 6,746 | 0.05 | 5,608 | 0.06 |
| 2009 | 1.97 | 6,393 | 0.05 | 5,677 | 0.05 |
| 2010 | 1.67 | 6,150 | 0.05 | 5,336 | 0.05 |
| 2011 | 1.20 | 5,691 | 0.05 | 5,061 | 0.05 |
| 2012 | 2.01 | 5,150 | 0.05 | 4,757 | 0.05 |
| 2013 | 2.46 | 5,029 | 0.06 | 4,182 | 0.05 |
| 2014 | 1.86 | 5,249 | 0.06 | 3,778 | 0.06 |
| 2015 | 1.96 | 5,400 | 0.06 | 3,954 | 0.06 |
| 2016 | 1.73 | 5,627 | 0.07 | 4,203 | 0.07 |
| 2017 | 1.93 | 5,791 | 0.07 | 4,499 | 0.07 |
| 2018 | 2.27 | 5,947 | 0.09 | 4,682 | 0.07 |
| 2019 | 2.16 | 6,077 | 0.12 | 4,722 | 0.09 |
| 2020 | 2.23 | 6,438 | 0.15 | 4,767 | 0.11 |
| 2021 | 2.06 |  |  |  |  |

Table 20. Annual abundance estimates of model recruits (millions of crab), legal male biomass ( t ) with coefficient of variations (CV), and mature male biomass ( $t$ ) with CV for model 21.1a for golden king crab in the WAG. Legal male biomass was estimated on July 1 (start of fishing year) of fishing year $y$. Mature male biomass for fishing year y was estimated on February 15 of year $\mathrm{y}+1$, after the year y fishery total catch removal. Recruits estimates for 1961 to 2021 are restricted to 1985-2021. Equilibrium $\mathrm{MMB}_{\mathrm{eq}}$ and $\mathrm{MMB}_{35 \%}$ are also listed.

| Year | Recruits to the Model ( $\geq$ 101 mm CL) | Mature Male Biomass $(\geq 111 \mathrm{~mm} \mathrm{CL})$ | CV | Legal Size Male <br> Biomass ( $\geq 136$ mm CL) | CV |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{gathered} \mathrm{MMB}_{\mathrm{eq}}=18,243 \\ M M B_{35 \%}=5,290 \end{gathered}$ |  |  |  |
| 1985 | 4.00 | 10,589 | 0.05 | 9,037 | 0.10 |
| 1986 | 3.46 | 8,140 | 0.05 | 8,478 | 0.08 |
| 1987 | 2.65 | 7,472 | 0.04 | 5,995 | 0.06 |
| 1988 | 1.87 | 6,390 | 0.04 | 5,602 | 0.05 |
| 1989 | 2.53 | 4,397 | 0.04 | 4,943 | 0.04 |
| 1990 | 1.93 | 4,102 | 0.05 | 3,109 | 0.06 |
| 1991 | 1.62 | 3,871 | 0.05 | 2,846 | 0.05 |
| 1992 | 2.02 | 4,008 | 0.04 | 2,793 | 0.05 |
| 1993 | 1.59 | 4,578 | 0.03 | 2,947 | 0.05 |
| 1994 | 1.95 | 3,876 | 0.03 | 3,525 | 0.03 |
| 1995 | 1.89 | 3,863 | 0.04 | 2,843 | 0.03 |
| 1996 | 1.73 | 3,871 | 0.04 | 2,778 | 0.04 |
| 1997 | 1.85 | 3,940 | 0.04 | 2,815 | 0.04 |
| 1998 | 1.89 | 4,256 | 0.04 | 2,896 | 0.04 |
| 1999 | 2.24 | 4,282 | 0.04 | 3,177 | 0.03 |
| 2000 | 2.51 | 4,433 | 0.04 | 3,111 | 0.04 |
| 2001 | 2.56 | 4,871 | 0.05 | 3,112 | 0.04 |
| 2002 | 2.51 | 5,432 | 0.05 | 3,438 | 0.05 |
| 2003 | 1.76 | 5,769 | 0.05 | 3,962 | 0.05 |
| 2004 | 2.28 | 5,886 | 0.06 | 4,464 | 0.05 |
| 2005 | 2.34 | 6,183 | 0.06 | 4,666 | 0.06 |
| 2006 | 2.47 | 6,720 | 0.05 | 4,846 | 0.06 |
| 2007 | 1.74 | 6,916 | 0.05 | 5,299 | 0.06 |
| 2008 | 1.52 | 6,746 | 0.05 | 5,608 | 0.06 |
| 2009 | 1.97 | 6,393 | 0.05 | 5,676 | 0.05 |
| 2010 | 1.67 | 6,150 | 0.05 | 5,336 | 0.05 |
| 2011 | 1.20 | 5,691 | 0.05 | 5,061 | 0.05 |
| 2012 | 2.01 | 5,150 | 0.05 | 4,757 | 0.05 |
| 2013 | 2.46 | 5,029 | 0.06 | 4,181 | 0.05 |
| 2014 | 1.86 | 5,249 | 0.06 | 3,778 | 0.06 |
| 2015 | 1.96 | 5,399 | 0.06 | 3,953 | 0.06 |
| 2016 | 1.73 | 5,626 | 0.07 | 4,203 | 0.07 |
| 2017 | 1.93 | 5,790 | 0.07 | 4,499 | 0.07 |
| 2018 | 2.27 | 5,945 | 0.09 | 4,682 | 0.07 |
| 2019 | 2.16 | 6,076 | 0.12 | 4,721 | 0.09 |
| 2020 | 2.23 | 6,436 | 0.15 | 4,765 | 0.11 |
| 2021 | 2.06 |  |  |  |  |

Table 21. Annual abundance estimates of model recruits (millions of crab), legal male biomass ( t ) with coefficient of variations (CV), and mature male biomass ( $t$ ) with CV for model 21.1 b for golden king crab in the WAG. Legal male biomass was estimated on July 1 (start of fishing year) of fishing year y . Mature male biomass for fishing year y was estimated on February 15 of year $y+1$, after the year y fishery total catch removal. Recruits estimates for 1961 to 2021 are restricted to 1985-2021. Equilibrium $\mathrm{MMB}_{\mathrm{eq}}$ and $\mathrm{MMB}_{35 \%}$ are also listed.

| Year | Recruits to the Model ( $\geq$ 101 mm CL) | Mature Male Biomass $(\geq 111 \mathrm{~mm} \mathrm{CL})$ | CV | Legal Size Male <br> Biomass ( $\geq 136$ mm CL) | CV |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{aligned} & \mathrm{MMB}_{\mathrm{eq}}=18,232 \\ & M M B_{35 \%}=5,279 \end{aligned}$ |  |  |  |
| 1985 | 3.99 | 10,594 | 0.05 | 9,040 | 0.10 |
| 1986 | 3.46 | 8,143 | 0.05 | 8,482 | 0.08 |
| 1987 | 2.65 | 7,473 | 0.04 | 5,997 | 0.06 |
| 1988 | 1.86 | 6,390 | 0.04 | 5,602 | 0.05 |
| 1989 | 2.53 | 4,398 | 0.04 | 4,942 | 0.04 |
| 1990 | 1.93 | 4,103 | 0.05 | 3,108 | 0.06 |
| 1991 | 1.62 | 3,872 | 0.05 | 2,845 | 0.05 |
| 1992 | 2.02 | 4,009 | 0.04 | 2,793 | 0.05 |
| 1993 | 1.59 | 4,578 | 0.03 | 2,947 | 0.05 |
| 1994 | 1.95 | 3,875 | 0.03 | 3,524 | 0.03 |
| 1995 | 1.89 | 3,862 | 0.03 | 2,841 | 0.03 |
| 1996 | 1.73 | 3,870 | 0.04 | 2,776 | 0.04 |
| 1997 | 1.85 | 3,939 | 0.04 | 2,813 | 0.04 |
| 1998 | 1.89 | 4,255 | 0.03 | 2,894 | 0.04 |
| 1999 | 2.24 | 4,281 | 0.04 | 3,175 | 0.03 |
| 2000 | 2.51 | 4,433 | 0.04 | 3,109 | 0.04 |
| 2001 | 2.56 | 4,870 | 0.05 | 3,110 | 0.04 |
| 2002 | 2.51 | 5,431 | 0.05 | 3,436 | 0.05 |
| 2003 | 1.77 | 5,770 | 0.05 | 3,960 | 0.05 |
| 2004 | 2.29 | 5,890 | 0.06 | 4,462 | 0.05 |
| 2005 | 2.32 | 6,185 | 0.06 | 4,666 | 0.06 |
| 2006 | 2.47 | 6,719 | 0.05 | 4,849 | 0.06 |
| 2007 | 1.74 | 6,918 | 0.05 | 5,298 | 0.06 |
| 2008 | 1.52 | 6,750 | 0.05 | 5,607 | 0.06 |
| 2009 | 1.98 | 6,397 | 0.05 | 5,678 | 0.05 |
| 2010 | 1.67 | 6,153 | 0.05 | 5,339 | 0.05 |
| 2011 | 1.20 | 5,694 | 0.05 | 5,063 | 0.05 |
| 2012 | 2.01 | 5,154 | 0.05 | 4,760 | 0.05 |
| 2013 | 2.44 | 5,028 | 0.06 | 4,184 | 0.05 |
| 2014 | 1.82 | 5,230 | 0.06 | 3,780 | 0.06 |
| 2015 | 1.97 | 5,366 | 0.06 | 3,947 | 0.06 |
| 2016 | 1.71 | 5,591 | 0.07 | 4,177 | 0.07 |
| 2017 | 1.94 | 5,752 | 0.07 | 4,466 | 0.07 |
| 2018 | 2.27 | 5,915 | 0.09 | 4,644 | 0.07 |
| 2019 | 2.23 | 6,071 | 0.13 | 4,686 | 0.09 |
| 2020 | 2.24 | 6,469 | 0.17 | 4,740 | 0.12 |
| 2021 | 2.06 |  |  |  |  |

Table 22. Annual abundance estimates of model recruits (millions of crab), legal male biomass ( t ) with coefficient of variations (CV), and mature male biomass ( $t$ ) with CV for model 21.1c for golden king crab in the WAG. Legal male biomass was estimated on July 1 (start of fishing year) of fishing year $y$. Mature male biomass for fishing year y was estimated on February 15 of year $\mathrm{y}+1$, after the year y fishery total catch removal. Recruits estimates for 1961 to 2021 are restricted to 1985-2021. Equilibrium $\mathrm{MMB}_{\mathrm{eq}}$ and $\mathrm{MMB}_{35 \%}$ are also listed.

| Year | Recruits to the Model ( $\geq$ 101 mm CL) | Mature Male Biomass $(\geq 111 \mathrm{~mm} \mathrm{CL})$ | CV | $\begin{gathered} \text { Legal Size Male } \\ \text { Biomass }(\geq 136 \\ \text { mm CL) } \end{gathered}$ | CV |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{aligned} & \mathrm{MMB}_{\mathrm{eq}}=18,014 \\ & M M B_{35 \%}=5,226 \end{aligned}$ |  |  |  |
| 1985 | 3.99 | 10,599 | 0.05 | 9,028 | 0.10 |
| 1986 | 3.46 | 8,152 | 0.05 | 8,481 | 0.08 |
| 1987 | 2.65 | 7,486 | 0.04 | 6,001 | 0.06 |
| 1988 | 1.86 | 6,401 | 0.04 | 5,609 | 0.05 |
| 1989 | 2.53 | 4,404 | 0.04 | 4,951 | 0.04 |
| 1990 | 1.92 | 4,106 | 0.04 | 3,114 | 0.05 |
| 1991 | 1.62 | 3,872 | 0.05 | 2,848 | 0.05 |
| 1992 | 2.00 | 4,002 | 0.04 | 2,792 | 0.05 |
| 1993 | 1.57 | 4,559 | 0.03 | 2,943 | 0.05 |
| 1994 | 1.97 | 3,852 | 0.03 | 3,512 | 0.03 |
| 1995 | 1.89 | 3,847 | 0.03 | 2,819 | 0.03 |
| 1996 | 1.72 | 3,858 | 0.04 | 2,757 | 0.03 |
| 1997 | 1.87 | 3,932 | 0.04 | 2,799 | 0.04 |
| 1998 | 1.90 | 4,259 | 0.03 | 2,881 | 0.04 |
| 1999 | 2.25 | 4,294 | 0.04 | 3,171 | 0.03 |
| 2000 | 2.51 | 4,451 | 0.04 | 3,115 | 0.04 |
| 2001 | 2.54 | 4,887 | 0.05 | 3,122 | 0.04 |
| 2002 | 2.49 | 5,436 | 0.05 | 3,450 | 0.05 |
| 2003 | 1.75 | 5,759 | 0.05 | 3,968 | 0.05 |
| 2004 | 2.31 | 5,875 | 0.06 | 4,458 | 0.05 |
| 2005 | 2.39 | 6,198 | 0.06 | 4,650 | 0.06 |
| 2006 | 2.52 | 6,775 | 0.05 | 4,839 | 0.06 |
| 2007 | 1.72 | 6,994 | 0.05 | 5,321 | 0.06 |
| 2008 | 1.46 | 6,804 | 0.05 | 5,667 | 0.05 |
| 2009 | 1.92 | 6,402 | 0.05 | 5,745 | 0.05 |
| 2010 | 1.64 | 6,115 | 0.05 | 5,372 | 0.05 |
| 2011 | 1.19 | 5,627 | 0.05 | 5,049 | 0.05 |
| 2012 | 1.99 | 5,069 | 0.05 | 4,707 | 0.05 |
| 2013 | 2.41 | 4,927 | 0.06 | 4,110 | 0.05 |
| 2014 | 1.80 | 5,113 | 0.06 | 3,692 | 0.06 |
| 2015 | 1.84 | 5,207 | 0.06 | 3,843 | 0.06 |
| 2016 | 1.55 | 5,339 | 0.06 | 4,056 | 0.06 |
| 2017 | 1.66 | 5,357 | 0.07 | 4,287 | 0.07 |
| 2018 | 1.88 | 5,294 | 0.09 | 4,366 | 0.07 |
| 2019 | 1.90 | 5,184 | 0.13 | 4,244 | 0.09 |
| 2020 | 2.16 | 5,413 | 0.16 | 4,063 | 0.12 |
| 2021 | 2.03 |  |  |  |  |

Table 23. Negative log-likelihood values of the fits for models 19.1 (last year's accepted model with additional 2020/21 data), 20.1a, 21.1 b, and 21.1 c for golden king crab in the WAG. Likelihood components with zero entry in the entire rows are omitted. RetdcatchB $=$ retained catch biomass

| Likelihood Component | Model 19.1 | Model 21.1a | Model 21.1b | Model 21.1c | 21.1a-19.1 | 21.1b-19.1 | 21.1c-19.1 |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  |  |  |
| Number of free parameters | 152 | 152 | 154 | 152 |  |  |  |
| Retlencomp | -1300.4500 | -1300.4200 | -1301.3600 | -1301.5800 | 0.03 | -0.91 | -1.13 |
| Totallencomp | -1606.4100 | -1606.4300 | -1607.5300 | -1612.6400 | -0.02 | -1.12 | -6.23 |
| Observer cpue | -14.6181 | -14.6180 | -14.6192 | -10.6439 | 0.00010 | -0.0011 | 3.9742 |
| RetdcatchB | 5.0494 | 5.0493 | 5.0588 | 5.2239 | -0.0001 | 0.0094 | 0.1745 |
| TotalcatchB | 45.3462 | 45.3465 | 45.4377 | 45.4480 | 0.0003 | 0.0915 | 0.1018 |
| GdiscdcatchB | 0.0014 | 0.0014 | 0.0014 | 0.0013 | 0 | 0 | -0.0001 |
| Rec_dev | 4.4273 | 4.4274 | 4.4465 | 4.8170 | 0.00010 | 0.0192 | 0.3897 |
| Pot $\bar{F}$ dev | 0.0266 | 0.0266 | 0.0266 | 0.0265 | 0 | 0 | -0.00010 |
| Gbyc_F_dev | 0.0435 | 0.0435 | 0.0435 | 0.0436 | 0 | 0 | 0.0001 |
| Tag | 2694.9900 | 2694.9900 | 2695.0400 | 2695.1300 | 0 | 0.05 | 0.14 |
| Fishery cpue | -5.4124 | -5.4131 | -5.3625 | -5.5391 | -0.0007 | 0.0499 | -0.1267 |
| RetcatchN | 0.0024 | 0.0024 | 0.0024 | 0.0019 | 0 | 0 | -0.0005 |
| Total | -177.0030 | -177.0010 | -178.8120 | -179.7090 | 0.002 | -1.809 | -2.706 |

Figures


Figure 1. Aleutian Islands, Area O, red and golden king crab management area (from Nichols et al. 2021).


Figure 2. Adak (Area R) and Dutch Harbor (Area O) king crab registration area and districts, 1984/85-1995/96 seasons (Leon et al. 2017).


Figure 3. Percent of total 1981/82-1995/96 golden king crab retained catch weight (harvest) from one-degree longitude intervals in the Aleutian Islands, with dotted line denoting the border at $171^{\circ}$ W longitude used during the 1984/85-1995/96 seasons to divide fishery management between the Dutch Harbor Area (east of $171^{\circ} \mathrm{W}$ longitude) and the Adak Area (west of $171^{\circ} \mathrm{W}$ longitude) and solid line denoting the border at $174^{\circ} \mathrm{W}$ longitude used since the 1996/97 season to manage crab east and west of $174^{\circ} \mathrm{W}$ longitude (adapted from Figure 4-2 in Morrison et al. 1998).


Figure 4. Retained catch ( t ) of golden king crab within one-degree longitude intervals in the Aleutian Islands during the 2000/01 through 2020/21 commercial fishery seasons; solid line denotes the border at $174^{\circ} \mathrm{W}$ longitude that has been used since the 1996/97 season to manage Aleutian Island golden king crab as separate stocks east and west of $174^{\circ} \mathrm{W}$ longitude and dashed line denotes the border at $171^{\circ} \mathrm{W}$ longitude used during the 1984/85-1995/96 seasons to divide fishery management between the Dutch Harbor Area (east of $171^{\circ} \mathrm{W}$ longitude) and the Adak Area (west of $171^{\circ} \mathrm{W}$ longitude).


Figure 5. Average golden king crab CPUE ( $\mathrm{kg} / \mathrm{nm} 2$ ) for tows, number of tows, and average depth of tows from one-degree longitude intervals during the 2002, 2004, 2006, 2010, and 2012 NMFS Aleutian Islands bottom trawl surveys; preliminary summary of data obtained on 1 April 2013 from http://www.afsc.noaa.gov/RACE/groundfish/survey_data/default.htm.


Figure 6. Historical commercial harvest (from fish tickets; metric tons), total allowable catch (TAC), and catch-per-unit effort (CPUE, number of crab per pot lift) of golden king crab in the EAG, 1985/86-2020/21 fisheries (note: 1985 refers to the 1985/86 fishing year).


Figure 7. Historical commercial harvest (from fish tickets; metric tons) ), total allowable catch (TAC), and catch-per-unit effort (CPUE, number of crab per pot lift) of golden king crab in the WAG, 1985/86-2019/20 fisheries (note: 1985 refers to the 1985/86 fishing year).


Figure 8. Catch distribution by statistical area.in 2020/21.

## Model 21.1a: Observed vs predicted Retained Catch Size Composition



Figure 9a. Predicted vs. observed retained catch relative length frequency distributions under model 21.1a for golden king crab in the EAG, 1985/86 to 2020/21. Each year has a pair of plots with the front plot for observed and the back plot for predicted proportions.

## Model 21.1b: Observed vs predicted Retained Catch Size Composition



Figure 9b. Predicted vs. observed retained catch relative length frequency distributions under model 21.1b for golden king crab in the EAG, 1985/86 to 2020/21. Each year has a pair of plots with the front plot for observed and the back plot for predicted proportions.

Model 21.1c: Observed vs predicted Retained Catch Size Composition


Figure 9c. Predicted vs. observed retained catch relative length frequency distributions under model 21.1c for golden king crab in the EAG, 1985/86 to 2020/21. Each year has a pair of plots with the front plot for observed and the back plot for predicted proportions.

 Carapace Size (mm)
Figure 10a. Predicted vs. observed total catch relative length frequency distributions under model 21.1a for golden king crab in the EAG, 1990/91 to 2020/21. Each year has a pair of plots with the front plot for observed and the back plot for predicted proportions.

Model 21.1b: Observed vs predicted Total Catch Size Composition


Figure 10b. Predicted vs. observed total catch relative length frequency distributions under model 21.1b for golden king crab in the EAG, 1990/91 to 2020/21. Each year has a pair of plots with the front plot for observed and the back plot for predicted proportions.


Figure 10 c . Predicted vs. observed total catch relative length frequency distributions under model 21.1c for golden king crab in the EAG, 1990/91 to 2020/21. Each year has a pair of plots with the front plot for observed and the back plot for predicted proportions.

