# Aleutian Islands Golden King Crab Stock Assessment 

May 2023 Crab SAFE DRAFT REPORT
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## Executive Summary

1. Stock

Golden king crab, Lithodes aequispinus, Aleutian Islands, east of $174^{\circ} \mathrm{W}$ longitude (Eastern Aleutian Islands; EAG) and west of $174^{\circ} \mathrm{W}$ longitude (Western Aleutian Islands; WAG).

## 2. Catch

The Aleutian Islands golden king crab (AIGKC) commercial fishery has been prosecuted every year since $1981 / 82$. Retained catch peaked in 1986/87 at 2,686 t $(5,922,425 \mathrm{lb})$ and $3,999 \mathrm{t}(8,816,319 \mathrm{lb})$, respectively, for EAG and WAG, but the retained catch dropped sharply from 1989/90 to 1990/91. The fishery has been managed separately east (EAG) and west (WAG) of $174^{\circ}$ W longitude since 1996/97, and Guideline Harvest Levels (GHLs) of $1,452 \mathrm{t}(3,200,000 \mathrm{lb})$ for EAG and $1,225 \mathrm{t}$ $(2,700,000 \mathrm{lb})$ for WAG were introduced into management. The GHL was subsequently reduced to $1,361 \mathrm{t}(3,000,000 \mathrm{lb})$ beginning in 1998/99 for EAG. The reduced harvest levels remained at $1,361 \mathrm{t}(3,000,000 \mathrm{lb})$ for EAG and $1,225 \mathrm{t}$ $(2,700,000 \mathrm{lb})$ for WAG through 2007/08 but were increased to $1,429 \mathrm{t}(3,150,000 \mathrm{lb})$ for EAG and $1,294 \mathrm{t}(2,835,000 \mathrm{lb})$ for WAG beginning with the 2008/09 fishing season following an Alaska Board of Fisheries (BOF) decision. The management specification changed from GHL to Total Allowable Catch (TAC) with adoption of the Crab Rationalization Program in 2005/06 (NPFMC 2007b). The TACs were increased by another BOF decision to $1,501 \mathrm{t}(3,310,000 \mathrm{lb})$ for EAG and $1,352 \mathrm{t}(2,980,000 \mathrm{lb})$ for WAG beginning with the 2012/13 fishing season. The below par fishery performance in WAG in 2014/15 and 2015/16 lead to reduction in TAC to $1,014 \mathrm{t}$ $(2,235,000 \mathrm{lb})$, which reflected a $25 \%$ reduction in the TAC for WAG, while the TAC for EAG was kept at the same level, $1,501 \mathrm{t}(3,310,000 \mathrm{lb})$ for the $2016 / 17$ through 2017/18 fishing seasons. With the improved fishery performance and stock status in 2017/18, the TACs were further increased to $1,134 \mathrm{t}(2,500,000 \mathrm{lb})$ for WAG and 1,749 $\mathrm{t}(3,856,000 \mathrm{lb})$ for EAG beginning with the $2018 / 19$ fishing season. With the implementation of a revised state harvest strategy in 2019, the TACs were further increased to $1,302 \mathrm{t}(2,870,000 \mathrm{lb})$ for WAG and $1,955 \mathrm{t}(4,310,000 \mathrm{lb})$ for EAG. Based on the model estimated abundances, the 2022/23 fishing season TACs were adjusted to $785 \mathrm{t}(1,730,000 \mathrm{lb})$ for WAG and $1,506 \mathrm{t}(3,320,000 \mathrm{lb})$ for EAG at the time of this assessment.

Retained catch has been regularly under the GHL/TAC but close to allowable levels since 1996/97. These TAC levels were set below the Allowable Biological Catch (ABC) levels determined under Tier 5 criteria (considering 1991-1995 mean catch for the whole Aleutian Islands region, $3,145 \mathrm{t}[6,933,822 \mathrm{lb}]$, as the catch limit) under the most recent crab management plan. A new harvest strategy based on model estimated mature male abundance was accepted by the BOF in March 2019, specifying a $15 \%$ maximum harvest rate for EAG and $20 \%$ maximum harvest rate for WAG, and implemented during the 2019/20 fishery. In addition to the retained catch allotted as TAC, there was retained catch in a cost-recovery fishery towards a $\$ 300,000$ goal in 2013/14 and 2014/15 to fund an onboard observer program, and towards a $\$ 500,000$ goal between 2015/16 - 2022/23 to fund an onboard observer program and cooperative stock survey.

Total mortality of Aleutian Islands golden king crab includes retained catch in the directed and the cost recovery fisheries, discards in the directed fisheries, and bycatch in fixed-gear and trawl groundfish fisheries, though bycatch in other fisheries is low compared to mortality in the directed fishery. Total retained catch in the postrationalized fishery (2005/06-2022/23) has ranged from $2,296 \mathrm{t}(5,060,806 \mathrm{lb})$ to 3,319 $\mathrm{t}(7,316,853 \mathrm{lb})$. Total mortality ranged from $2,506 \mathrm{t}(5,525,000 \mathrm{lb})$ to $3,729 \mathrm{t}$ $(8,222,000 \mathrm{lb})$ for the same period. Total retained catch in $2022 / 23$ was $2,369 \mathrm{t}$ $(5,759,039 \mathrm{lb}): 1,585 \mathrm{t}(3,494,050 \mathrm{lb})$ from the EAG fishery (which included costrecovery catch), and $784 \mathrm{t}(1,729,215 \mathrm{lb})$ from the WAG fishery. Discarded (nonretained) catch occurs mainly during the directed fishery. Although low levels of discarded catch can occur during other crab fisheries, there have been no such fisheries prosecuted locally since 2004/05, except as surveys for red king crab conducted under an Alaska Department of Fish and Game (ADF\&G) Commissioner's Permit (and no golden king crab were caught during the cooperative red king crab survey performed by industry and ADF\&G in the Adak area in September 2015; Hilsinger et al. 2016). Estimates of the bycatch mortality during crab fisheries decreased during 1995/962005/06, both in absolute value and relative to the retained catch weight and stabilized during 2005/06-2014/15. Total estimated bycatch mortality during crab fisheries in 2022/23 was $125 \mathrm{t}(276,263 \mathrm{lb})$ for EAG and $112 \mathrm{t}(246,151 \mathrm{lb})$ for WAG. Discarded catch also occurs during fixed-gear and trawl groundfish fisheries but is small relative to the directed fishery. Groundfish fisheries were a minor contributor to total fishery discard mortality, $5 \mathrm{t}(10,351 \mathrm{lb})$ for EAG and $1 \mathrm{t}(3,007 \mathrm{lb})$ for WAG in 2022/23.

Catch per unit effort (CPUE, i.e., catch per pot lift) of retained legal males decreased from the 1980s into the mid-1990s, but increased after 1994/95, particularly with the initiation of the Crab Rationalization Program in 2005/06. Although CPUE for the two areas showed similar trends through 2010/11, during 2011/12-2014/15 CPUE trends diverged (increasing for EAG and decreasing for WAG).

A cooperative golden king crab survey has been conducted by the Aleutian Islands King Crab Foundation (an industry group) and ADF\&G using vessels that are commercial fishing (i.e., each vessel fishing an allotted share of total allowable catch) in EAG (beginning in 2015) and WAG (beginning in 2018). The cooperative survey
was not conducted in 2020 (2020/21 fishery) due to COVID-19 but resumed since 2021 (2021/22 fishery). Bycatch mortality during the survey is accounted for by expanding reported discards for the season's entire fishery, including fishing activities conducted as part of the survey.

## 3. Stock biomass

Estimated mature male biomass (MMB) for EAG under all scenarios decreased from the 1980s to the 1990s, then increased during the 2000s and systematically increased from 2016 to 2020 and slightly decreased since 2021. Estimated MMB for WAG decreased during the late 1980s and 1990s, increased during the 2000s, decreased for several years since 2011. The low levels of MMB for EAG were observed in 19951997 and in 1990s for WAG. Stock trends reflected the fishery standardized CPUE trends in both regions.

## 4. Recruitment

Recruits to the model size groups under all scenarios have fluctuated in both EAG and WAG. For EAG, model recruitment was high during 2014-2019, highest in 2017; and lowest in 1985. The model recruitment for WAG was high during 1983 to 1986, highest in 1985, and lowest in 2010. A slightly increasing trend in recruitment was observed during 2021-2022 in EAG and in 2022 in WAG.

## 5. Management performance

King and Tanner crab stocks in the Bering Sea and Aleutian Islands are cooperatively managed by the State of Alaska through a federal king and Tanner crab fishery management plan (FMP). Under the FMP, management measures are divided into three categories: (1) fixed in the FMP, (2) frame worked in the FMP, and (3) discretion of the State of Alaska. The State of Alaska is responsible for determining and establishing the GHL/TAC under the management measure frame worked in the FMP. Aleutian Islands golden king crab stocks are managed under this FMP measure.

The size-based assessment model was accepted at the September 2016 CPT and October 2016 SSC meetings for OFL determination for the 2017/18 fishery cycle. In addition, the CPT in January 2017 and SSC in February 2017 recommended using the Tier 3 method to compute OFL and ABC. The assessment model was first used for setting OFL and ABC for the 2017/18 fishing season. The CPT in May 2017 and SSC in June 2017 accepted the authors' recommendation of using scenario 9 (i.e., model using the knife-edge maturity size of 111 mm carapace length to determine MMB) for OFL and ABC calculation. During the May 2017 meeting, the CPT noted that a single OFL and ABC are defined for Aleutian Islands golden king crab (AIGKC), however; separate models are available by area. Hence, following previous assessments, OFLs and ABCs by area were summed to determine OFL and ABC for the entire stock.

In this assessment simple sum of the area-specific OFLs and ABCs was used to calculate area-wide OFL and ABC although on an area-specific basis, the EAG stock was found to be in Tier 3a while the WAG stock was in Tier 3b. This simple summation was done because authors recommended model 22.1 e 2 was area specific and implemented in GMACS for OFL and ABC calculation.

All models for EAG and WAG were updated with the previous season's fishery information [i.e., 2022/23 ongoing fisheries (at the time of this assessment on 17 March 2023) in EAG and WAG]. For OFL and ABC presentation for CPT review, the 2022/23 MSST, MMB, OFL, and ABC were determined when both fisheries were not concluded [above $90 \%$ of TAC taken (as of the 17 March 2023 reported summary)]. Following the CPT/SSC suggestion, the EAG and WAG assessments in May 2023 considered a hypothetical completed fishery retained and total catches and an average of the last three years' groundfish bycatch removals. The EAG and WAG retained catches were set to respective TACs (TAC in lb was converted to number of crab by dividing by the average retained crab weight in the 2022/23 fishery). The total catches were estimated based on the predicted total effort (by dividing the TAC by the current CPUE) and the current nominal total CPUE. Model 22.1e2 was recommended for the EAG and WAG OFL and ABC calculation. Maturity size analysis (Siddeek et al., 2021) provided a strong justification to increase the previously used maturity size by one size bin. Model 22.1 e 2 considered a fixed period, 1987-2017, for the mean number of recruit calculation for reference points estimation, standardization of observer and fishery CPUE by the negative binomial generalized linear model, the knife-edge maturity size of 116 mm carapace length (CL) for mature male biomass (MMB) estimation, and three catchability periods for fitting three sets of CPUE indices.

Following January/February 2023 CPT and SSC suggestions, several models were presented for evaluation. Model 22.1 e 2 was the base model (accepted model in 2022) with three catchability periods, knife-edge male maturity size at 116 mm CL, M of 0.22 yr-1, selection of a fixed period, 1987-2017, for mean number of recruit calculation for reference points estimation, and addition of new data for 2022/23. Models 22.1f, 22.1g and 21.1 h were modifications from the base model. All models were implemented in GMACS. The GMACS input parameter values were estimated by model 22.9 c , which was the modification of model 22.1e2 for GMACS implementation. Model 22.1f was the same as model 22.1 e 2 but CPUE standardization was done considering Year:Area interaction. Models 22.1 g and 22.1 h were modifications from 22.1 e 2 and 22.1f, respectively, that included cooperative survey indices.

Status and catch specifications ( 1000 t) of Aleutian Islands golden king crab

| Year | MSST | Biomass <br> (MMB) | TAC | Retained <br> Catch | Total <br> Catch $^{\mathbf{a}}$ | OFL | ABC $^{\mathbf{b}}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $2019 / 20$ | 5.915 | 16.386 | 3.257 | 3.319 | 3.729 | 5.249 | 3.937 |
| $2020 / 21$ | 6.014 | 15.442 | 2.999 | 3.000 | 3.520 | 4.798 | 3.599 |
| $2021 / 22$ | 5.715 | 13.581 | 2.690 | 2.699 | 3.056 | 4.817 | 3.372 |
| $2022 / 23$ | $5.832^{\text {d }}$ | $13.600^{\text {d }}$ | 2.291 | $2.369^{*}$ | $2.612^{*}$ | $3.761^{\text {c }}$ | $2.821^{\text {c }}$ |
| $2023 / 24$ |  | $12.069^{\text {d }}$ |  |  |  | $4.182^{\text {d }}$ | $3.137^{\text {d }}$ |

Status and catch specifications (million lb) of Aleutian Islands golden king crab

| Year | MSST | Biomass <br> (MMB) | TAC | Retained <br> Catch | Total <br> Catch $^{\text {a }}$ | OFL | ABC $^{\text {b }}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $2019 / 20$ | 13.041 | 36.124 | 7.180 | 7.317 | 8.222 | 11.572 | 8.679 |
| $2020 / 21$ | 13.259 | 34.043 | 6.610 | 6.614 | 7.759 | 10.579 | 7.934 |
| $2021 / 22$ | 12.598 | 29.940 | 5.930 | 5.951 | 6.737 | 10.620 | 7.434 |
| $2022 / 23$ | $12.857^{\text {d }}$ | $29.984^{\text {d }}$ | 5.050 | $5.223^{*}$ | $5.758^{*}$ | $8.291^{\text {c }}$ | $6.219^{\text {c }}$ |
| $2023 / 24$ |  | $26.607^{\text {d }}$ |  |  |  | $9.220^{\text {d }}$ | $6.916^{\text {d }}$ |

*Incomplete fishery
a. Total catch was sum of retained catch and estimated bycatch mortality of discarded bycatch during crab fisheries and groundfish fisheries.
b. $25 \%$ buffer was applied to total catch OFL to determine ABC except 2021/22, during which $30 \%$ buffer was applied.
c. OFL, and ABC were estimated by the accepted model 21.1e2 in May 2022 assessment when the WAG fisheries was not completed.
d. MSST, MMB, OFL, and ABC were estimated by authors recommended model 22.1e2 in May 2023 assessment when the EAG and WAG fisheries were not completed.

## 6. Basis for the OFL

The length-based model developed for the Tier 3 analysis estimated mature male biomass (MMB) on February 15 each year for the period 1960 through 2022. The terminal year mature male biomass was projected by an additional year to determine OFL and ABC for the 2023/24 season. The Tier 3 approach uses a constant annual natural mortality $(M)$, a knife-edge maturity size, and a mean number of recruits for a selected period for OFL and ABC calculation. An $M$ of $0.22 \mathrm{yr}^{-1}$ was used (Siddeek et al. 2022).

Stock status, OFL, and ABC estimates based on incomplete fisheries data are provided for five models $22.9 \mathrm{c}, 22.1 \mathrm{e} 2$, $22.1 \mathrm{f}, 22.1 \mathrm{~g}$, and 22.1 h for EAG and three models $22.9 \mathrm{c}, 22.1 \mathrm{e} 2$, and 22.1 f for WAG in the following four tables. All models accept 22.9 c were implemented in GMACS. Model 22.9 c was run outside GMACS as the status quo model for comparison of results. The stock statuses of all models were above MMB $_{35 \%}$ (Tier 3a) for EAG, but below MMB $_{35 \%}$ (Tier 3b) for WAG in 2022/23.

## EAG (Tier 3):

Basis for the OFL from the completed fishery data: Biomass, total OFL, and ABC for the next fishing season in 1000 t . Current MMB = MMB on 15 Feb. 2024.

| Model | Tier | MMB $35 \%$ | Current <br> MMB | $\begin{gathered} \text { MMB/ } \\ M M B_{35 \%} \end{gathered}$ | $F_{\text {OFL }}$ | Natural |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  | Recruitment Years to Define $M M B_{35 \%}$ | $F_{35 \%}$ | Mortality | OFL | $\begin{gathered} \mathrm{ABC} \\ (0.75 * \mathrm{OFL}) \end{gathered}$ |
| EAG22.9c | 3a | 6.66545 | 7.48711 | 1.12 | 0.59 | 1987-2017 | 0.59 | 0.22 | 2.951716 | 2.213787 |
| EAG22.1e2 | 3 a | 6.68206 | 7.49393 | 1.12 | 0.59 | 1987-2017 | 0.59 | 0.22 | 2.939008 | 2.204256 |
| EAG22.1f | 3a | 6.69082 | 7.48936 | 1.12 | 0.58 | 1987-2017 | 0.58 | 0.22 | 2.898605 | 2.173954 |
| EAG22.1g | 3 a | 6.61224 | 6.78185 | 1.03 | 0.58 | 1987-2017 | 0.58 | 0.22 | 2.519652 | 1.889739 |
| EAG22.1h | 3a | 6.63743 | 6.71808 | 1.01 | 0.58 | 1987-2017 | 0.58 | 0.22 | 2.484526 | 1.863395 |

Basis for the OFL from the completed fishery data: Biomass, total OFL, and ABC for the next fishing season in millions of pounds. Current MMB $=$ MMB on 15 Feb. 2024.

| Model | Tier | $M_{M B}{ }_{35 \%}$ | Current MMB | $\begin{gathered} \text { MMB/ } \\ M M B_{35 \%} \end{gathered}$ | $F_{\text {OFL }}$ | Recruitment Years to define $M M B_{35 \%}$ | $F_{35 \%}$ | Natural <br> Mortality | OFL | $\begin{gathered} \hline \mathrm{ABC} \\ (0.75 * \mathrm{OFL}) \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| EAG22.9c | 3a | 14.69480 | 16.50625 | 1.12 | 0.59 | 1987-2017 | 0.59 | 0.22 | 6.507 | 4.881 |
| EAG22.1e2 | 3a | 14.73142 | 16.52129 | 1.12 | 0.59 | 1987-2017 | 0.59 | 0.22 | 6.479 | 4.860 |
| EAG22.1f | 3 a | 14.75073 | 16.51121 | 1.12 | 0.58 | 1987-2017 | 0.58 | 0.22 | 6.390 | 4.793 |
| EAG22.1g | 3a | 14.57749 | 14.95142 | 1.03 | 0.58 | 1987-2017 | 0.58 | 0.22 | 5.555 | 4.166 |
| EAG22.1h | 3a | 14.63303 | 14.81083 | 1.01 | 0.58 | 1987-2017 | 0.58 | 0.22 | 5.477 | 4.108 |

## WAG (Tier 3):

Basis for the OFL from the completed fishery data: Biomass, total OFL and ABC for the next fishing season in 1000 t . Current MMB = MMB on 15 Feb .2024.

| Model | Tier | MMB35\% | Current MMB | MMB / <br> $M M B ~_{35 \%}$ | $F_{\text {OFL }}$ | Recruitment Years <br> to Define $M M B_{35 \%}$ | $F_{35 \%}$ | Natural <br> Mortality | OFL | $\begin{gathered} \hline \mathrm{ABC} \\ (0.75 * \mathrm{OFL}) \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| WAG22.9c | 3b | 4.95983 | 4.53168 | 0.914 | 0.4972 | 1987-2017 | 0.55 | 0.22 | 1.232116 | 0.924087 |
| WAG22.1e2 | 3 b | 4.98184 | 4.57482 | 0.918 | 0.4957 | 1987-2017 | 0.55 | 0.22 | 1.243453 | 0.932590 |
| WAG22.1f | 3 b | 4.98048 | 4.44438 | 0.892 | 0.4732 | 1987-2017 | 0.54 | 0.22 | 1.130834 | 0.848125 |

Basis for the OFL from the completed fishery data: Biomass, total OFL, and ABC for the next fishing season in millions of pounds. Current $\mathrm{MMB}=\mathrm{MMB}$ on 15 Feb. 2024.

| Model | Tier | $M_{\text {M }}^{35 \%}$ | Current <br> MMB | $\begin{gathered} \text { MMB/ } \\ M M B_{35 \%} \end{gathered}$ | $F_{\text {OFL }}$ | Recruitment Years to Define$M M B_{35 \%}$ | Natural |  |  | $\begin{gathered} \hline \mathrm{ABC} \\ (0.75 * \mathrm{OFL}) \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  |  | $F_{35 \%}$ | Mortality | OFL |  |
|  |  |  |  |  |  |  |  |  |  |  |
| WAG22.9c | 3a | 10.93455 | 9.99064 | 0.914 | 0.4972 | 1987-2017 | 0.55 | 0.22 | 2.716 | 2.037 |
| WAG22.1e2 | 3 b | 10.98308 | 10.08575 | 0.918 | 0.4957 | 1987-2017 | 0.55 | 0.22 | 2.741 | 2.056 |
| WAG22.1f | 3 a | 10.98008 | 9.79818 | 0.892 | 0.4732 | 1987-2017 | 0.54 | 0.22 | 2.493 | 1.870 |

## 7. Basis for the ABC recommendation

A ' $x$ ' proportion buffer on the OFL, i.e., $\mathrm{ABC}=(1.0-\mathrm{x}) * O F L$, where the authors recommended an $x$ of 0.25 . This proportion was accepted in 2022.

Please see also the section $G$ on ABC .
8. A summary of the results of any rebuilding analysis:

Not applicable.

## A. Summary of Major Changes

1. Changes (if any) to management of the fishery

- None.

2. Changes to input data

- Commercial fisheries data were updated with values from the most recent observer and fish ticket data for 2022/23: retained catch for the directed fishery and discarded catch estimates for the directed fishery, non-directed crab fisheries, and groundfish fisheries. Thus, the time series of data used in the models are retained catch (1981/82-2022/23), total catch (1990/91-2022/23), and groundfish bycatch (1989/90-2022/23) biomass and size compositions.
- Fish ticket retained CPUE were standardized by the generalized linear model (GLM) with the negative binomial link function for the 1985/86-1998/98 period.
- Observer pot sample legal size crab CPUE data were standardized by the GLM with the negative binomial link function with predictor variable selection by CAIC (modified AIC) followed by the R square criterion, separately for 1995/962004/05 (pre-rationalization) and 2005/06-2022/23 (post-rationalization) periods. A Year and Area interaction factor was considered in one model (22.1f) to estimate CPUE indices, separately for pre- and post-rationalization periods. Cooperative survey indices were calculated by the random effects model and incorporate in models 22.1 g and 22.1 h . The habitat areas were determined from observer historical pot locations as fishing footprints (Appendix B).

3. Changes to assessment methodology

None.
4. Changes to assessment results

As expected, the addition of 2022/23 data (incomplete fishery) changed the OFL and ABC estimates, but changes in parameter or abundance estimates were not dramatic.

## B. Response to January 2023 CPT comments

Comment\#1. The confidence intervals for the total catch need to be corrected as they appear to be incorrectly plotted in the assessment document and do not appear to match the assumed CV of 0.2.

Response: We rectified the error below. Total catch confidence intervals are variable because of variable standard deviation of $\ln$ (total catch):


Figure C.1. Observed (open circle) vs. predicted (solid line) log retained catch (top left), $\log$ total catch (top right), and log groundfish bycatch (bottom left) with $95 \%$ confidence intervals of golden king crab under model 22.1 e 2 fit to the EAG data.


Figure C.2. Observed (open circle) vs. predicted (solid line) log retained catch (top left), log total catch (top right), and log groundfish bycatch (bottom left) with $95 \%$ confidence intervals of golden king crab under model 22.1 e 2 fit to the WAG data.

Comment\#2. The time-period for setting the years that define average recruitment should be justified, for example using a plot of years versus the variances of the recruitment deviations. This type of analysis should be included in all future assessments.

Response:
The time for setting the years that define average recruitment was brought up by the SSC in February 2022 and we responded to this question by showing that there were very little differences in the MMB trends and reference point estimates between two hypothetical periods.

The variance analysis is a good suggestion. However, because of limited time available we postpone this analysis to the next assessment cycle.

Comment\#3. The fits to the three CPUE series should be reported on separate plots.
Response: Done. See Figures 19, 20, and 33.

Comment\#4. The combined model (i.e., fitting the data for the EAG and WAG as a singlearea model) led to an OFL that is similar to the sum of those for the assessments of the EAG and WAG separately for the model 21.1 e 2 specifications. However, no fit diagnostics were provided for the combined model so the 2023 assessment should include an appendix with the fit diagnostics.

Response: Because of limited time available we did not take up this analysis in this assessment cycle.

Comment\#5. The rationale for considering model 21.1 f should be included in the assessment document, along with plots that show the extent to which the trend in CPUE varies among locations.

Response: We have provided the rationale for including the Year:Area interaction CPUE model in Appendix B. Because of limited time between January and May, we did not explore the extent to which the trend in CPUE varies among location. This can be done in the next assessment cycle.

Comment\#6. The assessment document should include information on the likely connectivity between the EAG and WAG as this appears to be very limited, justifying separate EAG and WAG assessments.

Response: Done. Information is included in the main document in section C.8.
Comment\#7. The assessment authors provided bridging models to assess the extent to which the assessment of AIGKC can be moved to being conducted using GMACS. The current (bespoke) model and the GMACS implementations provide very similar estimates of the time-series of numbers-at-size and MMB, except during the early (pre-data) period
and the first few years with data. The CPT provided reasons for some discrepancies in the outputs (see January 2023 CPT minutes) and recommended that the May 2023 assessment be conducted using GMACS.

Response: done.
Comment\#8. Recommendation for 2024 assessment:
Models 21.1e2CPUE5Wt and 21.1fCPUE5Wt fit the CPUE data for the EAG much better than the base model (as expected) but without an obvious visual change in the fit to the size-composition data. Models that are forced to achieve better fits to the CPUE indices should be explored; in particular it is necessary to conduct analyses to identify the data sources that preclude the model fitting the CPUE index data well.

Response: Plan to do.
Comment\#9. The models recommended by the CPT are:

- Model 21.1e2: The base model from the May 2022 assessment, except that the pre-specified value of M was changed from $0.21 \mathrm{yr}^{-1}$ to $0.22 \mathrm{yr}^{-1}$ based on a re-analysis of historical tagging data.
- Model 21.1f: As for model 21.1e2, plus observer CPUE data standardized including Year: Block interactions.
- Model 21.1g: As for model 21.1e2, but with the EAG cooperative survey standardized CPUE included.
- A model similar to 21.1 g but with 21.1 f as the base rather than 21.1 e 2 .

Response: done

## Response to February 2023 SSC comments:

Comment\#1. The SSC approves the above CPT recommended models for the May 2023 CPT presentation.

## Response: Implemented in this assessment..

Comment\#2. The SSC appreciates the work by the assessment author and others to transition this assessment to the GMACS framework as well as the detailed bridging analysis. The SSC agrees with the CPT that the May 2023 assessment be conducted using GMACS only and that the legacy model not be brought forward for the May assessment. However, the SSC requests that the base GMACS model EAG21.9c (modified 21.e2), which closely follows the legacy model, be included to facilitate comparisons with the previous bridging exercises.

Response: In this assessment, results of model 22.9c (modified 22.1e2) were compared with the GMACS output.

Comment\#3. The SSC agrees that the 1987-2017 recruitment period be used for this assessment, but that for future assessments the authors continue to consider other recruitment
time periods (including routinely adding a year to the series as is done in other assessments) and provide justification for the final choice.

Response: Our response to CPT comment\#2 touches this issue.

## C. Introduction

## 1. Scientific name:

Golden king crab, Lithodes aequispinus J.E. Benedict, 1895.

## 2. Distribution:

General distribution of golden king crab is summarized by NMFS (2004). Golden king crab, also called brown king crab, occur from the Sea of Japan to the northern Bering Sea (ca. $61^{\circ} \mathrm{N}$ latitude), around the Aleutian Islands, generally in high-relief habitat such as inter-island passes, on various sea mounts, and as far south as northern British Columbia (Alice Arm) (Jewett et al. 1985). They are typically found on the continental slope at depths of $300-1,000 \mathrm{~m}$ on extremely rough bottom. They are frequently found on coral bottom.

The Aleutian Islands king crab stock boundary is defined by the boundaries of the Aleutian Islands king crab Registration Area O (Figure 1). In this chapter, "Aleutian Islands Area" means the area described by the current definition of Aleutian Islands king crab Registration Area O. Nichols et al. (2021) define the boundaries of Aleutian Islands king crab Registration Area O:

The Aleutian Islands king crab Registration Area $O$ eastern boundary is the longitude of Scotch Cap Light ( $164^{\circ} 44.72^{\prime}$ W long); the northern boundary is a line from Cape Sarichef ( $54^{\circ} 36^{\prime} \mathrm{N}$ lat) to $171^{\circ} \mathrm{W}$ long, north to $55^{\circ} 30^{\prime} \mathrm{N}$ lat; and the western boundary the United States-Russia Maritime Boundary Line of 1990.