Model 21.1a: Observed vs predicted Groundfish Bycatch Size Composition


Figure 11. Predicted vs. observed groundfish discarded bycatch relative length frequency distributions under model 21.1a for golden king crab in the EAG, 1989/90 to 2020/21. Each year has a pair of plots with the front plot for observed and the back plot for predicted proportions.


Figure 12a. Estimated total (black solid line) and retained selectivity (red dotted line) for pre- and post- rationalization periods under models 19.1, 21.1a, 21.1b, and 21.1c fits to golden king crab data in the EAG. The post-rationalization total selectivity for 21.1 b corresponds to first part (20052015) of the post-rationalization period.


Figure 12b. Estimated total (black solid line) and retained selectivity (red dotted line) for second part (2016-2020) of the post- rationalization period under model 21.1b fit to golden king crab data in the EAG.


Figure 13. Observed (open circles) vs. predicted (solid line) tag recaptures by size bin for years 1 to 6 post tagging under model 21.1a for EAG golden king crab.


Figure 14. Estimated number of male recruits (crab size $\geq 101 \mathrm{~mm} \mathrm{CL}$ ) to the assessment model under models 19.1, 21.1a, 21.1b, and 21.1c fits to EAG golden king crab data, 1961-2021.


Figure 15. Recruit size distribution to the assessment model under models 19.1, 21.1a, 21.1b, and 21.1c fits to EAG golden king crab data.


Figure 16. Estimated molt probability vs. carapace length of golden king crab under models 19.1, 21.1a, 21.1b, and 21.1c fits to EAG golden king crab data.


Figure 17. Observed (open circle) vs. predicted (solid line) retained catch (top left), total catch (top right in), and groundfish bycatch (bottom left) of golden king crab under models 19.1, 21.1a, 21.1b, and 21.1c fits in EAG, 1981/82-2020/21.


Figure 18. Observed (open circle) vs. predicted (solid line) retained catch of golden king crab for models 19.1, 21.1a, 21.1b, and 21.1c fits in the EAG, 1981/82-1984/85. Note: Input retained catches to the model during pre-1985 fishery period was in number of crabs.


Figure 19. Bubble plot of standardized residuals of retained catch length composition for model 21.1a fit for EAG golden king crab, 1985/86-2020/21. Green circles are the positive and pink circles are the negative standardized residuals. The area of the circle is the relative magnitude of the residual.

EAG 21.1c Retained Catch Size Composition Standardized Residuals


Figure 20. Bubble plot of standardized residuals of retained catch length composition for model 21.1c fit for EAG golden king crab, 1985/86-2020/21. Green circles are the positive and pink circles are the negative standardized residuals. The area of the circle is the relative magnitude of the residual.


Figure 21. Bubble plot of standardized residuals of total catch length composition for model 21.1a fit for EAG golden king crab, 1990/91-2020/21. Green circles are the positive and pink circles are the negative standardized residuals. The area of the circle is the relative magnitude of the residual.

EAG 21.1c Retained Catch Size Composition Standardized Residuals



Figure 22. Bubble plot of standardized residuals of total catch length composition for model 21.1c fit for EAG golden king crab, 1990/91-2020/21. Green circles are the positive and pink circles are the negative standardized residuals. The area of the circle is the relative magnitude of the residual.


Figure 23. Retrospective fits of MMB by the model following removal of terminal year data under models 21.1a, 21.1b, and 21.1c for golden king crab in the EAG (left) and WAG (right), 1960/61-2020/21.


Figure 24. Comparison of input CPUE indices (open circles with $+/-2$ SE for model 19.1) with predicted CPUE indices (colored solid lines) under 19.1, 21.1a, 21.1b, and 21.1c for EAG golden king crab data, 1985/86-2020/21. Model estimated additional standard error was added to each input standard error.

## EAG



WAG


Figure 25. Trends in pot fishery full selection total fishing mortality of golden king crab for models 19.1, 21.1a, 21.1b, and 21.1c fits in the EAG (left) and WAG (right) data, 1981/82-2020/21.


Figure 26. Trends in golden king crab mature male biomass for models 19.1, 21.1a, 21.1b, and 21.1c fits to EAG (left) and WAG (right) data, 1961-2021. Model 21.1a estimate has two standard error confidence limits. Year 2021 refers to 2020/21 fishing season.


Figure 27a. Predicted vs. observed retained catch relative length frequency distributions under model 21.1a for golden king crab in the WAG, 1985/86 to 2020/21. Each year has a pair of plots with the front plot for observed and the back plot for predicted proportions.


Figure 27b. Predicted vs. observed retained catch relative length frequency distributions under model 21.1b for golden king crab in the WAG, 1985/86 to 2020/21. Each year has a pair of plots with the front plot for observed and the back plot for predicted proportions.


Figure 27c. Predicted vs. observed retained catch relative length frequency distributions under model 21.1c for golden king crab in the WAG, 1985/86 to 2020/21. Each year has a pair of plots with the front plot for observed and the back plot for predicted proportions.

Model 21.1a: Observed vs predicted Total Catch Size Composition


Figure 28a. Predicted vs. observed total catch relative length frequency distributions under model 21.1a for golden king crab in the WAG, 1990/91 to 2020/21. Each year has a pair of plots with the front plot for observed and the back plot for predicted proportions.

Model 21.1b: Observed vs predicted Total Catch Size Composition


Figure 28 b . Predicted vs. observed total catch relative length frequency distributions under model 21.1b for golden king crab in the WAG, 1990/91 to 2020/21. Each year has a pair of plots with the front plot for observed and the back plot for predicted proportions.

Model 21.1c: Observed vs predicted Total Catch Size Composition


Figure 28c. Predicted vs. observed total catch relative length frequency distributions under model 21.1c for golden king crab in the WAG, 1990/91 to 2020/21. Each year has a pair of plots with the front plot for observed and the back plot for predicted proportions.

Model 21.1a: Observed vs predicted Groundfish Bycatch Size Composition


Figure 29. Predicted vs. observed groundfish discarded bycatch relative length frequency distributions under model 21.1a for golden king crab in the WAG, 1989/90 to 2020/21. Each year has a pair of plots with the front plot for observed and the back plot for predicted proportions.


Figure 30a. Estimated total (black solid line) and retained selectivity (red dotted line) for pre- and post- rationalization periods under models 19.1, 21.1a, 21.1b, and 21.1c fits to golden king crab data in the WAG. The post-rationalization total selectivity for 21.1b corresponds to first part (2005-2015) of the post-rationalization period.


Figure 30b. Estimated total (black solid line) and retained selectivity (red dotted line) for second part (2016-2020) of the post- rationalization period under model 21.1b fit to golden king crab data in the WAG.


Figure 31. Observed (open circles) vs. predicted (solid line) tag recaptures by size bin for years 1 to 6 post tagging under model 21.1a fit to WAG golden king crab data.


Figure 32. Estimated number of male recruits (crab size $\geq 101 \mathrm{~mm} \mathrm{CL}$ ) to the assessment model under models 19.1, 21.1a, 21.1b, and 21.1c fits to WAG golden king crab data, 1961-2021.


Figure 33. Recruit size distribution to the assessment model under models 19.1, 21.1a, 21.1b, and 21.1c fits to WAG golden king crab data.


Figure 34. Estimated molt probability vs. carapace length of golden king crab for models 19.1, 21.1a, 21.1b, and 21.1c fits to WAG golden king crab data.


Figure 35 . Observed (open circle) vs. predicted (solid line) retained catch (top left), total catch (top right), and groundfish bycatch (bottom left) of golden king crab for models 19.1, 21.1a, 21.1b, and 21.1c fits to WAG data, 1981/82-2020/21.


Figure 36. Observed (open circle) vs. predicted (solid line) retained catch of golden king crab for models 19.1, 21.1a, 21.1b, and 21.1c fits to WAG data, 1981/82-1984/85. Note: Input retained catches to the model during pre-1985 fishery period was in number of crabs.


Figure 37. Bubble plot of standardized residuals of retained catch length composition for model 21.1a fit to WAG golden king crab data, 1985/86-2020/21. Green circles are the positive and pink circles are the negative standardized residuals. The area of the circle is the relative magnitude of the residual.


Figure 38. Bubble plot of standardized residuals of retained catch length composition for model 21.1c fit to WAG golden king crab data, 1985/86-2020/21. Green circles are the positive and pink circles are the negative standardized residuals. The area of the circle is the relative magnitude of the residual.


Figure 39. Bubble plot of standardized residuals of total catch length composition for model 21.1a fit to WAG golden king crab data, 1990/91-2020/21. Green circles are the positive and pink circles are the negative standardized residuals. The area of the circle is the relative magnitude of the residual.


Figure 40. Bubble plot of standardized residuals of total catch length composition for model 21.1c fit to WAG golden king crab data, 1990/91-2020/21. Green circles are the positive and pink circles are the negative standardized residuals. The area of the circle is the relative magnitude of the residual.


Figure 41. Comparison of input CPUE indices (open circles with $+/-2$ SE for model 19.1) with model predicted CPUE indices (colored solid lines) under models 19.1, 21.1a, 21.1b, and 21.1c fits to WAG golden king crab data, 1985/86-2020/21. Model estimated additional standard error was added to each input standard error.


Figure 42. Relationships between full fishing mortalities for the directed pot fishery and mature male biomass during 1981/82-2020/21 under models, 21.1a, 21.1b, and 21.1c fits to EAG and WAG data. $F$ in $2020 / 21$ (red) and 1981/82 (black) are shown in the plots.

## Appendix A: Integrated model

Aleutian Islands Golden King Crab (Lithodes aequispinus) Stock Assessment Model Development- east of $174^{\circ} \mathrm{W}$ (EAG) and west of $174^{\circ} \mathrm{W}$ (WAG) Aleutian Island stocks

## Basic population dynamics

The annual [male] abundances by size are modeled using the equation:
$N_{t+1, j}=\sum_{i=1}^{j}\left[N_{t, i} e^{-M}-\left(\hat{C}_{t, i}+\widehat{D}_{t, i}+\widehat{\operatorname{Tr}}_{t, i}\right) e^{\left(y_{t}-1\right) M}\right] X_{i, j}+R_{t+1, j}$
where $N_{t, i}$ is the number of [male] crab in length class $i$ on 1 July (start of fishing year) of year $t$; $\hat{C}_{t, i}, \widehat{D}_{t, i}$, and $\widehat{T} r_{t, i}$ are respectively the predicted fishery retained, pot fishery discard dead, and groundfish fishery discard dead catches in length class $i$ during year $t ; \widehat{D}_{t, i}$ is estimated from the intermediate total ( $\hat{T}_{t, i \text { temp }}$ ) catch and the retained ( $\hat{C}_{t, i}$ ) catch by Equation A.2c. $X_{i, j}$ is the probability of length-class $i$ growing into length-class $j$ during the year; $y_{t}$ is elapsed time period from 1 July to the midpoint of fishing period in year $t ; M$ is instantaneous rate of natural mortality; and $R_{t+1, j}$ recruitment to length class $j$ in year $t+1$.

The catches are predicted using the equations
$\widehat{T}_{t, j, \text { temp }}=\frac{F_{t} s_{t, j}^{T}}{Z_{t, j}} N_{t, j} e^{-y_{t} M}\left(1-e^{-Z_{t, j}}\right)$
$\hat{C}_{t, j}=\frac{F_{t} S_{t, j}^{T} s_{t, j}^{r}}{z_{t, j}} N_{t, j} e^{-y_{t} M}\left(1-e^{-Z_{t, j}}\right)$
$\widehat{D}_{t, j}=0.2\left(\widehat{T}_{t, j, t e m p}-\hat{C}_{t, j}\right)$
$\widehat{T r}_{t, j}=0.65 \frac{F_{t}^{T r} s_{j}^{T r}}{Z_{t, j}} N_{t, j} e^{-y_{t} M}\left(1-e^{-Z_{t, j}}\right)$
$\widehat{T}_{t, j}=\hat{C}_{t, j}+\widehat{D}_{t, j}$
where $Z_{t, j}$ is total fishery-related mortality on animals in length-class $j$ during year $t$ :

$$
\begin{equation*}
Z_{t, j}=F_{t} s_{t, j}^{T} s_{t, j}^{r}+0.2 F_{t} s_{t, j}^{T}\left(1-s_{t, j}^{r}\right)+0.65 F_{t}^{T r} s_{j}^{T r} \tag{A.3}
\end{equation*}
$$

$F_{t}$ is the full selection fishing mortality in the pot fishery, $F_{t}^{T r}$ is the full selection fishing mortality in the trawl fishery, $s_{t, j}^{T}$ is the total selectivity for animals in length-class $j$ by the pot fishery during year $t, s_{j}^{T r}$ is the selectivity for animals in length-class $j$ by the trawl fishery, $s_{t, j}^{r}$ is the probability of retention for animals in length-class $j$ by the pot fishery during year $t$. Pot bycatch mortality of
0.2 and groundfish bycatch mortality of 0.65 (average of trawl [0.8] and groundfish pot [0.5] mortality) were assumed.

## Initial abundance

The initial conditions are computed as the equilibrium initial condition using the following relations:

The equilibrium stock abundance is
$N=X . S . N+R$
The equilibrium abundance in $1960, N_{1960}$, is
$\underline{N}_{1960}=(I-X S)^{-1} \underline{R}$
where $X$ is the growth matrix, $S$ is a matrix with diagonal elements given by $e^{-M}, I$ is the identity matrix, and $\underline{R}$ is the product of average recruitment and relative proportion of total recruitment to each size-class.

We used the mean number of recruits from 1987 to 2012 in equation (A.5) to obtain the equilibrium solution under only natural mortality in year 1960, and then projected the equilibrium abundance under natural mortality with recruitment estimated for each year after 1960 up to 1985 with removal of retained catches during 1981/82-1984/85.

## Growth Matrix

The growth matrix $X$ is modeled as follows:
$X_{i, j}= \begin{cases}0 & \text { if } j<i \\ P_{i, j}+\left(1-m_{i}\right) & \text { if } j=i \\ P_{i, j} & \text { if } j>i\end{cases}$
where:

$$
P_{i, j}=m_{i}\left\{\begin{array}{rr}
\int_{-\infty}^{j_{2}-L_{i}} N\left(x \mid \mu_{i}, \sigma^{2}\right) d x & \text { if } j=i \\
\int_{j_{1}-L_{i}}^{j_{2}-L_{i}} N\left(x \mid \mu_{i}, \sigma^{2}\right) d x & \text { if } i<j<n, \\
\int_{j_{1}-L_{i}}^{\infty} N\left(x \mid \mu_{i}, \sigma^{2}\right) d x & \text { if } i=n
\end{array}, \begin{array}{c} 
\\
N\left(x \mid \mu_{i}, \sigma^{2}\right)=\frac{1}{\sqrt{2 \pi \sigma^{2}}} e^{-\left(\frac{x-\mu_{i}}{\sqrt{2} \sigma}\right)^{2}}, \text { and }
\end{array}\right.
$$

$\mu_{i}$ is the mean growth increment for crab in size-class $i$ :
$\mu_{i}=\omega_{1}+\omega_{2} * \bar{L}_{i}$.
$\omega_{1}, \omega_{2}, \quad$ and $\sigma$ are estimable parameters, $\mathrm{j}_{1}$ and $j_{2}$ are the lower and upper limits of the receiving length-class $j$ (in mm CL), and $\bar{L}_{i}$ is the mid-point of the contributing length interval $i$. The quantity $m_{i}$ is the molt probability for size-class $i$ :
$m_{i}=\frac{1}{1+e^{c\left(\tau_{i}-d\right)}}$
where $\tau_{i}$ is the mid-length of the $i$-th length-class, $c$ and $d$ are parameters.

## Selectivity and retention

Selectivity and retention are both assumed to be logistic functions of length. Selectivity depends on the fishing period for the directed pot fishery:
$S_{i}=\frac{1}{1+e^{\left[-\ln \left(19 \frac{\tau_{i}-\theta_{50}}{\theta_{95}-\theta_{50}}\right]\right.}}$
where $\theta_{95}$ and $\theta_{50}$ are the parameters of the selectivity/ retention pattern (Mark Maunder, unpublished generic crab model). In our program, we re-parameterized the denominator ( $\theta_{95}-\theta_{50}$ ) to $\log (\operatorname{delta} \theta)$ so that the difference is always positive and transformed $\theta_{50}$ to $\log \left(\theta_{50}\right)$ to keep the estimate always positive.

## Recruitment

Recruitment to length-class i during year $t$ is modeled as $R_{t, i}=\bar{R} e^{\epsilon_{i}} \Omega_{i}$ where $\Omega_{i}$ is a normalized gamma function

$$
\begin{equation*}
\operatorname{gamma}\left(x \mid \alpha_{r}, \beta_{r}\right)=\frac{x^{\alpha_{r}-1} e^{\frac{x}{\beta_{r}}}}{\beta_{r}^{\alpha_{r}} \Gamma_{\left(\alpha_{r}\right)}} \tag{A.10}
\end{equation*}
$$

with $\alpha_{r}$ and $\beta_{r}$ (restricted to the first five length classes).

## Parameter estimation

Table A1 lists the parameters of the model indicating which are estimated and which are prespecified. The objective function includes contributions related to the fit of the model to the available data and penalties (priors on various parameters).

Tables A2 lists parameter values (with the corresponding coefficient of variations in parentheses) used to weight the components of the objective functions for EAG and WAG.

## Likelihood components

## Catches

The contribution of the catch data (retained, total, and groundfish discarded) to the objective function is given by:

$$
\begin{align*}
& L L_{r}^{\text {catch }}=\lambda_{r} \sum_{t}\left\{\ln \left(\sum_{j} \hat{C}_{t, j} w_{j}+c\right)-\ln \left(\sum_{j} C_{t, j} w_{j}+c\right)\right\}^{2}  \tag{A.11a}\\
& L L_{T}^{\text {catch }}=\lambda_{T} \sum_{t}\left\{\ln \left(\sum_{j} \widehat{T}_{t, j} w_{j}+c\right)-\ln \left(\sum_{j} T_{t, j} w_{j}+c\right)\right\}^{2}  \tag{A.11b}\\
& L L_{G D}^{\text {catch }}=\lambda_{G D} \sum_{t}\left\{\ln \left(\sum_{j} \widehat{T r}_{t, j} w_{j}+c\right)-\ln \left(\sum_{j} T r_{t, j} w_{j}+c\right)\right\}^{2} \tag{A.11c}
\end{align*}
$$

where $\lambda_{r}, \lambda_{T}$, and $\lambda_{G D}$ are weights assigned to likelihood components for the retained, pot total, and groundfish discard catches; $w_{j}$ is the average mass of a crab is length-class $j ; C_{t, j}, T_{t, j}$, and $T r_{t, j}$ are, respectively, the observed numbers of crab in size class $j$ for retained, pot total, and groundfish fishery discarded crab during year $t$, and $c$ is a small constant value. We assumed $c=$ 0.001 .

An additional retained catch likelihood (using Equation A.11a without $w$ ) for the retained catch in number of crabs during 1981/82-1984/85 was also considered in all scenarios.

## Catch-rate indices

The catch-rate indices are assumed to be lognormally distributed about the model prediction. Account is taken of variation in addition to that related to sampling variation:
$\left.\left.L L_{r}^{C P U E}=\lambda_{r, C P U E}\left\{0.5 \sum_{t} \ln \left[2 \pi\left(\sigma_{r, t}^{2}+\sigma_{e}^{2}\right)\right]+\sum_{t} \frac{\left(\ln \left(C P U E_{t}^{r}+c\right)-\ln (C \widehat{P U E} r\right.}{r}+c\right)\right)^{2}\right\}$
where $C P U E_{t}^{r}$ is the standardized retain catch-rate index for year $t, \sigma_{r, t}$ is standard error of the logarithm of $C P U E_{t}^{r}$, and $\widehat{P U E}_{t}^{r}$ is the model-estimate of $C P U E_{t}^{r}$ :

$$
\begin{equation*}
\widehat{C P U E}{ }_{t}^{r}=q_{k} \sum_{j} S_{j}^{T} S_{j}^{r}\left(N_{t, j}-0.5\left[\widehat{C_{t, j}}+\widehat{D_{t, j}}+\widehat{T r_{t, j}}\right]\right) e^{-y_{t} M} \tag{A.13}
\end{equation*}
$$

in which $q_{k}$ is the catchability coefficient during the $k$-th period (e.g., pre- and post-rationalization time periods), $\sigma_{e}$ is the extent of over-dispersion, $c$ is a small constant to prevent zero values (we assumed $c=0.001$ ), and $\lambda_{r, C P U E}$ is the weight assigned to the catch-rate data. We used the same likelihood formula (A.12) for fish ticket and cooperative survey retained catch rate indices. However, for cooperative survey catch rate prediction we used a different catchability parameter.

Following Burnham et al. (1987), we computed the $\ln$ (CPUE) variance by:

$$
\begin{equation*}
\sigma_{r, t}^{2}=\ln \left(1+C V_{r, t}^{2}\right) \tag{A.14}
\end{equation*}
$$

## Length-composition data

The length-composition data are included in the likelihood function using the robust normal for proportions likelihood, i.e., generically:
$L L_{r}^{L F}=0.5 \sum_{t} \sum_{j} \ln \left(2 \pi \sigma_{t, j}^{2}\right)-\sum_{t} \sum_{j} \ln \left[\exp \left(-\frac{\left(P_{t, j}-\hat{P}_{t, j}\right)^{2}}{2 \sigma_{t, j}}\right)+0.01\right]$
where $P_{t, j}$ is the observed proportion of crabs in length-class j in the catch during year $\mathrm{t}, \hat{P}_{t, j}$ is the model-estimate corresponding to $P_{t, j}$, i.e.:
$\hat{\mathrm{L}}_{\mathrm{t}, \mathrm{j}}^{\mathrm{r}}=\frac{\widehat{\mathrm{C}}_{\mathrm{t}, \mathrm{j}}}{\sum_{\mathrm{j}}^{\mathrm{n}} \mathrm{C}_{\mathrm{t}, \mathrm{j}}}$
$\hat{L}_{\mathrm{t}, \mathrm{j}}^{\mathrm{T}}=\frac{\widehat{T}_{\mathrm{t}, \mathrm{j}}}{\sum_{\mathrm{j}}^{\mathrm{n}} \mathrm{T}_{\mathrm{t}, \mathrm{j}}}$
$\hat{\mathrm{L}}_{\mathrm{t}, \mathrm{j}}^{\mathrm{GF}}=\frac{\mathrm{Tr}_{\mathrm{t}, \mathrm{j}}}{\sum_{\mathrm{j}}^{\mathrm{n}} \mathrm{Tr}_{\mathrm{t}, \mathrm{j}}}$
$\sigma_{t, j}^{2}$ is the variance of $P_{t, j}$ :
$\sigma_{t, j}^{2}=\left[\left(1-P_{t, j}\right) P_{t, j}+\frac{0.1}{n}\right] / S_{t}$
and $S_{t}$ is the effective sample size for year $t$ and $n$ is the number of size classes.

Note: The likelihood calculation for retained length composition starts from length-class 6 (mid length 128 mm CL ) because the length-classes 1 to 5 mostly contain zero data.

## Tagging data

Let $V_{j, t, y}$ be the number of tagged male crab that were released during year $t$ that were in size-class $j$ when they were released and were recaptured after $y$ years, and $\rho_{j, t, y}$ be the vector of recaptures by size-class from the males that were released in year $t$ that were in size-class $j$ when they were released and were recaptured after $y$ years. The log-likelihood corresponding to the multinomial distribution for the tagging data is then:
$\ln L=\lambda_{y, t a g} \sum_{j} \sum_{t} \sum_{y} \sum_{i} \rho_{j, t, y, i} \ln \hat{\rho}_{j, t, y, i}$
where $\lambda_{y, \operatorname{tag}}$ is the weight assigned to the tagging data for recapture year $y, \hat{\rho}_{j, t y, i}$ is the proportion in size-class $i$ of the recaptures of males that were released during year $t$ that were in size-class $j$ when they were released and were recaptured after $y$ years:
$\underline{\hat{\rho}}_{j, t, y} \propto \underline{s}^{T}[\mathbf{X}]^{y} \underline{Z}^{(j)}$
where $Z^{(j)}$ is a vector with $V_{j, t, y}$ at element $j$ and 0 otherwise, and $S^{T}$ is the vector of total selectivity for tagged male crab by the pot fishery. This log-likelihood function is predicated on the assumption that all recaptures are in the pot fishery and the reporting rate is independent of the size of crab.

## Penalties

Penalties are imposed on the deviations of annual pot fishing mortality about mean pot fishing mortality, annual trawl fishing mortality about mean trawl fishing mortality, recruitment about mean recruitment, and the posfunction (fpen):
$P_{1}=\lambda_{F} \sum_{t}\left(\ell \mathrm{n} F_{t}-\ell \mathrm{n} \bar{F}\right)^{2}$
$P_{2}=\lambda_{F^{T r}} \sum_{t}\left(\ell \mathrm{n} F_{t}^{T r}-\ell \mathrm{n} \bar{F}^{T r}\right)^{2}$
$P_{3}=\lambda_{R} \sum_{t}\left(\ln \varepsilon_{t}\right)^{2}$
$P_{5}=\lambda_{\text {posfn }} *$ fpen

Standardized Residual of Length Composition

$$
\begin{equation*}
\text { Std. } \operatorname{Res}_{t, j}=\frac{P_{t, j}-\widehat{P_{t, j}}}{\sqrt{2 \sigma_{t, j}^{2}}} \tag{A.24}
\end{equation*}
$$

## Output Quantities

## Harvest rate

Total pot fishery harvest rate:

$$
\begin{equation*}
E_{t}=\frac{\sum_{j=1}^{n}\left(\hat{C}_{j, t}+\widehat{D}_{j, t}\right)}{\sum_{j=1}^{n} N_{j, t}} \tag{A.25}
\end{equation*}
$$

Exploited legal male biomass at the start of year $t$ :
$L M B_{t}=\sum_{j=\text { legal size }}^{n} s_{j}^{T} s_{j}^{r} N_{j, t} w_{j}$
where $w_{j}$ is the weight of an animal in length-class $j$.
Mature male biomass on 15 February spawning time (NPFMC 2007a, b) in the following year:
$M M B_{t}=\sum_{j=\text { mature size }}^{n}\left\{N_{j, t} e^{-y^{\prime} M}-\left(\hat{C}_{j, t}+\widehat{D}_{j, t}+\widehat{T r}_{j, t}\right) e^{\left(y_{t}-y^{\prime}\right) M}\right\} w_{j}$
where $y^{\prime}$ is the elapsed time from 1 July to 15 February in the following year.
For estimating the next year limit harvest levels from current year stock abundances, an $F_{O F L}$ value is needed. The current crab management plan specifies five different Tier formulas for different stocks depending on the strength of information available for a stock, for computing $F_{O F L}$ (NPFMC $2007 \mathrm{a}, \mathrm{b}$ ). For the golden king crab, the following Tier 3 formula is applied to compute $F_{O F L}$ :

If,
$M M B_{\text {current }}>M M B_{35 \%}, F_{\text {OFL }}=F_{35 \%}$
If,
$M M B_{\text {current }} \leq M M B_{35 \%}$ and $M M B_{\text {current }}>\beta M M B_{35 \%}$,
$F_{O F L}=F_{35 \%} \frac{\left(\frac{M M B_{\text {current }}}{M M B_{35 \%}}-\alpha\right)}{(1-\alpha)}$
If,
$M M B_{\text {current }} \leq \beta M M B_{35 \%}$,
$F_{O F L}=0$.
where
$\beta=$ a parameter with a restriction that $0 \leq \beta<1$. A default value of 0.25 is used,
$\alpha=$ a parameter with a restriction that $0 \leq \alpha \leq \beta$. A default value of 0.1 is used,
$\mathrm{MMB}_{\text {current }}$ = the mature male biomass in the current year, and
$M M B_{35 \%}=$ a proxy $M M B_{M S Y}$ for Tier 3 stocks.
Because projected $\mathrm{MMB}_{\mathrm{t}}$ (i.e., $\mathrm{MMB}_{\text {current }}$ ) depends on the intervening retained and discard catch (i.e., $\mathrm{MMB}_{\mathrm{t}}$ is estimated after the fishery), an iterative procedure is applied using Equations A. 27 and A. 28 with retained and discard catch predicted from Equations A.2b-d. The next year limit harvest catch is estimated using Equations A.2b-d with the estimated $F_{O F L}$ value.

Table A1. Pre-specified and estimated parameters of the population dynamics model

| Parameter | Number of parameters |
| :---: | :---: |
| Fishing mortalities: |  |
| Pot fishery, $F_{t}$ | 1981-2020 (estimated) |
| Mean pot fishery fishing mortality, $\bar{F}$ | 1 (estimated) |
| Groundfish fishery, $F_{t}{ }^{\text {rr }}$ | 1989-2020 (the mean F for 1989 to 1994 was used to estimate groundfish discards back to 1981 (estimated) |
| Mean groundfish fishery fishing mortality, $\bar{F}^{T r}$ | 1 (estimated) |
| Selectivity and retention: |  |
| Pot fishery total selectivity, $\theta_{50}^{\mathrm{T}}$ | 2 (1981-2004; 2005+) (estimated) |
| Pot fishery total selectivity difference, delta $\theta^{\text {T }}$ | 2 (1981-2004; 2005+) (estimated) |
| Pot fishery retention, $\theta_{50}^{\mathrm{r}}$ | 1 (1981+) (estimated) |
| Pot fishery retention selectivity difference, delta $\theta^{r}$ | 1 (1981+) (estimated) |
| Groundfish fishery selectivity | fixed at 1 for all size-classes |
| Growth: |  |
| Expected growth increment, $\omega_{1}, \omega_{2}$ | 2 (estimated) |
| Variability in growth increment, $\sigma$ | 1 (estimated) |
| Molt probability (size transition matrix with tag data), a | 1 (estimated) |
| Molt probability (size transition matrix with tag data), b | 1 (estimated) |
| Natural mortality, $M$ | 1 (pre-specified, $0.21 \mathrm{yr}^{-1}$ ) |
| Recruitment: |  |
| Number of recruiting length-classes | 5 (pre-specified) |
| Mean recruit length | 1 (pre-specified, 110 mm CL ) |
| Distribution to length-class, $\beta_{\mathrm{r}}$ | 1 (estimated) |
| Median recruitment, $\overline{\mathrm{R}}$ | 1 (estimated) |
| Recruitment deviations, $\boldsymbol{E}_{t}$ |  |
| Fishery catchability, q | 2 (1985-2004; 2005+) (estimated) |
| Additional CPUE indices standard deviation, $\sigma_{\mathrm{e}}$ | 1 (estimated) |
| Likelihood weights (coefficient of variation) | Pre-specified, varies by scenario |

Table A2. Specifications for the weights with corresponding coefficient of variations* in parentheses for each model for EAG and WAG.

| Weight | Models 19.1, 21.1a, 21.1b, and 21.1c |
| :---: | :---: |
| Catch: |  |
| Retained catch for 1981-1984 and/or 1985-2020, $\lambda_{r}$ | 500 (0.032) |
| Total catch for 1990-2020, $\lambda_{T}$ | Number of sampled pots scaled to a max 250 |
| Groundfish bycatch for 1989-2020, $\lambda_{G D}$ | 0.2 (3.344) |
| Catch-rate: <br> Observer legal size crab catch-rate for 1995-2020, $\lambda_{r, \text { CPUE }}$ |  |
| Fish ticket retained crab catch-rate for 1985-1998, $\lambda_{r, \text { CPUE }}$ | 1 (0.805) |
| Penalty weights: |  |
| $\begin{aligned} & \text { Pot fishing mortality dev, } \lambda_{F} \\ & \text { Groundfish fishing mortality dev, } \lambda_{F^{T r}} \end{aligned}$ | Initially 1000 , relaxed to 0.001 at phases $\geq$ select. phase <br> Initially 1000 , relaxed to 0.001 at phases $\geq$ select. phase |
| Recruitment, $\lambda_{R}$ <br> Posfunction (to keep abundance estimates always positive), $\lambda_{\text {posfn }}$ | $\begin{aligned} & 2(0.533) \\ & 1000(0.022) \end{aligned}$ |
| Tagging likelihood | EAG individual tag returns |

* Coefficient of Variation, CV $=\sqrt{\exp \left[\frac{1}{2 W}\right]-1}, \quad w=$ weight


## Appendix B: Catch and CPUE data

The commercial catch and length frequency distribution were estimated from ADF\&G landing records and dockside sampling (Bowers et al. 2008, 2011). The annual retained catch, total catch, and groundfish (or trawl) discarded mortality are provided in Tables 1, 2, and 2 b for EAG and WAG. The weighted length frequency data were used to distribute the catch into 5mm size intervals. The length frequency data for a year were weighted by each sampled vessel's catch as follows. The $i$-th length-class frequency was estimated as:

$$
\begin{equation*}
\sum_{j=1}^{k} C_{j} \frac{L F_{j, i}}{\sum_{i=1}^{n} L F_{j, i}} \tag{B.1}
\end{equation*}
$$

where $k=$ number of sampled vessels in a year, $L F_{j, i}=$ number of crabs in the $i$-th length-class in the sample from $j$-th vessel, $n=$ number of size classes, $C_{j}=$ number of crabs caught by $j$-th vessel. Then the relative frequency for the year was calculated and applied to the annual retained catch (in number of crabs) to obtain retained catch by length-class.

The annual total catch (in number of crabs) was estimated by the observer nominal (unstandardized) total CPUE considering all vessels multiplied by the total fishing effort (number of pot lifts). The weighted length frequency of the observer samples across the fleet was estimated using Equation B.1. Observer measurement of crab ranged from 20 to 220 mm CL. To restrict the total number of crabs to the model assumed size range ( $101-185+\mathrm{mm} \mathrm{CL}$ ), the proportion of observer total relative length frequency corresponding to this size range was multiplied by the total catch (number of crabs). This total number of crabs was distributed into length-classes using the weighted relative length frequency. Thus, crab sizes $<101 \mathrm{~mm}$ CL were excluded from the model. In addition, all crab $>185 \mathrm{~mm}$ CL were pooled into a plus length class. Note that the total crab catches by size that went into the model did not consider retained and discard components separately. However, once the model estimated the annual total catch, then retained catch was deducted from this total and multiplied by handling mortality [we used a $20 \%$ handling mortality (Siddeek et al. 2005) to obtain the directed fishery discarded (dead) catch].

Observer data have been collected since 1988 (Moore et al. 2000; Barnard et al. 2001; Barnard and Burt 2004; Gaeuman 2011), but data were not comprehensive in the initial years, so a shorter time series of data for the period 1990/91-2020/21 was selected for this analysis. During 1990/91-1994/95, observers were only deployed on catcher-processor vessels. During 1995/96-2004/05, observers were deployed on all fishing vessels during fishing activity. Observers have been deployed on all fishing vessels since 2005/06, but catcher-only vessels are only required to carry observers for a minimum of $50 \%$ of their fishing activity during a season; catcher-processor vessels are still required to carry observers during all fishing activity. Onboard observers sample seven pots per day (may be different numbers of pots per string) and count and measure all crabs caught and categorize catch as females, sublegal males, retained legal males, and non-retained legal males in a sampled pot. Prior to the 2009/10 season, depending on season, area, and type of fishing vessel, observers were also instructed to sample additional pots in which all crab were only counted and categorized as females, sublegal males, retained legal males, and non-retained legal males, but were not measured. Annual mean nominal CPUEs of retained and total crabs were estimated considering all
sampled pots within each season (Table 3). The observer CPUE data collection improved over the years and the data since 1995/96 are more reliable. Thus, for model fitting, the observer CPUE time series was restricted to 1995/96-2020/21. The 1990/91-2020/21 observer database consists of 118,025 records and that of 1995/96-2020/21 contains 113,746 records, For CPUE standardization, these data were further reduced by $5 \%$ cutoff of Soak time and $1 \%$ cutoff of Depth on both ends of the variable range to remove unreliable data or data from dysfunctional pot operations and restricting to vessels which have made five trips per year for at least three years during 1985/86-2020/21.

We detected some computational errors in raw size frequency summary data preparation (observer and fish ticket sampling) for 2016-2019 and rectified errors in relative retained and total size frequency computations in the current analysis. The correction of errors did not affect retained catch crab distribution by size bins but caused minor changes to allocation of total catch crab into size bins.

Length-specific CPUE data collected by observers provides information on a wider size range of the stock than did the commercial catch length frequency data obtained from mostly legalsized landed males.

There were significant changes in fishing practice due to changes in management regulations (e.g., since 1996/97 constant TAC and since 2005/06 crab rationalization), pot configuration (escape web on the pot door increased to 9 " since 1999), and improved observer recording in Aleutian Islands golden king crab fisheries since 1998. These changes prompted us to consider two separate observer CPUE time series, 1995/96-2004/05 and 2005/06-2020/21, to estimate CPUE indices for model input.

To include a long time series of CPUE indices for stock abundance contrast, we also considered the 1985/86-1998/99 legal size standardized CPUE as a separate likelihood component in all scenarios. Because of the lack of soak time data before 1990, we estimated the CPUE index considering a limited set of explanatory variables (e.g., vessel, captain, area, month) and fitting the negative binomial GLM model to fish ticket data (Tables 4 and 14).

When using CPUE indices in the model fit, we compared the predicted with the observed legal male CPUE in the observer CPUE likelihoods because legal male (retained plus non-retained) data are more reliable than total in the observer samples.

The CPUE standardization followed the GLM fitting procedure (Maunder and Punt 2004; Starr 2012; Siddeek et al. 2018). Following a suggestion from the CIE reviewers in June 2018 we reduced the number of gear codes in the database after consulting with the fishing industry (Rip Carlton, Chad Hoefer, and Scott Goodman, personal communication December 2018; Table B1). Following an SSC suggestion in October 2018, we used a hybrid procedure: First, we selected a scope of variables set by Akike Information Criterion, AIC (Burnham and Anderson 2002). An increase of more than 2 units in the AIC was used to identify the variable to be included successively (stepAIC program, R Core Team 2020). Then, the model parsimony was improved further by successively removing the term that explained the least
proportion of deviance ( $\mathrm{R}^{2}<0.01$ ) (stepCPUE R function was used, Siddeek et al. 2018). Feenstra, et al. (2019) used a similar hybrid approach.

Table B.1. Updated gear codes for observer data analysis. Only gear code \#5, 6, 7, 8, and 13 were considered following crab industry suggestion. Note: Identical codes were given to those gear codes with similar catchability/selectivity. X indicates gear codes that were ignored.

| Original Gear code | Pot gear description | Mark X against the code that can be ignored | Number encountered by observers during 1990-2016 | Updated gear code |
| :---: | :---: | :---: | :---: | :---: |
| 1 | Dungeness crab pot, small \& round | X | 2 | X |
| 2 | Pyramid pot, tunnel openings usually on sides, stackable | X | 2121 | X |
| 3 | Conical pot, opening at top of cone, stackable | X | 2000 | X |
| 4 | 4' X 4' rectangular pot |  | 60 | X |
| 5 | 5' X 5' rectangular pot |  | 18032 | 5 |
| 6 | 6' X 6' rectangular pot |  | 17508 | 6 |
| 7 | 7' X 7' rectangular pot |  | 23806 | 7 |
| 8 | $8^{\prime} \mathrm{X} \mathrm{8} 8$ 'rectangular pot |  | 1936 | 8 |
| 9 | $51 / 2^{\prime}$ X $51 / 2^{\prime}$ rectangular pot |  | 6934 | 5 |
| 10 | $61 / 2^{\prime}$ X $61 / 2^{\prime}$ rectangular pot |  | 22085 | 6 |
| 11 | $71 / 2^{\prime}$ X $71 / 2^{\prime}$ rectangular pot |  | 387 | 7 |
| 12 | Round king crab pot, enlarged version of Dungeness crab pot |  | 8259 | X |
| 13 | $10^{\prime} \mathrm{X} \mathrm{10'} \mathrm{rectangular} \mathrm{pot}$ |  | 466 | 13 |
| 14 | $9^{\prime} \mathrm{X} 9^{\prime}$ rectangular pot | X | 1 | X |
| 15 | $81 / 2^{\prime} \mathrm{X} 81 / 2^{\prime}$ rectangular pot | X | 1 | X |
| 16 | $91 / 2^{\prime} \mathrm{X} 91 / 2^{\prime}$ rectangular pot | X | Not used | X |
| 17 | $8^{\prime} \mathrm{X} 9{ }^{\prime}$ rectangular pot | X | 1 | X |
| 18 | $8^{\prime} \mathrm{X} 10{ }^{\prime}$ rectangular pot | X | 1 | X |
| 19 | 9' X 10' rectangular pot |  | Not used | X |
| 20 | 7' X 8' rectangular pot | X | 252 | X |
| 21 | Hair crab pot, longlined and small, stackable |  | Not used | X |
| 22 | snail pot | X | 1 | X |
| 23 | Dome-shaped pot, tunnel opening on top, often longlined in deep-water fisheries <br> ADF\&G shellfish research $7^{\prime}$ X $7^{\prime}$ X34" | X | 6756 | X |
| 24 | rectangular pot with 2.75 " stretch mesh and no escapement rings or mesh |  | Research pot | X |
| 80 | Historical: Cod pot, any shape pot targeting cod, usually with tunnel fingers | X | 711 | X |
| 81 | Historical: Rectangular pot, unknown size, with escape rings | X | 1123 | X |

All scenarios used CPUE indices estimated by the hybrid GLM method. Following a January 2019 CPT request, we considered a Year:Area interaction factor as a special case for a CPUE standardization scenario.