During 1984/85-1995/96, the Aleutian Islands king crab populations had been managed using the Adak and Dutch Harbor Registration Areas, which were divided at $171^{\circ} \mathrm{W}$ longitude (Figure 2), but from the 1996/97 season to present the fishery has been managed using a division at $174^{\circ} \mathrm{W}$ longitude (Figure 1). In March 1996, the Alaska Board of Fisheries (BOF) replaced the Adak and Dutch Harbor areas with the newly created Aleutian Islands Registration Area O and directed ADF\&G to manage the golden king crab fishery in the areas east and west of $174^{\circ} \mathrm{W}$ longitude as two distinct stocks. That re-designation of management areas was intended to reflect golden king crab stock distribution, congruent with the longitudinal pattern in fishery production prior to 1996/97 (Figure 3). The longitudinal pattern in fishery production relative to $174^{\circ} \mathrm{W}$ longitude since 1996/97 is like that observed prior to the change in management area definition, although there have been some changes in the longitudinal pattern in fishery production within the areas east and west of $174^{\circ} \mathrm{W}$ longitude (Figure 4).

Commercial fishing for golden king crab in the Aleutian Islands Area typically occurs at depths of 100-275 fathoms (183-503 m) (Gaeuman 2014). Pots sampled by at-sea fishery observers during 1990/91-2022/23 were fished at an average depth of 181 fathoms ( $331 \mathrm{~m} ; \mathrm{N}=57,792$ ) in the area east of $174^{\circ} \mathrm{W}$ longitude and 178 fathoms ( $326 \mathrm{~m} ; \mathrm{N}=62,062$ ) for the area west of $174^{\circ} \mathrm{W}$ longitude.

## 3. Evidence of stock structure:

Given the expansiveness of the Aleutian Islands Area and the existence of deep ( $>1,000 \mathrm{~m}$ ) canyons between some islands, at least some weak structuring of the stock within the area would be expected. Data for making inferences on stock structure of golden king crab within the Aleutian Islands are largely limited to the geographic distribution of commercial fishery catch and effort. Catch data by statistical area from fish tickets and catch data by location from pots sampled by observers suggest that habitat for legal-sized males may be continuous throughout the waters adjacent to the islands in the Aleutian chain. However, regions of low fishery catch suggest that availability of suitable habitat, in which golden king crab are present at only low densities, may vary longitudinally. Catch has been low in the fishery in the area between $174^{\circ} \mathrm{W}$ longitude and $176^{\circ} \mathrm{W}$ longitude (the Adak Island area, Figures 3 and 4) in comparison to adjacent areas, a pattern that is consistent with low CPUE for golden king crab between $174^{\circ} \mathrm{W}$ longitude and $176^{\circ} \mathrm{W}$ longitude (Figure 5) during the 2002, 2004, 2006, 2010, and 2012 NMFS Aleutian Islands bottom trawl surveys (von Szalay et al. 2011, 2017). In addition to longitudinal variation in density, there is also a gap in fishery catch and effort between the Petrel Bank-Petrel Spur area and the Bowers Bank area; both of those areas, which are separated by Bowers Canyon, have reported effort and catch. Recoveries during commercial fisheries of golden king crab tagged during ADF\&G surveys (Blau and Pengilly 1994; Blau et al. 1998; Watson and Gish 2002; Watson 2004 , 2007) provided no evidence of substantial movements by crab in the size classes that were tagged (males and females $\geq 90-\mathrm{mm}$ carapace length [CL]). Maximum straight-line distance between release and recovery location of 90 golden king crab released prior to the 1991/92 fishery and recovered through the 1992/93 fishery was 61.2 km (Blau and Pengilly 1994). Of the 4,567 recoveries reported through April 12, 2016, for the male and female golden king crab tagged and released between $170.5^{\circ} \mathrm{W}$ longitude and $171.5^{\circ} \mathrm{W}$ longitude during the 1991 , 1997, 2000, 2003, and 2006 ADF\&G Aleutian Island golden king pot surveys, none of the 3,807 with recovery locations specified by latitude and longitude were recovered west of $173^{\circ} \mathrm{W}$ longitude and only fifteen were recovered west of $172^{\circ}$ W longitude (V. Vanek, ADF\&G, Kodiak, pers. comm.). Similarly, of 139 recoveries in which only the statistical area of recovery was reported, none were recovered in statistical areas west of $173^{\circ} \mathrm{W}$ longitude and only one was in a statistical area west of $172^{\circ} \mathrm{W}$ longitude. Thus, little mixing of Dutch and Adak areas golden king crab provide a reason to undertaking separate stock assessment in each area.
4. Life history characteristics relevant to management:

There is a paucity of information on golden king crab life history characteristics due in part to the deep depth distribution ( $\sim 200-1000 \mathrm{~m}$ ) and the asynchronous nature of life history events (Otto and Cummiskey 1985; Somerton and Otto 1986). The reproductive cycle is thought to last approximately 24 months and at any time of year ovigerous females can be found carrying egg clutches in highly disparate developmental states (Otto and Cummiskey 1985). Females carry large, yolk-rich, eggs, which hatch into lecithotrophic (i.e., the larvae can develop successfully to juvenile crab without eating; Shirley and Zhou 1997) larvae that are negatively phototactic (Adams and Paul 1999). Molting and mating are also asynchronous and protracted (Otto and Cummiskey 1985; Shirley and Zhou 1997) with some indications of seasonality (Hiramoto 1985). Molt increment for large males (adults) in Southeast Alaska is 16.3 mm CL per molt (Koeneman and Buchanan 1985) and was estimated at 14.4 mm CL for legal males in the eastern Aleutian Islands (Watson et al. 2002). Annual molting probability of males decreases with increasing size, which results in a protracted inter-molt period and creates difficulty in determining annual molt probability (Watson et al. 2002). Male size-at-maturity varies among stocks (Webb 2014) and declines with increasing latitude from about 130 mm CL in the Aleutian Islands to 92 mm CL in Saint Matthew Island section (Somerton and Otto 1986). Along with a lack of annual survey data, limited stockspecific life history stock information prevents development of a comprehensive length-based assessment model.

## 5. Summary of management history:

A complete summary of the management history through 2015/16 is provided in Leon et al. (2017). The first commercial landing of golden king crab in the Aleutian Islands was in 1975/76 but directed fishing did not occur until 1981/82.

The Aleutian Islands golden king crab fishery was restructured beginning in 1996/97 to replace the Adak and Dutch Harbor areas with the newly created Aleutian Islands Registration Area O and golden king crab in the areas east and west of $174^{\circ} \mathrm{W}$ longitude were managed separately as two stocks (ADF\&G 2002). Hereafter, the east of $174^{\circ} \mathrm{W}$ longitude stock segment is referred to as EAG and the west of $174^{\circ} \mathrm{W}$ longitude stock segment is referred to as WAG. Table 1 provides the historical summary of number of vessels, GHL/TAC, harvest, effort, CPUE, and average weight of crab in the Aleutian Islands golden king crab fishery.

The fisheries in 1996/97-1997/98 were managed with GHLs of 1,452 t (3,200,000 lb) in EAG and $1,225 \mathrm{t}(2,700,000 \mathrm{lb})$ in WAG (Table 1). During 1998/99-2004/05 the fisheries were managed with GHLs of $1,361 \mathrm{t}(3,000,000 \mathrm{lb})$ for EAG and 1,225 $\mathrm{t}(2,700,000 \mathrm{lb})$ for WAG. During 2005/06-2007/08 the fisheries were managed with a total allowable catch (TAC) of $1,361 \mathrm{t}(3,000,000 \mathrm{lb})$ for EAG and a TAC of $1,225 \mathrm{t}(2,700,000 \mathrm{lb})$ for WAG. By state regulation (5 AAC 34.612), TAC for the Aleutian Islands golden king crab fishery during 2008/09-2011/12 was 1,429 t $(3,150,000 \mathrm{lb})$ for EAG and $1,286 \mathrm{t}(2,835,000 \mathrm{lb})$ for WAG. In March 2012, the BOF changed 5 AAC 34.612 so that the TAC beginning in 2012/13 would be 1,501
$\mathrm{t}(3,310,000 \mathrm{lb})$ for EAG and $1,352 \mathrm{t}(2,980,000 \mathrm{lb})$ for WAG. Additionally, the BOF added a provision to 5 AAC 34.612 that allows ADF\&G to lower the TAC below the specified level if conservation concerns arise. The TAC for 2016/17 (and $2017 / 18$ ) was reduced by $25 \%$ for WAG to $1,014 \mathrm{t}(2,235,000 \mathrm{lb})$ while keeping the TAC for EAG at the same level as the previous season.

During 1996/97-2022/23 the annual retained catch during commercial fishing (including cost-recovery fishing that occurred during 2013/14-2022/23) has averaged $2 \%$ below the annual GHL/TACs but has ranged from as much as $13 \%$ below (1998/99) to $6 \%$ above (2000/01) the GHL/TAC.

A summary of other relevant State of Alaska fishery regulations and management actions pertaining to the Aleutian Islands golden king crab fishery is provided below:

Beginning in 2005/06, the Aleutian Islands golden king crab fishery has been prosecuted under the Crab Rationalization Program. Accompanying the adoption of crab rationalization program was implementation of a community development quota (CDQ) fishery for golden king crab in the eastern Aleutians (i.e., EAG) and the Adak Community Allocation (ACA) fishery for golden king crab in the western Aleutians (i.e., WAG; Hartill 2012; Nichols et al. 2021). The CDQ fishery in the eastern Aleutians is allocated $10 \%$ of the golden king crab TAC for the area east of $174^{\circ} \mathrm{W}$ longitude and the ACA fishery in the western Aleutians is allocated $10 \%$ of the golden king crab TAC for the area west of $174^{\circ} \mathrm{W}$ longitude. The CDQ fishery and the ACA fishery are managed by ADF\&G and prosecuted concurrently with the individual fishing quota (IFQ) fishery.

Golden king crab may be commercially fished only with king crab pots (defined in state regulation 5 AAC 34.050). Pots used to fish for golden king crab in the Aleutian Islands Area must be longlined and, since 1996, each pot must have at least four escape rings of five and one-half inches minimum inside diameter installed on the vertical plane or at least one-third of one vertical surface of the pot composed of not less than nine-inch stretched mesh webbing to permit escapement of undersized golden king crab [5 AAC 34.625 (b)]. Prior to the regulation requiring an escape mechanism on pots, some participants in the Aleutian Islands golden king crab fishery voluntarily sewed escape rings (typically 139 mm [5.5 inches]) into their gear or, more rarely, included panels with escape mesh (Beers 1992). Regarding the gear used since the establishment of 5 AAC 34.625 (b) in 1996, Linda Kozak, a representative of the industry, reported in a 19 September 2008 email to the Crab Plan Team, "... the golden king crab fleet has modified their gear to allow for small crab sorting," and provided a written statement from Lance Nylander, of Dungeness Gear Works (DGW) in Seattle, who "believes he makes all the gear for the golden king crab harvesting fleet," saying that "Since 1999, DGW has installed 9 [-inch] escape web on the door of over $95 \%$ of Golden Crab pot orders manufactured." A study to estimate the contact-selection curve for male golden king crab was conducted aboard one vessel commercial fishing for golden king crab
during the 2012/13 season and found gear and fishing practices used by that vessel were highly effective in reducing bycatch of sublegal-sized males and females (Vanek et al. 2013). In March 2011 (effective for 2011/12), the BOF amended 5 AAC 34.625 (b) to relax the "biotwine" specification for pots used in the Aleutian Islands golden king crab fishery relative to the requirement in 5 AAC 39.145 that "(1) a sidewall ...of all shellfish and bottom fish pots must contain an opening equal to or exceeding 18 inches in length... The opening must be laced, sewn, or secured together by a single length of untreated, 100 percent cotton twine, no larger than 30 thread." Regulation 5 AAC 34.625 (b)(1) allows the opening described in 5 AAC 39.145 (1) to be "laced, sewn, or secured together by a single length of untreated, 100 percent cotton twine, no larger than 60 [rather than 30] thread."

Regulation (5 AAC 34.610 (b)) sets the commercial fishing season for golden king crab in the Aleutian Islands Area as 1 August through 30 April. That regulatory fishing season became effective in 2015/16 (the commercial fishing season was set in regulation as 15 August through 15 May during 2005/06-2014/15).

Current regulations ( 5 AAC $39.645(d)(4)(A))$ stipulate that onboard observers are required on catcher vessels during the time that at least $50 \%$ of the retained catch is captured in each of the three trimesters of the 9 -month fishing season. Onboard observers are required for $100 \%$ of fishing activity on catcher-processor vessels during the crab fishing season.

In addition, the commercial golden king crab fishery in the Aleutian Islands Area may only retain males at least 6.0 -inches $(152.4 \mathrm{~mm})$ carapace width ( CW ), including spines ( 5 AAC $34.620(b)$ ), which is at least one annual molt increment larger than the $50 \%$ maturity length of 120.8 mm CL for males as estimated by Otto and Cummiskey (1985). A carapace length (CL) $\geq 136 \mathrm{~mm}$ is used to identify legalsize males when CW measurements are not available (Table 3-5 in NPFMC 2007b). Note that the size limit for golden king crab has been 6-inches ( 152.4 mm ) CW for the entire Aleutian Islands Area since the 1985/86 season. Prior to the 1985/86 season, the legal-size limit was 6.5 -inches $(165.1 \mathrm{~mm}) \mathrm{CW}$ for at least one of the now-defunct Adak or Dutch Harbor Registration Areas.

The male maturity size using 1991 pot survey measurements of carapace length and chela height in EAG and 1984 NMFS measurements in WAG were re-evaluated (Siddeek et al. 2018). Bootstrap analysis of chela height and carapace length data provided the median $50 \%$ male maturity length estimates of 107.02 mm CL in EAG and 107.85 mm CL in WAG. The knife-edge maturity size of 111.0 mm CL, which is the lower limit of the next upper size bin, has been used for mature male biomass (MMB) estimation. Recently collected (2018 to 2020) chela height and carapace length data were analyzed and proposed a higher knife-edge maturity length of 116.0 mm CL for MMB calculation, which was accepted by the CPT/SSC in 2022.

Daily catch and catch-per-unit effort (CPUE) are determined in-season to monitor fishery performance and progress towards the respective TACs. Figures 6 to 8
provide the 1985/86-2022/23 time series of catch, CPUE, and the geographic distribution of catch during the 2022/23 fishing season. Increases in CPUE were observed during the late 1990s through the early 2000s, and with the implementation of crab rationalization in 2005. This was likely due to changes in gear configurations in the late 1990s (crab harvesters, personal communication, 1 July 2008) and, after rationalization due to increased soak time (Siddeek et al. 2015), and decreased competition owing to the reduced number of vessels fishing. Decreased competition could allow crab vessels to target only the most productive fishing areas. Trends in fishery nominal CPUE within the areas EAG and WAG generally paralleled each other during 1985/86-2010/11 but diverged thereafter (EAG CPUE exceeded one and half times of that in WAG). A moderate decreasing trend in CPUE was observed since 2014 in EAG and since 2019 in WAG (Figures 6 and 7).

## 6. Brief description of the annual ADF \&G harvest strategy:

In March 2019, the BOF adopted a revised harvest strategy (Daly et al. 2019). The annual TAC is set by state regulation, 5 AAC 34.612 (Harvest Levels for Golden King Crab in Registration Area O), per:
(a) In that portion of the Registration Area O east of $174^{\circ} \mathrm{W}$. long., the total allowable catch level shall be established as follows:
(1) if $\mathrm{MMA}_{\mathrm{E}}$ is less than 25 percent of $\mathrm{MMA}_{\mathrm{E},(1985-2017) \text {, the fishery will not }}$ open;
(2) if $\mathrm{MMA}_{\mathrm{E}}$ is at least 25 percent but not greater than 100 percent of MMA $_{\mathrm{E},(1985-2017)}$, the number of legal male golden king crab available for harvest will be computed as (0.15) x $\left(\mathrm{MMA}_{\mathrm{E}} / \mathrm{MMA}_{\mathrm{E},(1985-2017)}\right)$ $\mathrm{x}\left(\mathrm{MMA}_{\mathrm{E}}\right)$ or 25 percent of $\mathrm{LMA}_{\mathrm{E}}$, whichever is less; and
(3) if $\mathrm{MMA}_{\mathrm{E}}$ is greater than 100 percent of $\mathrm{MMA}_{\mathrm{E},(1985-2017), \text {, the number of }}$ legal male golden king crab available for harvest will be computed as $(0.15) \times\left(\mathrm{MMA}_{\mathrm{E}}\right)$ or 25 percent of $\mathrm{LMA}_{\mathrm{E}}$, whichever is less.
(b) In that portion of the Registration Area O west of $174^{\circ} \mathrm{W}$. long., the total allowable catch level shall be established as follows:
(1) if MMAW is less than 25 percent of $\mathrm{MMA}_{W}$, (1985-2017), the fishery will not open;
(2) if MMAw is at least 25 percent but not greater than 100 percent of MMAW, (1985-2017), the number of legal male golden king crab available for harvest will be computed as (0.20) x (MMAw/MMA ${ }_{\mathrm{W},(1985-2017)}$ ) $\mathrm{x}\left(\mathrm{MMAw}_{\mathrm{w}}\right.$ ) or 25 percent of LMAw, whichever is less; and
(3) if MMAW is greater than 100 percent of $\mathrm{MMA}_{\mathrm{w},(1985-2017),}$, the number of legal male golden king crab available for harvest will be computed as ( 0.20 ) $\mathrm{x}\left(\mathrm{MMA}_{w}\right)$ or 25 percent of $\mathrm{LMA}_{w}$, whichever is less.
(c) In implementing this harvest strategy, the department shall consider the reliability of estimates of golden king crab, the manageability of the fishery, and other factors the department determines necessary to be consistent with sustained yield principles and to use the best scientific information available and consider all sources of uncertainty as necessary to avoid overfishing.
(d) In this section,
(1) $\mathrm{MMA}_{\mathrm{E}}$ means the abundance of male golden king crab in the portion of the Aleutian Islands Management Area O east of $174^{\circ} \mathrm{W}$. long that are greater than or equal to 111 millimeters in carapace length estimated by the stock assessment model for the time prior to the start of the fishery.
(2) $\mathrm{MMA}_{\mathrm{E},(1985-2017)}$ means the mean value of the abundance of male golden king crab in the portion of the Aleutian Islands Management Area O east of $174^{\circ} \mathrm{W}$. long that are greater than or equal to 111 millimeters in carapace length estimated by the stock assessment model for the time prior to the start of the fishery for the period $1985-2017$.
(3) $\mathrm{LMA}_{\mathrm{E}}$ means the abundance of male golden king crab in the portion of the Aleutian Islands Management Area O east of $174^{\circ} \mathrm{W}$. long that are greater than or equal to 136 millimeters in carapace length estimated by the stock assessment model for the time prior to the start of the fishery.
(4) MMAw means the abundance of male golden king crab in the portion of the Aleutian Islands Management Area O west of $174^{\circ} \mathrm{W}$. long that are greater than or equal to 111 millimeters in carapace length estimated by the stock assessment model for the time prior to the start of the fishery.
(5) MMAw, (1985-2017) means the mean value of the abundance of male golden king crab in the portion of the Aleutian Islands Management Area O west of $174^{\circ} \mathrm{W}$. long that are greater than or equal to 111 millimeters in carapace length estimated by the stock assessment model for the time prior to the start of the fishery for the period 1985 - 2017; and
(6) $\mathrm{LMA}_{\mathrm{w}}$ means the abundance of male golden king crab in the portion of the Aleutian Islands Management Area O west of $174^{\circ} \mathrm{W}$. long that are greater than or equal to 136 millimeters in carapace length estimated by the stock assessment model for the time prior to the start of the fishery.

In addition to the retained catch that is limited by the TAC established by ADF\&G under 5 AAC 34.612, ADF\&G has authority to annually receive receipts up to $\$ 500,000$ through cost-recovery fishing on Aleutian Islands golden king crab. The retained catch from that cost-recovery fishing is not counted against attainment of the annually established TAC.
7. Summary of the history of the basis and estimates of $M M B_{M S Y}$ or proxy MMB $_{\text {MSY: }}$
The $M M B_{35 \%}$ is estimated as a proxy for $M M B_{M S Y}$ using the Tier 3 estimation procedure, which is explained in a subsequent section.
8. Justification for assessing Aleutian Islands golden king crab as two sub stocks:

We modelled EAG and WAG stocks separately for several reasons:
(a) Fishery catch data (e.g., CPUE magnitude and CPUE temporal trends) suggest that the productivity is different between the two areas.
(b) WAG has wider area of stock distribution compared to limited area distribution in EAG.
(c) The fishing areas are spatially separated with an area gap between EAG and WAG (see Figure 8 in the main text). Regions of low fishery catch suggest that availability of suitable habitat may vary longitudinally.
(d) Tagging studies have shown little mixing between the two areas (Watson and Gish 2002).
(e) Currents are known to be strong around the Aleutian Islands, thus larval mixing between the two regions may occur. Yet needed data to confirm larval drift trajectories or horizontal displacement are lacking. Unlike other king crabs, golden king crab females carry large, yolk-rich, eggs, which hatch into lecithotrophic (nonfeeding) larvae that do not require a pelagic distribution for encountering food items. Depth at larval release, the lecithotrophic nature of larvae, and swimming inactivity in lab studies implies benthic distributions, which may limit larval drift between areas if horizontal current velocities are reduced at depth.
(f) Integrating contrasting data in one single model may provide parameter estimates in between the two extremes which would not be applicable to either (Richards 1991; Schnute and Hilborn 1993).
(g) Area specific assessment is superior to a holistic approach for this stock because of patchy nature of golden king crab distribution.
h) Alaska Board of Fisheries decided to manage the two areas with separate total allowable catches.
i) Genetic analysis shows no significant differentiation between areas within the Aleutian Island population (Grant and Siddon 2018), thus there is no genetic support for subdividing this population; however, above listed factors support separate stock assessments in the two regions.

## D. Data

1. Summary of new information:
(a) Commercial fishery retained catch by size, estimated total catch by size, groundfish male discard catch by size, observer CPUE index, and commercial fishery CPUE index were updated to include 2022/23 information. Available data by year are shown below.


## 2. Data presented as time series:

## a. Total Catch:

Fish ticket data on retained catch weight, catch numbers, effort (pot lifts), CPUE, and average weight of retained catch for 1981/82-2022/23 are provided in Table 1. Estimated total catch weights for 1990/91-2022/23 are listed in Table 2a.

## b. Bycatch and discards:

Retained catch, bycatch mortality (male and female) separated by the crab fishery and groundfish fishery, and total fishery mortality for 1981/82-2022/23 are provided in Table 2. Crab fishery discards are available after observer sampling was established in 1988/89. Observer data for the 1988/89-1989/90 seasons are not considered reliable. Table 2 provides crab fishery discards and groundfish fishery bycatch for the 1991/92-2022/23 seasons.

## c. Catch-per-unit-effort:

- Pot fishery and observer nominal retained and total CPUE, pot fishery effort, observer sample size, and estimated observer CPUE index delineated by EAG and WAG for 1985/86-2022/23 are provided in Table 3.
- Estimated commercial fishery CPUE index with coefficient of variation are provided in Table 4 for EAG and Table 10 for WAG. The estimation methods and CPUE fits are described in Appendix B.


## d. Abundance-at-length:

Information on length compositions of abundance (N-Matrix) are provided in Figures 9 and 10 for EAG; and 23 and 24 for WAG for models 22.1 e 2 and 21.1f, respectively.

## e. Survey biomass estimates:

Cooperative pot survey estimate of biomass indices (2015/16 to 2019/20, and 2021/22 to 2022/23) were considered in the stock assessment models in this presentation.

## f. Survey catch-at-length:

Data are available but not processed for this presentation.
g. Other time series data: None.

## 3. Data which may be aggregated over time:

- Molt and size transition matrix: Tag release - recapture -time at liberty records from 1991, 1997, 2000, 2003, and 2006 male tag crab releases were aggregated by year at liberty to determine the molt increment and size transition matrix by the integrated model.
- Weight-at-length: Male length-weight relationship: $\mathrm{W}=\mathrm{aL}^{\mathrm{b}}$ where $a=1.445^{*} 10^{-4}, b=3.28113\left[\sigma_{a}=0.00737\right.$ (bias correction for $a$ was not required because of the very small value of $\sigma_{a}$ ); updated estimates from WAG data].
- Natural mortality: Tag-recapture model (Siddeek et al., 2022) estimated fixed natural mortality value of $0.22 \mathrm{yr}^{-1}$ was used in the assessment.


## 4. Information on any data sources that were available, but were excluded from the assessment:

Data from triennial ADF\&G pot surveys for Aleutian Islands golden king crab in a limited area in EAG (between $170^{\circ} 21^{\prime}$ and $171^{\circ} 33^{\prime} \mathrm{W}$ longitude) that were performed during 1997 (Blau et al. 1998), 2000 (Watson and Gish 2002), 2003 (Watson 2004), and 2006 (Watson 2007) are available, but were not used in this assessment. However, the tag release and recapture data from these surveys were used.

## E. Analytic Approach

1. History of modeling approaches for this stock:

A size structured assessment model based on only fisheries data was under development for several years for the EAG and WAG golden king crab stocks and accepted in 2016 for OFL and ABC setting for the 2017/18 season. The CPT in January 2017 and SSC in February 2017 recommended using the Tier 3 procedure to set the OFL and ABC. They also suggested using the maturity data to estimate the male mature biomass (MMB). These suggestions were followed in this assessment to estimate OFL and ABC.

## 2. Model Description:

a. Description of overall modeling approach:

The underlying population dynamics model was male-only and length-based (Appendix A). This model combined commercial retained catch, total catch, groundfish fishery discarded catch, standardized observer legal size catch-per-unit-effort (CPUE) indices, fishery retained catch size composition, total catch size composition, and tag recaptures by release-recapture length to estimate stock assessment parameters. The tagging data were used to calculate the size transition matrix. To estimate the MMB, a knife-edge maturity size based on chela height and carapace length data was used. To include a long time series of CPUE indices for stock abundance contrast, the 1985/86-1998/99 legal size standardized fishery CPUE indices were used in a separate likelihood component in all models (Table T1). All models but model 22.9c were
implemented in GMACS. Model 22.9c is the modified 22.1 e 2 model for GMACS implementation and used to compare its results with those of GMACS.

There were significant changes in fishing practice associated with changes in management regulations (e.g., constant GHL/TAC since 1996/97 and crab rationalization since 2005/06), pot configuration (escape web on the pot door increased to 9 -inch since 1999), and improved observer recording in Aleutian Islands golden king crab fisheries since 1998. These changes prompted us to consider two sets of total selectivity parameters with only one set of retention parameters for the periods 1985/86-2004/05 and 2005/06-2022/23. Since CPUE standardization was carried out based on three independent sets of data (fish ticket CPUE for the period 1985 - 1998; and observer CPUE for the two separate periods, 1995 - 2004 and 2005 - 2022), several models (22.1e2, 22.1f, 22.1 g , and 22.1 h ) with three catchability and additional CV parameters were considered for GMACS implementation, which seems more justifiable (Table T1).

The observer and commercial fishery CPUE indices with standard errors (estimated by GLM) and an additional assessment model estimable constant variance were fitted in all models. The assessment model predicted total and retained CPUE indices. However, only predicted retained CPUE indices were compared with input legal size crab CPUE indices in the likelihood functions because recordings of legal-size crab seem reliable.

The data series ranges used for WAG are the same as those for EAG.
b. Software:

AD Model Builder (Fournier et al. 2012).
c. -f . Details are given in Appendix A.
g. Critical assumptions and consequences of assumption failures:

Because of the lack of an annual stock survey, assessment models heavily relied on standardized CPUE indices (Appendix B) and catch and size composition information to determine stock abundance trends in both regions. For this analysis, the observer and fish ticket CPUE indices were assumed to be linearly related to exploitable abundance. The $M$ was kept constant at $0.22 \mathrm{yr}^{-1}$ and the knife-edge maturity size set at 116 mm CL, based on recent chela height analysis, was considered for MMB estimation (Siddeek et al. 2018, 2021, 2022). In addition, directed pot fishery discard mortality was set at $0.20 \mathrm{yr}^{-1}$, overall groundfish fishery mortality set at $0.65 \mathrm{yr}^{-1}$ (mean of groundfish pot fishery mortality $\left[0.5 \mathrm{yr}^{-1}\right]$ and groundfish trawl fishery mortality $\left[0.8 \mathrm{yr}^{-1}\right]$ ), and groundfish fishery selectivity set at full selection for all length classes (selectivity = 1.0). Any discard of legal-size males in the directed pot fishery was not considered in this analysis. These fixed values invariably reduced the number of model parameters to be estimated and helped in convergence.

Different $q$ 's (scaling parameter for standardized CPUE in the models, Equation A.13) and logistic selectivity patterns (Equation A.9) for different periods for the pot fishery were considered.
h. Changes to any of the above since the previous assessment: None.
i. Model code has been checked and validated.

GMACS and status quo models' codes have been checked at various times by independent reviewers and the current codes are available from the first author.