Thus we estimated two sets of observer CPUE indices for model input, 19.1 (reduced number of gear codes), and 21.1c (reduced number of gear codes and Year:Area interaction).

## Observer CPUE index by GLM

## a. Non-interaction GLM model

The CPUE standardization followed the GLM fitting procedure (Maunder and Punt 2004; Starr 2012; Siddeek et al. 2016b). We considered the negative binomial GLM on positive and zero catches to select the explanatory variables. The response variable CPUE is the observer sample catch record for a pot haul. The negative binomial model uses the log link function for the GLM fit.

For the non-interaction model, we assumed the null model to be:

$$
\begin{equation*}
\ln \left(\text { CPUE }_{i}\right)=\text { Year }_{y_{i}} \tag{B.2}
\end{equation*}
$$

where Year is a factorial variable.
The maximum set of model terms offered to the stepwise selection procedure was:
$\ln \left(\right.$ CPUE $\left._{\mathrm{I}}\right)=$ Year $_{\mathrm{y}_{\mathrm{i}}}+\mathrm{ns}\left(\right.$ Soak $\left._{\text {si }}, \mathrm{df}\right)+$ Month $_{\mathrm{m}_{\mathrm{i}}}+$ Vessel $_{\mathrm{vi}}+$ Captain $_{\mathrm{ci}}+$ Area $_{\mathrm{ai}}+$ Gear $_{\mathrm{gi}}+\mathrm{ns}\left(\right.$ Depth $\left._{\text {di }}, \mathrm{df}\right)$,
where Soak is in unit of days and is numeric; Month, Area (Block) code, Vessel code, Captain code, and Gear code are factorial variables; Depth in fathom is a numeric variable; ns=cubic spline, and $\mathrm{df}=$ degree of freedom.

We used a log link function and a dispersion parameter ( $\theta$ ) in the GLM fitting process. We used the $\mathrm{R}^{2}$ criterion for predictor variable selection (Siddeek et al. 2016b).

We calculated appropriate degrees of freedom and dispersion parameters by calculating AICs for a range of values and locating the best values at the minimum AIC (Figures B1 and B.2, respectively).


Figure B.1. AIC vs degrees of freedom for soak time and depth during pre- and postrationalization periods for EAG (top) and WAG (bottom). Vertical lines identify the optimum degrees of freedom values chosen for CPUE standardization.


Figure B.2. AIC vs theta (dispersion parameter) during pre- and post-rationalization periods for EAG (top) and WAG (bottom). Vertical lines identify the optimum theta values chosen for CPUE standardization.

Instead of using the traditional AIC ( $-2 \log$ _likelihood +2 p ) we used the Consistent Akaike Information Criteria (CAIC) (Bozdogan 1987) $\left\{-2 \log _{2}\right.$ likelihood $\left.+[\ln (\mathrm{n})+1]^{*} \mathrm{p}\right\}$ for variable selection by StepAIC, where $n=$ number of observations and $p=$ number of parameters to be estimated. The number of selected variables were further reduced for parsimony, if feasible, by the $\mathrm{R}^{2}$ criterion using the StepCPUE function. i.e., a hybrid selection procedure (Feenstra et al. 2019).

Example R codes used for main effect GLM fitting are as follows:
For EAG 1995_04 CPUE indices:
library(MASS)
library(splines)
Step 1:
glm.object $<-$ glm(Legals $\sim$ Year,family $=$ negative.binomial(1.38),data=datacore)
epotsampleoutAIC<-stepAIC(glm.object,scope=list(upper $=$
$\sim($ Year $+n s($ SoakDays, $d f=4)+$ Month + Vessel + Captain + Area + Gear $+n s($ Depth, $d f=9)$ ), lower $=$
$\sim$ Year),family $=$ negative.binomial(1.38), direction $=$ "forward", trace $=9, k=\log$ (nrow(datacore))
+1.0 )
Step 2:
glm.object $<-$ glm(Legals $\sim$ Year,family $=$ negative.binomial(1.38),data=datacore)
epotsampleout<-
stepCPUE(glm.object,scope $=$ list(upper $=\sim($ Year + Gear + Captain $+n s($ SoakDays, $d f=4)+$
Month + Area),lower $=\sim$ Year),family =negative.binomial(1.38),direction="forward",trace=9,r 2.change $=0.01$ )

The final main effect models for EAG were:
Model 19.1:
Initial selection by stepAIC:
$\ln ($ CPUE $)=$ Year + Gear + Captain + ns $($ Soak, 4$)+$ Month + Block
AIC=203,808
Final selection by stepCPUE:
$\ln ($ CPUE $)=$ Year + Captain + Gear + ns $($ Soak, 4$)+$ Month
for the 1995/96-2004/05 period $\left[\theta=1.38, \mathrm{R}^{2}=0.2205\right]$
Initial selection by stepAIC:

```
\(\ln (\) CPUE \()=\) Year + Captain + Gear + ns \((\) Soak, 3\()+\) Month
AIC=77,173
```

Final selection by stepCPUE:
$\ln$ (CPUE) $=$ Year + Captain + Gear + ns (Soak, 3)
for the 2005/06-2020/21 period $\left[\theta=2.32, R^{2}=0.1082\right]$.

The final models for WAG were:
Model 19.1:
Initial selection by stepAIC:
$\ln ($ CPUE $)=$ Year + Captain $+\mathrm{ns}($ Soak, 7$)+$ Gear + Area + Month $+\mathrm{ns}($ Depth, 5$)+$ Vessel
AIC=190,897
Final selection by stepCPUE:
$\ln$ (CPUE) $=$ Year + Captain + ns $($ Soak, 7$)+$ Gear
for the 1995/96-2004/05 period [ $\theta=0.97, \mathrm{R}^{2}=0.1681$ ]
Initial selection by stepAIC:
$\ln ($ CPUE $)=$ Year + Captain + Gear + Month $+\mathrm{ns}($ Soak, 18)
AIC=116,552
Final selection by stepCPUE:
$\ln ($ CPUE $)=$ Year + Captain + Gear $+\mathrm{ns}($ Soak, 18)
for the 2005/06-2020/21 period [ $\theta=1.13, R^{2}=0.0776$, Soak forced in].

## b. Year:Area interaction GLM:

For year and area interaction analysis, we designed the areas in to $1 \mathrm{nmi} \times 1 \mathrm{nmi}$ grids enmeshed in 10 larger blocks as follows. The number of blocks was restricted to a few to prevent GLM fitting problems (Figure B. 3 and Table B.2).


Figure B.3. The 1995/96-2020/21 observer pot samples enmeshed in 10 blocks for the Aleutian Islands golden king crab.
The blocks were determined from visually exploring each year's pot distribution locations (available with the first author). The blocks contain observed patches of crab distribution during this period.

Table B.2. Number of $1 \mathrm{nmi} \times 1 \mathrm{nmi}$ grids containing observer sample locations within each block by fishing year for the Aleutian Islands golden king crab, 1995/96-2020/21 data. Blocks 1-4 belong to EAG and 5-10 to WAG. Sum of ever fished number of grids for each block is listed at the bottom row.

| FMP Season | Block_1 | Block_2 | Block_3 | Block_4 | Block_5 | Block_6 | Block_7 | Block_8 | Block_9 | Block_10 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1995 | 125 | 529 | 748 | 379 | 218 | 373 | 112 | 722 | 166 | 122 |
| 1996 | 149 | 814 | 761 | 372 | 89 | 473 | 359 | 799 | 200 | 35 |
| 1997 | 116 | 530 | 755 | 257 | 202 | 443 | 104 | 568 | 274 | 0 |
| 1998 | 78 | 581 | 453 | 236 | 18 | 318 | 157 | 251 | 132 | 0 |
| 1999 | 123 | 593 | 454 | 231 | 163 | 476 | 182 | 627 | 193 | 145 |
| 2000 | 72 | 540 | 754 | 301 | 187 | 440 | 195 | 555 | 547 | 47 |
| 2001 | 123 | 507 | 507 | 329 | 45 | 369 | 288 | 634 | 256 | 9 |
| 2002 | 97 | 387 | 584 | 271 | 71 | 341 | 205 | 335 | 242 | 37 |
| 2003 | 43 | 492 | 530 | 299 | 111 | 347 | 212 | 465 | 150 | 61 |
| 2004 | 81 | 289 | 377 | 216 | 77 | 319 | 150 | 359 | 172 | 116 |
| 2005 | 0 | 205 | 221 | 118 | 8 | 220 | 83 | 261 | 54 | 0 |
| 2006 | 0 | 154 | 248 | 122 | 15 | 191 | 58 | 220 | 39 | 0 |
| 2007 | 0 | 111 | 177 | 110 | 24 | 228 | 78 | 173 | 20 | 0 |
| 2008 | 0 | 111 | 203 | 93 | 12 | 181 | 67 | 196 | 0 | 0 |
| 2009 | 0 | 59 | 146 | 60 | 6 | 137 | 95 | 220 | 25 | 0 |
| 2010 | 0 | 81 | 141 | 85 | 1 | 115 | 73 | 260 | 39 | 0 |
| 2011 | 0 | 126 | 117 | 33 | 3 | 83 | 73 | 266 | 9 | 0 |
| 2012 | 0 | 146 | 110 | 56 | 7 | 91 | 85 | 312 | 53 | 0 |
| 2013 | 2 | 149 | 129 | 51 | 12 | 144 | 105 | 293 | 86 | 0 |
| 2014 | 1 | 138 | 96 | 41 | 39 | 120 | 114 | 319 | 37 | 0 |
| 2015 | 0 | 135 | 147 | 61 | 46 | 163 | 106 | 280 | 16 | 48 |
| 2016 | 0 | 145 | 231 | 63 | 26 | 134 | 89 | 210 | 106 | 0 |
| 2017 | 0 | 97 | 170 | 110 | 11 | 87 | 79 | 198 | 118 | 0 |
| 2018 | 0 | 91 | 158 | 95 | 7 | 69 | 82 | 204 | 121 | 0 |
| 2019 | 1 | 112 | 171 | 101 | 0 | 0 | 89 | 316 | 138 | 0 |
| 2020 | 4 | 109 | 193 | 95 | 0 | 0 | 76 | 287 | 91 | 36 |


| Ever Fished: |  |  |  |  |  |  |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| AIGKC All Seasons | Block_1 | Block_2 | Block_3 | Block_4 | Block_5 | Block_6 | Block_7 | Block_8 | Block_9 | Block_10 |
| $1995-2020-$ Sum of $1 \times 1$ cells | 381 | 1402 | 1792 | 917 | 459 | 1028 | 796 | 2012 | 1021 | 334 |

We assumed the null model to be

$$
\begin{equation*}
\ln \left(\text { CPUE }_{\mathrm{i}}\right)=\text { Year }_{\mathrm{y}_{\mathrm{i}}}: \text { Area }_{\mathrm{ai}} \tag{B.8}
\end{equation*}
$$

The maximum set of model terms offered to the stepwise selection procedure was:
$\ln \left(\right.$ CPUE $\left._{\mathrm{I}}\right)=$ Year $_{\mathrm{yi}_{\mathrm{i}}}:$ Area $_{a i}+\mathrm{ns}\left(\right.$ Soak $_{\text {si }}$, df $)+$ Month $_{\mathrm{m}_{\mathrm{i}}}+$ Vessel $_{\mathrm{vi}}+$ Captain $_{\mathrm{ci}}+$
Area $_{\mathrm{ai}}+$ Gear $_{\mathrm{gi}}+\mathrm{ns}\left(\right.$ Depth $_{\text {di }}$, df).

Example R codes used for interaction effect GLM fitting are as follows:
For WAG 1995_04 CPUE indices:
library(MASS)

## library(splines)

Step 1:
glm.object $<-$ glm(Legals $\sim$ Year:Area,family $=$ negative.binomial(0.97),data=datacore)
wpotsampleoutAIC<-stepAIC(glm.object,scope=list(upper=
$\sim($ Year:Area + ns $($ SoakDays, $\mathrm{df}=7)+$ Month + Vessel + Captain + Area + Gear +
$\mathrm{ns}($ Depth, $\mathrm{df}=5$ )),lower=~Year:Area),family=
negative.binomial(0.97), direction="forward",trace $=9, \mathrm{k}=\log ($ nrow $($ datacore $))+1.0)$
Step 2:
glm.object<- glm(Legals $\sim$ Year:Area,family $=$ negative.binomial(0.97),data=datacore)
wpotsampleout<-stepCPUE(glm.object,scope=list(upper=
$\sim($ Captain $+\mathrm{ns}($ SoakDays, $\mathrm{df}=7)+$ Gear + Area + Month + Year:Area),lower $=\sim$ Year:Area),family $=$ negative.binomial(0.97),direction="forward",trace=9,r2.change=0.01)

The final interaction effect models for EAG were:
Model 21.1c:
Initial selection by stepAIC:
$\ln ($ CPUE $)=$ Gear + Captain + ns $($ Soak, 4$)+$ Month + Year: Area
AIC=203,851

Final selection by stepCPUE:
$\ln ($ CPUE $)=$ Gear + Captain + ns $($ Soak, 4$)+$ Year: Area
for the 1995/96-2004/05 period [ $\theta=1.38, \mathrm{R}^{2}=0.2235$ ]
Initial selection by stepAIC:
$\ln ($ CPUE $)=$ Vessel + Gear $+\mathrm{ns}($ Soak, 3$)+$ Month + Year: Area AIC=72,343

Final selection by stepCPUE:
$\ln ($ CPUE $)=$ Vessel + Gear $+\mathrm{ns}($ Soak, 3$)+$ Year: Area
for the 2005/06-2020/21 period $\left[\theta=2.32, \mathrm{R}^{2}=0.1169\right]$.

The final interaction effect models for WAG were:

Model 21.1c:
Initial selection by stepAIC:
$\ln ($ CPUE $)=$ Vessel + ns $($ Soak, 7$)+$ Gear + Month + Year: Area AIC=191,018

Final selection by stepCPUE:
$\ln ($ CPUE $)=$ Vessel + ns(Soak, 7) + Gear + Year: Area
for the 1995/96-2004/05 period [ $\theta=0.97, \mathrm{R}^{2}=0.1719$ ]
Initial selection by stepAIC:
$\ln ($ CPUE $)=$ Gear + Vessel + Month + Year: Area $+n s($ Soak, 18 $)$
AIC=116,859
Final selection by stepCPUE:
$\ln ($ CPUE $)=$ Gear + Year: Area $+n s($ Soak, 18)
for the 2005/06-2020/21 period $\left[\theta=1.13, R^{2}=0.0818\right.$, Soak forced $\left.i n\right]$.

## Steps:

1. Block-scale analysis:

The bias corrected estimate of CPUE index for each Year-Area (Area=Block) interaction was first obtained as:
$C P U E_{i j}=e^{Y B_{i j}+\sigma_{i j}^{2} / 2}$
where $C P U E_{i j}$ is the CPUE index in the ith year and jth block, $Y B_{i j}$ is the coefficient of the $i$ th year and $j$ th block interaction, and $\sigma_{i j}$ is the biased correction standard error for expected CPUE value.

The number of $1 \mathrm{nmi} \times 1 \mathrm{nmi}$ grids in each block can change from year to year; so, we considered using the number of grids ever fished in a block, $N_{\text {ever } j}$ [this is equivalent to assuming that the grids fished in any year randomly sample the stock in that block (Campbell, 2004)].

The abundance index for $j$ th block in $i$ th year is
$B_{i j}=N_{\text {ever }_{j}} C P U E_{i j}$
Notice in Table B. 2 that no or very few observer samplings occurred in certain years for a whole block. We filled the $B_{i j}$ index gaps resulting from Year:Area CPUE standardization model fit as follows:
$\widehat{B_{l, j}}=e^{A_{i}+C_{j}}$
fitted by GLM [i.e., fitting a log-linear model, $\ln \left(\widehat{B}_{i, j}\right)=A_{i}+C_{j}$ ],
where $B_{i, j}$ is the available index of biomass for year i and block $j, A_{i}$ is a year factor, and $C_{j}$ is a block factor, and used this model to predict the unavailable biomass index for blocks $x$ years with no (or very limited) data.

An example set of R codes used to predict the missing biomass index is as follows:

## library(MASS)

To fit the log-linear model (Equation B.16):

## glm.fit<- $\operatorname{glm}\left(\log \left(\mathrm{B}_{\mathrm{ij}}\right) \sim\right.$ Year $_{\mathrm{i}}+$ Block $_{\mathrm{j}}$, data=Bindex $)$

where the data frame "Bindex" contains available $\mathrm{B}_{\mathrm{i} j}$, Year $_{\mathrm{i}}$, and Block $_{\mathrm{j}}$ column values.
To predict the missing biomass index Y :

## Y<- predict.glm (glm.fit, BindexFillpredict, se.fit=TRUE)

where the new data frame "BindexFillpredict" contains Year ${ }_{i}$ and Block $_{j}$ column values for which $\mathrm{B}_{\mathrm{ij}}$ indices are needed and contains an empty $\mathrm{B}_{\mathrm{ij}}$ column for fill in.

By setting se.fit=TRUE, the standard errors, $\sigma_{i j}$, of predictions are also estimated.
Bias correction was made to each predicted biomass index by $B_{i, j}=e^{\hat{Y}_{i, j}+\sigma_{i j}^{2} / 2}$ where $\sigma_{i j}$ is the standard error of predicted $Y_{i, j}$ value, which is on the scale of the linear predictor (i.e., log transformed $\mathrm{B}_{\mathrm{ij}}$ ). The standard error for each year and area combination is estimated as follows.

If we denote the covariance matrix of the fitted "glm.fit" as $\Sigma$ and write the coefficients for linear combination of a set of predictors in a vector form as $C$, then the standard error of prediction for that combination is $\sqrt{\mathrm{C}^{\prime} \Sigma \mathrm{C}}$, where $C^{\prime}$ is the transpose of vector $C$.

Annual biomass index, $B_{i}$, was estimated as,
$B_{i}=\sum_{j} B_{i j}$

The variance of the total biomass index was computed as:
$\operatorname{Var}\left(B_{i}\right)=\sum_{j} N_{\text {ever }, j}{ }^{2} \operatorname{var}\left(\operatorname{CPUE}_{i, j}\right)$
where $\boldsymbol{N}_{\text {ever }, j}$ is the total number of 1 mnix 1 mni cells ever fished in block $j$, and $C P U E_{i, j}$ is the CPUE index for year $i$ and block $j$.

To compare with other CPUE index estimates (Figures 24 for EAG and 41 for WAG) as well as to use in the assessment model 21.1c, we rescaled the $B_{i}$ indices by the geometric mean of estimated $B_{i}$ values (Equation B.17) separately for the pre- and postrationalization periods. The corresponding standard error $(\sim \mathrm{CV})$ of $B_{i}$ was estimated by

$$
\begin{equation*}
\sqrt{\frac{\operatorname{Var}\left(B_{i}\right)}{\left(B_{i}\right)^{2}}} \tag{B.19}
\end{equation*}
$$

The rescaled biomass indices with standard errors are listed in Table B. 3 for EAG and Table B. 4 for WAG.