## 3. Model Selection and Evaluation:

## a. Description of alternative model configurations:

Five and three alternative models were considered for EAG and WAG, respectively (see Table T1 for description of alternative models). The OFL and $A B C$ results are presented for all models separately for EAG and WAG in the executive summary tables. The OFL and ABC are also provided for authors preferred model 22.1 e 2 for AI. Model 22.1 e 2 was considered as the base model. It considers:
i) Initial abundance by the equilibrium condition considering the mean number of recruits for 1987-2017: The equilibrium abundance was determined for 1960 (Equations A. 4 and A.5), projected forward with only $M$ and annual recruits until 1980, then retained catch removed during 19811984 and projected to obtain the initial abundance in 1985.
ii) Observer CPUE indices for 1995/96-2022/23.
iii) Fishery CPUE indices for 1985/86-1998/99.
iv) Initial (Stage-1) weighting of effective sample sizes: number of vessel-days for retained and total catch size compositions, and number of fishing trips for groundfish discard size composition (the groundfish size composition was not used in model fitting); and (Stage-2) iterative re-weighting of effective sample sizes by the Francis method.
v) Three catchabilities and two sets of logistic total selectivities for the periods 1985/86-1998/99, 1985/86-2004/05, and 2005/06-2022/23, and a single set of logistic retention curve parameters.
vi) Full selectivity (selectivity $=1.0$ ) for groundfish fishery bycatch.
vii) Knife-edge maturity size of 116 mm CL.
viii) Stock dynamics $M=0.22 \mathrm{yr}^{-1}$, pot fishery handling mortality $=0.2 \mathrm{yr}^{-1}$, and mean groundfish bycatch handling mortality $=0.65 \mathrm{yr}^{-1}$.
ix) Size transition matrix using tagging data estimated by the normal probability function with the logistic molt probability sub-model. The tagrecaptures were treated as Bernoulli trials (i.e., Stage-1 weighting).
x) The period, 1987-2017, was used to determine the mean number of recruits for $M M B_{35 \%}$ (a proxy for $M M B_{M S Y}$ ) estimation under Tier 3.

The salient features and variations from the base scenario of all other scenarios are listed in Table T1. The list of fixed and estimable parameters is provided in Table A1 and detail weights with coefficient of variations (CVs) assigned to each type of data are listed in Table A2.

Best estimates of parameter values for a representative model, 22.1e2, were jittered to confirm model global convergence. The results indicated that global convergence was achieved (Appendix D).

## b. Progression of results:

The OFL and ABC estimates are like those estimates made previously, 2021/22.
c. Label the approved model from the previous year as model:

The notation 22.1 e 2 for the base model came from the previous year (i.e., May/June 2022) accepted assessment model, 21.1e2. The model 22.1e2 includes additional 2022/23 data (unfinished fisheries) over the previous model.
d. Evidence of search for balance between realistic and simpler models:

Unlike annually surveyed stocks, Aleutian Islands golden king crab stock biomass is difficult to track, and several biological parameters are assumed based on knowledge from red king crab (e.g., handling mortality rate of 0.2 $\mathrm{yr}^{-1}$ ) due to a lack of species/stock specific information. Several model parameters were fixed after initially running the model with free parameters to reduce the number of parameters to be estimated (e.g., groundfish bycatch selectivity parameters were fixed). In CPUE standardization, instead of using the traditional AIC the Consistent Akaike Information Criteria (Bozdogan 1987) was used that considers number of parameters and data points used for fitting models when selecting the final model. The final GLM model predictor variables were tested for collinearity by the GIF statistics (Fox and Weisberg 2011) and found them to be non-collinear. The assessment models also considered different configuration of parameters to select parsimonious models. The detailed results of all models are provided in tables and figures.
e. Convergence status and criteria:

ADMB default convergence criteria were used.
f. Table of the sample sizes assumed for the size compositional data:

The initial input sample sizes (i.e., Stage-1) were estimated either as number of vessel-days for retained and total catch size compositions or number of fishing trips for groundfish size composition (note: the groundfish size composition was not used in model fitting) for all model scenarios. Then the Stage-2 effective sample sizes were estimated iteratively from Stage-1 sample sizes using the Francis' $(2011,2017)$ mean length-based method.

The initial input sample sizes (Stage-1) and Stage-2 effective sample sizes are provided for the bespoke model 22.9c in Table 5 for EAG and Table 11 for WAG.
g. Provide the basis for data weighting, including whether the input effective sample sizes are tuned, and the survey CV adjusted:
Described previously (f).
h. Do parameter estimates make sense and are they credible?

The estimated parameter values are within the bounds and various plots suggest that the parameter values are reasonable for the golden king crab stocks with the current sets of data.
i. Model selection criteria:

Several diagnostic criteria were used to select appropriate models for recommendation: CPUE fits, observed vs. predicted tag recapture numbers by time at large and release size, retained and total catch, and groundfish bycatch fits. Figures are provided for all model scenarios in the Results section.
j. Residual analysis:

Residual fits are illustrated by bubble plots for retained and total catch size composition predictions in various figures in the Results section.

## k. Model evaluation:

Only one base model with several model variations is presented and the evaluations are presented in the Results and Graphs sections below.

## 4. Results:

## 1. List of effective sample sizes and weighting factors:

The Stage-1 and Stage-2 effective sample sizes are listed for model 22.9c in Table 5 for EAG and Table 11 for WAG. The weights, with the CV specifications, for different data sets are provided in Table A2 for various models for both EAG and WAG. These weights (with the corresponding CV) adequately fitted the length compositions, and no further changes were examined.

Weighting factors were used for catch biomass, recruitment deviation, pot fishery $F$, and groundfish fishery $F$. The retained catch biomass weight was set to an arbitrarily large value 500.0 (corresponding to a CV of 0.0316 ) because retained catch data are more reliable than any other data sets. The total catch biomass weight was scaled in accordance with the observer annual sample sizes (number of pots) with a maximum of 250.0 (corresponding to variable CV). Total catch estimates were derived from observer nominal total CPUE and effort. In some years, observer sample sizes were low (Tables 3). A small groundfish bycatch weight ( 0.5 corresponding to a CV of 1.3108) was chosen based on the September 2015 CPT suggestion. The groundfish bycatch weight was slightly increased from previous
0.2 to 0.5 to obtain global convergence fits for all models (see the jitter tables in Appendix D). The best fit criteria was chosen to choose the lower weight for the groundfish bycatch. Groundfish bycatch of Aleutian Islands golden king crab is very low (Table 2). The CPUE weights were set to 1.0 for all models. A constant (model estimated) variance was included in addition to input CPUE variance for the CPUE fit. The Burnham et al. (1987) suggested formula was used for $\ln$ (CPUE) (and $\ln (\mathrm{MMB})$ ) variance estimation (Equation A.14). Note that the estimated additional variance values were small for both observer and fish ticket CPUE indices for the two regions. Nevertheless, the CPUE index variances estimated from the negative binomial GLM were adequate to fit the model, as confirmed by the fit diagnostics (Fox and Weisberg 2011). Parameter estimates are provided in Table 6 for EAG and Table 12 for WAG for models $22.9 \mathrm{c}, 22.1 \mathrm{e} 2,21.1 \mathrm{f}, 22.1 \mathrm{~g}$, and 22.1 h . The numbers of estimable parameters are listed in Table A1.

## 2. Include tables showing differences in likelihood:

Tables 9 and 15 list the total and component negative log likelihood values for representative models, $22.9 \mathrm{c}, 22.1 \mathrm{e} 2,22.1 \mathrm{f}, 22.1 \mathrm{~g}$, and 22.1 h for EAG and $22.9 \mathrm{c}, 22.1 \mathrm{e} 2$, and 22.1 f WAG, respectively.

## 3. Tables of estimates:

a. Parameter estimates for models 22.9c, 22.1e2, 22.1f, 22.1 g , and 22.1 h are summarized in Table 8 for EAG and for models 22.9c, 22.1e2, and 22.1f in Table 12 for WAG. The boundaries for parameter searches are also provided for the bespoke model 22.9 c in those tables. All parameter estimates were within the bounds of bespoke model.
b. All models considered molt probability parameters in addition to the linear growth increment and normally distributed growth variability parameters to determine the size transition matrix.
c. Mature male and legal male abundance time series for models $22.9 \mathrm{c}, 22.1 \mathrm{e} 2$, $22.1 \mathrm{f}, 22.1 \mathrm{~g}$, and 22.1 h are summarized in Tables 7 and 8 for EAG and for models $22.9 \mathrm{c}, 22.1 \mathrm{e} 2$, and 22.1f in Tables 13 and 14 for WAG.
d. The recruitment estimates for those models are summarized in Tables 7 and 8 for EAG and Tables 13 and 14 for WAG.
e. The negative log-likelihood component values and total negative loglikelihood values for models $22.9 \mathrm{c}, 22.1 \mathrm{e} 2,22.1 \mathrm{f}, 22.1 \mathrm{~g}$, and 22.1 h are summarized in Table 9 for EAG and for models 22.9c, 22.1e2, and 22.1f in Table 15 for WAG. Loglikelihood values of different models are not comparable because of data weighting (i.e., different magnitude of effective sample sizes) and different optimization processes between bespoke and GMACS models.

## 5. Graphs of estimates:

a. Selectivity:

Total selectivity and retention curves of the pre- and post-rationalization periods for models $22.9 \mathrm{c}, 22.1 \mathrm{e} 2,22.1 \mathrm{f}, 22.1 \mathrm{~g}$, and 22.1 h are illustrated in Figure 11 for EAG and for models 22.9c, 22.1e2, and 22.1f in Figure 25 for WAG. Total selectivity for the pre-rationalization period was used in fitting the tagging model. The groundfish bycatch selectivity appeared flat in the preliminary analysis, indicating that all size groups were vulnerable to the gear. Thus, the groundfish bycatch selectivity was set to 1.0 for all lengthclasses in the subsequent analysis.

## b. Mature male biomass:

Mature male biomass time series for models $22.9 \mathrm{c}, 22.1 \mathrm{e} 2,22.1 \mathrm{f}, 22.1 \mathrm{~g}$, and 22.1 h for EAG and models $22.9 \mathrm{c}, 22.1 \mathrm{e} 2$, and 22.1 f for WAG are depicted in Figures 22a (representing the long time series, 1960 to 2022) and 22b (representing the short time series, 2013 to 2022). Mature male biomass tracked the CPUE trends well for selected models for EAG and WAG. The biomass variance was estimated using the Burnham et al. (1987) suggested formula (Equation A.14). The mature male biomass values were determined on 15 February each year.

A fixed period time series of recruits (Table T1) was considered for estimating mean number of recruits for the $M M B_{35 \%}$ calculation under the Tier 3 approach.
c. Fishing mortality:

The full selection pot fishery $F$ values over time for models $22.9 \mathrm{c}, 22.1 \mathrm{e} 2$, $22.1 \mathrm{f}, 22.1 \mathrm{~g}$, and 22.1 h for EAG and models $22.9 \mathrm{c}, 22.1 \mathrm{e} 2$, and 22.1 f for WAG are shown in Figure 21. The $F$ peaked in late 1980s and early to mid1990s and systematically declined in EAG. Slight increases in $F$ were observed from 2014 to 2019, followed by a decline in EAG. On the other hand, the $F$ in WAG peaked in late 1980 s to mid-1990s, and early 2000s, declined in late 2000s, and slightly increased in 2013-2015 and again in 2019-2021.
d. F vs. MMB:

The plots for models 22.9c, 22.1e2, and 22.1f for EAG and WAG are provided in Figure 34. The 2022/23 $F$ was below the overfishing levels for all models in EAG but around the overfishing level for WAG. However, the 2022/23 F may change when completed fisheries data are available in both regions.
e. Stock-Recruitment relationship: Not analyzed.

## f. Recruitment:

Temporal changes in total number of recruits to the modeled population are illustrated in Figure 13 for EAG (models 22.9c, 22.1e2, 22.1f, 22.1g, and 22.1h) and Figure 27 for WAG (models 22.9c, 22.1e2, and 22.1f). The recruitment distribution to the model size group ( $101-185 \mathrm{~mm} \mathrm{CL}$ ) is shown in Figures 14 and 28 for EAG and WAG, for the respective models.

## 6. Evaluation of the fit to the data:

## g. Fits to catch:

The fishery retained and total catch, and groundfish bycatch (observed vs. estimated) plots are illustrated in Figure 17 for EAG (models 22.9c, 22.1e2, 22.1f, 22.1g, and 22.1h) and in Figure 33 for WAG (models 22.9c, 22.1e2, and 22.1 f ). All predicted fits were very close to observed values, especially for retained catch and groundfish bycatch mortality. However, pre-1995 total catch data did not fit well.
h. Survey data plot:

The cooperative survey plots are not provided in this report.
i. CPUE index data:

The comparison of predicted CPUE with input indices (open circles with $95 \%$ confidence intervals) for models $22.9 \mathrm{c}, 22.1 \mathrm{e} 2$, and 22.1 f are shown in Figures 19 and for models 22.1 g and 22.1 h (which included cooperative survey indices) in Figure 20 for EAG and for models 22.9c, 22.1e2, and 22.1f in Figure 33 for WAG. The CPUE variance was estimated using the Burnham et al. (1987) suggested formula (Equation A.14). These figures illustrate varying matches of CPUE predictions with input values by different models.
j. Tagging data:

The predicted vs. observed tag recaptures by length-class for years 1 to 6 post tagging are depicted in Figure 12 for EAG and Figure 26 for WAG. The predictions appeared reasonable. Note that EAG tagging data were used for size transition matrix estimation for both stocks (EAG and WAG). The size transition matrices estimated using EAG tagging data in EAG, and WAG models were similar.
k. Molt probability:

The predicted molt probabilities vs. CL are depicted in Figure 15 for EAG (models 22.9c, 22.1e2, 22.1f, 22.1g, and 22.1h) and in Figure 29 for WAG (models 22.9c, 22.1e2, and 22.1f). The fitted curves appear to be satisfactory.

1. Fit to size compositions:

Since the models were implemented in GMACS for reference points determination, we provide the time series (1981 to 2022) of estimated abundance by size in Figures 9 and 10 for EAG and in Figures 23 and 24 for WAG for models 22.1 e 2 and 22.1f, respectively. The selected period reflects the catch reporting period.

The standardized residual plots are illustrated as bubble plots of size composition over time for retained catch (Figures 17 for EAG and 31 for WAG) and for total catch (Figures 18 for EAG and 32 for WAG) for bespoke model 22.9 c . The retained catch bubble plots do not appear to exhibit major pronounced patterns among residuals.
m . Marginal distributions for the fits to the composition data:
This plot is not provided in this report.
n. Plots of implied versus input effective sample sizes and time series of implied effective sample sizes:
These plots are not provided in this report. However, the Stage-1 and the reweighted Stage-2 effective sample sizes are provided in Tables 5 for EAG and in Tables 11 for WAG, for the model 22.9c.
o. Tables of RMSEs for the indices:

This table is not provided in this report.

## p. Quantile-quantile (Q-Q) plots:

These plots for model fits are not provided in this report.

## 7. Retrospective and historical analysis:

We did not provide the retrospective analysis in this report because of limited time available. However, several previous, including that reported at the January 2023 CPT meeting, indicated some retrospective patterns in the EAG assessment but not in the WAG assessment. Since the status quo model was not changed from previous iterations except for addition of one more year's data, we anticipate similar retrospective pattern in the respective areas.

## 8. Uncertainty and sensitivity analysis:

The main task was to determine a plausible size transition matrix to project the population over time. In a previous study, the sensitivity of the model to determining the size transition matrix by incorporating or not incorporating a molt probability function was investigated (Siddeek et al. 2016a). The model fit improved when a molt probability model was included. Therefore, a molt probability sub-model was included for size transition matrix calculation in all models.

## 9. Conduct 'jitter analysis':

A jitter analysis was conducted on the plausible, representative, model, 22.1e2 outside the GMACS framework (Appendix D). The results indicated that global convergence was achieved for most runs in both EAG and WAG.

## F. Calculation of the OFL

## 1. Specification of the Tier level:

In the following section, the Tier 3 method to estimate OFL and ABC is provided.

Aleutian Islands golden king crab is currently placed in Tier 3a for EAG and Tier $3 b$ for WAG.
2. List of parameter and stock size estimates (or best available proxies thereof) required by limit and target control rules specified in the fishery management plan:

The critical assumptions for $M M B_{M S Y}$ reference points ( $M M B_{35 \%}$ and $F_{35 \%}$ ) estimation of Aleutian Islands golden king crab are:
a. Natural mortality is constant.
b. A fixed growth transition matrix and a molt probability sub-model are adequately estimated from tagging data.
c. Total fishery selectivity and retention curves are length-dependent and the 2005/06-2022/23 period selectivity estimates are applicable.
d. Groundfish bycatch fishery selectivity is kept constant at 1.0 for all length groups.
e. Model estimated recruits (in millions of crab) are valid for different periods considered in the selected models.
f. Model estimated groundfish bycatch mortality values are appropriately averaged for the period 2013/14-2022/23 (10 years).
g. The knife-edge maturity size used for MMB estimation is correct.

## 3. Specification of the OFL:

a. Provide the equations (from Amendment 24) on which the OFL is to be based:

The reference points are estimated by mature male biomass-per-recruit analysis using model estimated parameters.

The Tier 3 control rule formula for $F_{O F L}$ calculation is as follows:

$$
\begin{aligned}
& \text { If, } \\
& M M B_{\text {current }}>M M B_{35 \%}, F_{O F L}=F_{35 \%} \\
& \text { If, } \\
& M M B_{\text {current }} \leq M M B_{35 \%} \text { and } M M B_{\text {current }}>\beta M M B_{35 \%}, \\
& F_{O F L}=F_{35 \%} \frac{\left(\frac{M M B_{\text {current }}}{M M B_{35 \%}}-\alpha\right)}{(1-\alpha)} \\
& \text { If, } \\
& M M B_{\text {current }} \leq \beta M M B_{35 \%}, \\
& F_{O F L}=0 .
\end{aligned}
$$

where
$\beta=$ a parameter with a restriction that $0 \leq \beta<1$. A default value of 0.25 is used, $\alpha=$ a parameter with a restriction that $0 \leq \alpha \leq \beta$. A default value of 0.1 is used,
$\mathrm{MMB}_{\text {current }}=$ the mature male biomass in the current year, and
$M M B_{35 \%}=$ a proxy $M M B_{M S Y}$ for Tier 3 stocks.
Because projected $M M B_{t}$ (i.e., $M M B ~_{\text {current }}$ ) depends on the intervening retained and discard catch (i.e., $\mathrm{MMB}_{\mathrm{t}}$ is estimated after the fishery), an iterative procedure is applied with predicted retained and discard catch. The $F_{O F L}$ and OFL were estimated using MCMC in the GMACS models but the bisection method in the status quo model, 22.9c.
b. Basis for projecting MMB to the time of mating: The NPFMC (2007a) guideline was followed.
c. Specification of Fofs, OFL, and other applicable measures (if any) relevant to determining whether the stock is overfished or if overfishing is occurring:

The control rule is used for stock status determination. If total catch exceeds OFL, then "overfishing" occurs. If MMB equals or declines below $0.5 \mathrm{MMB}_{35 \%}$ (i.e., MSST), the stock is "overfished." If $\mathrm{MMB} / \mathrm{MMB}_{35 \%}$ equals or declines below $\beta$, then the stock productivity is severely depleted, and the directed fishery is closed.

The 2022/23 incomplete fishery data indicated that overfishing did not occur. Total fishery mortality in 2022/23 was 2.612 kt ( 5.758 million lb), which was less than the OFL of 3.761 kt ( 8.291 million lb). The OFL and ABC values for $2022 / 23$ in the tables below are authors-recommended values for consideration.

Status and catch specifications (1000 t) of Aleutian Islands golden king crab (model 22.1e2)

| Year | MSST | Biomass <br> (MMB) | TAC | Retained <br> Catch | Total <br> Catch $^{\mathbf{a}}$ | OFL | ABC $^{\text {b }}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $2019 / 20$ | 5.915 | 16.386 | 3.257 | 3.319 | 3.729 | 5.249 | 3.937 |
| $2020 / 21$ | 6.014 | 15.442 | 2.999 | 3.000 | 3.520 | 4.798 | 3.599 |
| $2021 / 22$ | 5.715 | 13.581 | 2.690 | 2.699 | 3.056 | 4.817 | 3.372 |
| $2022 / 23$ | $5.832^{\text {d }}$ | $13.600^{\text {d }}$ | 2.291 | $2.369^{*}$ | $2.612^{*}$ | $3.761^{\text {c }}$ | $2.821^{\text {c }}$ |
| $2023 / 24$ |  | $12.069^{\text {d }}$ |  |  |  | $4.182^{\text {d }}$ | $3.137^{\mathrm{d}}$ |

Status and catch specifications (million lb) of Aleutian Islands golden king crab (model 22.1e2)

| Year | MSST | Biomass <br> (MMB) | TAC | Retained <br> Catch | Total <br> Catch $^{\mathbf{a}}$ | OFL | ABC $^{\mathbf{b}}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $2019 / 20$ | 13.041 | 36.124 | 7.180 | 7.317 | 8.222 | 11.572 | 8.679 |
| $2020 / 21$ | 13.259 | 34.043 | 6.610 | 6.614 | 7.759 | 10.579 | 7.934 |
| $2021 / 22$ | 12.598 | 29.940 | 5.930 | 5.951 | 6.737 | 10.620 | 7.434 |
| $2022 / 23$ | $12.857^{\text {d }}$ | $29.984^{\text {d }}$ | 5.050 | $5.223^{*}$ | $5.758^{*}$ | $8.291^{\text {c }}$ | $6.219^{\text {c }}$ |
| $2023 / 24$ |  | $26.607^{\text {d }}$ |  |  |  | $9.220^{\text {d }}$ | $6.916^{\text {d }}$ |

*Incomplete fishery
a. Total catch was sum of retained catch and estimated bycatch mortality of discarded bycatch during crab fisheries and groundfish fisheries.
b. $25 \%$ buffer was applied to total catch OFL to determine ABC except $2021 / 22$, during which $30 \%$ buffer was applied.
c. OFL, and ABC were estimated by the accepted model 21.1e2 in May 2022 assessment when the WAG fisheries was not completed.
d. MSST, MMB, OFL, and ABC were estimated by the authors preferred model 22.1e2 in May 2023 assessment when the EAG and WAG fisheries were not completed.

## G. Calculation of ABC

Based on CPT/SSC suggestion in 2022, a $25 \%$ buffer on OFL was used to determine the ABC for the fishery (i.e., $\mathrm{ABC}=0.75^{*} \mathrm{OFL}$ ). The ABC estimates for approved models in previous years up to 2021/22 and authors recommended model 22.1 e 2 in 2022/23 are listed in the above two tables.

Alternative reference points including ABC based on model 22.1 f are also provided below if Year:Area interaction model is selected by the CPT/SSC. The differences between the two models' estimates are very small, however.

Status and catch specifications (1000 t) of Aleutian Islands golden king crab (model 22.1f)

| Year | MSST | Biomass <br> (MMB) | TAC | Retained <br> Catch | Total <br> Catch $^{\text {a }}$ | OFL | ABC $^{\text {b }}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $2022 / 23$ | $5.836^{\text {d }}$ | $13.269^{\mathrm{d}}$ | 2.291 | $2.369^{*}$ | $2.612^{*}$ | $3.761^{\text {c }}$ | $2.821^{\text {c }}$ |
| $2023 / 24$ |  | $11.934^{\mathrm{d}}$ |  |  |  | $4.029^{\mathrm{d}}$ | $3.022^{\mathrm{d}}$ |

Status and catch specifications (million lb) of Aleutian Islands golden king crab (model 22.1f)

| Year | MSST | Biomass <br> (MMB) | TAC | Retaine <br> d Catch | Total <br> Catch $^{\text {a }}$ | OFL | ABC $^{\mathbf{b}}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $2022 / 23$ | $12.865^{\mathrm{d}}$ | $29.253^{\mathrm{d}}$ | 5.050 | $5.223^{*}$ | $5.758^{*}$ | $8.291^{\text {c }}$ | $6.219^{\mathrm{c}}$ |
| $2023 / 24$ |  | $26.309^{\mathrm{d}}$ |  |  |  | $8.883^{\mathrm{d}}$ | $6.662^{\mathrm{d}}$ |

## 1. List of variables related to scientific uncertainty:

- Models rely largely on fisheries data.
- Observer and fisheries CPUE indices played a major role in the assessment model.
- The period to compute the average number of recruits relative to the assumption that this represents "a period determined to be representative of the production potential of the stock."
- Fixed bycatch mortality rates were used in each fishery (crab fishery and the groundfish fishery) that discarded golden king crab.
- Discarded catch and bycatch mortality for each fishery in which bycatch occurred during 1981/82-1989/90 were not available.


## 2. List of additional uncertainties for alternative sigma-b:

A buffer of $25 \%$ to account for additional uncertainties was used ,which was recommended in 2022.

## 3. Author recommended ABC :

Authors recommended one ABC based on $25 \%$ buffer on the OFL under model 22.1 e 2 with alternative 22.1f.

## H. Rebuilding Analysis

Not applicable. This stock has not been declared overfished.

## I. Data Gaps and Research Priorities

1. Recruit abundances were tied to commercial catch sampling data. The implicit assumption in the analysis was that the estimated recruits come solely from the same exploited stock through growth and mortality. The current analysis did not consider that additional recruitment may occur through immigration from neighboring areas and possibly separate substocks. The analysis also did not consider emigration from the study area, which would result in an assumption of increased $M$ or a reduced estimate of recruits. Extensive tagging experiments or resource surveys are needed to investigate stock distributions.
2. An extensive tagging study may provide independent estimates of molting probability and growth. Historical tagging data were used to determine the size transition matrix.
3. An arbitrary $20 \%$ handling mortality rate on discarded males was used, which was obtained from the red king crab literature (Kruse et al. 2000; Siddeek 2002). An experimentally based independent estimate of handling mortality is needed for Aleutian Islands golden king crab.
4. ADF\&G and the Aleutian King Crab Research Foundation recently initiated a cooperative crab survey program in the Aleutian Islands. This program needs to be strengthened and continued for golden king crab research to fill some of the data gaps and establish a fishery independent data source.
5. It is unclear how the recent changes in environmental conditions in the Bering Sea will affect golden king crab growth and survival. Limited length-weight data from the cooperative survey and independent biological sampling in 2018 and 2020 from WAG were used in the current assessment; however, more measurements are needed from both regions to increase the sample size to refine the length-weight model.
6. Morphometric maturity information was used to determine MMB. The ADF\&G observer sampling, dock side sampling, and cooperative survey programs collected male maturity data during 2018/19 through 2022/23. Preliminary analysis on these data was presented in the previous assessment. The CPT previously recommended to collect additional data on small size crab (sublegal) to improve maturity fit. The maturity data collection needs to be continued to accumulate more measurements on small size male crab.
7. Morphometric measurements provide size at maturity. Ideally, an experimental study under natural environment conditions is needed to collect male size at functional maturity data to determine functional maturity size.