Table B.3. Steps to estimate biomass-based abundance indices with standard errors for 1995/962020/21 in EAG. GMScaled B_index and B_Index SE were used as CPUE index and its standard error.

| Year | B_Index | GMScaled B_Index | $\operatorname{Var}(\mathrm{B}$ _index) | $\operatorname{Var}\left(\mathrm{B} \_\right.$Index $) /\left(B_{I} \text { Index }\right)^{2}$ | B_Index SE |
| :--- | ---: | ---: | ---: | ---: | ---: |
| 1995 | 1646.045 | 0.772 | 31119.811 | 0.011 | 0.107 |
| 1996 | 1664.192 | 0.781 | 30961.039 | 0.011 | 0.106 |
| 1997 | 1657.073 | 0.777 | 26973.953 | 0.010 | 0.099 |
| 1998 | 1983.401 | 0.930 | 28416.977 | 0.007 | 0.085 |
| 1999 | 1889.406 | 0.886 | 26998.598 | 0.008 | 0.087 |
| 2000 | 1829.271 | 0.858 | 48321.122 | 0.014 | 0.120 |
| 2001 | 2644.434 | 1.240 | 159513.439 | 0.023 | 0.151 |
| 2002 | 2685.133 | 1.260 | 33738.827 | 0.005 | 0.068 |
| 2003 | 2403.298 | 1.127 | 33864.001 | 0.006 | 0.077 |
| 2004 | 3651.543 | 1.713 | 114213.956 | 0.009 | 0.093 |
| 2005 | 11479.919 | 1.051 | 2445608.389 | 0.019 | 0.136 |
| 2006 | 8686.932 | 0.796 | 2442560.854 | 0.032 | 0.180 |
| 2007 | 9615.569 | 0.881 | 2441515.275 | 0.026 | 0.163 |
| 2008 | 9592.846 | 0.879 | 2445626.553 | 0.027 | 0.163 |
| 2009 | 8430.385 | 0.772 | 2454692.100 | 0.035 | 0.186 |


| 2010 | 8569.358 | 0.785 | 2452577.255 | 0.033 | 0.183 |
| :--- | ---: | ---: | ---: | :--- | :--- |
| 2011 | 11695.405 | 1.071 | 2455608.856 | 0.018 | 0.134 |
| 2012 | 11150.360 | 1.021 | 2451865.327 | 0.020 | 0.140 |
| 2013 | 11544.882 | 1.057 | 2516532.727 | 0.019 | 0.137 |
| 2014 | 14396.446 | 1.319 | 2451276.698 | 0.012 | 0.109 |
| 2015 | 13446.866 | 1.232 | 2445059.570 | 0.014 | 0.116 |
| 2016 | 11681.802 | 1.070 | 2443192.818 | 0.018 | 0.134 |
| 2017 | 10484.499 | 0.960 | 2447524.594 | 0.022 | 0.149 |
| 2018 | 12931.639 | 1.184 | 2447048.616 | 0.015 | 0.121 |
| 2019 | 12126.070 | 1.111 | 2569011.349 | 0.017 | 0.132 |
| 2020 | 10966.012 | 1.004 | 2507915.925 | 0.021 | 0.144 |

Table B.4. Steps to estimate biomass-based abundance indices with standard errors for 1995/962020/21 in WAG. GMScaled B_index and B_Index SE were used as CPUE index and its standard error.

| Year | B_Index | GMScaled B_Index | Var(B_index) | Var(B_Index)/(B_Index) | B_Index SE |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 1995 | 4171.339 | 1.133 | 108954.723 | 0.006 | 0.079 |
| 1996 | 3700.400 | 1.005 | 59000.363 | 0.004 | 0.066 |
| 1997 | 3793.778 | 1.030 | 62175.636 | 0.004 | 0.066 |
| 1998 | 3890.218 | 1.056 | 83518.738 | 0.006 | 0.074 |
| 1999 | 3419.423 | 0.928 | 56573.751 | 0.005 | 0.070 |
| 2000 | 3235.253 | 0.878 | 57888.952 | 0.006 | 0.074 |
| 2001 | 2947.962 | 0.800 | 130461.360 | 0.015 | 0.123 |
| 2002 | 3411.078 | 0.926 | 62878.499 | 0.005 | 0.074 |
| 2003 | 4145.339 | 1.126 | 56898.996 | 0.003 | 0.058 |
| 2004 | 4371.503 | 1.187 | 63567.812 | 0.003 | 0.058 |
| 2005 | 12519.564 | 1.069 | 336101.196 | 0.002 | 0.046 |
| 2006 | 12648.627 | 1.080 | 262528.825 | 0.002 | 0.041 |
| 2007 | 12145.212 | 1.037 | 276246.442 | 0.002 | 0.043 |
| 2008 | 13834.526 | 1.182 | 294798.093 | 0.002 | 0.039 |
| 2009 | 18360.125 | 1.568 | 423396.136 | 0.001 | 0.035 |
| 2010 | 12742.844 | 1.088 | 271918.180 | 0.002 | 0.041 |
| 2011 | 12819.358 | 1.095 | 325347.368 | 0.002 | 0.044 |
| 2012 | 12472.968 | 1.065 | 247042.061 | 0.002 | 0.040 |
| 2013 | 8698.067 | 0.743 | 308718.510 | 0.004 | 0.064 |
| 2014 | 8667.031 | 0.740 | 250151.136 | 0.003 | 0.058 |
| 2015 | 8566.046 | 0.732 | 262916.058 | 0.004 | 0.060 |
| 2016 | 10509.029 | 0.898 | 247759.318 | 0.002 | 0.047 |
| 2017 | 12543.516 | 1.071 | 258282.878 | 0.002 | 0.041 |
| 2018 | 15738.039 | 1006.568 | 0.344 | 298711.438 | 0.001 |

## c. Commercial fishery CPUE index by non-interaction model

We fitted the negative binomial GLM model for fish ticket retained CPUE time series 1985/86 1998/99 offering Year, Month, Vessel, Captain, and Area as explanatory variables and applying the hybrid selection method. Reduced area resolution (grouped ADF\&G codes to AreaGP) was used for model fitting.

The final model for EAG was:

```
Initial selection by stepAIC:
ln(CPUE) = Year + Vessel + Month
AIC=16,996
```

Final selection by stepCPUE:
$\ln ($ CPUE $)=$ Year + Vessel + Month
for the 1985/86-1998/99 period $\left[\theta=10.40, \mathrm{R}^{2}=0.3327\right]$
and that for WAG was:
Initial selection by stepAIC:
$\ln ($ CPUE $)=$ Year + Vessel + Area
AIC=31,701
Final selection by stepCPUE:
$\ln ($ CPUE $)=$ Year + Vessel + Area
for the 1985/86-1998/99 period [ $\theta=6.67, \mathrm{R}^{2}=0.3569$ ]

## Appendix C. Cooperative Survey

## 1. Summary of the survey method

The ADF\&G and industry collaborative pot survey was initiated in 2015 in the EAG and has continued since then. The survey was extended to WAG in 2018. A stratified two-stage sampling design has been implemented in a $2 \mathrm{nmi} \times 2 \mathrm{nmi}$ grids within 1000 m depth covering the entire golden king crab fishing area. The $2 \mathrm{nmi} \times 2 \mathrm{nmi}$ choice was the best compromise between scale of fishing gear, accuracy of defining habitat, and number of possible stations (Figure C1).


Figure C.1. Survey design: 2 nmi x 2 nmi grids overlaid on observer pot sample locations (green squares) in EAG.

There are nearly 1100 grids in the EAG divided into three equal size strata for selecting random pot sampling locations (Figures C. 2 and C.3).


Figure C.2. Survey design: $2 \mathrm{nmi} \times 2 \mathrm{nmi}$ grids stratified by three equal sizes for selecting random pot sampling locations in EAG.


Figure C.3. Random sample of 22 cells selected in each of three sub strata in EAG during the 2019 fishery.

Surveys occur during the first month of each fishing season with one to two ADF\&G biologists onboard the fishing vessel to collect fishery and biological data. Fishing operation takes place in a randomly selected set of grids in each stratum with long-line pots. The number of pots per string ranges from 30 to $40,200 \mathrm{~m}$ apart, and a vessel carries on average 35 strings. Pot sizes range from $5.5 \mathrm{ft} \times 5.5 \mathrm{ft}$ to $7 \mathrm{ft} \times 7 \mathrm{ft}$ with large mesh sizes for retention of legal-sized king crab. A few small mesh size research pots are also deployed for special studies. Fishing operation is not standardized for depth or soak time to allow normal fishing practices.

There are multiple pots (typically about 5 pots) sampled for each long-line string with approximately 35 crab measurement made per pot. For example, if 100 crabs are caught in a sampled pot, the biologist measures every third crab. The following snapshot of an observation record provides an example of what stock assessment data are collected.

| fishery | year | vessel | skipper | String\# | pot_size | mesh_size | bait | subsample_rate | species_code | sex | size | legal |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| EAG | 2015 | 20556 | Chad_Hoefer | 1 | 5x5 | king(large) | halibut | 2 | 923 | 1 | 187 | 1 |


| Pot\# | date_in | time_in | depth_start | start_lat | start_lon | depth_out | end_lat | end_lon | date_out | time_out | comments | soak_time |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | $8 / 4 / 2015$ | $17: 00$ | 132 | 52.74133 | -170.692 | 133 | 52.7515 | -170.675 | $8 / 17 / 2015$ | $3: 00$ |  |  |

## 2. Standardization of cooperative survey CPUE

## Data

A unique property of the cooperative survey is that multiple pots from multiple strings are sampled. All sample measurements were taken in EAG except for 2018/19, during which measurements were also taken from WAG. There was no survey during 2020/21 due to COVID related restriction.

There are 27,255 records from five years (2015-2019) of surveys. After cleaning up for missing entries, the number of records reduced to 27,122 golden king crab.

## Method

Data preparation for CPUE standardization:
i.) Created two new columns by concatenating Vessel Code with String\# as well as with String\# and Pot\# because String\# and Pot\# are not unique numbers to each vessel. The new column names were identified as VesString and VesStingPot. For example, a Vessel Code 20556 with a String\# 3 was concatenated to be 205563 in a new column VesString, and a Vessel Code 20556 with a String\#17 and a Pot\# 5 was concatenated to be 20556175 in a new column VesStringPot.
ii.) Raised the Catch in each record by the Sample Rate.
iii.) Subset the data by large mesh king crab pot [Mesh ID not equal to 2 (i.e., small mesh pot)], legal size ( $>135 \mathrm{~mm}$ CL), and EAG (EAGWAG=1). The female $(S e x=2)$ and unclassified catch without any male crab $(S e x=1)$ in a crab pot was set to 0 to account for the possibility of zero catch for expected male CPUE determination.
iv.) Further subset the data by $5 \%$ to $95 \%$, trimmed Soak time, and $1 \%$ to $99 \%$ trimmed Depth. This is to exclude catches from any unusual pot operations.
v.) Summed up the catch across sizes for each Pot\# and labelled it as SumCatch. Thus, each Pot\# has a single catch number.

The sampling design (sampling crab from a pot within a string within a vessel) begged for application of a mixed effects model to analyze data, which was recommended by the CPT. However, we explored different model structures before finalizing on a model: a fixed effect model and two versions of a random effects model. The dispersion parameter value for the negative binomial error model and the degrees of freedom for cubic splines for soak time and depth were borrowed from the observer final GLM model estimates for EAG for the post rationalization period, 2005-2020.

## Results

1. Fixed effect model:

Sum Catch $=\mathrm{Y}$, family $=$ negative binomial $(\theta=2.32)$

The maximum set of model terms offered to the stepwise selection procedure was:
Sum Catch $=\mathrm{Y}+\mathrm{ns}($ Soak, $\mathrm{df}=3)+$ VesselStringPot + Captain + Block $+\mathrm{ns}($ Depth, $\mathrm{df}=10)$, family $=$ negative binomial $(\theta=2.32)$.

Final model:
Sum Catch $=\mathrm{Y}+$ VesselStringPot $+\mathrm{ns}($ Depth, $\mathrm{df}=10)+$ Block $+\mathrm{ns}($ Soak, $\mathrm{df}=3)$
$R^{2}=0.6088$ (Soak forced in).
2. Random intercept model (model 1):

Sum Catch $=\mathrm{Y}+\mathrm{ns}($ Depth, $\mathrm{df}=10)+\mathrm{ns}($ Soak, $\mathrm{df}=3)+$
$(1 \mid$ Vessel/VesStringPot $)+(1 \mid$ Block/VesselString $)$ family $=$ negative binomial $(\theta=2.32)$.
We selected relevant fixed effect components from the final fixed effect model for the random intercept models 1 and 2. We used the "lme4" library in R (R Core Team, 2020) with the "glmer()" function for model fitting. The glmer() function allows use of any type of error model to fit the data. The random intercept model 1 resulted in a singular fit (i.e., Vessel and VesStringPot:Vessel group variances were (very close to) zero):

Table C.1. Random intercept model 1 output.

| Groups | Name | Variance | Std.Dev. |
| :--- | :--- | :---: | :---: |
| VesStringPot:Vessel | (Intercept) | 0.00000 | 0.00000 |
| VesString:Block | (Intercept) | 0.35685 | 0.59737 |
| Block | (Intercept) | 0.00059 | 0.02439 |
| Vessel | (Intercept) | 0.00000 | 0.00000 |

3. Therefore, we used the following simpler form of the random intercept model (model 2):

Sum Catch $=\mathrm{Y}+\mathrm{ns}($ Depth, $\mathrm{df}=10)+\mathrm{ns}($ Soak, $\mathrm{df}=3)+(1 \mid$ Block $/$ VesselString $)$ family $=$ negative binomial $(\theta=2.32)$.

The random intercept model 2 converged with the following output:
Table C.2. Random intercept model 2 parameter estimates.
Random Effects:
Groups Name Variance Std.Dev.
VesString:Block (Intercept) 0.35690 .5974
Block (Intercept) 0.0006

Fixed Effects:

|  | Estimate | Std. Error | $z_{-}$value | $\operatorname{Pr}(\|z\|)$ |
| :--- | :--- | :--- | :--- | :--- |
| Intercept | 3.0426 | 0.2776 | 10.959 | 0.0000 |
| Year2016 | -0.2952 | 0.1005 | -2.937 | 0.0033 |
| Year2017 | -0.0621 | 0.1107 | -0.561 | 0.5748 |


| Year2018 | 0.0963 | 0.1060 | 0.909 | 0.3634 |
| :--- | :--- | :--- | :--- | :--- |
| Year2019 | -0.3591 | 0.1052 | -3.415 | 0.0006 |
| ns(Depth, $D F=10) 1$ | 0.6944 | 0.2399 | 2.895 | 0.0038 |
| ns(Depth, DF=10)2 | 0.3130 | 0.3152 | 0.993 | 0.3206 |
| ns(Depth, DF=10)3 | 0.0967 | 0.2740 | 0.353 | 0.7241 |
| ns(Depth, DF=10)4 | 0.4526 | 0.3268 | 1.385 | 0.1661 |
| ns(Depth, DF=10)5 | 0.0541 | 0.3188 | 0.170 | 0.8653 |
| ns(Depth, $\mathrm{DF}=10) 6$ | 0.0146 | 0.3219 | 0.045 | 0.9639 |
| ns(Depth, DF=10)7 | 0.7676 | 0.3270 | 2.348 | 0.0189 |
| ns(Depth, DF=10)8 | -0.0894 | 0.2574 | -0.347 | 0.7285 |
| ns(Depth, DF=10)9 | 0.3136 | 0.6123 | 0.512 | 0.6085 |
| ns(Depth, DF=10)10 | 0.9769 | 0.2983 | 3.275 | 0.0011 |
| ns(Soak, DF=3)1 | 0.2857 | 0.1638 | 1.744 | 0.0811 |
| ns(Soak, DF=3)2 | 0.3628 | 0.4337 | 0.836 | 0.4029 |
| ns(Soak, DF=3)3 | 0.7725 | 0.2448 | 3.156 | 0.0016 |

Inadequate time series (2015-2019) with fewer random effect levels (only three levels for Vessel and, somewhat better, four levels for Block) prevented us from exploring expanded model structures, such as a random intercept with random slope model. Categorical variable levels above 5 is recommended to be ideal for determining variances of the distribution of random effect factors (Gelman and Hill 2007). Comparison of the random intercept model 2 (C.3) with that of the fixed effects model (C.1) by the Hausman's (1978) model selection test resulted in rejecting the null hypothesis that random effect model is consistent with the data (Chi Square $=1124.4, \mathrm{df}=18, \mathrm{p}<$ $2.2 \mathrm{e}^{-16}$ ). However, because of limiting factors discussed above that could spoil any statistical test, we based our selection of random effects model 2 on the sampling design (i.e., multi-level sampling) implemented in data collection.

There is a plan to continue the cooperative survey in 2021/22, which will increase the time series of data to 6 years. Note that we do not have a flexibility to increase the number of Vessel levels from three but do have flexibility to increase the number of Block levels from four. Therefore, we intend to increase the number of Block levels by defining smaller areas in EAG for the next round of analysis.

## Diagnostic test

The QQ plot for the fit assured that model 2 assumptions were correct (Figure C.4).

## Random Effects Model 2 Fit, Cooperative Survey 2015-2019



Figure C.4. Studentized residual plot for the mixed random effects model fit using the 20152019 EAG data.

Comparison of standardized CPUE from cooperative survey data (2015-19) for EAG and the corresponding years' observer CPUE indices indicated a similar pattern except for 2019 (Figure C.5).


Figure C.5. Comparison of cooperative survey random effects model 2 CPUE indices (green) and observer non interaction factor model CPUE indices (red, 19.1) for EAG. The confidence limits are determined with $\pm 2$ SE. Model estimated additional standard error was added to SE.

We standardized the yearly mean of predicted survey CPUEs for 2015-2019 by the geometric mean to obtain the CPUE indices for input to the assessment model (Table C.3).

Table C.3. The cooperative survey predicted legal male standardized (by geometric mean) CPUE indices by the mixed random effects model 2, standard errors (SE), and lower- and upper- $95 \%$ confidence limits with added model estimated additional standard error for EAG, 2015-2019 data.

| Year | Predicted CPUE <br> index | SE | Lower <br> Limit | Upper <br> Limit | Sample <br> size |
| :---: | :---: | :---: | :---: | :---: | ---: |
| 2015 | 1.11164 | 0.02666 | 0.60593 | 2.03943 | 335 |
| 2016 | 0.94664 | 0.02656 | 0.51610 | 1.73636 | 304 |
| 2017 | 1.10698 | 0.04148 | 0.58576 | 2.09200 | 206 |
| 2018 | 1.17588 | 0.03565 | 0.62952 | 2.19642 | 199 |
| 2019 | 0.73004 | 0.03749 | 0.38940 | 1.36867 | 289 |

We added a likelihood function with the 2015-2019 survey indices using Equations A. 12 and A. 13 to the likelihoods of observer indices (1995-2014) and fishery indices (1985-1998) and formulated a new model 21.1d. We maintained the same post-rationalization fishery catchability, total and retained selectivity for fitting survey indices. The reference points estimates were like those of 21.1a but a little lower.

## EAG (Tier 3):

Biomass, total OFL, and ABC for the next fishing season in millions of pounds. Current MMB $=$ MMB on 15 Feb. 2022

| Model | Tier | $M M B_{35 \%}$ | Current <br> MMB | MMB/ <br> $M M B_{35 \%}$ | $F_{\text {OFL }}$ | Recruitment Years <br> to define $M M B_{35 \%}$ | $F_{35 \%}$ | OFL | ABC <br> $\left(\mathrm{P}^{*}=0.49\right)$ | ABC <br> $(0.75 * O F L)$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| EAG21.1d | 3 a | 14.686 | 17.823 | 1.21 | 0.62 | $1987-2017$ | 0.62 | 5.822 | 5.791 | 4.367 |

Biomass in 1000 t ; total OFL and ABC for the next fishing season in t .

| Model | Tier | $M M B_{35 \%}$ | Current <br> MMB | $\begin{gathered} M M B / \\ M M B_{35 \%} \end{gathered}$ | $F_{\text {OFL }}$ | Recruitment Years <br> to Define MMB35\% | $F_{35 \%}$ | OFL | $\begin{gathered} A B C \\ \left(P^{*}=0.49\right) \end{gathered}$ | $\begin{gathered} A B C \\ (0.75 * O F L) \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| EAG21.1d | 3a | 6661.73 | 8084.57 | 1.21 | 0.62 | 1987-2017 | 0.62 | 2,641.064 | 2,626.844 | 1,980.798 |

Figure C. 6 provides the long-term trends in MMB by model 21.1d with the state quo knife-edge maturity size of 111 mm CL.


Figure C.6. Comparison of trends in golden king crab mature male biomass between models 21.1a and 21.1d for EAG, 1961-2021. Year 2021 refers to 2020/21 fishing season.

## Appendix D: Male Maturity

## Introduction

Sexual maturity is associated with alterations in both external morphology and internal physiology, on which bases different types of maturity can be defined: physiological, morphometric, and functional maturity. Although functional maturity is the true way of determining maturity, it requires elaborate lab or field experiments. Hence, crab researchers often adapt an indirect detection technique via morphometric measurement for maturity determination. Chelae allometry has been used to determine morphometric male size-at-maturity among several king crab (Lithodidae) stocks. Male golden king crab provides a better discrimination of chelae height against size at onset of maturity than other king crab stocks (Somerton and Otto 1986). Table D. 1 lists the literature reported estimates of size-at-maturity of male golden king crab (Lithodes aequispins) stocks in Alaska. Breakpoint analysis has been used to estimate maturity in majority of cases.

Table D.1. Review of estimates of male size-at-maturity of golden (Lithodes aequispins) king crabs by regions in Alaska. Numbers in parentheses are standard errors (SE).

| Species | Sex | Size-atMaturity (mm CL) | Method | Area | Sources |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Lithodes aequispins | Male | 114 (11.4) | Breakpoint analysis on $\log$ (chela height) vs. $\log$ (carapace length) | British Columbia, Canada | Jewett et al. 1985 |
|  |  | 92 (2.4) | Breakpoint analysis on | St. Matthew Is. | Somerton and Otto |
|  |  | 107 (4.6) | $\log$ (chela height) vs. | District | 1986 |
|  |  | 130 (4.0) | $\log$ (carapace length) | Pribilof Is District Eastern Aleutian Is |  |
|  |  | $\begin{aligned} & 117.9 \text { to } \\ & 158.0 \end{aligned}$ | Breakpoint analysis on $\log$ (chela height) vs. $\log$ (carapace length) | Various water inlets in southeast Alaska | Olson 2016 |
|  |  | 108.6 (2.6) | Breakpoint analysis on | Bowers Ridge | Otto and |
|  |  | 120.8 (2.9) | $\log$ (chela height) vs. $\log$ (carapace length) | Seguam Pass | Cummiskey 1985 |
|  |  | 110 | Minimum size of successful mating (lab observation) | Prince William Sound | Paul and Paul 2001 |

## Method

We used the carapace length ( mm CL ) and chela height (up to one-tenth of a mm CH ) data collected by the observer, retained catch, and cooperative surveys sampling during the 2018/19, 2019/20, and 2020/21 fishing seasons for male maturity investigation. We determined bend points and corresponding two segmented lines for different groups of data outside the assessment model using the 'segmented regression' package available in R version 3.6.3 (R Core Team 2020).

First, we fitted a linear regression model to the data pair using the R package as follows:
$\ln (C H / C L)=\beta_{0}+\beta_{1} C L$
where $\beta_{0}$ and $\beta_{1}$ are regression parameters
The procedure of 'segmented regression' uses maximum likelihood to fit a somewhat different parameterization of the linear model. It can be approximated as $\ln (C H / C L)=\beta_{0}+\beta_{1} C L+\beta_{2}[C L-c]+\gamma I[C L>c]$
where $\beta_{2}$ is a regression parameter and c is the break-point, and $\gamma I[C L>c]$ is a dummy variable. When $C L<c$, the model reduces to,
$\ln (C H / C L)=\beta_{0}+\beta_{1} C L+\beta_{2}[C L-c]$
The $\gamma$ term is a measure of the distance between the end of the first segment and the beginning of the next. The model converges when $\gamma$ is minimized, thus this method constrains the segments to be (nearly) continuous.

We further refined the estimates by bootstrapping each data set (ln (CH/CL), CL pairs) 1000 times and applying 'segmented regression' to each bootstrapped sample. We used the bootstrap median bend point, intercept, and slope estimates to establish the two segmented lines (i.e., left hand line 1 was for immature and right-hand line 2 was for mature crab).

Finally, we categorized the observed $\ln (\mathrm{CH} / \mathrm{CL})$ vs CL pair in to mature (code 1$)$ if the $\ln (\mathrm{CH} / \mathrm{CL})$ value was on or above line 2 or immature (code 0 ) if this value was below line 2 for a given CL. Then we fitted the following logistic model by GLM to the response variable (mature 1, immature 0 , converted to mature proportion $P$ ) vs. the independent variable (CL) to obtain a maturity curve.
$P_{i}=\frac{e^{\alpha+\beta C L_{i}}}{1+e^{\alpha+\beta C L_{i}}}$
where $P$ is the proportion mature, $\alpha$ and $\beta$ are intercept and slope parameters of the maturity curve, and CL is carapace length.

## Data

We used the following data sets (Table D.2) for current maturity analysis. We restricted the size range to 85.0 mm CL to 142.0 mm CL (a plausible morphometric male maturity size range for golden king crab in the Bering Sea and Aleutian Islands (i.e., 92.0 mm CL - 3SE to 130 mm CL +3 SE, Table D.1) for 'segmented regression' fit and maturity determination from segmented line 2 for subsequent logistic model fit.

Table D.2. Golden king crab male carapace length and chela height data collected during 2018/19 - 2020/21 fishing seasons in the Aleutian Islands.

| Measurement type | Source and season of data collection | Group |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  |  | Aleutian Islands (AI) 2018/192020/21 | $\begin{aligned} & \hline \text { EAG } \\ & 2018 / 19- \\ & 2019 / 20 \end{aligned}$ | $\begin{aligned} & \hline \text { WAG } \\ & 2018 / 19- \\ & 2020 / 21 \end{aligned}$ |
|  | Co-operative survey (2018/19, 2019/20) <br> Observer sampling (2018/19, 2019/20) <br> Retained catch sampling (2018/19, 2019/20) <br> Special sampling WAG (2020/21) |  |  |  |
| Carapace length and chela height records (all sizes) |  | 10760 | 5433 | 5327 |
| Carapace length and chela height records ( 85 mm CL-142 mm CL) |  | 4025 | 1901 | 2124 |

## Results

The median breakpoint ranged from 117.800 to 119.984 mm CL for the three 2018-2020 data sets (AI, EAG, and WAG). These values are one to two 5 mm CL bins higher than the WAG 1984 and EAG 1991 estimates considered previously (Table D.3). The focus of the current analysis is to establish separate maturity curves for AI, EAG, and WAG for MMB determination. Table D. 4 lists the logistic maturity curve parameter estimates for AI, EAG, and WAG. The estimates for the three data sets are highly significant. We considered three options for MMB estimation: $\geq 111$ mm CL (status quo knife edge maturity), $\geq 116 \mathrm{~mm}$ CL (a cautionary choice based on the current determination of $50 \%$ maturity length by the logistic model for AI), and the AI logistic maturity curve common to both EAG and WAG.

Figures D.1, D.2, and D. 3 show the fitted segmented regression lines overlayed on observed log (CH/CL) vs CL data for AI, EAG, and WAG, respectively. Figures D.4, D.5, and D. 6 show the fitted logistic curves overlayed on observed maturity proportion vs CL for AI, EAG, and WAG, respectively. The observed mature (1) and immature (0) counts at each 2 mm CL interval were used to calculate the observed mature proportion by size. The observed maturity data were hard cutoff to be immature below 85 mm CL and mature above 142 mm CL (i.e., outside the data range considered for segmented regression fitting) for logistic model fitting. Although arbitrary, but reasonable, data massaging occurred before logistic model fitting. We considered that the fitted
logistic curve (with highly significant parameter estimates) was appropriate for MMB determination.

Table D.3. Segment regression fit to 2018-2020 $\log (C H / C L)$ vs. CL data pairs and median estimates from 1000 bootstrap samples including breakpoints for EAG, WAG, and combined (AI) data sets. The data sets were truncated to $85.0-142.0 \mathrm{~mm}$ CL range for segmented regression fits. We also provide re-estimated parameters for the 1991 EAG and 1984 WAG data with the same size restriction for comparison. Intercept of line 2 was determined for each data set by solving the two lines at the bend point.

|  | Estimate | SE | t -value | $\operatorname{Pr}(>\|\mathrm{t}\|)$ | Breakpoint <br> $(\mathrm{mm} \mathrm{CL})$ | SE | Remarks |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| AI 2018- |  |  |  |  |  | Fit to original data, 85.0- <br> 20: |  |
|  |  |  |  | 142.0 mm CL range |  |  |  |


| EAG 1991: |  |  |  |  |  |  | Fit to $85.0-142.0 \mathrm{~mm} \mathrm{CL}$ range <br> Estimates from 2457 measurements |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept |  |  |  |  | 107.000 | 1.915 |  |
| line 1 | -1.60166 | 0.02286 | -70.04911 | 0.00000 |  |  |  |
| Slope line 1 | 0.00070 | 0.00026 | 2.71486 | 0.00668 |  |  |  |
| Slope to |  |  |  |  |  |  |  |
| add | 0.00424 | 0.00029 | 14.45235 | 0.00000 |  |  |  |
| Slope line 20.00494 |  |  |  |  |  |  |  |
| WAG |  |  |  |  |  |  | Fit to $85.0-142.0 \mathrm{~mm} \mathrm{CL}$ |
| 1984: |  |  |  |  |  |  | range |
| Intercept |  |  |  |  | 105.824 | 4.650 | Estimates from 341 |
| line 1 | -1.67570 | 0.09222 | -18.17129 | 0.00000 |  |  | measurements |
| Slope line 1 | 0.00126 | 0.00097 | 1.29613 | 0.19582 |  |  |  |
| Slope to |  |  |  |  |  |  |  |
| add | 0.00332 | 0.00106 | 3.12254 | 0.00195 |  |  |  |
| Slope line 2 | 0.00458 |  |  |  |  |  |  |

Table D.4. Logistic fit to 2018-2020 maturity data for AI, EAG, and WAG.

|  | Estimate | SE | z -value | $\operatorname{Pr}(>\|\mathrm{z}\|)$ | $\mathrm{L}_{50}(\mathrm{~mm}$ <br> $\mathrm{CL})$ | Remarks |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| AI 2018-20: |  |  |  |  | Logistic model fit to <br> $52.6-198.7$ mm CL range, |  |
|  |  |  |  |  | 10760 measurements |  |

Figures D.1, D. 2 and D. 3 provide the segment regression lines fitted to the $\log (\mathrm{CH} / \mathrm{CL})$ vs. CL data pairs and D.4, D.5, and D. 6 depict the logistic curves fitted to mature proportion vs. CL data pairs for 2018-2020 in AI, EAG, and WAG, respectively:


Figure D.1. Segmented linear regression fit to $\log (\mathrm{CH} / \mathrm{CL})$ vs. CL data of male golden king crab for 2018-2020 in AI.


Figure D.2. Segmented linear regression fit to $\log (\mathrm{CH} / \mathrm{CL})$ vs. CL data of male golden king crab for 2018-2020 in EAG.


Figure D.3. Segmented linear regression fit to $\log (\mathrm{CH} / \mathrm{CL})$ vs. CL data of male golden king crab for 2018-2020 in WAG.


Figure D.4. Logistic fit to mature proportion of male golden king crab for 2018-2020 in AI.


Figure D.5. Logistic fit to mature proportion of male golden king crab for 2018-2020 in EAG.


Figure D.6. Logistic fit to mature proportion of male golden king crab for 2018-2020 in WAG.

## Implication on mature male biomass estimation:

Figure D. 7 provides the long-term trends in MMB by models 21.1a, 21.1b, and 21.1c with varying maturity assumptions: status quo knife-edge maturity size of $111 \mathrm{~mm} \mathrm{CL}(\ldots 1 \mathrm{a}, \ldots 1 \mathrm{~b}, \ldots 1 \mathrm{c})$, higher maturity size of $116 \mathrm{~mm} \mathrm{CL}(\ldots 1 \mathrm{a} 1, \ldots 1 \mathrm{bl}, \ldots 1 \mathrm{c} 1)$, and logistic maturity curve (...1a2, ...1b2, ...1c2). Changes from status quo maturity assumption generally result in lower MMB values.


Figure D.7. Trends in golden king crab mature male biomass for models 21.1a, 21.1a1, 21.1a2, 21.1b1, 21.1b2, 21.1c1, and 21.1c2 fits to EAG (left) and WAG (right) data, 1961-2021. Model 21.1a estimate has two standard error confidence limits. Year 2021 refers to 2020/21 fishing season.

## Appendix E: Jittering

## Jittering of model 19.1 parameter estimates

We followed the Stock Synthesis approach to do 100 jitter runs of model 19.1 parameter estimates to use as initial parameter values (as .PIN file in ADMB) to assess model stability and to determine whether a global, as opposed to local, minima has been reached by the search algorithm:

Following CPT suggestion, we increased the jittering to $50 \%$ from previously used $30 \%$. A Jitter factor of 0.5 was multiplied by a random normal deviation $r d e v=N(0,1)$ to create a transformed parameter value based upon the predefined parameter:

$$
\begin{equation*}
\text { temp }=0.5 * \text { rdev* Jitterfactor } * \ln \left(\frac{P_{\max }-P_{\min }+0.0000002}{P_{v a l}-P_{\min }+0.0000001}-1\right) \tag{E.1}
\end{equation*}
$$

with the final jittered initial parameter value back transformed as:

$$
\begin{equation*}
P_{\text {new }}=P_{\min }+\frac{P_{\max }-P_{\min }}{1.0+\exp (-2.0 \text { temp })} \tag{E.2}
\end{equation*}
$$

where $P_{\max }$ and $P_{\min }$ are upper and lower bounds of parameter search space and $P_{\text {val }}$ is the estimated parameter value before the jittering.

The jitter results are summarized for scenario 19.1 in Tables E. 1 and E. 2 for EAG and WAG, respectively. All runs converged to the highest $\log$ likelihood values except for nonconvergent runs. We concluded from jitter results that optimization of 19.1 model achieved global minima for both EAG and WAG.

Table E.1. Results from 100 jitter runs for scenario 19.1 for EAG. Jitter run 0 corresponds to the original optimized estimates. NA: model did not converge for run\#3.

| Jitter Run | $\begin{array}{l}\text { Objective } \\ \text { Function }\end{array}$ |  | $\begin{array}{l}\text { Maximum } \\ \text { Gradient }\end{array}$ |  | $\mathrm{B}_{35 \%}(\mathrm{t})$ |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | \(\left.$$
\begin{array}{rrrrr}\text { OFL (t) }\end{array}
$$ \begin{array}{l}Current <br>

MMB (t)\end{array}\right]\)

| 13 | -94.08746 | 0.000152 | 6,696 | 2,935 | 8,722 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 14 | -94.08746 | 0.000075 | 6,696 | 2,935 | 8,722 |
| 15 | -94.08746 | 0.000166 | 6,696 | 2,935 | 8,722 |
| 16 | -94.08746 | 0.000060 | 6,696 | 2,935 | 8,722 |
| 17 | -94.08746 | 0.000024 | 6,696 | 2,935 | 8,722 |
| 18 | -94.08746 | 0.000153 | 6,696 | 2,935 | 8,722 |
| 19 | -94.08746 | 0.000202 | 6,696 | 2,935 | 8,722 |
| 20 | -94.08746 | 0.000020 | 6,696 | 2,935 | 8,722 |
| 21 | -94.08746 | 0.000078 | 6,696 | 2,935 | 8,722 |
| 22 | -94.08746 | 0.001453 | 6,696 | 2,935 | 8,722 |
| 23 | -94.08746 | 0.000478 | 6,696 | 2,935 | 8,722 |
| 24 | -94.08746 | 0.000241 | 6,696 | 2,935 | 8,722 |
| 25 | -94.08746 | 0.000309 | 6,696 | 2,935 | 8,722 |
| 26 | -94.08746 | 0.000128 | 6,696 | 2,935 | 8,722 |
| 27 | -94.08746 | 0.000054 | 6,696 | 2,935 | 8,722 |
| 28 | -94.08746 | 0.000259 | 6,696 | 2,935 | 8,722 |
| 29 | -94.08746 | 0.000169 | 6,696 | 2,935 | 8,722 |
| 30 | -94.08746 | 0.000145 | 6,696 | 2,935 | 8,722 |
| 31 | -94.08746 | 0.000079 | 6,696 | 2,935 | 8,722 |
| 32 | -94.08746 | 0.000033 | 6,696 | 2,935 | 8,722 |
| 33 | -94.08746 | 0.001445 | 6,696 | 2,935 | 8,722 |
| 34 | -94.08746 | 0.000158 | 6,696 | 2,935 | 8,722 |
| 35 | -94.08746 | 0.000159 | 6,696 | 2,935 | 8,722 |
| 36 | -94.08746 | 0.000389 | 6,696 | 2,935 | 8,722 |
| 37 | -94.08746 | 0.000111 | 6,696 | 2,935 | 8,722 |
| 38 | -94.08746 | 0.000243 | 6,696 | 2,935 | 8,722 |
| 39 | -94.08746 | 0.000041 | 6,696 | 2,935 | 8,722 |
| 40 | -94.08746 | 0.000098 | 6,696 | 2,935 | 8,722 |
| 41 | -94.08746 | 0.000045 | 6,696 | 2,935 | 8,722 |
| 42 | -94.08746 | 0.000689 | 6,696 | 2,935 | 8,722 |
| 43 | -94.08746 | 0.000013 | 6,696 | 2,935 | 8,722 |
| 44 | -94.08746 | 0.000250 | 6,696 | 2,935 | 8,722 |
| 45 | -94.08746 | 0.000222 | 6,696 | 2,935 | 8,722 |
| 46 | -94.08746 | 0.000114 | 6,696 | 2,935 | 8,722 |
| 47 | -94.08746 | 0.000306 | 6,696 | 2,935 | 8,722 |
| 48 | -94.08746 | 0.000029 | 6,696 | 2,935 | 8,722 |
| 49 | -94.08746 | 0.000175 | 6,696 | 2,935 | 8,722 |
| 50 | -94.08746 | 0.000231 | 6,696 | 2,935 | 8,722 |
| 51 | -94.08746 | 0.000163 | 6,696 | 2,935 | 8,722 |
| 52 | -94.08746 | 0.000145 | 6,696 | 2,935 | 8,722 |
| 53 | -94.08746 | 0.000192 | 6,696 | 2,935 | 8,722 |
| 54 | -94.08746 | 0.000051 | 6,696 | 2,935 | 8,722 |
| 55 | -94.08746 | 0.000281 | 6,696 | 2,935 | 8,722 |
| 56 | -94.08746 | 0.000124 | 6,696 | 2,935 | 8,722 |


| 57 | -94.08746 | 0.000162 | 6,696 | 2,935 | 8,722 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 58 | -94.08746 | 0.000141 | 6,696 | 2,935 | 8,722 |
| 59 | -94.08746 | 0.000254 | 6,696 | 2,935 | 8,722 |
| 60 | -94.08746 | 0.000045 | 6,696 | 2,935 | 8,722 |
| 61 | -94.08746 | 0.000005 | 6,696 | 2,935 | 8,722 |
| 62 | -94.08746 | 0.000354 | 6,696 | 2,935 | 8,722 |
| 63 | -94.08746 | 0.000068 | 6,696 | 2,935 | 8,722 |
| 64 | -94.08746 | 0.000124 | 6,696 | 2,935 | 8,722 |
| 65 | -94.08746 | 0.000057 | 6,696 | 2,935 | 8,722 |
| 66 | -94.08746 | 0.000015 | 6,696 | 2,935 | 8,722 |
| 67 | -94.08746 | 0.000088 | 6,696 | 2,935 | 8,722 |
| 68 | -94.08746 | 0.000195 | 6,696 | 2,935 | 8,722 |
| 69 | -94.08746 | 0.000204 | 6,696 | 2,935 | 8,722 |
| 70 | -94.08746 | 0.000623 | 6,696 | 2,935 | 8,722 |
| 71 | -94.08746 | 0.000090 | 6,696 | 2,935 | 8,722 |
| 72 | -94.08746 | 0.000083 | 6,696 | 2,935 | 8,722 |
| 73 | -94.08746 | 0.000031 | 6,696 | 2,935 | 8,722 |
| 74 | -94.08746 | 0.000336 | 6,696 | 2,935 | 8,722 |
| 75 | -94.08746 | 0.000113 | 6,696 | 2,935 | 8,722 |
| 76 | -94.08746 | 0.000029 | 6,696 | 2,935 | 8,722 |
| 77 | -94.08746 | 0.000065 | 6,696 | 2,935 | 8,722 |
| 78 | -94.08746 | 0.000238 | 6,696 | 2,935 | 8,722 |
| 79 | -94.08746 | 0.000034 | 6,696 | 2,935 | 8,722 |
| 80 | -94.08746 | 0.000134 | 6,696 | 2,935 | 8,722 |
| 81 | -94.08746 | 0.000152 | 6,696 | 2,935 | 8,722 |
| 82 | -94.08746 | 0.000230 | 6,696 | 2,935 | 8,722 |
| 83 | -94.08746 | 0.000141 | 6,696 | 2,935 | 8,722 |
| 84 | -94.08746 | 0.000142 | 6,696 | 2,935 | 8,722 |
| 85 | -94.08746 | 0.000147 | 6,696 | 2,935 | 8,722 |
| 86 | -94.08746 | 0.000151 | 6,696 | 2,935 | 8,722 |
| 87 | -94.08746 | 0.000146 | 6,696 | 2,935 | 8,722 |
| 88 | -94.08746 | 0.000093 | 6,696 | 2,935 | 8,722 |
| 89 | -94.08746 | 0.000118 | 6,696 | 2,935 | 8,722 |
| 90 | -94.08746 | 0.000073 | 6,696 | 2,935 | 8,722 |
| 91 | -94.08746 | 0.000041 | 6,696 | 2,935 | 8,722 |
| 92 | -94.08746 | 0.000050 | 6,696 | 2,935 | 8,722 |
| 93 | -94.08746 | 0.000034 | 6,696 | 2,935 | 8,722 |
| 94 | -94.08746 | 0.000038 | 6,696 | 2,935 | 8,722 |
| 95 | -94.08746 | 0.000108 | 6,696 | 2,935 | 8,722 |
| 96 | -94.08746 | 0.000006 | 6,696 | 2,935 | 8,722 |
| 97 | -94.08746 | 0.000079 | 6,696 | 2,935 | 8,722 |
| 98 | -94.08746 | 0.000236 | 6,696 | 2,935 | 8,722 |
| 99 | -94.08746 | 0.000039 | 6,696 | 2,935 | 8,722 |
| 100 | -94.08746 | 0.000179 | 6,696 | 2,935 | 8,722 |

Table E. 2 Results from 100 jitter runs for scenario 19.1 for WAG. Jitter run 0 corresponds to the original optimized estimates. NA: model did not converge for runs \#9 and 30.


| 39 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 40 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 41 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 42 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 43 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 44 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 45 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 46 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 47 | -177.0032 | 0.000280 | 5,782 | 1,960 | 6,331 |
| 48 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 49 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 50 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 51 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 52 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 53 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 54 | -177.0032 | 0.000280 | 5,810 | 1,964 | 6,334 |
| 55 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 56 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 57 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 58 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 59 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 60 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 61 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 62 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 63 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 64 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 65 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 66 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 67 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 68 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 69 | -177.0032 | 0.000280 | 5,782 | 1,960 | 6,331 |
| 70 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 71 | -177.0032 | 0.000280 | 5,872 | 1,962 | 6,378 |
| 72 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 73 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 74 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 75 | -177.0032 | 0.000280 | 5,901 | 1,958 | 6,338 |
| 76 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 77 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 78 | -177.0032 | 0.000280 | 5,872 | 1,962 | 6,378 |
| 79 | -177.0032 | 0.000280 | 5,872 | 1,962 | 6,378 |
| 80 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 81 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 82 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |


| 83 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| ---: | ---: | ---: | ---: | :--- | :--- |
| 84 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 85 | -177.0032 | 0.000280 | 5,872 | 1,962 | 6,378 |
| 86 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 87 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 88 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 89 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 90 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 91 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 92 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 93 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 94 | -177.0032 | 0.000280 | 5,782 | 1,960 | 6,331 |
| 95 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 96 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 97 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 98 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 99 | -177.0032 | 0.000280 | 5,300 | 1,884 | 6,099 |
| 100 | -177.0032 | 0.000280 | 5,872 | 1,962 | 6,378 |

## Appendix F: Progress in Gmacs

## Introduction

Implementation of Aleutian Islands golden king crab stock assessment in gmacs started in 2020 and the effort is continuing.