## J. Acknowledgments

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Table T1. Features of all model scenarios: Initial condition was estimated in year 1960 by the equilibrium condition; two sets of logistic total selectivity curves were used for the pre- and post-rationalization periods; a single retention curve was used for the whole period; and a common $M$ of $0.22 \mathrm{yr}^{-1}$ was used. The effective sample sizes for size compositions were estimated in two stages: Stage-1: as the number of vessel days or trips and Stage-2: as the Francis re-iteration method.

| Model | CPUE Data Type and Maturity Option | Period for Mean Number of Recruit Calculation for (a) Initial Equilibrium Abundance and (b) Reference Points Estimations; and Remarks |
| :---: | :---: | :---: |
| 22.9c (modified version of the accepted model $22 \_1 \mathrm{e} 2$ in 2022; implemented with up to 2022/23 data)-status quo model | Observer data from 1995/96-2022/23; Fish ticket data from 1985/86-1998/99; Observer and fish ticket CPUE standardization by the negative binomial model; the knifeedge maturity size of 116 mm CL ; and three catchability and additional CVs during 1985-1998; 1995-2004; and 20052022. | 1987-2017; CPT/SSC suggested status quo model. |
| 22.1e2 | $22.9 \mathrm{c}+$ GMACS implementation. |  |
| 22.1f | $22.1 \mathrm{e} 2+$ the observer CPUE data standardized including <br> Year:Area interactions + GMACS implementation. |  |
| 22.1 g | 22.1e2+ Observer indices from 1995/96-2014/15 + cooperative survey indices from 2015/16-2022/23 (without 2020/21 missing data) + GMACS implementation |  |
| 22.1h | $22.1 \mathrm{~g}+$ the observer CPUE data standardized including Year:Area interactions + cooperative survey indices + GMACS implementation. |  |

Table 1. Commercial fishery history for the Aleutian Islands golden king crab fishery 1981/82-2021/22: number of vessels, guideline harvest level (GHL; established in lb, converted to t) for 1996/97-2004/05, total allowable catch (TAC; established in lb, converted to $t$ ) for 2005/06-2022/23, weight of retained catch (harvest; $t$ ), number of retained crab, pot lifts, fishery catch-per-unit- effort (CPUE; retained number of crab per pot lift), and average weight $(\mathrm{kg})$ of landed crab. The values are separated by EAG and WAG beginning in 1996/97.

| Crab Fishing Season | Vessels | GHL/TAC | Harvest ${ }^{\text {a }}$ | Crab | Pot Lifts | CPUE ${ }^{\text {b }}$ | Average Weight ${ }^{\text {c }}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1981/82 | 14-20 | - | 599 | 240,458 | 27,533 | 9 | $2.5{ }^{\text {d }}$ |
| 1982/83 | 99-148 | - | 4,169 | 1,737,109 | 179,472 | 10 | $2.4{ }^{\text {d }}$ |
| 1983/84 | 157-204 | - | 4,508 | 1,773,262 | 256,393 | 7 | $2.5{ }^{\text {d }}$ |
| 1984/85 | 38-51 | - | 2,132 | 971,274 | 88,821 | 11 | $2.2{ }^{\text {e }}$ |
| 1985/86 | 53 | - | 5,776 | 2,816,313 | 236,601 | 12 | $2.1{ }^{\text {f }}$ |
| 1986/87 | 64 | - | 6,685 | 3,345,680 | 433,870 | 8 | $2.0{ }^{\text {f }}$ |
| 1987/88 | 66 | - | 4,199 | 2,177,229 | 307,130 | 7 | $1.9{ }^{\text {f }}$ |
| 1988/89 | 76 | - | 4,820 | 2,488,433 | 321,927 | 8 | $1.9{ }^{\text {f }}$ |
| 1989/90 | 68 | - | 5,453 | 2,902,913 | 357,803 | 8 | $1.9{ }^{\text {f }}$ |
| 1990/91 | 24 | - | 3,153 | 1,707,618 | 215,840 | 8 | $1.9{ }^{\text {f }}$ |
| 1991/92 | 20 | - | 3,494 | 1,847,398 | 234,857 | 8 | $1.9{ }^{\text {f }}$ |
| 1992/93 | 22 | - | 2,854 | 1,528,328 | 203,221 | 8 | $1.9{ }^{\text {f }}$ |
| 1993/94 | 21 | - | 2,518 | 1,397,530 | 234,654 | 6 | $1.8{ }^{\text {f }}$ |


| Crab <br> Fishing Season | Vessels |  | GHL/TAC |  | Harvest ${ }^{\text {a }}$ |  | Crab |  | Pot Lifts |  | CPUE ${ }^{\text {b }}$ |  | Average Weight ${ }^{\text {c }}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1994/95 | 35 |  | - |  | 3,687 |  | 1,924,271 |  | 386,593 |  | 5 |  | $1.9{ }^{\text {f }}$ |  |
| 1995/96 | 28 |  | - |  | 3,157 |  | 1,582,333 |  | 293,021 |  | 5 |  | $2.0{ }^{\text {f }}$ |  |
|  | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG |
| 1996/97 | 14 | 13 | 1,452 | 1,225 | 1,493 | 1,145 | 731,909 | 602,968 | 113,460 | 99,267 | 7 | 6 | $2.04{ }^{\text {f }}$ | $1.91{ }^{\text {f }}$ |
| 1997/98 | 13 | 9 | 1,452 | 1,225 | 1,588 | 1,109 | 780,610 | 569,550 | 106,403 | 86,811 | 7 | 7 | $2.04{ }^{\text {f }}$ | $1.95{ }^{\text {f }}$ |
| 1998/99 | 14 | 3 | 1,361 | 1,225 | 1,473 | 768 | 740,011 | 410,018 | 83,378 | 35,975 | 9 | 11 | $2.00{ }^{\text {f }}$ | $1.86{ }^{\text {f }}$ |
| 1999/00 | 15 | 15 | 1,361 | 1,225 | 1,392 | 1,256 | 709,332 | 676,558 | 79,129 | 107,040 | 9 | 6 | $1.95{ }^{\text {f }}$ | $1.86{ }^{\text {f }}$ |
| 2000/01 | 15 | 12 | 1,361 | 1,225 | 1,422 | 1,308 | 704,702 | 705,613 | 71,551 | 101,239 | 10 | 7 | $2.00{ }^{\text {f }}$ | $1.86{ }^{\text {f }}$ |
| 2001/02 | 19 | 9 | 1,361 | 1,225 | 1,442 | 1,243 | 730,030 | 686,738 | 62,639 | 105,512 | 12 | 7 | $2.00{ }^{\text {f }}$ | $1.81{ }^{\text {f }}$ |
| 2002/03 | 19 | 6 | 1,361 | 1,225 | 1,280 | 1,198 | 643,886 | 664,823 | 52,042 | 78,979 | 12 | 8 | $2.00^{\text {f }}$ | $1.81{ }^{\text {f }}$ |
| 2003/04 | 18 | 6 | 1,361 | 1,225 | 1,350 | 1,220 | 643,074 | 676,633 | 58,883 | 66,236 | 11 | 10 | $2.09{ }^{\text {f }}$ | $1.81{ }^{\text {f }}$ |
| 2004/05 | 19 | 6 | 1,361 | 1,225 | 1,309 | 1,219 | 637,536 | 685,465 | 34,848 | 56,846 | 18 | 12 | $2.04{ }^{\text {f }}$ | $1.77^{\text {f }}$ |
| 2005/06 | 7 | 3 | 1,361 | 1,225 | 1,300 | 1,204 | 623,971 | 639,368 | 24,569 | 30,116 | 25 | 21 | $2.09{ }^{\text {f }}$ | $1.91{ }^{\text {f }}$ |
| 2006/07 | 6 | 4 | 1,361 | 1,225 | 1,357 | 1,030 | 650,587 | 527,734 | 26,195 | 26,870 | 25 | 20 | $2.09{ }^{\text {f }}$ | $1.95{ }^{\text {f }}$ |
| 2007/08 | 4 | 3 | 1,361 | 1,225 | 1,356 | 1,142 | 633,253 | 600,595 | 22,653 | 29,950 | 28 | 20 | $2.13{ }^{\text {f }}$ | $1.91{ }^{\text {f }}$ |
| 2008/09 | 3 | 3 | 1,361 | 1,286 | 1,426 | 1,150 | 666,946 | 587,661 | 24,466 | 26,200 | 27 | 22 | $2.13{ }^{\text {f }}$ | $1.95{ }^{\text {f }}$ |


|  | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2009/10 | 3 | 3 | 1,429 | 1,286 | 1,429 | 1,253 | 679,886 | 628,332 | 29,298 | 26,489 | 26 | 24 | $2.09^{\text {f }}$ | $2.00^{\text {f }}$ |
| 2010/11 | 3 | 3 | 1,429 | 1,286 | 1,428 | 1,279 | 670,983 | 626,246 | 25,851 | 29,994 | 26 | 21 | $2.13{ }^{\text {f }}$ | $2.04{ }^{\text {f }}$ |
| 2011/12 | 3 | 3 | 1,429 | 1,286 | 1,429 | 1,276 | 668,828 | 616,118 | 17,915 | 26,326 | 37 | 23 | $2.13{ }^{\text {f }}$ | $2.09^{\text {f }}$ |
| 2012/13 | 3 | 3 | 1,501 | 1,352 | 1,504 | 1,339 | 687,666 | 672,916 | 20,827 | 32,716 | 33 | 21 | $2.18{ }^{\text {f }}$ | $2.00^{\text {f }}$ |
| 2013/14 | 3 | 3 | 1,501 | 1,352 | 1,546 | 1,347 | 720,220 | 686,883 | 21,388 | 41,835 | 34 | 16 | $2.13{ }^{\text {f }}$ | $1.95{ }^{\text {f }}$ |
| 2014/15 | 3 | 2 | 1,501 | 1,352 | 1,554 | 1,217 | 719,064 | 635,312 | 17,002 | 41,548 | 42 | 15 | $2.18{ }^{\text {f }}$ | $1.91{ }^{\text {f }}$ |
| 2015/16 | 3 | 2 | 1,501 | 1,352 | 1,590 | 1,139 | 763,604 | 615,355 | 19,376 | 41,108 | 39 | 15 | $2.09^{\text {f }}$ | $1.85{ }^{\text {f }}$ |
| 2016/17 | 3 | 3 | 1,501 | 1,014 | 1,578 | 1,015 | 793,983 | 543,796 | 24,470 | 38,118 | 32 | 14 | $1.99{ }^{\text {f }}$ | $1.87{ }^{\text {f }}$ |
| 2017/18 | 3 | 3 | 1,501 | 1,014 | 1,571 | 1,014 | 802,610 | 519,051 | 25,516 | 30,885 | 31 | 17 | $1.96{ }^{\text {f }}$ | $1.95{ }^{\text {f }}$ |
| 2018/19 | 3 | 3 | 1,749 | 1,134 | 1,830 | 1,135 | 940,336 | 578,221 | 25,553 | 29,156 | 37 | 20 | $1.95{ }^{\text {f }}$ | $1.96{ }^{\text {f }}$ |
| 2019/20 | 3 | 3 | 1,955 | 1,302 | 2,031 | 1,288 | 1,057,464 | 649,832 | 30,998 | 42,924 | 34 | 15 | $1.92{ }^{\text {f }}$ | $1.98{ }^{\text {f }}$ |
| 2020/21 | 3 | 3 | 1,656 | 1,343 | 1,733 | 1,267 | 902,122 | 682,107 | 30,072 | 46,701 | 30 | 15 | $1.92{ }^{\text {f }}$ | $1.86{ }^{\text {f }}$ |
| 2021/22 | 3 | 3 | 1,637 | 1052 | 1,706 | 993 | 863,269 | 538,064 | 30,948 | 46,161 | 28 | 12 | $1.98{ }^{\text {f }}$ | $1.85{ }^{\text {f }}$ |
| 2022/23 | 3 | 3 | 1,506 | 785 | 1,585 | 784 | 811,282 | 427,696 | 21,600 | 32,786 | 38 | 13 | $1.95{ }^{\text {f }}$ | $1.83{ }^{\text {f }}$ |

Note:
a. Includes deadloss.
b. Number of crab per pot lift.
c. Average weight of landed crab, including dead loss.
d. Managed with $6.5^{\prime \prime}$ carapace width (CW) minimum size limit.
e. Managed with $6.5^{\prime \prime} \mathrm{CW}$ minimum size limit west of $171^{\circ} \mathrm{W}$ longitude and $6.0^{\prime \prime}$ minimum size limit east of $171^{\circ} \mathrm{W}$ longitude.
f. Managed with $6.0^{\prime \prime}$ minimum size limit.
g. Catch and effort data include cost recovery fishery

Table 2. Annual weight of total fishery mortality to Aleutian Islands golden king crab, 1981/82 - 2022/23, partitioned by source of mortality: retained catch, bycatch mortality during crab fisheries, and bycatch mortality during groundfish fisheries. For bycatch in the federal groundfish fisheries, historical data (19912008) are not available for areas east and west of 174 W , and are listed for federal groundfish reporting areas 541,542 , and 543 combined. The 2009 - present data are available by separate EAG and WAG fisheries and are listed as such. A mortality rate of $20 \%$ was applied for crab fisheries bycatch, and a mortality rate of $50 \%$ for groundfish pot fisheries and $80 \%$ for trawl fisheries were applied.

| Season |  |  | Bycatch Mortality by Fishery Type (t) |  |  |  | Total Fishery Mortality (t)EAG WAG Entire AI |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Retained Catch <br> (t) |  | Crab |  | Groundfish |  |  |  |  |
|  |  |  |  |  |  |  |  |  |  |
|  | EAG | WAG | EAG | WAG | EAG | WAG |  |  |  |
| 1981/82 | 490 | 95 |  |  |  |  |  |  | 585 |
| 1982/83 | 1,260 | 2,655 |  |  |  |  |  |  | 3,914 |
| 1983/84 | 1,554 | 2,991 |  |  |  |  |  |  | 4,545 |
| 1984/85 | 1,839 | 424 |  |  |  |  |  |  | 2,263 |
| 1985/86 | 2,677 | 1,996 |  |  |  |  |  |  | 4,673 |
| 1986/87 | 2,798 | 4,200 |  |  |  |  |  |  | 6,998 |
| 1987/88 | 1,882 | 2,496 |  |  |  |  |  |  | 4,379 |
| 1988/89 | 2,382 | 2,441 |  |  |  |  |  |  | 4,823 |
| 1989/90 | 2,738 | 3,028 |  |  |  |  |  |  | 5,766 |
| 1990/91 | 1,623 | 1,621 |  |  |  |  |  |  | 3,244 |
| 1991/92 | 2,035 | 1,397 | 515 | 344 |  | 0 |  |  | 4,291 |
| 1992/93 | 2,112 | 1,025 | 1,206 | 373 |  | 0 |  |  | 4,716 |
| 1993/94 | 1,439 | 686 | 383 | 258 |  | 4 |  |  | 2,770 |
| 1994/95 | 2,044 | 1,540 | 687 | 823 |  | 1 |  |  | 5,095 |
| 1995/96 | 2,259 | 1,203 | 725 | 530 |  | 2 |  |  | 4,719 |
| 1996/97 | 1,738 | 1,259 | 485 | 439 |  | 5 |  |  | 3,926 |
| 1997/98 | 1,588 | 1,083 | 441 | 343 |  | 1 |  |  | 3,455 |
| 1998/99 | 1,473 | 955 | 434 | 285 |  | 1 |  |  | 3,149 |
| 1999/00 | 1,392 | 1,222 | 313 | 385 |  | 3 |  |  | 3,316 |
| 2000/01 | 1,422 | 1,342 | 82 | 437 |  | 2 |  |  | 3,285 |
| 2001/02 | 1,442 | 1,243 | 74 | 387 |  | 0 |  |  | 3,146 |
| 2002/03 | 1,280 | 1,198 | 52 | 303 |  | 8 |  |  | 2,850 |
| 2003/04 | 1,350 | 1,220 | 53 | 148 |  | 0 |  |  | 2,792 |
| 2004/05 | 1,309 | 1,219 | 41 | 143 |  | 1 |  |  | 2,715 |
| 2005/06 | 1,300 | 1,204 | 22 | 73 |  | 2 |  |  | 2,601 |
| 2006/07 | 1,357 | 1,022 | 28 | 81 |  | 8 |  |  | 2,506 |
| 2007/08 | 1,356 | 1,142 | 24 | 114 |  | 9 |  |  | 2,695 |
| 2008/09 | 1,426 | 1,150 | 61 | 102 |  | 3 |  |  | 2,772 |
| 2009/10 | 1,429 | 1,253 | 111 | 108 | 18 | 5 | 1,558 | 1,366 | 2,923 |
| 2010/11 | 1,428 | 1,279 | 123 | 124 | 49 | 3 | 1,600 | 1,407 | 3,006 |
| 2011/12 | 1,429 | 1,276 | 106 | 117 | 25 | 4 | 1,560 | 1,398 | 2,957 |
| 2012/13 | 1,504 | 1,339 | 118 | 145 | 9 | 6 | 1,631 | 1,491 | 3,122 |
| 2013/14 | 1,546 | 1,347 | 113 | 174 | 5 | 7 | 1,665 | 1,528 | 3,192 |
| 2014/15 | 1,554 | 1,217 | 127 | 175 | 9 | 5 | 1,691 | 1,397 | 3,088 |
| 2015/16 | 1,590 | 1,139 | 165 | 157 | 23 | 2 | 1,778 | 1,298 | 3,076 |
| 2016/17 | 1,578 | 1,015 | 203 | 145 | 95 | 4 | 1,877 | 1,164 | 3,041 |
| 2017/18 | 1,571 | 1,014 | 219 | 126 | 46 | 2 | 1,836 | 1,142 | 2,978 |
| 2018/19 | 1,830 | 1,135 | 240 | 140 | 24 | 3 | 2,094 | 1,278 | 3,372 |
| 2019/20 | 2,031 | 1,288 | 275 | 112 | 17 | 6 | 2,323 | 1,406 | 3,729 |
| 2020/21 | 1,733 | 1,267 | 241 | 147 | 125 | 6 | 2,100 | 1,420 | 3,519 |
| 2021/22 | 1,706 | 993 | 169 | 173 | 17 | 1 | 1,892 | 1,168 | 3,060 |
| 2022/23 | 1,585 | 784 | 125 | 112 | 5 | 1 | 1,715 | 897 | 2,612 |

Table 2a. Time series of estimated total male catch (weight of crab on the deck without applying any handling mortality) for the EAG and WAG golden king crab stocks (1990/91$2022 / 23$ ). The crab weights are for the size range $\geq 101 \mathrm{~mm}$ CL and a length-weight formula was used to predict weight at the mid-point of each size bin. NA: no observer sampling to compute catch.

| Year | Total Catch Biomass (t) <br> EAG | Total Catch Biomass (t) <br> WAG |
| :---: | :---: | :---: |
| $1990 / 91$ | 3,981 | 3,982 |
| $1991 / 92$ | 6,597 | 2,118 |
| $1992 / 93$ | 5,436 | 1,039 |
| $1993 / 94$ | NA | 3,601 |
| $1994 / 95$ | 3,444 | 5,054 |
| $1995 / 96$ | 4,641 | 2,619 |
| $1996 / 97$ | 2,563 | 1,972 |
| $1997 / 98$ | 2,977 | 1,892 |
| $1998 / 99$ | 3,141 | 1,107 |
| $199 / 00$ | 2,606 | 2,178 |
| $2000 / 01$ | 2,760 | 2,273 |
| $2001 / 02$ | 2,238 | 2,155 |
| $2002 / 03$ | 1,916 | 1,900 |
| $2003 / 04$ | 1,902 | 1,867 |
| $2004 / 05$ | 1,695 | 1,886 |
| $2005 / 06$ | 1,742 | 1,796 |
| $2006 / 07$ | 1,647 | 1,551 |
| $2007 / 08$ | 1,820 | 1,614 |
| $2008 / 09$ | 1,824 | 1,733 |
| $2009 / 10$ | 1,770 | 1,690 |
| $2010 / 11$ | 1,757 | 1,605 |
| $2011 / 12$ | 1,781 | 1,517 |
| $2012 / 13$ | 1,947 | 1,839 |
| $2013 / 14$ | 1,852 | 1,919 |
| $2014 / 15$ | 1,967 | 1,592 |
| $2015 / 16$ | 2,136 | 1,565 |
| $2016 / 17$ | 2,234 | 1,570 |
| $2017 / 18$ | 2,339 | 1,437 |
| $2018 / 19$ | 2,735 | 1,637 |
| $2019 / 20$ | 3,033 | 1,714 |
| $2020 / 21$ | 2,608 | 1,844 |
| $2021 / 22$ | 2,427 | 1,612 |
| $2022 / 23 *$ | 1,732 | 1,068 |
|  |  |  |

*Incomplete fishery data used

Table 3. Time series of nominal annual pot fishery retained, observer retained, and observer total catch-per-unit-effort (CPUE, number of crab per pot lift), total pot fishing effort (number of pot lifts), observer sample size (number of sampled pots), and GLM estimated observer CPUE Index (for non-interaction model) for the EAG and WAG golden king crab stocks, 1985/86-2022/23. Observer nominal retained CPUE includes retained and non-retained legal-size crab.

| Year | Pot Fishery Nominal Retained CPUE |  | Obs. Nominal Retained CPUE |  | Obs. Nominal Total CPUE |  | Pot Fishery Effort (no.pot lifts) |  | Obs. Sample Size (no.pot lifts) |  | Obs. CPUE Index |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG | EAG | WAG |
| 1985/86 | 11.90 | 11.90 |  |  |  |  | 117,718 | 118,563 |  |  |  |  |
| 1986/87 | 8.42 | 7.32 |  |  |  |  | 155,240 | 277,780 |  |  |  |  |
| 1987/88 | 7.03 | 7.15 |  |  |  |  | 146,501 | 160,229 |  |  |  |  |
| 1988/89 | 7.52 | 7.93 |  |  |  |  | 155,518 | 166,409 |  |  |  |  |
| 1989/90 | 8.49 | 7.83 |  |  |  |  | 155,262 | 202,541 |  |  |  |  |
| 1990/91 | 8.90 | 7.00 | 6.84 | 8.34 | 13.00 | 28.40 | 106,281 | 108,533 | 138 | 340 |  |  |
| 1991/92 | 8.20 | 7.40 | 9.84 | 6.14 | 36.91 | 15.48 | 133,428 | 101,429 | 377 | 857 |  |  |
| 1992/93 | 8.40 | 5.90 | 10.44 | 4.26 | 38.52 | 11.36 | 133,778 | 69,443 | 199 | 690 |  |  |
| 1993/94 | 7.80 | 4.40 | 5.91 | 12.75 | 20.81 | 21.25 | 106,890 | 127,764 | 31 | 174 |  |  |
| 1994/95 | 5.90 | 4.10 | 4.66 | 6.62 | 12.91 | 19.52 | 191,455 | 195,138 | 127 | 1,270 |  |  |
| 1995/96 | 5.90 | 4.70 | 6.03 | 6.03 | 16.98 | 17.30 | 177,773 | 115,248 | 6,388 | 5,598 | 1.00 | 1.16 |
| 1996/97 | 6.50 | 6.10 | 6.02 | 5.90 | 13.81 | 14.85 | 113,460 | 99,267 | 8,360 | 7,194 | 0.94 | 0.98 |
| 1997/98 | 7.30 | 6.60 | 7.99 | 6.72 | 18.25 | 15.54 | 106,403 | 86,811 | 4,670 | 3,985 | 0.87 | 0.98 |
| 1998/99 | 8.90 | 11.40 | 9.82 | 9.43 | 25.77 | 23.09 | 83,378 | 35,975 | 3,616 | 1,876 | 1.00 | 1.09 |
| 1999/00 | 9.00 | 6.30 | 10.28 | 6.09 | 20.77 | 14.83 | 79,129 | 107,040 | 3,851 | 4,523 | 0.92 | 0.91 |
| 2000/01 | 9.90 | 7.00 | 10.40 | 6.46 | 25.39 | 16.76 | 71,551 | 101,239 | 5,043 | 4,740 | 0.82 | 0.84 |
| 2001/02 | 11.70 | 6.50 | 11.73 | 6.04 | 22.48 | 14.70 | 62,639 | 105,512 | 4,626 | 4,454 | 1.04 | 0.82 |
| 2002/03 | 12.40 | 8.40 | 12.70 | 7.47 | 22.59 | 17.37 | 52,042 | 78,979 | 3,980 | 2,509 | 1.10 | 0.91 |
| 2003/04 | 10.90 | 10.20 | 11.34 | 9.33 | 19.43 | 18.21 | 58,883 | 66,236 | 3,960 | 3,334 | 0.97 | 1.16 |
| 2004/05 | 18.30 | 12.10 | 18.34 | 11.14 | 28.48 | 22.44 | 34,848 | 56,846 | 2,206 | 2,619 | 1.44 | 1.25 |
| 2005/06 | 25.40 | 21.20 | 29.52 | 23.68 | 38.55 | 35.87 | 24,569 | 30,116 | 1,193 | 1,365 | 0.97 | 1.24 |
| 2006/07 | 24.80 | 19.60 | 25.13 | 23.96 | 33.39 | 33.41 | 26,195 | 26,870 | 1,098 | 1,183 | 0.78 | 1.16 |
| 2007/08 | 28.00 | 20.00 | 31.10 | 21.04 | 40.38 | 32.46 | 22,653 | 29,950 | 998 | 1,082 | 0.87 | 1.04 |
| 2008/09 | 27.30 | 22.40 | 29.97 | 24.53 | 38.23 | 38.12 | 24,466 | 26,200 | 613 | 979 | 0.87 | 1.20 |
| 2009/10 | 25.90 | 23.70 | 26.60 | 26.53 | 35.88 | 34.05 | 26,298 | 26,489 | 408 | 892 | 0.72 | 1.26 |
| 2010/11 | 26.00 | 20.90 | 26.40 | 22.39 | 37.10 | 29.10 | 25,851 | 29,994 | 436 | 867 | 0.74 | 1.08 |
| 2011/12 | 37.30 | 23.40 | 39.48 | 23.71 | 52.04 | 31.04 | 17,915 | 26,326 | 361 | 837 | 1.06 | 1.15 |
| 2012/13 | 33.02 | 20.57 | 37.82 | 22.86 | 47.57 | 30.80 | 20,827 | 32,716 | 438 | 1,109 | 1.02 | 1.19 |


| 2013/14 | 33.67 | 16.42 | 35.94 | 16.93 | 46.16 | 24.95 | 21,388 | 41,835 | 499 | 1,223 | 1.00 | 0.85 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2014/15 | 42.29 | 15.29 | 47.01 | 15.28 | 60.00 | 22.64 | 17,002 | 41,548 | 376 | 1,137 | 1.32 | 0.81 |
| 2015/16 | 39.41 | 14.97 | 43.27 | 15.82 | 58.68 | 22.22 | 19,376 | 41,108 | 478 | 1,296 | 1.28 | 0.79 |
| 2016/17 | 32.45 | 14.29 | 36.89 | 16.73 | 52.82 | 24.31 | 24,470 | 38,118 | 617 | 1,060 | 1.05 | 0.90 |
| 2017/18 | 31.46 | 16.81 | 35.18 | 19.28 | 54.62 | 25.53 | 25,516 | 30,885 | 585 | 760 | 1.00 | 1.04 |
| 2018/19 | 36.80 | 19.83 | 41.57 | 22.84 | 62.97 | 30.61 | 25,553 | 29,156 | 475 | 688 | 1.22 | 1.30 |
| 2019/20 | 34.11 | 15.10 | 40.88 | 16.30 | 57.46 | 22.73 | 30,998 | 42,963 | 540 | 967 | 1.15 | 0.99 |
| 2020/21 | 30.00 | 14.61 | 36.40 | 15.66 | 51.58 | 22.73 | 30,072 | 46,701 | 567 | 1,137 | 1.04 | 0.92 |
| 2021/22 | 27.89 | 11.66 | 33.56 | 13.43 | 42.83 | 20.88 | 30,948 | 46,161 | 478 | 858 | 0.93 | 0.72 |
| 2022/23* | 36.49 | 13.22 | 46.94 | 14.19 | 56.79 | 21.04 | 20,875 | 32,786 | 174 | 805 | 1.27 | 0.70 |

* Incomplete fishery data used

Table 4. Time series of negative binomial GLM estimated CPUE indices and coefficient of variation (CV) for the fish ticket based retained catch-per-pot lift for the EAG golden king crab stock. The GLM was fitted to the 1985/86 to 1998/99 time series of data.

| Year | CPUE <br> Index | CV |
| :---: | :---: | :---: |
| $1985 / 86$ | 1.63 | 0.03 |
| $1986 / 87$ | 1.23 | 0.04 |
| $1987 / 88$ | 0.96 | 0.05 |
| $1988 / 89$ | 1.04 | 0.04 |
| $1989 / 90$ | 1.08 | 0.03 |
| $1990 / 91$ | 0.99 | 0.05 |
| $1991 / 92$ | 0.90 | 0.05 |
| $1992 / 93$ | 0.92 | 0.05 |
| $1993 / 94$ | 0.91 | 0.05 |
| $1994 / 95$ | 0.81 | 0.05 |
| $1995 / 96$ | 0.78 | 0.06 |
| $1996 / 97$ | 0.78 | 0.06 |
| $1997 / 98$ | 1.05 | 0.04 |
| $1998 / 99$ | 1.21 | 0.04 |

Table 5. Initial input number of vessel-days/trips and Stage-2 effective sample sizes iteratively estimated by the Francis method for retained, total, and groundfish discard catch size compositions of golden king crab for model 22.9 c fit to EAG data. NA: not available.

| Year | Initial <br> Input <br> Retained <br> Vessel- <br> Days <br> Sample <br> Size (no) | Stage-2 <br> Retained Effective Sample Size (no) | Initial <br> Input <br> Total <br> Vessel- <br> Days <br> Sample <br> Size <br> (no) | Stage-2 <br> Total <br> Effective <br> Sample <br> Size (no) | Initial Input Groundfish Trip Sample Size (no) | Stage-2 Groundfish Effective Sample Size (no) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1985/86 | 57 | 42 |  |  |  |  |
| 1986/87 | 11 | 8 |  |  |  |  |
| 1987/88 | 61 | 45 |  |  |  |  |
| 1988/89 | 352 | 260 |  |  |  |  |
| 1989/90 | 792 | 584 |  |  | 9 | 4 |
| 1990/91 | 163 | 120 | 22 | 12 | 13 | 6 |
| 1991/92 | 140 | 103 | 48 | 25 | NA | NA |
| 1992/93 | 49 | 36 | 41 | 22 | 2 | 1 |
| 1993/94 | 340 | 251 | NA | NA | 2 | 1 |
| 1994/95 | 319 | 235 | 34 | 18 | 4 | 2 |
| 1995/96 | 879 | 648 | 1,117 | 590 | 5 | 2 |
| 1996/97 | 547 | 403 | 509 | 269 | 4 | 2 |
| 1997/98 | 538 | 397 | 711 | 375 | 8 | 4 |
| 1998/99 | 541 | 399 | 574 | 303 | 15 | 7 |
| 1999/00 | 463 | 342 | 607 | 321 | 14 | 7 |
| 2000/01 | 436 | 322 | 495 | 261 | 16 | 8 |
| 2001/02 | 488 | 360 | 510 | 269 | 13 | 6 |
| 2002/03 | 406 | 299 | 438 | 231 | 15 | 7 |
| 2003/04 | 405 | 299 | 416 | 220 | 17 | 8 |
| 2004/05 | 280 | 207 | 299 | 158 | 10 | 5 |
| 2005/06 | 266 | 196 | 232 | 123 | 12 | 6 |
| 2006/07 | 234 | 173 | 143 | 76 | 14 | 7 |
| 2007/08 | 199 | 147 | 134 | 71 | 17 | 8 |
| 2008/09 | 197 | 145 | 113 | 60 | 15 | 7 |
| 2009/10 | 170 | 125 | 95 | 50 | 16 | 8 |
| 2010/11 | 183 | 135 | 108 | 57 | 26 | 13 |
| 2011/12 | 160 | 118 | 107 | 57 | 13 | 6 |
| 2012/13 | 187 | 138 | 99 | 52 | 18 | 9 |
| 2013/14 | 193 | 142 | 122 | 64 | 17 | 8 |
| 2014/15 | 168 | 124 | 99 | 52 | 16 | 8 |
| 2015/16 | 190 | 140 | 125 | 66 | 10 | 5 |
| 2016/17 | 247 | 182 | 155 | 82 | 12 | 6 |
| 2017/18 | 224 | 165 | 133 | 70 | 12 | 6 |
| 2018/19 | 256 | 189 | 234 | 124 | 9 | 4 |
| 2019/20 | 242 | 178 | 148 | 78 | 8 | 4 |
| 2020/21 | 227 | 167 | 155 | 82 | 6 | 3 |
| 2021/22 | 271 | 200 | 138 | 73 | 15 | 7 |
| 2022/23 | 161 | 119 | 44 | 23 | 8 | 4 |