## Method

As a first step, we tried to compare EAG19.1 assessment results with that of gmacs. Estimated parameters from a modified EAG19.1 model (known as modifiedEAG19.1) that was reparametrized for gmacs computational formulas were input to gmacs ctl file. Parallel EAG19.1 data and projection files were also created for gmacs runs (gmacsEAG19.1CatchNo.ctl, gmacsEAG19.1CatchNo.dat, and gmacsEAG19.1CatchNo.prj). We compared time series of abundance composition ( N - matrix), retained catch composition, and CPUE indices among originalEAG19.1, modifiedEAG19.1, and gmacsEAG19.1 for two options: (1) fixed parameters of modifiedEAG19.1 for gmacs run (run\#10) and (2) free parameters of modifiedEAG19.1 for gmacs run (run\#9).

## Results

The gmacs ctl, dat, and prj files for EAG19.1 are provided in Tables F.1, F.2, and F.3, respectively. The abundance and retained catch compositions compare well among the three versions of EAG19.1 (originalEAG19.1, modifiedEAG19.1, and gmacsEAG19.1) in Figures F1, and F.2). The CPUE trends also compare well among the three versions (Figure F.3).

We found some differences in likelihood and reference points estimates between the original EAG19.1 model and its gmacs version. We will address those discrepancies before going into gmacs full implementation.

Table F1. gmacsEAG19.1.ctl file.








Table F2. gmacs EAG19.1. dat file.

| \# | EAG19.1 |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| \# | Gmacs |  | Main |  | Data | File |  |  |  |
| \# | GEAR_ |  | INDEX |  | DESCRIPTION |  |  |  |  |
| \# |  | 1 | : |  | Pot | fishery | Retained | catch |  |
| \# |  | 2 | : |  | Pot | fishery | total | catch |  |
| \# |  | 3 | : |  | Trawl | bycatch |  |  |  |
| \# |  | 4 | : |  | Observer | CPUE |  |  |  |
| \# |  | 5 | : |  | Fishery | CPUE |  |  |  |
| \# | Fisheries: |  |  | 1 | Pot | Fishery, | 2 | Pot | Total |
| \# | Cooperative |  | Survey: |  |  |  |  |  |  |
| \# |  |  |  |  |  |  |  |  |  |
| 1960 | \# |  | start |  | year) |  |  |  |  |
| 2019 | \# |  | terminal |  | year |  |  |  |  |
| \#2020 | \# |  | Projection |  | year |  |  |  |  |
| 6 | \# |  | Number |  | of | seasons: |  |  |  |
| 2 | \# |  | Number |  | of | distinct | data | groups |  |
| 1 | \# |  | Number |  | of | sexes |  |  |  |


|  | 1 | \# |  | Number |  | of |  | shell | condition | types |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | \# |  | Number |  | of |  | maturity | types |  |  |  |  |  |  |  |
|  | 17 | \# |  | Number |  | of |  | size-classes |  |  |  |  |  |  |  |  |
|  | 6 | \# |  | Season |  | when |  | recruitment | occurs |  |  |  |  |  |  |  |
|  | 6 | \# |  | Season |  | when |  | molting | and | growth | occur |  |  |  |  |  |
|  | 5 | \# |  | Season |  | to |  | calculate | MMB |  |  |  |  |  |  |  |
|  | 1 | \# |  | Season |  | for |  | N | output |  |  |  |  |  |  |  |
| \# |  | maximum |  | size-class |  |  |  |  |  |  |  |  |  |  |  |  |
|  | 17 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| \# |  | size_breaks |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
|  | 100.5 |  | 105.5 |  | 110.5 |  | 115.5 | 120.5 | 125.5 | 130.5 | 135.5 | 140.5 | 145.5 | 150.5 | 155.5 | 160.5 |
|  | 165.5 |  | 170.5 |  | 175.5 |  | 180.5 | 185.5 |  |  |  |  |  |  |  |  |
| \# |  | Natural |  | mortality |  | per |  | season |  |  |  |  |  |  |  |  |
|  | 2 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| \# |  | Proportion |  | of |  | the |  | total | natural | mortality |  |  |  |  |  |  |
| \# |  |  | 1 | Start |  | biological |  | year | (Jul | 1) | instantaneous | N | estimation |  |  |  |
| \# |  |  | 2 | to |  | mid |  | fishing | time |  |  |  |  |  |  |  |
| \# |  |  | 3 | instantanous |  | C |  | removal |  |  |  |  |  |  |  |  |
| \# |  |  | 4 | to |  | spawning |  | time |  |  |  |  |  |  |  |  |
| \# |  |  | 5 | instantaneous |  | byc |  | removal | and | estimate | MMB |  |  |  |  |  |
| \# |  |  | 6 | Rest |  | of |  | the | period | of | non | fishing |  |  |  |  |
|  |  |  |  |  |  | Feb |  | 15 |  |  | 30 |  |  |  |  |  |
| \#Ins |  | N |  | Jul1-MidFish |  | Inst |  | C | $\begin{aligned} & \text { Jul1- } \\ & \text { 15Feb } \end{aligned}$ | Ins |  |  |  |  |  |  |
|  | 0 |  | 0.16667 |  | 0 |  | 0.463 | 0 | 0.3699 | \#1960 |  |  |  |  |  |  |
|  | 0 |  | 0.16667 |  | 0 |  | 0.463 | 0 | 0.3699 | \#1961 |  |  |  |  |  |  |
|  | 0 |  | 0.16667 |  | 0 |  | 0.463 | 0 | 0.3699 | \#1962 |  |  |  |  |  |  |
|  | 0 |  | 0.16667 |  | 0 |  | 0.463 | 0 | 0.3699 | \#1963 |  |  |  |  |  |  |
|  | 0 |  | 0.16667 |  | 0 |  | 0.463 | 0 | 0.3699 | \#1964 |  |  |  |  |  |  |
|  | 0 |  | 0.16667 |  | 0 |  | 0.463 | 0 | 0.3699 | \#1965 |  |  |  |  |  |  |
|  | 0 |  | 0.16667 |  | 0 |  | 0.463 | 0 | 0.3699 | \#1966 |  |  |  |  |  |  |
|  | 0 |  | 0.16667 |  | 0 |  | 0.463 | 0 | 0.3699 | \#1967 |  |  |  |  |  |  |
|  | 0 |  | 0.16667 |  | 0 |  | 0.463 | 0 | 0.3699 | \#1968 |  |  |  |  |  |  |


| 0 | 0.16667 | 0 | 0.463 | 0 | 0.3699 | \#1969 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0 | 0.16667 | 0 | 0.463 | 0 | 0.3699 | \#1970 |
| 0 | 0.16667 | 0 | 0.463 | 0 | 0.3699 | \#1971 |
| 0 | 0.16667 | 0 | 0.463 | 0 | 0.3699 | \#1972 |
| 0 | 0.16667 | 0 | 0.463 | 0 | 0.3699 | \#1973 |
| 0 | 0.16667 | 0 | 0.463 | 0 | 0.3699 | \#1974 |
| 0 | 0.16667 | 0 | 0.463 | 0 | 0.3699 | \#1975 |
| 0 | 0.16667 | 0 | 0.463 | 0 | 0.3699 | \#1976 |
| 0 | 0.16667 | 0 | 0.463 | 0 | 0.3699 | \#1977 |
| 0 | 0.16667 | 0 | 0.463 | 0 | 0.3699 | \#1978 |
| 0 | 0.16667 | 0 | 0.463 | 0 | 0.3699 | \#1979 |
| 0 | 0.16667 | 0 | 0.463 | 0 | 0.3699 | \#1980 |
| 0 | 0.43973 | 0 | 0.19 | 0 | 0.3699 | \#1981 |
| 0 | 0.48082 | 0 | $0.149$ | 0 | 0.3699 | \#1982 |
| 0 | 0.48082 | 0 | 0.149 | 0 | 0.3699 | \#1983 |
| 0 | 0.3137 | 0 | 0.316 | 0 | 0.3699 | \#1984 |
| 0 | 0.16575 | 0 | 0.464 | 0 | 0.3699 | \#1985 |
| 0 | 0.24932 | 0 | 0.381 | 0 | 0.3699 | \#1986 |
| 0 | 0.08493 | 0 | 0.545 | 0 | 0.3699 | \#1987 |
| 0 | 0.29726 | 0 | 0.333 | 0 | 0.3699 | \#1988 |
| 0 | 0.31233 | 0 | $0.318$ | 0 | 0.3699 | \#1989 |
| 0 | 0.26301 | 0 | 0.367 | 0 | 0.3699 | \#1990 |
| 0 | 0.27123 | 0 | 0.359 | 0 | 0.3699 | \#1991 |
| 0 | 0.27397 | 0 | 0.356 | 0 | 0.3699 | \#1992 |
| 0 | $0.46027$ | 0 | 0.17 | 0 | 0.3699 | \#1993 |
| 0 | 0.24795 | 0 | 0.382 | 0 | 0.3699 | \#1994 |
| 0 | 0.22192 | 0 | 0.408 | 0 | 0.3699 | \#1995 |
| 0 | 0.3274 | 0 | 0.303 | 0 | 0.3699 | \#1996 |
| 0 | 0.28493 | 0 | 0.345 | 0 | 0.3699 | \#1997 |
| 0 | 0.26301 | 0 | 0.367 | 0 | 0.3699 | \#1998 |
| 0 | 0.24521 | 0 | 0.385 | 0 | 0.3699 | \#1999 |
| 0 | 0.17808 | 0 | 0.452 | 0 | 0.3699 | \#2000 |


| 0 |  | 0.1589 |  | 0 |  | 0.471 |  | 0 | 0.3699 | \#2001 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0 |  | 0.15479 |  | 0 |  | 0.475 |  | 0 | 0.3699 | \#2002 |
| 0 |  | 0.15616 |  | 0 |  | 0.474 |  | 0 | 0.3699 | \#2003 |
| 0 |  | 0.14247 |  | 0 |  | 0.488 |  | 0 | 0.3699 | \#2004 |
| 0 |  | 0.43288 |  | 0 |  | 0.197 |  | 0 | 0.3699 | \#2005 |
| 0 |  | 0.33151 |  | 0 |  | 0.299 |  | 0 | 0.3699 | \#2006 |
| 0 |  | 0.36849 |  | 0 |  | 0.262 |  | 0 | 0.3699 | \#2007 |
| 0 |  | 0.30274 |  | 0 |  | 0.327 |  | 0 | 0.3699 | \#2008 |
| 0 |  | 0.3274 |  | 0 |  | 0.303 |  | 0 | 0.3699 | \#2009 |
| 0 |  | 0.29315 |  | 0 |  | 0.337 |  | 0 | 0.3699 | \#2010 |
| 0 |  | 0.26301 |  | 0 |  | 0.367 |  | 0 | 0.3699 | \#2011 |
| 0 |  | 0.27534 |  | 0 |  | 0.355 |  | 0 | 0.3699 | \#2012 |
| 0 |  | 0.2726 |  | 0 |  | 0.358 |  | 0 | 0.3699 | \#2013 |
| 0 |  | 0.24795 |  | 0 |  | 0.382 |  | 0 | 0.3699 | \#2014 |
| 0 |  | 0.22877 |  | 0 |  | 0.401 |  | 0 | 0.3699 | \#2015 |
| 0 |  | 0.42055 |  | 0 |  | 0.21 |  | 0 | 0.3699 | \#2016 |
| 0 |  | 0.40959 |  | 0 |  | 0.221 |  | 0 | 0.3699 | \#2017 |
| 0 |  | 0.34932 |  | 0 |  | 0.281 |  | 0 | 0.3699 | \#2018 |
| 0 |  | 0.3274 |  | 0 |  | 0.303 |  | 0 | 0.3699 | \#2019 |
| \# |  |  |  |  |  |  |  |  |  |  |
| \# | Fishing |  | fleet |  | names |  |  |  |  |  |
| Pot_Fishery | Trawl_Bycatch |  |  |  |  |  |  |  |  |  |
| \# | Survey |  | names |  |  |  |  |  |  |  |
| \# | Are | $1{ }^{\text {the }}$ |  |  | seasons | discrete-instantaneous |  |  |  |  |
| 1 |  |  |  | 1 |  | 1 |  | 1 | 1 |  |
| \# | Number |  | of |  | catch |  | data |  | frames |  |
| 3 |  |  |  |  |  |  |  |  |  |  |
| \# | Number |  | of |  | rows |  | in |  | each | data |
| \# |  | 1993 | total |  | catch |  | is |  | missing, |  |
| \# | retained |  | catch |  | 1981/82- | 2019/20 |  |  |  |  |
| 39 |  | 29 |  | 31 |  |  |  |  |  |  |



| 2006 | 3 | 1 | 1 | 650.6 | 0.032 | 1 | 2 | 1 | 0 | 0.2 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2007 | 3 | 1 | 1 | 633.3 | 0.032 | 1 | 2 | 1 | 0 | 0.2 |
| 2008 | 3 | 1 | 1 | 666.9 | 0.032 | 1 | 2 | 1 | 0 | 0.2 |
| 2009 | 3 | 1 | 1 | 679.9 | 0.032 | 1 | 2 | 1 | 0 | 0.2 |
| 2010 | 3 | 1 | 1 | 671 | 0.032 | 1 | 2 | 1 | 0 | 0.2 |
| 2011 | 3 | 1 | 1 | 668.8 | 0.032 | 1 | 2 | 1 | 0 | 0.2 |
| 2012 | 3 | 1 | 1 | 687.7 | 0.032 | 1 | 2 | 1 | 0 | 0.2 |
| 2013 | 3 | 1 | 1 | 720.2 | 0.032 | 1 | 2 | 1 | 0 | 0.2 |
| 2014 | 3 | 1 | 1 | 719.1 | 0.032 | 1 | 2 | 1 | 0 | 0.2 |
| 2015 | 3 | 1 | 1 | 763.6 | 0.032 | 1 | 2 | 1 | 0 | 0.2 |
| 2016 | 3 | 1 | 1 | 794 | 0.032 | 1 | 2 | 1 | 0 | 0.2 |
| 2017 | 3 | 1 | 1 | 802.6 | 0.032 | 1 | 2 | 1 | 0 | 0.2 |
| 2018 | 3 | 1 | 1 | 940.3 | 0.032 | 1 | 2 | 1 | 0 | 0.2 |
| 2019 | 3 | 1 | 1 | 1057 | 0.032 | 1 | 2 | 1 | 0 | 0.2 |





|  | 1 |  | 1985 | 3 | 1 | 1 | 0 | 1.6287 | 0.051 | 2 | 0.5 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 |  | 1986 | 3 | 1 | 1 | 0 | 1.2289 | 0.047 | 2 | 0.5 |
|  | 1 |  | 1987 | 3 | 1 | 1 | 0 | 0.9552 | 0.049 | 2 | 0.5 |
|  | 1 |  | 1988 | 3 | 1 | 1 | 0 | 1.0358 | 0.041 | 2 | 0.5 |
|  | 1 |  | 1989 | 3 | 1 | 1 | 0 | 1.0765 | 0.034 | 2 | 0.5 |
|  | 1 |  | 1990 | 3 | 1 | 1 | 0 | 0.9868 | 0.045 | 2 | 0.5 |
|  | 1 |  | 1991 | 3 | 1 | 1 | 0 | 0.9046 | 0.043 | 2 | 0.5 |
|  | 1 |  | 1992 | 3 | 1 | 1 | 0 | 0.9172 | 0.043 | 2 | 0.5 |
|  | 1 |  | 1993 | 3 | 1 | 1 | 0 | 0.9145 | 0.049 | 2 | 0.5 |
|  | 1 |  | 1994 | 3 | 1 | 1 | 0 | 0.8086 | 0.042 | 2 | 0.5 |
|  | 1 |  | 1995 | 3 | 1 | 1 | 0 | 0.7798 | 0.043 | 2 | 0.5 |
|  | 1 |  | 1996 | 3 | 1 | 1 | 0 | 0.7791 | 0.044 | 2 | 0.5 |
|  | 1 |  | 1997 | 3 | 1 | 1 | 0 | 1.0505 | 0.045 | 2 | 0.5 |
|  | 1 |  | 1998 | 3 | 1 | 1 | 0 | 1.2141 | 0.051 | 2 | 0.5 |
| \# | Observer | CPUE | index |  |  |  |  |  |  |  |  |
|  | 1 |  | 1995 | 3 | 1 | 1 | 0 | 1.0034 | 0.032 | 2 | 0.5 |
|  | 1 |  | 1996 | 3 | 1 | 1 | 0 | 0.9444 | 0.021 | 2 | 0.5 |
|  | 1 |  | 1997 | 3 | 1 | 1 | 0 | 0.8742 | 0.021 | 2 | 0.5 |
|  | 1 |  | 1998 | 3 | 1 | 1 | 0 | 1.0004 | 0.019 | 2 | 0.5 |
|  | 1 |  | 1999 | 3 | 1 | 1 | 0 | 0.9154 | 0.018 | 2 | 0.5 |
|  | 1 |  | 2000 | 3 | 1 | 1 | 0 | 0.8196 | 0.016 | 2 | 0.5 |
|  | 1 |  | 2001 | 3 | 1 | 1 | 0 | 1.0429 | 0.018 | 2 | 0.5 |
|  | 1 |  | 2002 | 3 | 1 | 1 | 0 | 1.1029 | 0.021 | 2 | 0.5 |
|  | 1 |  | 2003 | 3 | 1 | 1 | 0 | 0.9714 | 0.019 | 2 | 0.5 |
|  | 1 |  | 2004 | 3 | 1 | 1 | 0 | 1.4394 | 0.027 | 2 | 0.5 |
|  | 2 |  | 2005 | 3 | 1 | 1 | 0 | 0.9829 | 0.026 | 2 | 0.5 |
|  | 2 |  | 2006 | 3 | 1 | 1 | 0 | 0.8087 | 0.023 | 2 | 0.5 |
|  | 2 |  | 2007 | 3 | 1 | 1 | 0 | 0.9017 | 0.022 | 2 | 0.5 |
|  | 2 |  | 2008 | 3 | 1 | 1 | 0 | 0.8819 | 0.026 | 2 | 0.5 |
|  | 2 |  | 2009 | 3 | 1 | 1 | 0 | 0.7266 | 0.031 | 2 | 0.5 |
|  | 2 |  | 2010 | 3 | 1 | 1 | 0 | 0.7518 | 0.031 | 2 | 0.5 |


| \#\# |  | Number |  | of |  | length |  | frequency |  | matrices |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| \#3 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
|  | 2 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| \#\# |  | Number |  | of |  | rows |  | in |  | each | matrix |  |  |  |  |  |  |  |  |  |
|  | 35 |  | 29 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| \#30 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| \#\# |  | Number |  | of |  | bins |  | in |  | each | matrix | (columns |  | of | size | data) |  |  |  |  |
|  | 17 |  | 17 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| \#17 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| \#\# |  | SIZE |  | COMPOSITION |  | DATA |  | FOR |  | ALL | FLEETS |  |  |  |  |  |  |  |  |  |
| \#\# |  | SIZE |  | COMP |  | LEGEND |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| \#\# |  | Sex: |  |  | 1 | $=$ |  | male, |  | 2 | $=$ | female, |  | 0 | $=$ | sexes | combin |  |  |  |
| \#\# |  | Type |  | of |  | composition: |  |  | 1 | = | retained, |  | 2 | $=$ | discard, | 0 | $=$ | total |  |  |
| \#\# |  | Maturity |  | state: |  |  | 1 | $=$ |  | immature, | 2 |  |  | mature, | 0 | $=$ |  |  |  |  |
| \#\# |  | Shell |  | condition: |  |  | 0 | $=$ |  |  | shell |  |  | combined |  |  |  |  |  |  |
| \#\# |  | Type |  |  | 1 | effective |  | sample: |  | Nsamp |  |  |  |  |  |  |  |  |  |  |
| \#\# |  | Retain |  | catch |  | size |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| \#\#Year, |  | Seas, |  | Fleet, |  | Sex, |  | Type, |  | Shell, | Maturity, | Nsamp, |  | DataVec |  |  |  |  |  |  |
|  |  |  | 1985 |  | 3 |  | 1 |  | 1 | 1 | 0 |  | 0 | 47 | 0 | 0 | 0 | 0 | 0 | 0.002 |
|  |  |  | 1986 |  | 3 |  | 1 |  | 1 | 1 | 0 |  | 0 | 9 | 0 | 0 | 0 | 0 | 0 | 0.001 |
|  |  |  | 1987 |  | 3 |  | 1 |  | 1 | 1 | 0 |  | 0 | 50 | 0 | 0 | 0.004 | 0 | 0.00055 | 0.003 |


| 1988 | 3 | 1 | 1 | 1 | 0 | 0 | 291 | 0 | 0 | 0 | 0 | 0.00025 | 0.005 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1989 | 3 | 1 | 1 | 1 | 0 | 0 | 655 | 0 | 0 | 0 | 0.000 | 0.00019 | 0.008 |
| 1990 | 3 | 1 | 1 | 1 | 0 | 0 | 135 | 0 | 0.000 | 0.000 | 0 | 0.00034 | 0.006 |
| 1991 | 3 | 1 | 1 | 1 | 0 | 0 | 116 | 0 | 0 | 0 | 0 | 0.00029 | 0.006 |
| 1992 | 3 | 1 | 1 | 1 | 0 | 0 | 41 | 0 | 0 | 0.000 | 0.000 | 0.00045 | 0.005 |
| 1993 | 3 | 1 | 1 | 1 | 0 | 0 | 281 | 0 | 0 | 0 | 0 | 0.00127 | 0.006 |
| 1994 | 3 | 1 | 1 | 1 | 0 | 0 | 264 | 0 | 0 | 0 | 0 | 0 | 0.005 |
| 1995 | 3 | 1 | 1 | 1 | 0 | 0 | 727 | 0 | 0 | 0.000 | 0 | 0.00013 | 0.003 |
| 1996 | 3 | 1 | 1 | 1 | 0 | 0 | 452 | 0 | 0.001 | 0 | 0.003 | 0.00446 | 0.011 |
| 1997 | 3 | 1 | 1 | 1 | 0 | 0 | 445 | 0 | 0 | 0 | 0 | 0.00055 | 0.006 |
| 1998 | 3 | 1 | 1 | 1 | 0 | 0 | 447 | 0 | 0 | 0 | 0 | 0.00015 | 0.002 |
| 1999 | 3 | 1 | 1 | 1 | 0 | 0 | 383 | 0 | 0 | 0 | 0 | 0 | 0.003 |
| 2000 | 3 | 1 | 1 | 1 | 0 | 0 | 360 | 0 | 0 | 0 | 0 | 0 | 0.002 |
| 2001 | 3 | 1 | 1 | 1 | 0 | 0 | 403 | 0 | 0.000 | 0 | 0 | 0 | 0.002 |
| 2002 | 3 | 1 | 1 | 1 | 0 | 0 | 336 | 0.0004 | 0 | 0 | 0 | 0 | 0.001 |
| 2003 | 3 | 1 | 1 | 1 | 0 | 0 | 335 | 0 | 0 | 0 | 0 | 0.0001 | 0.001 |
| 2004 | 3 | 1 | 1 | 1 | 0 | 0 | 231 | 0 | 0 | 0 | 0 | 0 | 0.000 |
| 2005 | 3 | 1 | 1 | 1 | 0 | 0 | 220 | 0 | 0 | 0 | 0 | 0 | 0.001 |
| 2006 | 3 | 1 | 1 | 1 | 0 | 0 | 193 | 0 | 0 | 0 | 0 | 0 | 0.000 |
| 2007 | 3 | 1 | 1 | 1 | 0 | 0 | 165 | 0.0003 | 0 | 0 | 0 | 0.00062 | 0 |
| 2008 | 3 | 1 | 1 | 1 | 0 | 0 | 163 | 0 | 0 | 0 | 0 | 0 | 0.001 |
| 2009 | 3 | 1 | 1 | 1 | 0 | 0 | 141 | 0 | 0 | 0 | 0 | 0 | 0 |
| 2010 | 3 | 1 | 1 | 1 | 0 | 0 | 151 | 0 | 0 | 0 | 0 | 0 | 0.000 |
| 2011 | 3 | 1 | 1 | 1 | 0 | 0 | 132 | 0 | 0 | 0 | 0 | 0 | 0.000 |
| 2012 | 3 | 1 | 1 | 1 | 0 | 0 | 155 | 0 | 0 | 0 | 0 | 0 | 0.001 |
| 2013 | 3 | 1 | 1 | 1 | 0 | 0 | 160 | 0 | 0 | 0 | 0 | 0 | 0.002 |
| 2014 | 3 | 1 | 1 | 1 | 0 | 0 | 139 | 0 | 0 | 0 | 0 | 0 | 0.001 |
| 2015 | 3 | 1 | 1 | 1 | 0 | 0 | 157 | 0 | 0 | 0 | 0 | 0 | 0.000 |
| 2016 | 3 | 1 | 1 | 1 | 0 | 0 | 184 | 0 | 0 | 0 | 0 | 0 | 0.002 |
| 2017 | 3 | 1 | 1 | 1 | 0 | 0 | 176 | 0 | 0 | 0 | 0.000 | 0 | 0.000 |
| 2018 | 3 | 1 | 1 | 1 | 0 | 0 | 180 | 0 | 0 | 0 | 0.001 | 0 | 0.000 |
| 2019 | 3 | 1 | 1 | 1 | 0 | 0 | 177 | 0 | 0 | 0 | 0 | 0 | 0.001 |


| \#\# | Total |  | catch |  | size |  | comp |  |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| \#\#Year, | Seas, |  | Fleet, |  | Sex, |  | Type, |  | Shell, |  | Maturity, | Nsamp, |  |  |  |  |  |  |  |  |
|  |  | 1990 |  | 3 |  | 1 |  | 1 |  | 0 | 0 |  | 0 | 13 | 0.0942 | 0.086 | 0.084 | 0.093 | 0.12217 | 0.11 |
|  |  | 1991 |  | 3 |  | 1 |  | 1 |  | 0 | 0 |  | 0 | 28 | 0.0462 | 0.064 | 0.078 | 0.084 | 0.10383 | 0.117 |
|  |  | 1992 |  | 3 |  | 1 |  | 1 |  | 0 | 0 |  | 0 | 24 | 0.0703 | 0.085 | 0.084 | 0.108 | 0.11602 | 0.101 |
|  |  | 1994 |  | 3 |  | 1 |  | 1 |  | 0 | 0 |  | 0 | 20 | 0.1205 | 0.111 | 0.092 | 0.089 | 0.10685 | 0.123 |
|  |  | 1995 |  | 3 |  | 1 |  | 1 |  | 0 | 0 |  | 0 | 649 | 0.0365 | 0.049 | 0.069 | 0.088 | 0.10852 | 0.128 |
|  |  | 1996 |  | 3 |  | 1 |  | 1 |  | 0 | 0 |  | 0 | 296 | 0.033 | 0.045 | 0.061 | 0.08 | 0.1005 | 0.123 |
|  |  | 1997 |  | 3 |  | 1 |  | 1 |  | 0 | 0 |  | 0 | 413 | 0.0327 | 0.042 | 0.059 | 0.077 | 0.09885 | 0.121 |
|  |  | 1998 |  | 3 |  | 1 |  | 1 |  | 0 | 0 |  | 0 | 333 | 0.0296 | 0.042 | 0.061 | 0.084 | 0.1087 | 0.129 |
|  |  | 1999 |  | 3 |  | 1 |  | 1 |  | 0 | 0 |  | 0 | 352 | 0.0242 | 0.032 | 0.043 | 0.064 | 0.09039 | 0.121 |
|  |  | 2000 |  | 3 |  | 1 |  | 1 |  | 0 | 0 |  | 0 | 287 | 0.0222 | 0.032 | 0.045 | 0.063 | 0.07697 | 0.111 |
|  |  | 2001 |  | 3 |  | 1 |  | 1 |  | 0 | 0 |  | 0 | 296 | 0.0175 | 0.024 | 0.034 | 0.047 | 0.06295 | 0.092 |
|  |  | 2002 |  | 3 |  | 1 |  | 1 |  | 0 | 0 |  | 0 | 254 | 0.0204 | 0.026 | 0.032 | 0.039 | 0.05208 | 0.084 |
|  |  | 2003 |  | 3 |  | 1 |  | 1 |  | 0 | 0 |  | 0 | 242 | 0.0147 | 0.024 | 0.029 | 0.04 | 0.05596 | 0.09 |
|  |  | 2004 |  | 3 |  | 1 |  | 1 |  | 0 | 0 |  | 0 | 174 | 0.0127 | 0.017 | 0.024 | 0.033 | 0.055 | 0.08 |
|  |  | 2005 |  | 3 |  | 1 |  | 1 |  | 0 | 0 |  | 0 | 135 | 0.0066 | 0.009 | 0.01 | 0.016 | 0.02939 | 0.042 |
|  |  | 2006 |  | 3 |  | 1 |  | 1 |  | 0 | 0 |  | 0 | 83 | 0.005 | 0.009 | 0.011 | 0.014 | 0.02415 | 0.038 |
|  |  | 2007 |  | 3 |  | 1 |  | 1 |  | 0 | 0 |  | 0 | 78 | 0.0028 | 0.004 | 0.005 | 0.008 | 0.01826 | 0.033 |
|  |  | 2008 |  | 3 |  | 1 |  | 1 |  | 0 | 0 |  | 0 | 66 | 0.0039 | 0.005 | 0.008 | 0.012 | 0.02182 | 0.042 |
|  |  | 2009 |  | 3 |  | 1 |  | 1 |  | 0 | 0 |  | 0 | 55 | 0.004 | 0.005 | 0.009 | 0.015 | 0.02396 | 0.048 |
|  |  | 2010 |  | 3 |  | 1 |  | 1 |  | 0 | 0 |  | 0 | 63 | 0.0074 | 0.009 | 0.014 | 0.02 | 0.03743 | 0.057 |
|  |  | 2011 |  | 3 |  | 1 |  | 1 |  | 0 | 0 |  | 0 | 62 | 0.0049 | 0.007 | 0.01 | 0.014 | 0.02171 | 0.039 |
|  |  | 2012 |  | 3 |  | 1 |  | 1 |  | 0 | 0 |  | 0 | 57 | 0.0023 | 0.005 | 0.007 | 0.007 | 0.01942 | 0.037 |
|  |  | 2013 |  | 3 |  | 1 |  | 1 |  | 0 | 0 |  | 0 | 71 | 0.0038 | 0.007 | 0.01 | 0.016 | 0.0268 | 0.051 |
|  |  | 2014 |  | 3 |  | 1 |  | 1 |  | 0 | 0 |  | 0 | 57 | 0.0042 | 0.006 | 0.01 | 0.016 | 0.02694 | 0.047 |
|  |  | 2015 |  | 3 |  | 1 |  | 1 |  | 0 | 0 |  | 0 | 73 | 0.0053 | 0.009 | 0.014 | 0.022 | 0.03282 | 0.054 |
|  |  | 2016 |  | 3 |  | 1 |  | 1 |  | 0 | 0 |  | 0 | 90 | 0.0087 | 0.008 | 0.014 | 0.019 | 0.04065 | 0.066 |
|  |  | 2017 |  | 3 |  | 1 |  | 1 |  | 0 | 0 |  | 0 | 77 | 0.0058 | 0.007 | 0.013 | 0.02 | 0.03253 | 0.085 |
|  |  | 2018 |  | 3 |  | 1 |  | 1 |  | 0 | 0 |  | 0 | 136 | 0.0038 | 0.006 | 0.01 | 0.017 | 0.04301 | 0.08 |
|  |  | 2019 |  | 3 |  | 1 |  | 1 |  | 0 | 0 |  | 0 | 86 | 0.0007 | 0.002 | 0.005 | 0.009 | 0.02372 | 0.079 |

\#

| \#\# | Trawl |  | byc |  | size |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| \#\#Year, | Seas, |  | Fleet, |  | Sex, |  |
| \#1989 |  | 5 |  | 2 |  | 1 |
| \#1990 |  | 5 |  | 2 |  | 1 |
| \#1992 |  | 5 |  | 2 |  | 1 |
| \#1993 |  | 5 |  | 2 |  | 1 |
| \#1994 |  | 5 |  | 2 |  | 1 |
| \#1995 |  | 5 |  | 2 |  | 1 |
| \#1996 |  | 5 |  | 2 |  | 1 |
| \#1997 |  | 5 |  | 2 |  | 1 |
| \#1998 |  | 5 |  | 2 |  | 1 |
| \#1999 |  | 5 |  | 2 |  | 1 |
| \#2000 |  | 5 |  | 2 |  | 1 |
| \#2001 |  | 5 |  | 2 |  | 1 |
| \#2002 |  | 5 |  | 2 |  | 1 |
| \#2003 |  | 5 |  | 2 |  | 1 |
| \#2004 |  | 5 |  | 2 |  | 1 |
| \#2005 |  | 5 |  | 2 |  | 1 |
| \#2006 |  | 5 |  | 2 |  | 1 |
| \#2007 |  | 5 |  | 2 |  | 1 |
| \#2008 |  | 5 |  | 2 |  | 1 |
| \#2009 |  | 5 |  | 2 |  | 1 |
| \#2010 |  | 5 |  | 2 |  | 1 |
| \#2011 |  | 5 |  | 2 |  | 1 |
| \#2012 |  | 5 |  | 2 |  | 1 |
| \#2013 |  | 5 |  | 2 |  | 1 |
| \#2014 |  | 5 |  | 2 |  | 1 |
| \#2015 |  | 5 |  | 2 |  | 1 |
| \#2016 |  | 5 |  | 2 |  | 1 |
| \#2017 |  | 5 |  | 2 |  | 1 |
| \#2018 |  | 5 |  | 2 |  | 1 |



| 0 |
| :---: |
| 0 |
| 0 |
| 0 |
| 0 |
| 0 |
| 0 |
| 0 |
| 0 |
| 0 |
| 0 |
| 0 |
| 0 |
| 0 |
| 0 |
| 0 |
| 0 |
| 0 |
| 0 |
| 0 |
| 0 |
| 0 |
| 0 |
| 0 |
| 0 |
| 0 |
| 0 |
| 0 |

Maturity
Nsamp,

DataVec

| 4 | 0 | 0.0545 | 0.127 | 0.091 | 0.073 | 0.09091 | 0.127 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 6 | 0.074 | 0.0465 | 0.039 | 0.066 | 0.078 | 0.03876 | 0.101 |
| 1 | 0 | 0.1667 | 0 | 0.167 | 0.167 | 0.16667 | 0 |
| 1 | 0 | 0 | 0.25 | 0 | 0.25 | 0 | 0 |
| 2 | 0.167 | 0.2407 | 0.185 | 0.167 | 0.074 | 0.03704 | 0.019 |
| 2 | 0.037 | 0.037 | 0.037 | 0.148 | 0.111 | 0.11111 | 0.185 |
| 2 | 0 | 0 | 1 | 0 | 0 | 0 | 0 |
| 4 | 0.096 | 0.0769 | 0.058 | 0.135 | 0.115 | 0.13462 | 0.115 |
| 7 | 0.088 | 0.0949 | 0.066 | 0.08 | 0.153 | 0.12409 | 0.058 |
| 7 | 0.152 | 0.1714 | 0.057 | 0.086 | 0.076 | 0.08571 | 0.029 |
| 8 | 0.197 | 0.171 | 0.068 | 0.063 | 0.063 | 0.05386 | 0.063 |
| 6 | 0.076 | 0.0714 | 0.107 | 0.103 | 0.083 | 0.05804 | 0.06 |
| 7 | 0.225 | 0.2143 | 0.132 | 0.093 | 0.126 | 0.03297 | 0.055 |
| 8 | 0.301 | 0.1399 | 0.07 | 0.091 | 0.063 | 0.05594 | 0.056 |
| 5 | 0.095 | 0.0476 | 0.048 | 0.048 | 0.048 | 0 | 0 |
| 6 | 0.268 | 0.1959 | 0.082 | 0.082 | 0.052 | 0.05155 | 0.021 |
| 7 | 0.269 | 0.1346 | 0.115 | 0.096 | 0.067 | 0.06731 | 0.058 |
| 8 | 0.257 | 0.2163 | 0.122 | 0.073 | 0.082 | 0.07755 | 0.037 |
| 7 | 0.229 | 0.2048 | 0.117 | 0.083 | 0.052 | 0.05714 | 0.052 |
| 8 | 0.174 | 0.0413 | 0.025 | 0.058 | 0.107 | 0.05785 | 0.058 |
| 12 | 0.184 | 0.2175 | 0.154 | 0.103 | 0.091 | 0.06647 | 0.063 |
| 6 | 0.324 | 0.1639 | 0.172 | 0.107 | 0.049 | 0.05328 | 0.02 |
| 9 | 0.063 | 0.0833 | 0.083 | 0.083 | 0.042 | 0.02083 | 0.021 |
| 8 | 0 | 0 | 0 | 0 | 0.059 | 0 | 0.059 |
| 8 | 0.063 | 0 | 0.031 | 0.125 | 0.094 | 0.09375 | 0.063 |
| 5 | 0.116 | 0.1053 | 0.116 | 0.116 | 0.137 | 0.12632 | 0.105 |
| 6 | 0.039 | 0.0789 | 0.118 | 0.118 | 0.158 | 0.14474 | 0.066 |
| 6 | 0.304 | 0.1957 | 0.098 | 0.054 | 0.043 | 0.03261 | 0.043 |
|  | 0.286 | 0.119 | 0.119 | 0.04 | 0.071 | 0.0952 | 0.04 |

```
\begin{tabular}{|c|c|c|c|c|c|c|c|c|c|c|c|c|c|c|}
\hline \#2019 & 5 & 2 & 1 & 2 & 0 & 0 & 4 & & & 0.175 & & & & \\
\hline
\end{tabular}
#
##
#
# nobs_growth
0
#
Class-at-release; Sex;
\begin{tabular}{ll} 
Class-at- & Years-at- \\
recapture; & liberty;
\end{tabular}
number transition matrix;
\#not considered
\#\# eof
9999
```

Table F3. gmacs EAG19.1. prj file.

|  | 0 |  | Do | not | compute | MSY | (1=Yes) |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 1 | if | future | F | is | to | be | fixed |  |  |
|  | 1987 | 2017 | for | Rbar |  | calc, |  |  |  |  |  |
|  | 1985 | 2019 | First | and | last | year | for | average | sex | ratio |  |
|  | 2010 | 2019 | First | and | last | year | for | average | F | for | discards |
| \# | OFL | specifications |  |  |  |  |  |  |  |  |  |
|  | 0.35 | Target | SPR | ratio | for | Bmsy | proxy. |  |  |  |  |
|  | 3 | Tier |  |  |  |  |  |  |  |  |  |
|  | 0.1 | Alpha |  |  |  |  |  |  |  |  |  |
|  | 0.25 | Beta |  |  |  |  |  |  |  |  |  |
|  | 1 | Gamma |  |  |  |  |  |  |  |  |  |
|  | 0.75 | ABC-OFL | buffer |  |  |  |  |  |  |  |  |
|  | 0 | Produce | a | yield | curve or not |  |  |  |  |  |  |





Figure F1. Comparison of time series of abundance by size (N-matrix) of EAG golden king crab, 1960-2019 [blue: OrignalEAG19.1; red: ModifiedEAG19.1; dark green: gmacs run\#9 (free parameters); green: gmacs run\#10(fixed parameters)]


Figure F2. Comparison of time series of retained catch by size of EAG golden king crab, 1985-2019 (blue: OrignalEAG19.1; red: ModifiedEAG19.1; dark green: gmacs run\#9; green: gmacs run\#10)


Figure F3. Comparison of time series of CPUE indices for EAG golden king crab, 1995-2019 (blue: OrignalEAG19.1A; red: ModifiedEAG19.1A; green: gmacs run\#10AUpdate). For this comparison, the gmacs base model was slightly modified (update) to instantaneously remove catches at the middle of the fishing period.


[^0]:    *As of March 26, 2021, WAG fishery is ongoing.