Table 6. Comparison of parameter estimates and the 2022 MMB (MMB estimated on 15 Feb 2023 ) for models 22.9c, 22.1e2, 22.1f, 22.1 g , and 21.1 h for the golden king crab data from EAG, 1985/86-2022/23. Recruitment and fishing mortality deviations and initial size frequency determination parameters were omitted from this list. Parameter estimates for all models except 22.9c are from GMACS fits. Model 22.9c estimates are listed with the CVs.

|  | Model 22.9c |  | Model 22.1e2 | Model 22.1f | Model 22.1g | Model 22.1h | Model 22.9c |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Parameter | Estimate | CV | Estimate | Estimate | Estimate | Estimate | Limits |
| $\log _{-} \omega_{1}$ ( growth incr. intercept) | 2.513 | 0.007 | 2.513 | 2.518 | 2.518 | 2.518 | 1.0, 4.5 |
| $\omega_{2}$ (growth incr. slope) | -12.951 | 0.129 | -12.947 | -12.177 | -12.132 | -12.146 | -15.0, 5.0 |
| $\log _{-} \mathrm{a}$ (molt prob. slope) | -2.542 | 0.020 | -2.542 | -2.537 | -2.540 | -2.537 | -4.61, -1.39 |
| log_b (molt prob. L50) | 4.952 | 0.001 | 4.952 | 4.953 | 4.953 | 4.953 | 3.869, 5.05 |
| $\sigma$ (growth variability std) | 3.681 | 0.026 | 3.681 | 3.678 | 3.679 | 3.679 | 0.1, 12.0 |
| log_total sel delta $\theta$, 1985-04 | 4.238 | 0.029 | 4.237 | 4.137 | 4.128 | 4.132 | 0.0, 4.4 |
| $\log _{-}$total sel delta $\theta, 2005-22$ | 3.186 | 0.018 | 3.186 | 3.168 | 3.176 | 3.171 | 0.0, 4.4 |
| $\log _{-}$ret. sel delta $\theta, 1985-22$ | 1.867 | 0.019 | 1.867 | 1.863 | 1.863 | 1.863 | 0.0, 4.4 |
| log_tot sel $\theta_{50}, 1985-04$ | 4.798 | 0.009 | 4.798 | 4.786 | 4.783 | 4.786 | 4.0, 5.0 |
| log_tot sel $\theta_{50}, 2005-22$ | 4.917 | 0.002 | 4.917 | 4.914 | 4.917 | 4.915 | 4.0, 5.0 |
| $\log _{\text {_ret. }}$ sel $\theta_{50}, 1985-22$ | 4.916 | 0.000 | 4.916 | 4.916 | 4.916 | 4.916 | 4.0, 5.0 |
| $\log _{\_} \beta_{\mathrm{r}}$ (rec.distribution par.) | 0.480 | 0.650 | 0.480 | 0.394 | 0.397 | 0.392 | -12.0, 12.0 |
| $\operatorname{logq} 1$ (fishery catchability, 1985-98) | -0.469 | 0.187 | -0.469 | -0.478 | -0.479 | -0.478 | -9.0, 2.25 |
| $\operatorname{logq} 2$ (fishery/observer catchability, 1985-04) | -0.624 | 0.178 | -0.625 | -0.626 | -0.620 | -0.629 | -9.0, 2.25 |
| logq3 (observer catchability, 2005-22) | -0.806 | 0.142 | -0.805 | -0.804 | -0.814 | -0.812 | -9.0, 2.25 |
| log_mean_rec (mean rec.) | 0.883 | 0.046 | 1.008 | 1.006 | 0.990 | 0.994 | 0.01, 5.0 |
| $\log _{\text {_ }}$ mean_Fpot (Pot fishery F) | -1.005 | 0.070 | -1.005 | -1.017 | -0.991 | -1.003 | -15.0, -0.01 |
| log_mean_Fground (GF byc. F) | -8.431 | 0.082 | -8.431 | -8.431 | -8.404 | -8.412 | -15.0, -1.6 |
| $\log S E 1$ (fishery CPUE additional std, 1985-98) | -1.629 | 0.136 | -1.622 | -1.596 | -1.590 | -1.595 | -8.0, 1.0 |
| $\log S E 2$ (fishery/observer CPUE additional std, 1985-04) | -1.489 | 0.166 | -1.489 | -2.170 | -1.504 | -2.169 | -8.0, 0.15 |

$\log S E 3$ (observer CPUE additional std, 2005-22)
$\begin{array}{llllll}-1.427 & 0.140 & -1.428 & -1.600 & -1.299 & -1.351\end{array}$
$-8.0,0.15$
2022 MMB $\qquad$

Table 7. Annual abundance estimates of model recruits (millions of crab), legal male biomass ( t ) with coefficient of variations (CV), and mature male biomass ( t ) with CV for model 22.9c for golden king crab in EAG. Legal male biomass was estimated on July 1 (start of fishing year) of fishing year y . Mature male biomass for fishing year y was estimated on February 15 of year $\mathrm{y}+1$, after the year y fishery total catch removal. Recruits estimates for 1961 to 2023 are restricted to 1985-2023. Equilibrium $\mathrm{MMB}_{\text {eq }}$ and $\mathrm{MMB}_{35 \%}$ are also listed.

| Year | Recruits to the Model ( $\geq 101$ mm CL) | Mature Male Biomass $(\geq 116 \mathrm{~mm} \mathrm{CL})$ | CV | Legal Size Male Biomass ( $\geq 136$ mm CL) | CV |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{aligned} & \mathrm{MMB}_{\mathrm{eq}}=21,415 \\ & M M B_{35 \%}=6,665 \end{aligned}$ |  |  |  |
| 1985 | 2.13 | 9,350 | 0.04 | 10,380 | 0.05 |
| 1986 | 1.38 | 7,254 | 0.04 | 8,543 | 0.04 |
| 1987 | 2.18 | 6,242 | 0.04 | 6,629 | 0.04 |
| 1988 | 5.92 | 5,931 | 0.05 | 5,658 | 0.04 |
| 1989 | 1.55 | 5,607 | 0.06 | 4,854 | 0.05 |
| 1990 | 3.25 | 5,662 | 0.05 | 4,613 | 0.06 |
| 1991 | 3.30 | 5,765 | 0.04 | 4,908 | 0.05 |
| 1992 | 2.92 | 5,741 | 0.04 | 4,884 | 0.05 |
| 1993 | 2.32 | 6,161 | 0.04 | 4,871 | 0.04 |
| 1994 | 2.52 | 5,758 | 0.04 | 5,392 | 0.03 |
| 1995 | 2.67 | 5,192 | 0.04 | 5,052 | 0.03 |
| 1996 | 1.95 | 5,205 | 0.05 | 4,468 | 0.04 |
| 1997 | 3.18 | 5,356 | 0.05 | 4,539 | 0.05 |
| 1998 | 3.11 | 6,035 | 0.06 | 4,624 | 0.05 |
| 1999 | 2.86 | 6,810 | 0.06 | 5,157 | 0.06 |
| 2000 | 3.23 | 7,544 | 0.06 | 5,931 | 0.06 |
| 2001 | 1.95 | 8,036 | 0.07 | 6,644 | 0.06 |
| 2002 | 2.52 | 8,229 | 0.07 | 7,241 | 0.07 |
| 2003 | 1.97 | 8,315 | 0.07 | 7,523 | 0.07 |
| 2004 | 1.62 | 8,078 | 0.08 | 7,620 | 0.07 |
| 2005 | 3.20 | 7,989 | 0.08 | 7,479 | 0.08 |
| 2006 | 2.48 | 8,364 | 0.07 | 7,237 | 0.08 |
| 2007 | 2.03 | 8,484 | 0.07 | 7,473 | 0.08 |
| 2008 | 2.75 | 8,471 | 0.07 | 7,712 | 0.07 |
| 2009 | 2.41 | 8,616 | 0.07 | 7,697 | 0.07 |
| 2010 | 2.05 | 8,597 | 0.06 | 7,795 | 0.07 |
| 2011 | 2.02 | 8,350 | 0.06 | 7,851 | 0.06 |
| 2012 | 1.93 | 7,966 | 0.06 | 7,670 | 0.06 |
| 2013 | 1.86 | 7,488 | 0.06 | 7,313 | 0.06 |
| 2014 | 2.26 | 7,111 | 0.06 | 6,858 | 0.06 |
| 2015 | 3.02 | 7,033 | 0.07 | 6,446 | 0.06 |
| 2016 | 3.06 | 7,368 | 0.07 | 6,220 | 0.07 |
| 2017 | 3.37 | 7,957 | 0.08 | 6,432 | 0.07 |
| 2018 | 4.05 | 8,660 | 0.09 | 6,956 | 0.08 |
| 2019 | 3.12 | 9,246 | 0.11 | 7,530 | 0.09 |
| 2020 | 2.70 | 9,646 | 0.13 | 8,128 | 0.11 |
| 2021 | 1.86 | 9,465 | 0.16 | 8,700 | 0.13 |
| 2022 | 2.20 | 9,059 | 0.18 | 8,711 | 0.16 |
| 2023 | 2.42 |  |  |  |  |

Table 8. Annual abundance estimates of model recruits (millions of crab) and mature male biomass (t) from GMACS implementation of models $22.1 \mathrm{e} 2,22.1 \mathrm{f}, 22.1 \mathrm{~g}$, and 22.1h for golden king crab in EAG. Estimates are restricted to 1985-2022.

|  | 22.1e2 | 22.1 e 2 | 22.1 f | 22.1 f | 22.1g | 22.1g | 22.1h | 22.1h |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | Recruits to the Model ( $\geq$ 101 mm CL) | Mature Male Biomass $\begin{gathered} (\geq 116 \\ \mathrm{mm} \text { CL }) \end{gathered}$ | Recruits to the Model ( $\geq$ 101 mm CL) | Mature Male Biomass $\begin{gathered} (\geq 116 \\ \mathrm{mm} \mathrm{CL}) \end{gathered}$ | Recruits to the Model ( $\geq 101$ mm CL) | Mature Male Biomass $\begin{gathered} (\geq 116 \\ \mathrm{mm} \mathrm{CL}) \end{gathered}$ | Recruits to the Model ( $\geq 101$ mm CL) | Mature Male Biomass $\begin{gathered} (\geq 116 \\ \mathrm{mm} \mathrm{CL}) \end{gathered}$ |
| 1985 | 1.38 | 9,350 | 1.33 | 9,330 | 1.32 | 9,311 | 1.32 | 9,330 |
| 1986 | 2.18 | 7,254 | 2.19 | 7,235 | 2.18 | 7,223 | 2.18 | 7,235 |
| 1987 | 5.93 | 6,242 | 5.99 | 6,206 | 6.00 | 6,193 | 6.01 | 6,202 |
| 1988 | 1.55 | 5,930 | 1.54 | 5,895 | 1.53 | 5,884 | 1.53 | 5,894 |
| 1989 | 3.25 | 5,607 | 3.20 | 5,593 | 3.20 | 5,581 | 3.19 | 5,596 |
| 1990 | 3.30 | 5,663 | 3.44 | 5,635 | 3.45 | 5,621 | 3.45 | 5,635 |
| 1991 | 2.92 | 5,766 | 2.84 | 5,748 | 2.85 | 5,735 | 2.84 | 5,748 |
| 1992 | 2.32 | 5,743 | 2.26 | 5,758 | 2.28 | 5,751 | 2.26 | 5,760 |
| 1993 | 2.52 | 6,163 | 2.48 | 6,145 | 2.45 | 6,149 | 2.48 | 6,147 |
| 1994 | 2.67 | 5,761 | 2.65 | 5,709 | 2.64 | 5,717 | 2.66 | 5,709 |
| 1995 | 1.95 | 5,194 | 1.88 | 5,117 | 1.86 | 5,112 | 1.88 | 5,116 |
| 1996 | 3.18 | 5,208 | 3.22 | 5,106 | 3.20 | 5,089 | 3.22 | 5,107 |
| 1997 | 3.11 | 5,359 | 3.05 | 5,236 | 3.02 | 5,203 | 3.06 | 5,239 |
| 1998 | 2.86 | 6,038 | 2.91 | 5,912 | 2.86 | 5,860 | 2.91 | 5,918 |
| 1999 | 3.23 | 6,814 | 3.27 | 6,680 | 3.24 | 6,601 | 3.28 | 6,692 |
| 2000 | 1.95 | 7,547 | 1.96 | 7,445 | 1.93 | 7,336 | 1.97 | 7,465 |
| 2001 | 2.52 | 8,039 | 2.62 | 7,971 | 2.55 | 7,840 | 2.64 | 8,000 |
| 2002 | 1.97 | 8,233 | 2.04 | 8,208 | 2.04 | 8,048 | 2.06 | 8,248 |
| 2003 | 1.62 | 8,318 | 1.66 | 8,369 | 1.68 | 8,184 | 1.68 | 8,425 |
| 2004 | 3.20 | 8,080 | 3.09 | 8,191 | 3.15 | 8,022 | 3.13 | 8,263 |
| 2005 | 2.48 | 7,991 | 2.56 | 8,096 | 2.55 | 7,968 | 2.58 | 8,186 |
| 2006 | 2.03 | 8,366 | 2.06 | 8,433 | 2.04 | 8,350 | 2.06 | 8,542 |
| 2007 | 2.75 | 8,487 | 2.79 | 8,578 | 2.78 | 8,507 | 2.77 | 8,689 |
| 2008 | 2.41 | 8,474 | 2.36 | 8,582 | 2.32 | 8,513 | 2.31 | 8,680 |
| 2009 | 2.05 | 8,618 | 2.01 | 8,723 | 1.96 | 8,648 | 1.97 | 8,788 |
| 2010 | 2.02 | 8,599 | 2.07 | 8,664 | 2.01 | 8,569 | 2.02 | 8,681 |
| 2011 | 1.93 | 8,352 | 1.95 | 8,397 | 1.88 | 8,271 | 1.89 | 8,367 |
| 2012 | 1.86 | 7,969 | 1.82 | 8,030 | 1.77 | 7,862 | 1.78 | 7,947 |
| 2013 | 2.26 | 7,490 | 2.26 | 7,549 | 2.21 | 7,343 | 2.21 | 7,419 |
| 2014 | 3.02 | 7,113 | 3.03 | 7,150 | 2.97 | 6,915 | 2.97 | 6,986 |
| 2015 | 3.06 | 7,036 | 3.04 | 7,062 | 2.84 | 6,800 | 2.84 | 6,866 |
| 2016 | 3.37 | 7,370 | 3.33 | 7,386 | 3.11 | 7,058 | 3.09 | 7,118 |
| 2017 | 4.05 | 7,958 | 4.06 | 7,945 | 3.85 | 7,482 | 3.82 | 7,527 |
| 2018 | 3.12 | 8,660 | 3.04 | 8,620 | 2.72 | 8,006 | 2.68 | 8,027 |
| 2019 | 2.70 | 9,245 | 2.65 | 9,183 | 2.38 | 8,397 | 2.34 | 8,386 |
| 2020 | 1.86 | 9,644 | 1.92 | 9,535 | 1.80 | 8,558 | 1.78 | 8,506 |
| 2021 | 2.20 | 9,462 | 2.24 | 9,346 | 2.17 | 8,247 | 2.17 | 8,162 |
| 2022 | 2.42 | 9,055 | 2.41 | 8,981 | 2.38 | 7,864 | 2.38 | 7,765 |

Table 9. Negative log-likelihood values of the fits for model 22.9c (with additional 2022/23 data), and GMACS implementation of models $22.1 \mathrm{e} 2,22.1 \mathrm{f}, 22.1 \mathrm{~g}$, and 21.1 h for golden king crab in EAG. Likelihood components with zero entry in the entire rows are omitted. RetdcatchB $=$ retained catch biomass.

| Likelihood Component | Model 22.9c | Model 22.1e2 | Model 22.1f | $\mathbf{2 2 . 1 g}$ | 22.1h |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Retlencomp | 286.2230 | 286.2369 | 265.4302 | 262.7069 | 262.3774 |
| Totallencomp | 520.2600 | 520.2876 | 553.999 | 555.5594 | 554.3931 |
| Observer cpue | -26.7588 | -26.7606 | -32.6846 | -23.8624 | -28.4356 |
| Fishery cpue | -15.5853 | -15.5297 | -15.1827 | -15.1038 | -15.177 |
| RetdcatchB | -421.9470 | -421.953 | -422.049 | -422.125 | -422.053 |
| TotalcatchB | -40.9361 | -40.9455 | -41.384 | -41.4766 | -41.3155 |
| GdiscdcatchB | 30.3249 | 30.32492 | 30.3248 | 30.3248 | 30.3247 |
| Rec_dev | 22.7112 | 20.7514 | 20.8089 | 20.6410 | 20.6312 |
| Pot F_dev | 0.0135 |  |  |  |  |
| Gbyc_F_dev | 0.0239 |  |  |  |  |
| Sum (Pot F_dev+ Gbyc_F_dev) | 0.0374 | 0.0373 | 0.0371 | 0.0371 | 0.0373 |
| Tag | 2701.2600 | 2701.2579 | 2700.409 | 2700.569 | 2700.389 |
| Total | 3055.5900 | 3079.431812 | 3085.433 | 3092.9951 | 3086.8961 |

Table 10. Time series of negative binomial GLM estimated CPUE indices and coefficient of variations (CV) for the fish ticket based retained catch-per-pot lift for the WAG golden king crab stock. The GLM was fitted to the 1985/86 to 1998/99 time series of data. GLM predictor variables were selected by R square criteria.

| Year | CPUE <br> Index | CV |
| :---: | :---: | :---: |
| $1985 / 86$ | 2.07 | 0.02 |
| $1986 / 87$ | 1.59 | 0.03 |
| $1987 / 88$ | 1.22 | 0.04 |
| $1988 / 89$ | 1.41 | 0.02 |
| $1989 / 90$ | 1.15 | 0.02 |
| $1990 / 91$ | 0.87 | 0.04 |
| $1991 / 92$ | 0.76 | 0.05 |
| $1992 / 93$ | 0.61 | 0.07 |
| $1993 / 94$ | 0.76 | 0.07 |
| $1994 / 95$ | 0.83 | 0.04 |
| $1995 / 96$ | 0.90 | 0.04 |
| $1996 / 97$ | 0.84 | 0.04 |
| $1997 / 98$ | 0.76 | 0.04 |
| $1998 / 99$ | 1.06 | 0.03 |

Table 11. Initial input number of vessel-days/trips and Stage-2 effective sample sizes iteratively estimated by the Francis method for retained, total, and groundfish discard catch size compositions of golden king crab for model 22.9 c fit to WAG data. NA: not available.

| Year | Initial <br> Input Retained VesselDays Sample Size (no) | Stage-2 <br> Retained <br> Effective <br> Sample <br> Size (no) | Initial Input Total VesselDays Sample Size (no) | Stage-2 <br> Total Effective Sample Size (no) | Initial Input Groundfish Trip Sample Size (no) | Stage-2 Groundfish Effective Sample Size (no) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1985/86 | 45 | 20 |  |  |  |  |
| 1986/87 | 23 | 10 |  |  |  |  |
| 1987/88 | 8 | 4 |  |  |  |  |
| 1988/89 | 286 | 128 |  |  |  |  |
| 1989/90 | 513 | 229 |  |  | 7 | 3 |
| 1990/91 | 205 | 92 | 190 | 108 | 6 | 3 |
| 1991/92 | 102 | 46 | 104 | 59 | 1 | 1 |
| 1992/93 | 76 | 34 | 94 | 53 | 3 | 1 |
| 1993/94 | 378 | 169 | 62 | 35 | NA | NA |
| 1994/95 | 367 | 164 | 119 | 68 | 2 | 1 |
| 1995/96 | 705 | 315 | 907 | 516 | 5 | 2 |
| 1996/97 | 817 | 365 | 1,061 | 603 | 8 | 4 |
| 1997/98 | 984 | 440 | 1,116 | 635 | 6 | 3 |
| 1998/99 | 613 | 274 | 638 | 363 | 14 | 7 |
| 1999/00 | 915 | 409 | 1,155 | 657 | 18 | 9 |
| 2000/01 | 1,029 | 460 | 1,205 | 685 | 11 | 5 |
| 2001/02 | 898 | 402 | 975 | 555 | 11 | 5 |
| 2002/03 | 628 | 281 | 675 | 384 | 16 | 8 |
| 2003/04 | 688 | 308 | 700 | 398 | 8 | 4 |
| 2004/05 | 449 | 201 | 488 | 278 | 9 | 4 |
| 2005/06 | 337 | 151 | 220 | 125 | 6 | 3 |
| 2006/07 | 337 | 151 | 321 | 183 | 14 | 7 |
| 2007/08 | 276 | 123 | 257 | 146 | 17 | 8 |
| 2008/09 | 318 | 142 | 258 | 147 | 19 | 9 |
| 2009/10 | 362 | 162 | 292 | 166 | 24 | 12 |
| 2010/11 | 328 | 147 | 222 | 126 | 13 | 6 |
| 2011/12 | 295 | 132 | 252 | 143 | 14 | 7 |
| 2012/13 | 288 | 129 | 241 | 137 | 18 | 9 |
| 2013/14 | 327 | 146 | 236 | 134 | 17 | 8 |
| 2014/15 | 305 | 136 | 219 | 125 | 18 | 9 |
| 2015/16 | 287 | 128 | 243 | 138 | 10 | 5 |
| 2016/17 | 408 | 183 | 253 | 144 | 12 | 6 |
| 2017/18 | 309 | 138 | 222 | 126 | 10 | 5 |
| 2018/19 | 291 | 130 | 318 | 181 | 5 | 2 |
| 2019/20 | 363 | 162 | 224 | 127 | 6 | 3 |
| 2020/21 | 462 | 207 | 302 | 172 | 7 | 3 |
| 2021/22 | 446 | 199 | 247 | 140 | 4 | 2 |
| 2022/23 | 296 | 132 | 195 | 111 | 5 | 2 |

Table 12. Comparison of parameter estimates and the 2022 MMB (MMB estimated on 15 Feb 2023 ) for models 22.9 c , 22.1e2, and 22.1 f for the golden king crab data from WAG, 1985/86-2022/23. Recruitment and fishing mortality deviations and initial size frequency determination parameters were omitted from this list. Parameter estimates for all models except 22.9c are from GMACS fits. Model 22.9 c estimates are listed with the CVs.

|  | Model 22.9c |  | Model 22.1e2 | Model 22.1f | Model 22.9c |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Parameter | Estimate | CV | Estimate | Estimate | Limits |
| $\log _{-} \omega_{1}$ ( growth incr. intercept) | 2.506 | 0.007 | 2.506 | 2.518 | 1.0, 4.5 |
| $\omega_{2}$ (growth incr. slope) | -13.156 | 0.124 | -13.156 | -11.550 | -15.0, 5.0 |
| log_a (molt prob. slope) | -2.706 | 0.022 | -2.706 | -2.693 | -4.61, -1.39 |
| $\log _{-} \mathrm{b}$ (molt prob. L50) | 4.951 | 0.001 | 4.951 | 4.952 | 3.869, 5.05 |
| $\sigma$ (growth variability std) | 3.672 | 0.026 | 3.672 | 3.667 | 0.1, 12.0 |
| log_total sel delta $\theta$, 1985-04 | 3.979 | 0.015 | 3.978 | 3.857 | 0.0, 4.4 |
| $\log _{-}$total sel deltag, 2005-22 | 3.069 | 0.014 | 3.069 | 3.062 | 0.0, 4.4 |
| $\log _{-}$ret. sel delta $\theta, 1985-22$ | 1.708 | 0.024 | 1.708 | 1.705 | 0.0, 4.4 |
| log_tot sel $\theta_{50}, 1985-04$ | 4.909 | 0.005 | 4.909 | 4.885 | 4.0, 5.0 |
| log_tot sel $\theta_{50}, 2005-22$ | 4.904 | 0.001 | 4.904 | 4.902 | 4.0, 5.0 |
| log_ret. sel $\theta_{50}, 1985-22$ | 4.913 | 0.0002 | 4.913 | 4.913 | 4.0, 5.0 |
| $\log _{\_} \beta_{\mathrm{r}}$ (rec.distribution par.) | -0.074 | 2.104 | -0.074 | -0.211 | -12.0, 12.0 |
| $\operatorname{logq1} 1$ (fishery catchability, 1985-98) | 0.040 | 1.930 | 0.039 | -0.015 | -9.0, 2.25 |
| $\operatorname{logq} 2$ (fishery/observer catchability, 1985-04) | 0.089 | 1.160 | 0.087 | 0.045 | -9.0, 2.25 |
| $\operatorname{logq} 3$ (observer catchability, 2005-22) | -0.315 | 0.243 | -0.316 | -0.310 | -9.0, 2.25 |
| log_mean_rec (mean rec.) | 0.700 | 0.054 | 0.825 | 0.819 | 0.01, 5.0 |
| log_mean_Fpot (Pot fishery F) | -0.695 | 0.082 | -0.696 | -0.723 | -15.0, -0.01 |
| log_mean_Fground (GF byc. F) | -8.174 | 0.085 | -8.175 | -8.172 | -15.0, -1.6 |
| $\log S E 1$ (fishery CPUE additional std, 1985-98) | -1.938 | 0.134 | -1.955 | -1.964 | -8.0, 1.0 |
| $\log$ SE2 (fishery/observer CPUE additional std, 1985-04) | -1.496 | 0.165 | -1.494 | -1.587 | -8.0, 0.15 |
| $\log$ SE3 (observer CPUE additional std, 2005-22) |  |  |  |  |  |
| 2022 MMB | -2.135 4,495 | 0.100 0.13 | -2.124 4,545 | -2.047 4,288 | -8.0, 0.15 |

Table 13. Annual abundance estimates of model recruits (millions of crab), legal male biomass ( t ) with coefficient of variations (CV), and mature male biomass ( t ) with CV for model 22.9c for golden king crab in WAG. Legal male biomass was estimated on July 1 (start of fishing year) of fishing year y . Mature male biomass for fishing year y was estimated on February 15 of year $\mathrm{y}+1$, after the year y fishery total catch removal. Recruits estimates for 1961 to 2023 are restricted to 1985-2023. Equilibrium $\mathrm{MMB}_{\text {eq }}$ and $\mathrm{MMB}_{35 \%}$ are also listed.

| Year | $\begin{aligned} & \text { Recruits to the } \\ & \text { Model }(\geq 101 \\ & \text { mm CL) } \end{aligned}$ | Mature Male Biomass $(\geq 116 \mathrm{~mm} \mathrm{CL})$ | CV | Legal Size Male Biomass ( $\geq 136$ mm CL) | CV |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{aligned} & \mathrm{MMB}_{\mathrm{eq}}=17,757 \\ & M M B_{35 \%}=4,960 \end{aligned}$ |  |  |  |
| 1985 | 3.56 | 9,969 | 0.05 | 9,329 | 0.08 |
| 1986 | 3.85 | 7,482 | 0.05 | 8,761 | 0.06 |
| 1987 | 2.89 | 7,050 | 0.04 | 6,236 | 0.06 |
| 1988 | 1.95 | 6,163 | 0.04 | 5,877 | 0.04 |
| 1989 | 2.38 | 4,225 | 0.04 | 5,246 | 0.04 |
| 1990 | 1.83 | 3,893 | 0.04 | 3,473 | 0.05 |
| 1991 | 1.35 | 3,602 | 0.04 | 3,191 | 0.04 |
| 1992 | 1.58 | 3,516 | 0.04 | 3,024 | 0.04 |
| 1993 | 2.34 | 4,051 | 0.03 | 3,000 | 0.04 |
| 1994 | 1.91 | 3,687 | 0.03 | 3,421 | 0.03 |
| 1995 | 2.00 | 3,723 | 0.04 | 2,994 | 0.03 |
| 1996 | 1.68 | 3,727 | 0.04 | 3,056 | 0.03 |
| 1997 | 1.94 | 3,812 | 0.04 | 3,096 | 0.03 |
| 1998 | 1.95 | 4,166 | 0.04 | 3,195 | 0.04 |
| 1999 | 2.35 | 4,249 | 0.04 | 3,512 | 0.03 |
| 2000 | 2.53 | 4,441 | 0.05 | 3,536 | 0.04 |
| 2001 | 2.40 | 4,849 | 0.05 | 3,644 | 0.04 |
| 2002 | 2.70 | 5,374 | 0.05 | 4,017 | 0.05 |
| 2003 | 1.39 | 5,683 | 0.05 | 4,509 | 0.05 |
| 2004 | 2.25 | 5,588 | 0.06 | 4,926 | 0.05 |
| 2005 | 2.50 | 5,837 | 0.06 | 4,926 | 0.06 |
| 2006 | 2.29 | 6,389 | 0.05 | 5,040 | 0.06 |
| 2007 | 2.17 | 6,626 | 0.05 | 5,531 | 0.05 |
| 2008 | 1.51 | 6,634 | 0.05 | 5,796 | 0.05 |
| 2009 | 1.72 | 6,210 | 0.05 | 5,914 | 0.05 |
| 2010 | 1.82 | 5,871 | 0.05 | 5,584 | 0.05 |
| 2011 | 1.28 | 5,473 | 0.04 | 5,228 | 0.05 |
| 2012 | 1.74 | 4,882 | 0.05 | 4,874 | 0.04 |
| 2013 | 2.20 | 4,570 | 0.05 | 4,314 | 0.05 |
| 2014 | 1.95 | 4,696 | 0.05 | 3,896 | 0.05 |
| 2015 | 1.64 | 4,803 | 0.05 | 3,959 | 0.05 |
| 2016 | 1.66 | 4,893 | 0.05 | 4,121 | 0.05 |
| 2017 | 1.39 | 4,884 | 0.05 | 4,271 | 0.05 |
| 2018 | 1.57 | 4,633 | 0.05 | 4,311 | 0.05 |
| 2019 | 1.65 | 4,318 | 0.06 | 4,084 | 0.05 |
| 2020 | 1.76 | 4,041 | 0.08 | 3,742 | 0.06 |
| 2021 | 1.67 | 4,158 | 0.10 | 3,433 | 0.08 |
| 2022 | 1.64 | 4,495 | 0.13 | 3,539 | 0.10 |
| 2023 | 2.01 |  |  |  |  |

Table 14. Annual abundance estimates of model recruits (millions of crab) and mature male biomass ( t ) from GMACS implementation of models 22.1 e 2 and 22.1 f for golden king crab in WAG. Estimates are restricted to 1985-2022.

|  | 22.1e2 | 22.1 e 2 | 22.1 f | 22.1 f |
| :---: | :---: | :---: | :---: | :---: |
| Year | Recruits to the Model $(\geq 101 \mathrm{~mm} \mathrm{CL})$ | Mature Male Biomass $(\geq 116 \mathrm{~mm} \mathrm{CL})$ | Recruits to the Model $(\geq 101 \mathrm{~mm} \mathrm{CL})$ | Mature Male Biomass $(\geq 116 \mathrm{~mm} \text { CL })$ |
| 1985 | 3.85 | 9,937 | 3.81 | 9,966 |
| 1986 | 2.89 | 7,477 | 2.89 | 7,462 |
| 1987 | 1.95 | 7,052 | 1.91 | 7,017 |
| 1988 | 2.38 | 6,166 | 2.56 | 6,127 |
| 1989 | 1.84 | 4,229 | 1.76 | 4,210 |
| 1990 | 1.35 | 3,899 | 1.31 | 3,936 |
| 1991 | 1.58 | 3,609 | 1.62 | 3,622 |
| 1992 | 2.35 | 3,521 | 2.41 | 3,523 |
| 1993 | 1.91 | 4,054 | 1.89 | 4,075 |
| 1994 | 2.00 | 3,690 | 1.94 | 3,734 |
| 1995 | 1.68 | 3,726 | 1.72 | 3,751 |
| 1996 | 1.94 | 3,731 | 1.94 | 3,736 |
| 1997 | 1.95 | 3,816 | 1.86 | 3,830 |
| 1998 | 2.35 | 4,169 | 2.37 | 4,159 |
| 1999 | 2.53 | 4,252 | 2.45 | 4,196 |
| 2000 | 2.41 | 4,445 | 2.49 | 4,363 |
| 2001 | 2.70 | 4,853 | 2.75 | 4,750 |
| 2002 | 1.39 | 5,378 | 1.57 | 5,313 |
| 2003 | 2.25 | 5,687 | 2.49 | 5,704 |
| 2004 | 2.51 | 5,591 | 2.51 | 5,756 |
| 2005 | 2.29 | 5,840 | 2.33 | 6,130 |
| 2006 | 2.17 | 6,393 | 2.10 | 6,708 |
| 2007 | 1.51 | 6,629 | 1.37 | 6,932 |
| 2008 | 1.72 | 6,636 | 1.61 | 6,858 |
| 2009 | 1.82 | 6,213 | 1.82 | 6,313 |
| 2010 | 1.28 | 5,874 | 1.26 | 5,879 |
| 2011 | 1.74 | 5,476 | 1.70 | 5,446 |
| 2012 | 2.20 | 4,885 | 2.28 | 4,822 |
| 2013 | 1.95 | 4,574 | 1.92 | 4,497 |
| 2014 | 1.65 | 4,701 | 1.57 | 4,649 |
| 2015 | 1.67 | 4,808 | 1.52 | 4,736 |
| 2016 | 1.39 | 4,899 | 1.39 | 4,760 |
| 2017 | 1.58 | 4,891 | 1.59 | 4,684 |
| 2018 | 1.66 | 4,641 | 1.63 | 4,433 |
| 2019 | 1.78 | 4,333 | 1.73 | 4,131 |
| 2020 | 1.68 | 4,067 | 1.61 | 3,853 |
| 2021 | 1.65 | 4,199 | 1.68 | 3,958 |
| 2022 | 2.01 | 4,545 | 2.00 | 4,288 |

Table 15. Negative log-likelihood values of the fits for model 22.9c (with additional 2022/23 data), and GMACS implementation of models 22.1e2 and 22.1f for golden king crab in WAG. Likelihood components with zero entry in the entire rows are omitted. RetdcatchB = retained catch biomass.

| Likelihood Component | Model 22.9c | Model 22.1e2 | Model 22.1f |
| :--- | :---: | :---: | :---: |
| Retlencomp | 363.7120 | 363.8280 | 313.3108 |
| Totallencomp | 435.9380 | 436.0861 | 478.6189 |
| Observer cpue | -38.6873 | -38.5262 | -37.7272 |
| Fishery cpue | -19.6942 | -19.8406 | -19.9340 |
| RetdcatchB | -420.4380 | -420.436 | -420.458 |
| TotalcatchB | 14.1469 | 14.13333 | 12.9985 |
| GdiscdcatchB | 30.3262 | 30.32618 | 30.3258 |
| Rec_dev | 21.5391 | 19.5703 | 20.0221 |
| Pot F_dev | 0.0264 |  |  |
| Gbyc_F_dev | 0.0428 |  |  |
| Sum (Pot F_dev+ Gbyc_F_dev) | 0.0692 | 0.0692 | 0.0692 |
| Tag | 2705.5800 | 2705.561 | 2703.436 |
| Total | 3092.5000 | 3115.8015 | 3105.693 |



Figure 1. Aleutian Islands, Area O, red and golden king crab management area (from Nichols et al. 2021).


Figure 2. Adak (Area R) and Dutch Harbor (Area O) king crab registration area and districts, 1984/85-1995/96 seasons (Leon et al. 2017).


Figure 3. Percent of total 1981/82-1995/96 golden king crab retained catch weight (harvest) from one-degree longitude intervals in the Aleutian Islands, with dotted line denoting the border at $171^{\circ} \mathrm{W}$ longitude used during the 1984/85-1995/96 seasons to divide fishery management between the Dutch Harbor Area (east of $171^{\circ} \mathrm{W}$ longitude) and the Adak Area (west of $171^{\circ} \mathrm{W}$ longitude) and solid line denoting the border at $174^{\circ} \mathrm{W}$ longitude used since the 1996/97 season to manage crab east and west of $174^{\circ} \mathrm{W}$ longitude (adapted from Figure 4-2 in Morrison et al. 1998).


Figure 4. Retained catch ( t ) of golden king crab within one-degree longitude intervals in the Aleutian Islands during the 2000/01 through 2022/23 commercial fishery seasons; solid line denotes the border at $174^{\circ} \mathrm{W}$ longitude that has been used since the 1996/97 season to manage Aleutian Island golden king crab as separate stocks east and west of $174^{\circ} \mathrm{W}$ longitude and dashed line denotes the border at $171^{\circ} \mathrm{W}$ longitude used during the 1984/851995/96 seasons to divide fishery management between the Dutch Harbor Area (east of $171^{\circ} \mathrm{W}$ longitude) and the Adak Area (west of $171^{\circ} \mathrm{W}$ longitude).


Figure 5. Average golden king crab CPUE (kg/nm2) for tows, number of tows, and average depth of tows from one-degree longitude intervals during the 2002, 2004, 2006, 2010, and 2012 NMFS Aleutian Islands bottom trawl surveys; preliminary summary of data obtained on 1 April 2013 from http://www.afsc.noaa.gov/RACE/groundfish/survey_data/default.htm.


Figure 6. Historical commercial harvest (from fish tickets; metric tons), total allowable catch (TAC), and catch-per-unit effort (CPUE, number of crab per pot lift) of golden king crab in EAG, 1985/86-2022/23 (note: 1985 refers to the 1985/86 fishing year).


Figure 7. Historical commercial harvest (from fish tickets; metric tons)), total allowable catch (TAC), and catch-per-unit effort (CPUE, number of crab per pot lift) of golden king crab in WAG, 1985/86-2022/23 (note: 1985 refers to the 1985/86 fishing year).


Figure 8. Catch distribution by statistical area in 2022/23.



Figure 9. GMACS estimated N matrix under model 22.1 e 2 for golden king crab in EAG, 1981/82 to 2022/23.



Figure 10. GMACS estimated N matrix under model 22.1 f for golden king crab in EAG, 1981/82 to 2022/23.


Figure 11. Estimated total (solid line) and retained (dashed line) selectivity for pre- and post- rationalization periods under models 22.9 c (blue), 22.1e2 (red), 22.1f (green), 22.1 g (violet), and 22.1 h (orange) fits to golden king crab data in EAG.


Figure 12. Observed (open circles) vs. predicted (solid line) tag recaptures by size bin for years 1 to 6 post tagging under model 22.9c for EAG golden king crab.


Figure 13. Estimated number of male recruits (crab size $\geq 101 \mathrm{~mm} \mathrm{CL}$ ) to the assessment model under models 22.1e2 (red), 22.1f (green), 22.9c (blue), 22.1g (violet), and 22.1h (orange) fits to the EAG golden king crab data, 1960-2022.


Figure 14. Recruit size distribution to the assessment model under models 22.1 e 2 (red), 22.1 f (green), 22.9 c (blue), 22.1 g (violet), and 22.1 h (orange) fits to the EAG golden king crab data.


Figure 15. Estimated molt probability vs. carapace length of golden king crab under models 22.1 e 2 (red), 22.1 f (green), 22.9 c (blue), 22.1 g (violet), and 22.1h (orange) fits to the EAG golden king crab data.


Figure 16. Observed (open circle) vs. predicted (solid line) retained catch (top left), total catch (top right), and groundfish bycatch (bottom left) of golden king crab under models 22.9c (blue), 22.1e2 (red), 22.1f (green), 22.1 g (violet), and 22.1 h (orange) fits to the EAG data.


Figure 17. Bubble plot of standardized residuals of retained catch length composition for model 22.9c fit to the EAG golden king crab data, 1985/86-2022/23. Green circles are the positive and pink circles are the negative standardized residuals. The area of the circle is the relative magnitude of the residual.

EAG 22.9c Total Catch Size Composition Standardized Residuals


Figure 18. Bubble plot of standardized residuals of total catch length composition for model 22.9c fit to the EAG golden king crab data, 1990/91-2022/23. Green circles are the positive and pink circles are the negative standardized residuals. The area of the circle is the relative magnitude of the residual.


Figure 19. Comparison of input CPUE indices [open circles with $+/-2$ SE for model 22.1 e 2 (left) and model 22.1 f (right)] with predicted CPUE indices (colored solid lines) for 22.1 e 2 (red) and 22.9 c (blue)[left]; and 22.1 f (green) [right] fits to the EAG golden king crab data, 1995/96-2022/23. Fish ticket CPUE indices for 1985/86-1998/99 are superimposed on observer indices with predicted CPUE indices for 22.1 e 2 (purple) and 22.9 c (dark green)[left]. Model estimated additional standard error was added to each input standard error.


Figure 20. Comparison of input CPUE indices [open circles with $+/-2$ SE for model 22.1 g (left) and model 22.1 h (right)] with predicted CPUE indices (colored solid lines) for 22.1 g (violet) [left]; and 22.1h (orange) [right] fits to the EAG golden king crab data, 1995/962022/23. Model estimated additional standard error was added to each input standard error. Note: The CPUE indices for 2015/162022/23 are cooperative survey indices.

## EAG



WAG


Figure 21. Trends in pot fishery full selection total fishing mortality of golden king crab for models 22.9 c (blue), 22.1e2 (red), 22.1f (green), 22.1 g (violet), and 22.1 h (orange) fits to the EAG (left) and models 22.9 c (blue), 22.1 e 2 (red), and 22.1 f (green) fits to the WAG (right) data, 1981/82-2022/23.


Figure 22a. Long time series trends in golden king crab mature male biomass for models 22.9 c (blue), 22.1 e 2 (red), 22.1 f (green), 22.1 g (violet), and 22.1h (orange) fits to the EAG (left) and models 22.9c (blue), 22.1e2 (red), and 22.1f (green) fits to the WAG (right) data, 1960-2022.


Figure 22b. Short time series trends in golden king crab mature male biomass for models 22.9 c (blue), 22.1 e 2 (red), 22.1 f (green), 22.1 g (violet), and 22.1 h (orange) fits to the EAG (left) and models 22.9 c (blue), 22.1 e 2 (red), and 22.1 f (green) fits to the WAG (right) data, 2013-2022.





$$
2012
$$









Figure 23. GMACS estimated N matrix under model 22.1e2 for golden king crab in WAG, 1981/82 to 2022/23.



Figure 24. GMACS estimated N matrix under model 22.1f for golden king crab in WAG, 1981/82 to 2022/23


Figure 25. Estimated total (solid line) and retained (dashed line) selectivity for pre- and post- rationalization periods under models 22.9c (blue), 22.1e2 (red), and 22.1f (green)fits to golden king crab data in WAG.


Figure 26. Observed (open circles) vs. predicted (solid line) tag recaptures by size bin for years 1 to 6 post tagging under model 22.9c fit to the WAG golden king crab data.


Figure 27. Estimated number of male recruits (crab size $\geq 101 \mathrm{~mm} \mathrm{CL}$ ) to the assessment model under models 22.9 c (blue), 22.1 e 2 (red), and 22.1 f (green) fits to the WAG golden king crab data, 1960-2022.


Figure 28. Recruit size distribution to the assessment model under models 22.9 c (blue), 22.1 e 2 (red), and 22.1 f (green) fits to the WAG golden king crab data.


Figure 29. Estimated molt probability vs. carapace length of golden king crab for models 22.9 c (blue), 22.1 e 2 (red), and 22.1 f (green) fits to the WAG golden king crab data.

Retained Catch


Groundfish Bycatch


Total Catch


Figure 30. Observed (open circle) vs. predicted (solid line) retained catch (top left), total catch (top right), and groundfish bycatch (bottom left) of golden king crab under models 22.9 c (blue), 22.1 e 2 (red) and 22.1 f (green) fits to the WAG data.


Figure 31. Bubble plot of standardized residuals of retained catch length composition for model 22.9c fit to the WAG golden king crab data, 1985/86-2022/23. Green circles are the positive and pink circles are the negative standardized residuals. The area of the circle is the relative magnitude of the residual.


Figure 32. Bubble plot of standardized residuals of total catch length composition for model 22.9c fit to the WAG golden king crab data, 1990/91-2022/23. Green circles are the positive and pink circles are the negative standardized residuals. The area of the circle is the relative magnitude of the residual.


Figure 33. Comparison of input CPUE indices [open circles with $+/-2$ SE for model 22.1 e 2 (left) and model 22.1 f (right)] with predicted CPUE indices (colored solid lines) for 22.1 e 2 (red) and 22.9c (blue)[left]; and 22.1f (green) [right] fits to the WAG golden king crab data, 1985/86-2022/23. Fish ticket CPUE indices for 1985/86-1998/99 are superimposed on observer indices with predicted CPUE indices for 22.1 e 2 (purple) and 22.9 c (dark green)[left]. Model estimated additional standard error was added to each input standard error.


Figure 34. Relationships between full fishing mortalities for the directed pot fishery and mature male biomass during 1981/82-2022/23 under models, 22.9 (modified 22.1e2, non GMACS), 22.1e2, and 22.1f fits to the EAG and WAG data. $F$ in 2022/23 (red) and 1981/82 (black) are shown in the plots.

## Appendix A: Integrated model

Aleutian Islands Golden King Crab (Lithodes aequispinus) Stock Assessment Model Development- east of $174^{\circ} \mathrm{W}$ (EAG) and west of $174^{\circ} \mathrm{W}$ (WAG) Aleutian Island stocks

## Basic population dynamics

The annual [male] abundances by size are modeled using the equation:
$N_{t+1, j}=\sum_{i=1}^{j}\left[N_{t, i} e^{-M}-\left(\hat{C}_{t, i}+\widehat{D}_{t, i}+\widehat{T r}_{t, i}\right) e^{\left(y_{t}-1\right) M}\right] X_{i, j}+R_{t+1, j}$
where $\quad N_{t, i}$ is the number of [male] crab in length class $i$ on 1 July (start of fishing year) of year $t ; \quad \hat{C}_{t, i}, \widehat{D}_{t, i}$, and $\widehat{T} r_{t, i}$ are respectively the predicted fishery retained, pot fishery discard dead, and groundfish fishery discard dead catch estimates in length class $i$ during year $t ; \widehat{D}_{t, i}$ is estimated from the intermediate total ( $\widehat{T}_{t, i \text { temp }}$ ) catch and the retained ( $\widehat{C}_{t, i}$ ) catch by Equation A.2c. $X_{i, j}$ is the probability of length-class $i$ growing into length-class $j$ during the year; $y_{t}$ is elapsed time period from 1 July to the midpoint of fishing period in year $t ; M$ is instantaneous rate of natural mortality; and $R_{t+1, j}$ recruitment to length class $j$ in year $t+1$.

The catch are predicted using the equations

$$
\begin{align*}
& \hat{T}_{t, j, t e m p}=\frac{F_{t} s_{t, j}^{T}}{z_{t, j}} N_{t, j} e^{-y_{t} M}\left(1-e^{-Z_{t, j}}\right)  \tag{A.2a}\\
& \hat{C}_{t, j}=\frac{F_{t} s_{t, j}^{T} s_{t, j}^{r}}{z_{t, j}} N_{t, j} e^{-y_{t} M}\left(1-e^{-Z_{t, j}}\right)  \tag{A.2b}\\
& \widehat{D}_{t, j}=0.2\left(\widehat{T}_{t, j, t e m p}-\hat{C}_{t, j}\right)  \tag{A.2c}\\
& \widehat{\operatorname{Tr}}_{t, j}=0.65 \frac{F_{t}^{T r} s_{j}^{T r}}{Z_{t, j}} N_{t, j} e^{-y_{t} M}\left(1-e^{-Z_{t, j}}\right)  \tag{A.2d}\\
& \widehat{T}_{t, j}=\hat{C}_{t, j}+\widehat{D}_{t, j} \tag{A.2e}
\end{align*}
$$

where $Z_{t, j}$ is total fishery-related mortality on animals in length-class $j$ during year $t$ :

$$
\begin{equation*}
Z_{t, j}=F_{t} s_{t, j}^{T} s_{t, j}^{r}+0.2 F_{t} s_{t, j}^{T}\left(1-s_{t, j}^{r}\right)+0.65 F_{t}^{T r} s_{j}^{T r} \tag{A.3}
\end{equation*}
$$

$F_{t}$ is the full selection fishing mortality in the pot fishery, $F_{t}^{T r}$ is the full selection fishing mortality in the trawl fishery, $s_{t, j}^{T}$ is the total selectivity for animals in length-class $j$ by the pot fishery during year $t, s_{j}^{T r}$ is the selectivity for animals in length-class $j$ by the trawl fishery, $s_{t, j}^{r}$ is the probability of retention for animals in length-class $j$ by the pot fishery
during year $t$. Pot bycatch mortality of 0.2 and groundfish bycatch mortality of 0.65 (average of trawl [0.8] and groundfish pot [0.5] mortality) were assumed.

## Initial abundance

The initial conditions are computed as the equilibrium initial condition using the following relations:

The equilibrium stock abundance is
$N=X . S . N+R$
The equilibrium abundance in $1960, N_{1960}$, is
$\underline{N}_{1960}=(I-X S)^{-1} \underline{R}$
where $X$ is the growth matrix, $S$ is a matrix with diagonal elements given by $e^{-M}, I$ is the identity matrix, and $\underline{R}$ is the product of average recruitment and relative proportion of total recruitment to each size-class.

We used the mean number of recruits from 1987 to 2017 in equation (A.5) to obtain the equilibrium solution under only natural mortality in year 1960, and then projected the equilibrium abundance under natural mortality with recruitment estimated for each year after 1960 up to 1985 with removal of retained catch during 1981/82-1984/85.

## Growth Matrix

The growth matrix $X$ is modeled as follows:
$X_{i, j}= \begin{cases}0 & \text { if } j<i \\ P_{i, j}+\left(1-m_{i}\right) & \text { if } j=i \\ P_{i, j} & \text { if } j>i\end{cases}$
where:

$$
P_{i, j}=m_{i}\left\{\begin{array}{lr}
\int_{-\infty}^{j_{2}-L_{i}} N\left(x \mid \mu_{i}, \sigma^{2}\right) d x & \text { if } j=i \\
\int_{j_{1}-L_{i}}^{j_{2}-L_{i}} N\left(x \mid \mu_{i}, \sigma^{2}\right) d x & \text { if } i<j<n, \\
\int_{j_{1}-L_{i}}^{\infty} N\left(x \mid \mu_{i}, \sigma^{2}\right) d x & \text { if } i=n
\end{array} \quad \begin{array}{r} 
\\
N\left(x \mid \mu_{i}, \sigma^{2}\right)=\frac{1}{\sqrt{2 \pi \sigma^{2}}} e^{-\left(\frac{x-\mu_{i}}{\sqrt{2} \sigma}\right)^{2}}, \text { and }
\end{array}\right.
$$

$\mu_{i}$ is the mean growth increment for crab in size-class $i$ :
$\mu_{i}=\omega_{1}+\omega_{2} * \bar{L}_{i}$.
$\omega_{1}, \omega_{2}, \quad$ and $\sigma$ are estimable parameters, $\mathrm{j}_{1}$ and $j_{2}$ are the lower and upper limits of the receiving length-class $j$ (in mm CL), and $\bar{L}_{i}$ is the mid-point of the contributing length interval $i$. The quantity $m_{i}$ is the molt probability for size-class $i$ :
$m_{i}=\frac{1}{1+e^{c\left(\tau_{i}-d\right)}}$
where $\tau_{i}$ is the mid-length of the $i$-th length-class, $c$ and $d$ are parameters.

## Selectivity and retention

Selectivity and retention are both assumed to be logistic functions of length. Selectivity depends on the fishing period for the directed pot fishery:
$S_{i}=\frac{1}{1+e^{\left[-\ln (19) \frac{\tau_{i}-\theta_{50}}{\theta_{95}-\theta_{50}}\right]}}$
where $\theta_{95}$ and $\theta_{50}$ are the parameters of the selectivity/ retention pattern (Mark Maunder, unpublished generic crab model). In our program, we re-parameterized the denominator $\left(\theta_{95}-\theta_{50}\right)$ to $\log ($ delta $\theta)$ so that the difference is always positive and transformed $\theta_{50}$ to $\log \left(\theta_{50}\right)$ to keep the estimate always positive.

## Recruitment

Recruitment to length-class i during year $t$ is modeled as $R_{t, i}=\bar{R} e^{\epsilon_{t}} \Omega_{i}$ where $\Omega_{i}$ is a normalized gamma function
$\operatorname{gamma}\left(x \mid \alpha_{r}, \beta_{r}\right)=\frac{x^{\alpha_{r}-1} e^{\frac{x}{\beta_{r}}}}{\beta_{r}{ }^{\alpha_{r}} \Gamma_{\left(\alpha_{r}\right)}}$
with $\alpha_{r}$ and $\beta_{r}$ (restricted to the first five length classes).

## Parameter estimation

Table A1 lists the parameters of the model indicating which are estimated and which are pre-specified. The objective function includes contributions related to the fit of the model to the available data and penalties (priors on various parameters).

Tables A2 lists parameter values (with the corresponding coefficient of variations in parentheses) used to weight the components of the objective functions for EAG and WAG.

## Likelihood components

## Catch and discard

The contribution of the catch data (retained, total, and groundfish discarded) to the objective function is given by:

$$
\begin{align*}
& L L_{r}^{\text {catch }}=\lambda_{r} \sum_{t}\left\{\ln \left(\sum_{j} \hat{C}_{t, j} w_{j}+c\right)-\ln \left(\sum_{j} C_{t, j} w_{j}+c\right)\right\}^{2}  \tag{A.11a}\\
& L L_{T}^{\text {catch }}=\lambda_{T} \sum_{t}\left\{\ln \left(\sum_{j} \widehat{T}_{t, j} w_{j}+c\right)-\ln \left(\sum_{j} T_{t, j} w_{j}+c\right)\right\}^{2}  \tag{A.11b}\\
& L L_{G D}^{\text {catch }}=\lambda_{G D} \sum_{t}\left\{\ln \left(\sum_{j} \widehat{T r}_{t, j} w_{j}+c\right)-\ln \left(\sum_{j} T r_{t, j} w_{j}+c\right)\right\}^{2} \tag{A.11c}
\end{align*}
$$

where $\lambda_{r}, \lambda_{T}$, and $\lambda_{G D}$ are weights assigned to likelihood components for the retained, pot total, and groundfish discard catch estimates; $w_{j}$ is the average mass of a crab is length-
class $j ; C_{t, j}, T_{t, j}$, and $\operatorname{Tr}_{t, j}$ are, respectively, the observed numbers of crab in size class $j$ for retained, pot total, and groundfish fishery discarded crab during year $t$, and $c$ is a small constant value. A small value of 0.001 is assumed for $c$.

An additional retained catch likelihood (using Equation A.11a without w) for the retained catch in number of crab during 1981/82-1984/85 was also considered in all scenarios.

## Catch-rate indices

The catch-rate indices are assumed to be lognormally distributed about the model prediction. Account is taken of variation in addition to that related to sampling variation:

$$
\begin{equation*}
L L_{r}^{C P U E}=\lambda_{r, C P U E}\left\{0.5 \sum_{t} \ln \left[2 \pi\left(\sigma_{r, t}^{2}+\sigma_{e}^{2}\right)\right]+\sum_{t} \frac{\left(\ln \left(C P U E_{t}^{r}+c\right)-\ln \left(\left(\widehat{\text { PUE }}_{t}^{r}+c\right)\right)^{2}\right.}{2\left(\sigma_{r, t}^{2}+\sigma_{e}^{2}\right)}\right\} \tag{A.12}
\end{equation*}
$$

where $C P U E_{t}^{r}$ is the standardized retain catch-rate index for year $t, \sigma_{r, t}$ is standard error of the logarithm of $C P U E_{t}^{r}$, and $C \widehat{P U E}_{t}^{r}$ is the model-estimate of $C P U E_{t}^{r}$ :

$$
\begin{equation*}
\widehat{C P U E}_{t}^{r}=q_{k} \sum_{j} S_{j}^{T} S_{j}^{r}\left(N_{t, j}-0.5\left[\widehat{C_{t, j}}+\widehat{D_{t, j}}+\widehat{T r_{t, j}}\right]\right) e^{-y_{t} M} \tag{A.13}
\end{equation*}
$$

in which $q_{k}$ is the catchability coefficient during the $k$-th period (e.g., pre-, and postrationalization time periods), $\sigma_{e}$ is the extent of over-dispersion, $c$ is a small constant to prevent zero values (we assumed $c=0.001$ ), and $\lambda_{r, C P U E}$ is the weight assigned to the catch-rate data. We used the same likelihood formula (A.12) for fish ticket and cooperative survey retained catch rate indices. However, for cooperative survey catch rate prediction we used a different catchability parameter.

Following Burnham et al. (1987), we computed the $\ln$ (CPUE) variance by:

$$
\begin{equation*}
\sigma_{r, t}^{2}=\ln \left(1+C V_{r, t}^{2}\right) \tag{A.14}
\end{equation*}
$$

## Length-composition data

The length-composition data are included in the likelihood function using the robust normal for proportions likelihood, i.e., generically:
$L L_{r}^{L F}=0.5 \sum_{t} \sum_{j} \ln \left(2 \pi \sigma_{t, j}^{2}\right)-\sum_{t} \sum_{j} \ln \left[\exp \left(-\frac{\left(P_{t, j}-\hat{P}_{, j,}\right)^{2}}{2 \sigma_{t, j}^{2}}\right)+0.01\right]$
(A.15)
where $P_{t, j}$ is the observed proportion of crab in length-class j in the catch during year t , $\hat{P}_{t, j}$ is the model-estimate corresponding to $P_{t, j}$, i.e.:
$\hat{\mathrm{L}}_{\mathrm{t}, \mathrm{j}}^{\mathrm{r}}=\frac{\widehat{\mathrm{C}}_{\mathrm{t}, \mathrm{j}}}{\Sigma_{\mathrm{j}, \mathrm{C}}^{\mathrm{C}} \mathrm{C}_{\mathrm{j}}}$
$\hat{\mathrm{L}}_{\mathrm{t}, \mathrm{j}}^{\mathrm{T}}=\frac{\hat{\mathrm{T}}_{\mathrm{t}, \mathrm{j}}}{\sum_{\mathrm{j}}^{\mathrm{n}} \mathrm{T}_{\mathrm{t}, \mathrm{j}}}$

$$
\begin{align*}
& \hat{\mathrm{L}}_{\mathrm{t}, \mathrm{j}}^{\mathrm{GF}}=\frac{\widehat{\mathrm{Tr}}_{\mathrm{t}, \mathrm{j}}}{\sum_{\mathrm{j}}^{\mathrm{T}} \mathrm{Tr}_{\mathrm{t}, \mathrm{j}}}  \tag{A.16}\\
& \sigma_{t, j}^{2} \text { is the variance of } P_{t, j}: \\
& \sigma_{t, j}^{2}=\left[\left(1-P_{t, j}\right) P_{t, j}+\frac{0.1}{n}\right] / S_{t} \tag{A.17}
\end{align*}
$$

and $S_{t}$ is the effective sample size for year $t$ and $n$ is the number of size classes.

## Tagging data

Let $V_{j, t, y}$ be the number of tagged male crab that were released during year $t$ that were in size-class $j$ when they were released and were recaptured after $y$ years, and $\rho_{j, t, y}$ be the vector of recaptures by size-class from the males that were released in year $t$ that were in size-class $j$ when they were released and were recaptured after $y$ years. The log-likelihood corresponding to the multinomial distribution for the tagging data is then:
$\ln L=\lambda_{y, t a g} \sum_{j} \sum_{t} \sum_{y} \sum_{i} \rho_{j, t, y, i} \ln \hat{\rho}_{j, t, y, i}$
where $\lambda_{y, \operatorname{tag}}$ is the weight assigned to the tagging data for recapture year $y, \hat{\rho}_{j, t, y, i}$ is the proportion in size-class $i$ of the recaptures of males that were released during year $t$ that were in size-class $j$ when they were released and were recaptured after $y$ years:
$\underline{\hat{\rho}}_{j, t, y} \propto \underline{s}^{T}[\mathbf{X}]^{y} \underline{Z}^{(j)}$
where $Z^{(j)}$ is a vector with $V_{j, t, y}$ at element $j$ and 0 otherwise, and $S^{T}$ is the vector of total selectivity for tagged male crab by the pot fishery. This log-likelihood function is predicated on the assumption that all recaptures are in the pot fishery and the reporting rate is independent of the size of crab.

## Penalties

Penalties are imposed on the deviations of annual pot fishing mortality about mean pot fishing mortality, annual trawl fishing mortality about mean trawl fishing mortality, recruitment about mean recruitment, and the posfunction (fpen):

$$
\begin{align*}
& P_{1}=\lambda_{F} \sum_{t}\left(\ell \mathrm{n} F_{t}-\ell \mathrm{n} \bar{F}\right)^{2}  \tag{A.20}\\
& P_{2}=\lambda_{F^{T r}} \sum_{t}\left(\ell \mathrm{n} F_{t}^{T r}-\ell \mathrm{n} \bar{F}^{T r}\right)^{2}  \tag{A.21}\\
& P_{3}=\lambda_{R} \sum_{t}\left(\ell \ln \varepsilon_{t}\right)^{2}  \tag{A.22}\\
& P_{5}=\lambda_{\text {posfn }} * \text { fpen } \tag{A.23}
\end{align*}
$$

$$
\begin{equation*}
\text { Std. } \text { Res }_{t, j}=\frac{P_{t, j}-\widehat{P_{t, j}}}{\sqrt{2 \sigma_{t, j}^{2}}} \tag{A.24}
\end{equation*}
$$

## Output Quantities

Harvest rate
Total pot fishery harvest rate: $E_{t}=\frac{\sum_{j=1}^{n}\left(\hat{c}_{j, t}+\widehat{D}_{j, t}\right)}{\sum_{j=1}^{n} N_{j, t}}$
Exploited legal male biomass at the start of year $t$ :
$L M B_{t}=\sum_{j=\text { legal size }}^{n} s_{j}^{T} s_{j}^{r} N_{j, t} w_{j}$
where $w_{j}$ is the weight of an animal in length-class $j$.
Mature male biomass on 15 February spawning time (NPFMC 2007a, b) in the following year:
$M M B_{t}=\sum_{j=\text { mature size }}^{n}\left\{N_{j, t} e^{-y^{\prime} M}-\left(\hat{C}_{j, t}+\widehat{D}_{j, t}+\widehat{\operatorname{Tr}}_{j, t}\right) e^{\left(y_{t}-y^{\prime}\right) M}\right\} w_{j}$
where $\mathrm{y}^{\prime}$ is the elapsed time from 1 July to 15 February in the following year.
For estimating the next year limit harvest levels from current year stock abundances, an $F_{O F L}$ value is needed. The current crab management plan specifies five different Tier formulas for different stocks depending on the strength of information available for a stock, for computing $F_{O F L}$ (NPFMC 2007a, b). For the golden king crab, the following Tier 3 formula is applied to compute $F_{O F L}$ :

If,
$M M B_{\text {current }}>M M B_{35 \%}, F_{\text {OFL }}=F_{35 \%}$
If,
$M M B_{\text {current }} \leq M M B_{35 \%}$ and $M M B_{\text {current }}>\beta M M B_{35 \%}$,
$F_{O F L}=F_{35 \%} \frac{\left(\frac{M M B_{\text {current }}}{M M B_{35 \%}}-\alpha\right)}{(1-\alpha)}$
If,
$M M B_{\text {current }} \leq \beta M M B_{35 \%}$,
$F_{O F L}=0$.
where
$\beta=$ a parameter with a restriction that $0 \leq \beta<1$. A default value of 0.25 is used, $\alpha=$ a parameter with a restriction that $0 \leq \alpha \leq \beta$. A default value of 0.1 is used,
$\mathrm{MMB}_{\text {current }}=$ the mature male biomass in the current year, and
$M M B_{35 \%}=$ a proxy $M M B_{M S Y}$ for Tier 3 stocks.
Because projected $\mathrm{MMB}_{\mathrm{t}}$ (i.e., $\mathrm{MMB}_{\text {current }}$ ) depends on the intervening retained and discard catch (i.e., $\mathrm{MMB}_{\mathrm{t}}$ is estimated after the fishery), an iterative procedure is applied using Equations A. 27 and A. 28 with retained and discard catch predicted from Equations A.2b-d. The next year limit harvest catch is estimated using Equations A.2b-d with the estimated $F_{O F L}$ value.

Table A1. Pre-specified and estimated parameters of the population dynamics model

| Parameter | Number of parameters |
| :---: | :---: |
| Fishing mortalities: |  |
| Pot fishery, $F_{t}$ | 1981-2022 (estimated) |
| Mean pot fishery fishing mortality, $\bar{F}$ | 1 (estimated) |
| Groundfish fishery, $F_{t}{ }^{T r}$ | 1989-2022 (the mean F for 1989 to 1994 was used to estimate groundfish discards back to 1981 (estimated) |
| Mean groundfish fishery fishing mortality, $\bar{F}^{T r}$ | 1 (estimated) |
| Selectivity and retention: |  |
| Pot fishery total selectivity, $\theta_{50}^{\mathrm{T}}$ | 2 (1981-2004; 2005+) (estimated) |
| Pot fishery total selectivity difference, delta $\theta^{\text {T }}$ | 2 (1981-2004; 2005+) (estimated) |
| Pot fishery retention, $\theta_{50}^{\mathrm{r}}$ | 1 (1981+) (estimated) |
| Pot fishery retention selectivity difference, delta $\theta^{\text {r }}$ | 1 (1981+) (estimated) |
| Groundfish fishery selectivity | fixed at 1 for all size-classes |
| Growth: |  |
| Expected growth increment, $\omega_{1}, \omega_{2}$ | 2 (estimated) |
| Variability in growth increment, $\sigma$ | 1 (estimated) |
| Molt probability (size transition matrix with tag data), a | 1 (estimated) |
| Molt probability (size transition matrix with tag data), b | 1 (estimated) |
| Natural mortality, M | 1 (pre-specified, $0.22 \mathrm{yr}^{-1}$ ) |
| Recruitment: |  |
| Number of recruiting length-classes | 5 (pre-specified) |
| Mean recruit length | 1 (pre-specified, 110 mm CL ) |
| Distribution to length-class, $\beta_{\mathrm{r}}$ | 1 (estimated) |
| Median recruitment, $\overline{\mathrm{R}}$ | 1 (estimated) |
| Recruitment deviations, $\boldsymbol{E}_{t}$ |  |
| Fishery catchability, q | 2 (1985-2004; 2005+) (estimated) |
| Additional CPUE indices standard deviation, $\sigma_{\mathrm{e}}$ | 1 (estimated) |
| Likelihood weights (coefficient of variation) | Pre-specified, varies by scenario |

Table A2. Specifications for the weights with corresponding coefficient of variations* in parentheses for each model for EAG and WAG.

| Weight | Models $22.9 \mathrm{c}, 22.1 \mathrm{e} 2,22.1 \mathrm{f}, 22.1 \mathrm{~g} \text {, and } 22.1 \mathrm{~h}$ |
| :---: | :---: |
| Catch: |  |
| Retained catch for 1981-1984 and/or 1985-2021, $\lambda_{r}$ | 500 (0.0316) |
| Total catch for 1990-2021, $\lambda_{T}$ | Number of sampled pots scaled to a max 250 |
| Groundfish bycatch for 1989-2021, $\lambda_{G D}$ | 0.5 (1.3108) |
| Catch-rate: <br> Observer legal size crab catch-rate for 1995-2021, $\begin{equation*} \lambda_{r, C P U E} \tag{0.8054} \end{equation*}$ |  |
| Fish ticket retained crab catch-rate for 1985-1998, $\lambda_{r, \text { CPUE }}$ | 1 (0.8054) |
| Penalty weights: |  |
| Pot fishing mortality dev, $\lambda_{F}$ Groundfish fishing mortality dev, $\lambda_{F^{T r}}$ | Initially 1000 , relaxed to 0.001 at phases $\geq$ select. phase <br> Initially 1000 , relaxed to 0.001 at phases $\geq$ select. phase |
| Recruitment, $\lambda_{R}$ <br> Posfunction (to keep abundance estimates always positive), $\lambda_{\text {posfn }}$ | $\begin{aligned} & 2(0.5329) \\ & 1000(0.0224) \end{aligned}$ |
| Tagging likelihood | EAG individual tag returns |

* Coefficient of Variation, $\mathrm{CV}=\sqrt{\exp \left[\frac{1}{2 W}\right]-1}, \mathrm{w}=$ weight


## Appendix B: Catch and CPUE data

The commercial catch and length frequency distribution were estimated from ADF\&G landing records and dockside sampling (Bowers et al. 2008, 2011). The annual retained catch, total catch, and groundfish (or trawl) discarded mortality are provided in Tables 1, 2, and 2 b for EAG and WAG. The weighted length frequency data were used to distribute the catch into 5mm size intervals. The length frequency data for a year were weighted by each sampled vessel's catch as follows. The $i$-th length-class frequency was estimated as:

$$
\begin{equation*}
\sum_{j=1}^{k} C_{j} \frac{L F_{j, i}}{\sum_{i=1}^{n} L F_{j, i}} \tag{B.1}
\end{equation*}
$$

where $k=$ number of sampled vessels in a year, $L F_{j, i}=$ number of crab in the $i$-th length-class in the sample from $j$-th vessel, $n=$ number of size classes, $C_{j}=$ number of crab caught by $j$-th vessel. Then the relative frequency for the year was calculated and applied to the annual retained catch (in number of crab) to obtain retained catch by length-class.

The annual total catch (in number of crab) was estimated by the observer nominal (unstandardized) total CPUE considering all vessels multiplied by the total fishing effort (number of pot lifts). The weighted length frequency of the observer samples across the fleet was estimated using Equation B.1. Observer measurement of crab ranged from 20 to 220 mm CL . The proportion of observer total relative length frequency corresponding to this size range was multiplied by the total catch (number of crab) to get the crab number by size. Thus, the season total catch was distributed into length-classes using the weighted relative length frequency of B.1. To restrict the number of crab to model assumed size range (101-185+ mm CL), crab sizes < 101 mm CL were pooled into 101 length class and all crab >185 mm CL were pooled into a $185+$ length class. Note that the total crab catch by size that went into the model did not consider retained and discard components separately. However, once the model estimated the annual total catch, then retained catch was deducted from this total and multiplied by handling mortality [we used a $20 \%$ handling mortality (Siddeek et al. 2005) to obtain the directed fishery discarded (dead) catch].

Observer data have been collected since 1988 (Moore et al. 2000; Barnard et al. 2001; Barnard and Burt 2004; Gaeuman 2011), but data were not comprehensive in the initial years, so a shorter time series of data for the period 1990/91-2022/23 was selected for this analysis. During 1990/91-1994/95, observers were only deployed on catcher-processor vessels. During 1995/96-2004/05, observers were deployed on all fishing vessels during fishing activity. Observers have been deployed on all fishing vessels since 2005/06, but catcher-only vessels are only required to carry observers for a minimum of $50 \%$ of their fishing activity during a season; catcher-processor vessels are still required to carry observers during all fishing activity. Onboard observers sample seven pots per day (may be different numbers of pots per string) and count and measure all crab caught and categorize catch as females, sublegal males, retained legal males, and non-retained legal males in a sampled pot. Prior to the 2009/10 season, depending on season, area, and type of fishing vessel, observers were also instructed to sample additional pots in which all crab were only counted and categorized as females, sublegal males, retained legal males, and non-retained legal males, but were not measured. Annual mean nominal CPUEs of retained and total crab were estimated considering all sampled pots within
each season (Table 3). The observer CPUE data collection improved over the years and the data since 1995/96 are more reliable. Thus, for model fitting, the observer CPUE time series was restricted to 1995/96-2022/23. The 1990/91-2022/23 observer database consists of 121,003 records and that of 1995/96-2022/23 contains 116,724 records, For CPUE standardization, these data were further reduced by $5 \%$ cutoff of Soak time and $1 \%$ cutoff of Depth on both ends of the variable range to remove unreliable data or data from dysfunctional pot operations and restricting to vessels which have made five trips per year for at least three years during 1985/86-2022/23.

Length-specific CPUE data collected by observers provides information on a wider size range of the stock than did the commercial catch length frequency data obtained from mostly legalsized landed males.

There were significant changes in fishing practice due to changes in management regulations (e.g., since 1996/97 constant TAC and since 2005/06 crab rationalization), pot configuration (escape web on the pot door increased to 9 " since 1999), and improved observer recording in Aleutian Islands golden king crab fisheries since 1998. These changes prompted us to consider two separate observer CPUE time series, 1995/96-2004/05 and 2005/06-2022/23, to estimate CPUE indices for model input.

To include a long time series of CPUE indices for stock abundance contrast, we also considered the 1985/86-1998/99 legal size standardized CPUE as a separate likelihood component in all scenarios. Because of the lack of soak time data before 1990, we estimated the CPUE index considering a limited set of explanatory variables (e.g., vessel, captain, area, month) and fitting the negative binomial GLM model to fish ticket data (Tables 4 and 10).

When using CPUE indices in the model fit, we compared the predicted with the observed legal male CPUE in the observer CPUE likelihoods because legal male (retained plus non-retained) data are more reliable than total in the observer samples.

The CPUE standardization followed the GLM fitting procedure (Maunder and Punt 2004; Starr 2012; Siddeek et al. 2018). Following a suggestion from the CIE reviewers in June 2018 we reduced the number of gear codes in the database after consulting with the fishing industry (Rip Carlton, Chad Hoefer, and Scott Goodman, personal communication December 2018; Table B1). Following an SSC suggestion in October 2018, we used a hybrid procedure: First, we selected a scope of variables set by Akike Information Criterion, AIC (Burnham and Anderson 2002). A decrease of more than 2 units in the AIC was used to identify the variable to be included successively (stepAIC program, R Core Team 2021). Then, the model parsimony was improved further by successively removing the term that explained the least proportion of deviance ( $\mathrm{R}^{2}<0.01$ ) (stepCPUE R function was used, Siddeek et al. 2018). Feenstra, et al. (2019) used a similar hybrid approach.

Table B.1. Updated gear codes for observer data analysis. Only gear code \# 5, 6, 7, 8, and 13 were considered following crab industry suggestion. Note: Identical codes were given to those gear codes with similar catchability/selectivity. X indicates gear codes that were ignored.

| Original Gear code | Pot gear description | Mark X against the code that can be ignored | Number encountered by observers during 1990-2016 | Updated gear code |
| :---: | :---: | :---: | :---: | :---: |
| 1 | Dungeness crab pot, small \& round | X | 2 | X |
| 2 | Pyramid pot, tunnel openings usually on sides, stackable | X | 2121 | X |
| 3 | Conical pot, opening at top of cone, stackable | X | 2000 | X |
| 4 | 4' X 4' rectangular pot |  | 60 | X |
| 5 | $5^{\prime} \mathrm{X} 5{ }^{\prime}$ rectangular pot |  | 18032 | 5 |
| 6 | 6' ${ }^{\prime}$ 6' rectangular pot |  | 17508 | 6 |
| 7 | 7' X 7' rectangular pot |  | 23806 | 7 |
| 8 | 8' X 8' rectangular pot |  | 1936 | 8 |
| 9 | $51 / 2^{\prime} \mathrm{X} 51 / 2^{\prime}$ rectangular pot |  | 6934 | 5 |
| 10 | $61 / 2^{\prime} \mathrm{X} 61 / 2^{\prime}$ rectangular pot |  | 22085 | 6 |
| 11 | 7 1/2' X 7 1/2' rectangular pot |  | 387 | 7 |
| 12 | Round king crab pot, enlarged version of Dungeness crab pot |  | 8259 | X |
| 13 | $10^{\prime} \mathrm{X} 10^{\prime}$ rectangular pot |  | 466 | 13 |
| 14 | 9' X 9' rectangular pot | X | 1 | X |
| 15 | $81 / 2^{\prime} \mathrm{X} 81 / 2^{\prime}$ rectangular pot | X | 1 | X |
| 16 | $91 / 2^{\prime}$ X $91 / 2^{\prime}$ rectangular pot | X | Not used | X |
| 17 | 8' X 9' rectangular pot | X | 1 | X |
| 18 | 8' X 10' rectangular pot | X | 1 | X |
| 19 | 9' X 10' rectangular pot |  | Not used | X |
| 20 | 7' X 8' rectangular pot | X | 252 | X |
| 21 | Hair crab pot, longlined and small, stackable |  | Not used | X |
| 22 | snail pot | X | 1 | X |
| 23 | Dome-shaped pot, tunnel opening on top, often longlined in deep-water fisheries | X | 6756 | X |
| 24 | ADF\&G shellfish research 7' X $7^{\prime} \mathrm{X} 34^{\prime \prime}$ rectangular pot with 2.75 " stretch mesh and no escapement rings or mesh |  | Research pot | X |
| 80 | Historical: Cod pot, any shape pot targeting cod, usually with tunnel fingers | X | 711 | X |
| 81 | Historical: Rectangular pot, unknown size, with escape rings | X | 1123 | X |

All scenarios used CPUE indices estimated by the hybrid GLM method. Following a January 2019 CPT request, we considered a Year:Area interaction factor as a special case for a CPUE standardization scenario.

Thus we estimated two sets of observer CPUE indices for model input, 21.1a (reduced number of gear codes), and 21.1f (reduced number of gear codes and Year:Area interaction).

## Observer CPUE index by GLM

## a. Non-interaction GLM model

The CPUE standardization followed the GLM fitting procedure (Maunder and Punt 2004; Starr 2012; Siddeek et al. 2016b). We considered the negative binomial GLM on positive and zero catch to select the explanatory variables. The response variable CPUE is the observer sample catch record for a pot haul. The negative binomial model uses the log link function for the GLM fit.

For the non-interaction model, we assumed the null model to be:

$$
\begin{equation*}
\ln \left(\text { CPUE }_{i}\right)=\text { Year }_{y_{i}} \tag{B.2}
\end{equation*}
$$

where Year is a factorial variable.
The maximum set of model terms offered to the stepwise selection procedure was:
$\ln \left(\right.$ CPUE $\left._{\mathrm{I}}\right)=$ Year $_{\mathrm{y}_{\mathrm{i}}}+\mathrm{ns}\left(\right.$ Soak $_{\mathrm{si}}$, df $)+$ Month $_{\mathrm{m}_{\mathrm{i}}}+$ Vessel $_{\mathrm{vi}}+$ Captain $_{\mathrm{ci}}+$ Block $_{\mathrm{ai}}+$ Gear $_{\mathrm{gi}}+\mathrm{ns}\left(\right.$ Depth $\left._{\text {di }}, \mathrm{df}\right)$,
where Soak is in unit of days and is numeric; Month, Area (Block) code, Vessel code, Captain code, and Gear code are factorial variables; Depth in fathom is a numeric variable; ns=cubic spline, and $\mathrm{df}=$ degree of freedom.

We used a log link function and a dispersion parameter $(\theta)$ in the GLM fitting process. We used the $\mathrm{R}^{2}$ criterion for predictor variable selection (Siddeek et al. 2016b).

We calculated appropriate degrees of freedom and dispersion parameters by calculating AICs for a range of values and locating the best values at the minimum AIC (see Siddeek et al, 2021 SAFE report). We further reduced the spline number of degrees of freedom based on significant model fits.

Instead of using the traditional AIC ( $-2 \log$ likelihood +2 p ) we used the Consistent Akaike Information Criteria (CAIC) (Bozdogan 1987) $\left\{-2 \log _{1}\right.$ likelihood+[ln(n)+1] *p\} for variable selection by StepAIC, where $n=$ number of observations and $p=$ number of parameters to be estimated. The number of selected variables were further reduced for parsimony, if feasible, by the $\mathrm{R}^{2}$ criterion using the StepCPUE function. i.e., a hybrid selection procedure (Feenstra et al. 2019).

The final main effect models for EAG were:

Initial selection by stepAIC:
$\ln ($ CPUE $)=$ Year + Gear + Captain + ns $($ Soak, 4$)+$ Month
AIC=203,808
Final selection by stepCPUE:
$\ln ($ CPUE $)=$ Year + Captain + ns (Soak, 4$)+$ Month
for the 1995/96-2004/05 period $\left[\theta=1.38, \mathrm{R}^{2}=0.2205\right]$
Initial selection by stepAIC:
$\ln ($ CPUE $)=$ Year + Captain + Gear $+\mathrm{ns}($ Soak, 10 $)+$ Month
AIC=81,580
Final selection by stepCPUE:
$\ln ($ CPUE $)=$ Year + Captain + ns $($ Soak, 10$)+$ Gear
for the 2005/06-2022/23 period [ $\left.\theta=2.34, \mathrm{R}^{2}=0.1103\right]$.

The final models for WAG were:
Initial selection by stepAIC:
$\ln ($ CPUE $)=$ Year + Captain $+\mathrm{ns}($ Soak, 7$)+$ Gear + Block + Month + Vessel
AIC=191,025
Final selection by stepCPUE:
$\ln ($ CPUE $)=$ Year + Captain + ns $($ Soak, 7$)+$ Gear
for the 1995/96-2004/05 period $\left[\theta=0.97, \mathrm{R}^{2}=0.1681\right]$
Initial selection by stepAIC:
$\ln$ (CPUE) $=$ Year + Captain + Gear + Month + ns $($ Soak, 3$)$
AIC=130,731
Final selection by stepCPUE

$$
\begin{equation*}
\ln (\text { CPUE })=\text { Year }+ \text { Gear }+\mathrm{ns}(\text { Soak, } 2) \tag{B.7}
\end{equation*}
$$

for the 2005/06-2022/23 period [ $\theta=1.11, R^{2}=0.0749$, Soak forced in].

## b. Year:Area interaction GLM:

Justification: The interaction CPUE model was necessary because of non-uniform distribution of golden king crab (see the distribution section) in the Aleutian Islands as well as small number of vessels operating in the two fisheries (a maximum of three in each region in recent years).

All scenarios used CPUE indices estimated by the hybrid GLM method. Following a January 2019 CPT request, we considered a Year:Block (aka Year:Area) interaction factor as a special case for a CPUE standardization scenario.

For year and area interaction analysis, we divided the areas into 1 nmix 1 nmi grids enmeshed in 10 larger blocks as follows. The number of blocks was restricted to a few to prevent GLM fitting problems (Figure B. 1 and Table B.2).


Figure B.1. The 1995/96-2022/23 observer pot samples enmeshed in 10 blocks for the Aleutian Islands golden king crab.

The blocks were determined from visually exploring each year's pot distribution locations (available with the first author). The blocks contain observed patches of crab distribution during this period

Table B.2. Number of $1 \mathrm{nmi} x 1 \mathrm{nmi}$ grids containing observer sample locations within each block by fishing year for the Aleutian Islands golden king crab, 1995/96-2022/23 data. Blocks 1-4 belong to EAG and 5-10 to WAG. Sum of ever fished number of grids for each block is listed at the bottom row.

| FMP Season | Block_1 | Block_2 | Block_3 | Block_4 | Block_5 | Block_6 | Block_7 | Block_8 | Block_9 | Block_10 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1995 | 125 | 529 | 748 | 379 | 218 | 373 | 112 | 722 | 166 | 122 |
| 1996 | 149 | 814 | 761 | 372 | 89 | 473 | 359 | 799 | 200 | 35 |
| 1997 | 116 | 530 | 755 | 257 | 202 | 443 | 104 | 568 | 274 | 0 |
| 1998 | 78 | 581 | 453 | 236 | 18 | 318 | 157 | 251 | 132 | 0 |
| 1999 | 123 | 593 | 454 | 231 | 163 | 476 | 182 | 627 | 193 | 145 |
| 2000 | 72 | 540 | 754 | 301 | 187 | 440 | 195 | 555 | 547 | 47 |
| 2001 | 123 | 507 | 507 | 329 | 45 | 369 | 288 | 634 | 256 | 9 |
| 2002 | 97 | 387 | 584 | 271 | 71 | 341 | 205 | 335 | 242 | 37 |
| 2003 | 43 | 492 | 530 | 299 | 111 | 347 | 212 | 465 | 150 | 61 |
| 2004 | 81 | 289 | 377 | 216 | 77 | 319 | 150 | 359 | 172 | 116 |
| 2005 | 0 | 205 | 221 | 118 | 8 | 220 | 83 | 261 | 54 | 0 |
| 2006 | 0 | 154 | 248 | 122 | 15 | 191 | 58 | 220 | 39 | 0 |
| 2007 | 0 | 111 | 177 | 110 | 24 | 228 | 78 | 173 | 20 | 0 |
| 2008 | 0 | 111 | 203 | 93 | 12 | 181 | 67 | 196 | 0 | 0 |
| 2009 | 0 | 59 | 146 | 60 | 6 | 137 | 95 | 220 | 25 | 0 |
| 2010 | 0 | 81 | 141 | 85 | 1 | 115 | 73 | 260 | 39 | 0 |
| 2011 | 0 | 126 | 117 | 33 | 3 | 83 | 73 | 266 | 9 | 0 |
| 2012 | 0 | 146 | 110 | 56 | 7 | 91 | 85 | 312 | 53 | 0 |
| 2013 | 2 | 149 | 129 | 51 | 12 | 144 | 105 | 293 | 86 | 0 |
| 2014 | 1 | 138 | 96 | 41 | 39 | 120 | 114 | 319 | 37 | 0 |
| 2015 | 0 | 135 | 147 | 61 | 46 | 163 | 106 | 280 | 16 | 48 |
| 2016 | 0 | 145 | 231 | 63 | 26 | 134 | 89 | 210 | 106 | 0 |
| 2017 | 0 | 97 | 170 | 110 | 11 | 87 | 79 | 198 | 118 | 0 |


| 2018 | 0 | 91 | 158 | 95 | 7 | 69 | 82 | 204 | 121 |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 2019 | 1 | 112 | 171 | 101 | 0 | 0 | 89 | 316 | 138 |  |
| 2020 | 4 | 109 | 193 | 95 | 0 | 0 | 76 | 287 | 91 |  |
| 2021 | 0 | 83 | 156 | 113 | 0 | 0 | 66 | 289 | 14 |  |
| 2022 | 0 | 36 | 81 | 27 | 0 | 0 | 89 | 391 | 37 |  |
| Ever Fished: |  |  |  |  |  |  |  |  |  |  |
| AIGKC All Seasons | Block_1 | Block_2 | Block_3 | Block_4 | Block_5 | Block_6 | Block_7 | Block_8 | Block_9 | Block_10 |
| $1995-2022-$ Sum of 1x1 cells | 375 | 1364 | 1765 | 915 | 452 | 1026 | 812 | 2172 | 1042 | 334 |

In the following, we provide the results from up to 2022/23 observer data analysis:
We assumed the null model to be

$$
\begin{equation*}
\ln \left(\text { CPUE }_{\mathrm{i}}\right)=\text { Year }_{\mathrm{y}_{\mathrm{i}}}: \text { Block }_{\mathrm{ai}} \tag{B.8}
\end{equation*}
$$

The maximum set of model terms offered to the stepwise selection procedure was:
$\ln \left(\right.$ CPUE $\left._{\mathrm{I}}\right)=$ Year $_{\mathrm{y}_{\mathrm{i}}}:$ Block $_{a i}+\mathrm{ns}\left(\right.$ Soak $\left._{\mathrm{si}}, \mathrm{df}\right)+$ Month $_{\mathrm{m}_{\mathrm{i}}}+$ Vessel $_{\mathrm{vi}}+$ Captain $_{\text {ci }}+$ Gear $_{\mathrm{gi}}+\mathrm{ns}\left(\right.$ Depth $_{\text {di }}$, df).

The final interaction effect models for EAG were:
Initial selection by stepAIC:
$\ln ($ CPUE $)=$ Gear + Captain $+\mathrm{ns}($ Soak, 4$)+$ Month + Year $:$ Block
AIC=203,851
Final selection by stepCPUE:
$\ln ($ CPUE $)=$ Gear + Captain + ns $($ Soak, 4$)+$ Year: Block
for the 1995/96-2004/05 period [ $\theta=1.38, \mathrm{R}^{2}=0.2235$ ]
Initial selection by stepAIC:
$\ln ($ CPUE $)=$ Vessel + Gear $+\mathrm{ns}($ Soak, 10 $)+$ Month + Year: Block AIC=81,772

Final selection by stepCPUE:
$\ln ($ CPUE $)=$ Vessel + ns $($ Soak, 10$)+$ Gear + Year: Block
for the 2005/06-2022/23 period $\left[\theta=2.34, \mathrm{R}^{2}=0.1201\right]$.

The final interaction effect models for WAG were:
Initial selection by stepAIC:
$\ln ($ CPUE $)=$ Vessel + ns $($ Soak, 7$)+$ Gear + Month + Year: Block
AIC=191,060
Final selection by stepCPUE:
$\ln ($ CPUE $)=$ Vessel + ns(Soak, 7$)+$ Gear + Year: Block
for the 1995/96-2004/05 period [ $\theta=0.97, \mathrm{R}^{2}=0.1719$ ]
Initial selection by stepAIC:
$\ln ($ CPUE $)=$ Gear + Month + Vessel $+\mathrm{ns}($ Soak, 3$)+$ Year: Block AIC=131,060

Final selection by stepCPUE:
$\ln$ (CPUE) $=$ Gear + Month + Year: Block $+n s($ Soak, 3$)$
for the 2005/06-2022/23 period [ $\theta=1.11, R^{2}=0.0897$, Soak forced in].
The diagnostic plots (studentized residual plots) for EAG and WAG observer CPUE indices are provided below:

Negative Binomial Fit, EAG 1995/96-2004/05


Negative Binomial Fit, EAG 2005/06-2022/23


Figure B.2. Studentized residual plots for Year: Block interaction CPUE fits (top: 1995/96-2004/05 and bottom 2005/06-2022/23) for EAG golden king crab data.

Negative Binomial Fit, WAG 1995/96-2004/05


Figure B.3. Studentized residual plots for Year: Block interaction CPUE fits (top: 1995/96-2004/05 and bottom 2005/06-2022/23) for WAG golden king crab data.

Steps:

1. Block-scale analysis:

The bias corrected estimate of CPUE index for each Year-Block (Block = Area) interaction was first obtained as:

$$
\begin{equation*}
C P U E_{i j}=e^{Y B_{i j}+\sigma_{i j}^{2} / 2} \tag{B.14}
\end{equation*}
$$

where $C P U E_{i j}$ is the CPUE index in the ith year and jth block, $Y B_{i j}$ is the coefficient of the $i$ th year and $j$ th block interaction, and $\sigma_{i j}$ is the biased correction standard error for expected CPUE value.

The number of $1 \mathrm{nmi} \times 1 \mathrm{nmi}$ grids in each block can change from year to year; so, we considered using the number of grids ever fished in a block, $N_{\text {ever } j}$ [this is equivalent to assuming that the grids fished in any year randomly sample the stock in that block (Campbell, 2004)].

The abundance index for $j$ th block in $i$ th year is
$B_{i j}=N_{\text {ever }}^{j}$ CPUE $i j$
Notice in Table B. 2 that no or very few observer samplings occurred in certain years for a whole block. We filled the $B_{i j}$ index gaps resulting from Year:Block CPUE standardization model fit as follows:
$\widehat{B_{l, j}}=e^{A_{i}+C_{j}}$
fitted by GLM [i.e., fitting a log-linear model, $\ln \left(\widehat{B}_{i, j}\right)=A_{i}+C_{j}$ ],
where $B_{i, j}$ is the available index of biomass for year i and block $j, A_{i}$ is a year factor, and $C_{j}$ is a block factor, and used this model to predict the unavailable biomass index for blocks $x$ years with no (or very limited, <10) data.

An example set of R codes used to predict the missing biomass index is as follows:

## library (MASS)

To fit the log-linear model of Equation B.16:

```
glm.fit<- glm(log(B}\mp@subsup{\mathbf{B}}{\textrm{ij}}{})~\mp@subsup{\mathrm{ Yeari}}{i}{}+\mp@subsup{\mathrm{ Blockj}}{\mathbf{j}}{},\mathrm{ data=Bindex 
```

where the data frame "Bindex" contains available $\mathrm{B}_{\mathrm{ij}}$, Year $_{\mathrm{i}}$, and Block $\mathrm{k}_{\mathrm{j}}$ column values.

To predict the missing biomass index Y :

## Y<- predict.glm (glm.fit, BindexFillpredict, se.fit=TRUE)

where the new data frame "BindexFillpredict" contains Year $_{i}$ and Block $_{j}$ column values for which $\mathrm{B}_{\mathrm{ij}}$ indices are needed and contains an empty $\mathrm{B}_{\mathrm{ij}}$ column for fill in.

By setting se.fit=TRUE, the standard errors, $\sigma_{i j}$, of predictions are also estimated.
Bias correction was made to each predicted biomass index by $B_{i, j}=e^{\hat{Y}_{i, j}+\sigma_{i j}^{2} / 2}$
where $\sigma_{i j}$ is the standard error of predicted $Y_{i, j}$ value, which is on the scale of the linear predictor (i.e., $\log$ transformed $\mathrm{B}_{\mathrm{ij}}$ ). The standard error for each year and block combination is estimated as follows.

If we denote the covariance matrix of the fitted "glm.fit" as $\Sigma$ and write the coefficients for linear combination of a set of predictors in a vector form as $C$, then the standard error of prediction for that combination is $\sqrt{\mathrm{C}^{\prime} \Sigma \mathrm{C}}$, where $C^{\prime}$ is the transpose of vector $C$.

Annual biomass index, $B_{i}$, was estimated as, $B_{i}=\sum_{j} B_{i j}$

The variance of the total biomass index was computed as:
$\operatorname{Var}\left(B_{i}\right)=\sum_{j} N_{\text {ever }, j}{ }^{2} \operatorname{var}\left(C P U E_{i, j}\right)$
where $\boldsymbol{N}_{\text {ever }, j}$ is the total number of 1 mni x 1 mni cells ever fished in block $j$, and $C P U E_{i, j}$ is the CPUE index for year $i$ and block $j$.

To use in the assessment model, 22.1f, we rescaled the $B_{i}$ indices by the geometric mean of estimated $B_{i}$ values (Equation B.16) separately for the pre- and postrationalization periods. The corresponding standard error ( $\sim \mathrm{CV}$ ) of $B_{i}$ was estimated by

$$
\begin{equation*}
\sqrt{\frac{\operatorname{Var}\left(B_{i}\right)}{\left(B_{i}\right)^{2}}} \tag{B.19}
\end{equation*}
$$

The rescaled biomass indices with standard errors are listed in Tables B. 3 and B. 4 for EAG and WAG, respectively.

Table B.3. Steps to estimate biomass-based abundance indices with standard errors for 1995/96-2022/23 in EAG. GMScaled B_index and B_Index SE were used as CPUE index and its standard error.

| Year | GMScaled B_Index | B_Index SE |
| ---: | ---: | ---: |
| 1995 | 0.771 | 0.107 |
| 1996 | 0.779 | 0.106 |
| 1997 | 0.776 | 0.099 |
| 1998 | 0.931 | 0.085 |
| 1999 | 0.886 | 0.087 |
| 2000 | 0.859 | 0.120 |
| 2001 | 1.241 | 0.151 |
| 2002 | 1.260 | 0.068 |


| 2003 | 1.128 | 0.076 |
| :--- | ---: | :--- |
| 2004 | 1.714 | 0.093 |
| 2005 | 1.046 | 0.074 |
| 2006 | 0.767 | 0.100 |
| 2007 | 0.858 | 0.088 |
| 2008 | 0.866 | 0.090 |
| 2009 | 0.759 | 0.107 |
| 2010 | 0.774 | 0.104 |
| 2011 | 1.056 | 0.077 |
| 2012 | 0.998 | 0.080 |
| 2013 | 1.042 | 0.080 |
| 2014 | 1.299 | 0.062 |
| 2015 | 1.236 | 0.063 |
| 2016 | 1.057 | 0.073 |
| 2017 | 0.949 | 0.083 |
| 2018 | 1.179 | 0.066 |
| 2019 | 1.101 | 0.091 |
| 2020 | 1.008 | 0.079 |
| 2021 | 0.957 | 0.082 |
| 2022 | 1.289 | 0.069 |

Table B.4. Steps to estimate biomass-based abundance indices with standard errors for 1995/96-2022/23 in EAG. GMScaled B_index and B_Index SE were used as CPUE index and its standard error.

| Year | GMScaled B_Index | B_Index SE |
| ---: | ---: | ---: |
| 1995 | 1.134 | 0.079 |
| 1996 | 1.000 | 0.066 |
| 1997 | 1.028 | 0.066 |
| 1998 | 1.055 | 0.075 |
| 1999 | 0.929 | 0.070 |
| 2000 | 0.879 | 0.075 |
| 2001 | 0.801 | 0.121 |
| 2002 | 0.927 | 0.074 |
| 2003 | 1.125 | 0.058 |
| 2004 | 1.192 | 0.058 |
| 2005 | 1.099 | 0.035 |
| 2006 | 1.200 | 0.032 |
| 2007 | 1.117 | 0.036 |
| 2008 | 1.267 | 0.034 |
| 2009 | 1.942 | 0.021 |
| 2010 | 1.115 | 0.036 |
| 2011 | 1.137 | 0.041 |
| 2012 | 1.159 | 0.034 |


| 2013 | 0.830 | 0.045 |
| :--- | :--- | :--- |
| 2014 | 0.825 | 0.047 |
| 2015 | 0.780 | 0.051 |
| 2016 | 0.943 | 0.040 |
| 2017 | 1.020 | 0.037 |
| 2018 | 1.213 | 0.032 |
| 2019 | 0.836 | 0.051 |
| 2020 | 0.746 | 0.057 |
| 2021 | 0.661 | 0.068 |
| 2022 | 0.733 | 0.061 |

## c. Commercial fishery CPUE index by non-interaction model

We fitted the negative binomial GLM model for fish ticket retained CPUE time series 1985/86 - 1998/99 offering Year, Month, Vessel, Captain, and Area as explanatory variables and applying the hybrid selection method. Reduced area resolution (ADF\&G area codes were grouped to AreaGP) was used for model fitting. There were 24,075 records in the analysis.

The final model for EAG was:

> Initial selection by stepAIC:
> $\ln ($ CPUE $)=$ Year + Vessel + Month
> AIC $=16,997$

Final selection by stepCPUE: $\ln (C P U E)=$ Year + Vessel + Month for the 1985/86-1998/99 period $\left[\theta=10.45, \mathrm{R}^{2}=0.3328\right]$
and that for WAG was:
Initial selection by stepAIC:
$\ln ($ CPUE $)=$ Year + Vessel + Area
AIC=31,701
Final selection by stepCPUE: $\ln ($ CPUE $)=$ Year + Vessel + Area
for the 1985/86-1998/99 period [ $\left.\theta=6.67, \mathrm{R}^{2}=0.3569\right]$

## Appendix C: Cooperative survey

## 1. Summary of the survey method

The ADF\&G and industry collaborative pot survey was initiated in 2015 in the EAG and has continued since then. The survey was extended to WAG in 2018. A stratified two-stage sampling design has been implemented in a $2 \mathrm{nmi} \times 2 \mathrm{nmi}$ grids within 1000 m depth covering the entire golden king crab fishing area. The $2 \mathrm{nmi} \times 2 \mathrm{nmi}$ choice was the best compromise between scale of fishing gear, accuracy of defining habitat, and number of possible stations (Figure C1).


Figure C.1. Survey design: 2 nmi x 2 nmi grids overlaid on observer pot sample locations (green squares) in EAG.

There are nearly 1100 grids in the EAG divided into three equal size strata for selecting random pot sampling locations (Figures C. 2 and C.3).


Figure C.2. Survey design: $2 \mathrm{nmi} \times 2 \mathrm{nmi}$ grids stratified by three equal sizes for selecting random pot sampling locations in EAG.


Figure C.3. Random sample of 22 cells selected in each of three sub strata in EAG during the 2019 fishery.

Surveys occur during the first month of each fishing season with one to two ADF\&G biologists onboard the fishing vessel to collect fishery and biological data. Fishing operation takes place in a randomly selected set of grids in each stratum with long-line pots. The number of pots per string ranges from 30 to $40,200 \mathrm{~m}$ apart, and a vessel carries on average 35 strings. Pot sizes range from $5.5 \mathrm{ft} \times 5.5 \mathrm{ft}$ to $7 \mathrm{ft} \times 7 \mathrm{ft}$ with large mesh sizes for retention of legal-sized king crab. A few small mesh size research pots are also deployed for special studies. Fishing operation is not standardized for depth or soak time to allow normal fishing practices.

There are multiple pots (typically about 5 pots) sampled for each long-line string with approximately 35 crab measurement made per pot. For example, if 100 crabs are caught in a sampled pot, the biologist measures every third crab. The following snapshot of an observation record provides an example of what stock assessment data are collected.

| fishery | year | vessel | skipper | String\# | pot_size | mesh_size | bait | subsample_rate | species_code | sex | size | legal |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| EAG | 2015 | 20556 | Chad_Hoefer | 1 | 5×5 | king(large) | halibut | 2 | 923 | 1 | 187 | 1 |


| Pot\# | date_in | time_in | depth_start | start_lat | start_lon | depth_out | end_lat | end_lon | date_out | time_out | comments |
| :--- | :---: | :---: | :---: | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| soak_time |  |  |  |  |  |  |  |  |  |  |  |
| 1 | $8 / 4 / 2015$ | $17: 00$ | 132 | 52.74133 | -170.692 | 133 | 52.7515 | -170.675 | $8 / 17 / 2015$ | $3: 00$ |  |

## 2. Standardization of cooperative survey CPUE

## Data

A unique property of the cooperative survey is that multiple pots from multiple strings are sampled. All sample measurements were taken in EAG except for 2018/19, during which measurements were also taken from WAG. There was no survey during 2020/21 due to COVID related restriction. There are 47,075 records from seven years (20152019, 2021-2022) of surveys. After cleaning up for missing entries, the number of records reduced to 35,443 golden king crab comprised of 29,799 males and 5,644 females.

## Method

Data preparation for CPUE standardization:
i.) Created two new columns by concatenating Vessel code with String\# as well as with String\# and Pot\# because String\# and Pot\# are not unique numbers to each vessel. The new column names were identified as VesString and VesStingPot. For example, a Vessel Code 20556 with a String\# 3 was concatenated to be 205563 in a new column VesString, and a Vessel Code 20556 with a String\#17 and a Pot\# 5 was concatenated to be 20556175 in a new column VesStringPot.
ii.) Raised the Catch in each record by the Sample Rate.
iii.) Subset the data by large mesh king crab pot [Mesh ID not equal to 2 (i.e., small mesh pot)], legal size (> 135 mm CL), and EAG (EAGWAG=1). The female (Sex=2) and unclassified catch without any male crab (Sex=1) in a crab pot was set to 0 to account for the possibility of zero catch for expected male CPUE determination.
iv.) Further subset the data by $5 \%$ to $95 \%$, trimmed Soak time, and $1 \%$ to $99 \%$ trimmed Depth. This is to exclude catches from any unusual pot operations.
v.) Summed up the catch across sizes for each Pot\# and labelled it as SumCatch. Thus, each Pot\# has a single catch number.

The sampling design (sampling crab from a pot within a string within a vessel) begged for application of a mixed effects model to analyze the data, which was also recommended by the CPT.

The dispersion parameter value for the negative binomial error model and the degrees of freedom for cubic splines for soak time and depth variables were estimated by the fixed effect model using survey data. To apply the fixed effect model, a new response variable, Mean Pot Catch, was created by calculating a mean pot catch for a given set of year, vessel, captain, and string levels, ensuring a unique response variable value for a given set of predictor variable values.

## 4. Results

4.1. Appropriate degrees of freedom for splines and dispersion parameter for the negative binomial model were determined by calculating the minimum AIC of individual variable's GLM fit. The final splines degrees of freedom for Depth was 9 and that for Soak time was 3. The dispersion parameter was 3.01 .
4.2. Relevant fixed effect components from the final fixed effect model were selected for the random intercept model (C.1). The "lme4" library in R (R Core Team, 2022) with the "glmer()" function for model fitting was used in model fitting. The glmer() function allows use of any type of error model to fit the data.

Sum Catch $=\mathrm{Y}+\mathrm{ns}($ Depth, $\mathrm{df}=9)+\mathrm{ns}($ Soak, $\mathrm{df}=3)+$ Captain $+(1 \mid$ Block/VesselString $)$ family $=$ negative binomial $(\theta=3.01)$.

The random intercept model with the above parameterization converged.

### 4.3 Diagnostic test

The QQ plot for the fit assured that model assumptions were correct (Figure C.4).

Random Effects Model Fit, Cooperative Survey 2015-2022


Figure C.4. Studentized residual plot for the mixed random effects model fit using the 2015-2019, 2021-2022 EAG data.

### 4.4 CPUE indices

The fitted mixed effects model was used to predict row CPUEs, which were used for annual CPUE index, variance, and standard error calculations as follows:
a). Predicted CPUEs were log transformed, annual mean of log (CPUE) was calculated, and retransformed to row mean value with bias correction:
bias corrected mean $C P U E_{t} \approx e^{\text {mean }(\log (C P U E))} *\left(1+\frac{\operatorname{var}(\log (C P U E))}{2}\right)$
where $t$ is the year.
Then,
${\text { CPUE } \text { index }_{t}}=\frac{\text { mean } C P U E_{t}}{\text { Geomean }\left(\text { mean } C P U E_{t} s\right)}$
b). variance of $C P U E$ index $_{t}=\operatorname{var}(\log (C P U E)) / n$
where $n$ is the sample size.
c). standard error of $C P U E$ index $_{t}=\sqrt{\operatorname{var}(\log (C P U E)) / n}$

The CPUE indices, standard errors, samples sizes, and confidence limits are listed in Table C.1.

Table C.1. The cooperative survey predicted legal male standardized CPUE indices by the mixed random effects model, standard errors (SE), and lower- and upper- $95 \%$ confidence limits for EAG, 2015-2019, 2021-2022 data.

| Year | Predicted CPUE <br> index | SE | Lower <br> Limit | Upper <br> Limit | Sample <br> size |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 2015 | 1.24034 | 0.03186 | 1.16379 | 1.32194 | 263 |
| 2016 | 0.97640 | 0.02909 | 0.92121 | 1.03490 | 294 |
| 2017 | 1.17179 | 0.03954 | 1.08271 | 1.26821 | 205 |
| 2018 | 1.15031 | 0.03310 | 1.07663 | 1.22904 | 235 |
| 2019 | 0.69394 | 0.03674 | 0.64477 | 0.74685 | 258 |
| 2021 | 0.76192 | 0.03209 | 0.71456 | 0.81242 | 292 |
| 2022 | 1.15859 | 0.03835 | 1.07306 | 1.25094 | 274 |

A likelihood function with the 2015-2019, and 2021-2022 survey indices, using Equations A. 12 and A.13, was added to the likelihoods of observer indices (1995-2014) and fishery indices (1985-1998) and formulated a new model scenario 22.1g.

## Appendix D: Jittering

## Jittering of model $22.1 e 2$ parameter estimates

The Stock Synthesis approach was followed to do 100 jitter runs using model 22.1e2 parameter estimates as initial parameter values (i.e., as PIN file in ADMB) to assess model stability and to determine whether a global, as opposed to local, minimum has been reached by the search algorithm. Jittering was done outside the GMACS framework.

The $30 \%$ jittering was used to investigate the minimization process. In the jittering process, a Jitter factor of 0.3 was multiplied by a random normal deviation $r d e v=N(0,1)$ to create a transformed parameter value based upon the predefined parameter:

$$
\begin{equation*}
\text { temp }=0.5 * \text { rdev }^{*} \text { Jitterfactor } * \ln \left(\frac{P_{\max }-P_{\min }+0.0000002}{P_{v a l}-P_{\min }+0.0000001}-1\right) \tag{D.1}
\end{equation*}
$$

with the final jittered initial parameter value back transformed as:

$$
\begin{equation*}
P_{\text {new }}=P_{\min }+\frac{P_{\max }-P_{\min }}{1.0+\exp (-2.0 \text { temp })}, \tag{D.2}
\end{equation*}
$$

where $P_{\max }$ and $P_{\min }$ are upper and lower bounds of parameter search space and $P_{\text {val }}$ is the estimated parameter value before the jittering.

The jitter results are summarized for the model 22.1e2 in Tables D. 1 and D. 2 for EAG and WAG, respectively. The minimum negative log likelihood values realized in jitter runs coincided with optimized estimates by the model 22.1e2 runs for EAG and WAG, except for nonconvergent runs. We concluded from jitter results that optimization of model 22.1e2 achieved the global minima. Since GMACS formulation of models $22.1 \mathrm{e} 2,22.1 \mathrm{f}, 22.1 \mathrm{~g}$, and 21.1h have similar model structures leading to closer reference points estimation, it was concluded that their optimizations would have reached global minima.

Table D.1. Results from 100 jitter runs for scenario 22.1 e 2 for EAG. Jitter run 0 corresponds to the original optimized estimates. NA: model did not converge.

| Jitter Run | Objective Function | Maximum Gradient | $\mathrm{B}_{35 \%}$ ( t ) | OFL (t) | Current MMB (t) |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 0 | -940.3357 | 0.000327 | 6,516 | 2,843 | 7,404 |
| 1 | -818.6883 | 0.000066 | 7,105 | 3,389 | 8,450 |
| 2 | -940.3357 | 0.000066 | 6,516 | 2,843 | 7,404 |
| 3 | -818.8862 | 0.000103 | 7,072 | 3,402 | 8,439 |
| 4 | -817.9728 | 0.000099 | 7,104 | 3,453 | 8,472 |
| 5 | -818.8862 | 0.000084 | 7,072 | 3,402 | 8,439 |
| 6 | -818.6883 | 0.000212 | 7,105 | 3,389 | 8,450 |
| 7 | -940.3357 | 0.000137 | 6,516 | 2,843 | 7,404 |
| 8 | -818.6883 | 0.000641 | 7,105 | 3,389 | 8,450 |
| 9 | -940.3357 | 0.000215 | 6,516 | 2,843 | 7,404 |


| 10 |  | -940.3357 |  | 0.000202 | 6,516 | 2,843 | 7,404 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 11 |  | -940.3357 |  | 0.000107 | 6,516 | 2,843 | 7,404 |
| 12 |  | -940.3357 |  | 0.000056 | 6,516 | 2,843 | 7,404 |
| 13 |  | -940.3357 |  | 0.000103 | 6,516 | 2,843 | 7,404 |
| 14 | NA |  | NA |  | NA | NA | NA |
| 15 |  | -817.9728 |  | 0.000068 | 7,104 | 3,453 | 8,472 |
| 16 |  | -940.3357 |  | 0.000083 | 6,516 | 2,843 | 7,404 |
| 17 |  | -817.9728 |  | 0.000067 | 7,104 | 3,453 | 8,472 |
| 18 |  | -940.3357 |  | 0.000030 | 6,516 | 2,843 | 7,404 |
| 19 |  | -940.3357 |  | 0.000076 | 6,516 | 2,843 | 7,404 |
| 20 |  | -940.3357 |  | 0.000085 | 6,516 | 2,843 | 7,404 |
| 21 |  | -817.9728 |  | 0.000127 | 7,104 | 3,453 | 8,472 |
| 22 |  | -817.9728 |  | 0.000027 | 7,104 | 3,453 | 8,472 |
| 23 |  | -818.6883 |  | 0.000143 | 7,105 | 3,389 | 8,450 |
| 24 |  | -940.3357 |  | 0.000129 | 6,516 | 2,843 | 7,404 |
| 25 |  | -817.9728 |  | 0.000190 | 7,104 | 3,453 | 8,472 |
| 26 |  | -818.6883 |  | 0.000188 | 7,105 | 3,389 | 8,450 |
| 27 |  | -817.9728 |  | 0.000153 | 7,104 | 3,453 | 8,472 |
| 28 |  | -894.7376 |  | 0.000238 | 6,450 | 2,847 | 7,327 |
| 29 |  | -817.9728 |  | 0.000424 | 7,104 | 3,453 | 8,472 |
| 30 |  | -817.9728 |  | 0.000857 | 7,104 | 3,453 | 8,472 |
| 31 |  | -940.3357 |  | 0.000046 | 6,516 | 2,843 | 7,404 |
| 32 | NA |  | NA |  | NA | NA | NA |
| 33 |  | -817.9728 |  | 0.000037 | 7,104 | 3,453 | 8,472 |
| 34 |  | -940.3357 |  | 0.000300 | 6,516 | 2,843 | 7,404 |
| 35 |  | -940.3357 |  | 0.000309 | 6,516 | 2,843 | 7,404 |
| 36 |  | -817.9728 |  | 0.000067 | 7,104 | 3,453 | 8,472 |
| 37 |  | -940.3357 |  | 0.000009 | 6,516 | 2,843 | 7,404 |
| 38 |  | -894.7376 |  | 0.000178 | 6,450 | 2,847 | 7,327 |
| 39 |  | -818.6883 |  | 0.000330 | 7,105 | 3,389 | 8,450 |
| 40 |  | -818.8862 |  | 0.000181 | 7,072 | 3,402 | 8,439 |
| 41 |  | -817.9728 |  | 0.000202 | 7,104 | 3,453 | 8,472 |
| 42 |  | -817.9728 |  | 0.000104 | 7,104 | 3,453 | 8,472 |
| 43 |  | -817.9728 |  | 0.000120 | 7,104 | 3,453 | 8,472 |
| 44 |  | -817.9728 |  | 0.000398 | 7,104 | 3,453 | 8,472 |
| 45 |  | -817.9728 |  | 0.000390 | 7,104 | 3,453 | 8,472 |
| 46 |  | -818.8862 |  | 0.000089 | 7,072 | 3,402 | 8,439 |
| 47 |  | -817.9728 |  | 0.001069 | 7,104 | 3,453 | 8,472 |
| 48 |  | -894.7376 |  | 0.000098 | 6,450 | 2,847 | 7,327 |
| 49 |  | -817.9728 |  | 0.000700 | 7,104 | 3,453 | 8,472 |
| 50 |  | -818.6883 |  | 0.000194 | 7,105 | 3,389 | 8,450 |
| 51 |  | -817.9728 |  | 0.000260 | 7,104 | 3,453 | 8,472 |
| 52 |  | -817.9728 |  | 0.000246 | 7,104 | 3,453 | 8,472 |
| 53 |  | -940.3357 |  | 0.000129 | 6,516 | 2,843 | 7,404 |


| 54 |  | -817.9728 |  | 0.000014 | 7,104 | 3,453 | 8,472 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 55 |  | -818.8862 |  | 0.000034 | 7,072 | 3,402 | 8,439 |
| 56 | NA |  | NA |  | NA | NA | NA |
| 57 |  | -817.9728 |  | 0.000085 | 7,104 | 3,453 | 8,472 |
| 58 |  | -817.9728 |  | 0.000535 | 7,104 | 3,453 | 8,472 |
| 59 |  | -817.9728 |  | 0.000051 | 7,104 | 3,453 | 8,472 |
| 60 |  | -817.9728 |  | 0.000150 | 7,104 | 3,453 | 8,472 |
| 61 |  | -818.6883 |  | 0.000304 | 7,105 | 3,389 | 8,450 |
| 62 |  | -818.8862 |  | 0.000430 | 7,072 | 3,402 | 8,439 |
| 63 | NA |  | NA |  | NA | NA | NA |
| 64 | NA |  | NA |  | NA | NA | NA |
| 65 |  | -817.9728 |  | 0.000139 | 7,104 | 3,453 | 8,472 |
| 66 | NA |  | NA |  | NA | NA | NA |
| 67 |  | -940.3357 |  | 0.000018 | 6,516 | 2,843 | 7,404 |
| 68 |  | -817.9728 |  | 0.000241 | 7,104 | 3,453 | 8,472 |
| 69 |  | -817.9728 |  | 0.000053 | 7,104 | 3,453 | 8,472 |
| 70 |  | -817.9728 |  | 0.000060 | 7,104 | 3,453 | 8,472 |
| 71 |  | -940.3357 |  | 0.000476 | 6,516 | 2,843 | 7,404 |
| 72 |  | -940.3357 |  | 0.000094 | 6,516 | 2,843 | 7,404 |
| 73 |  | -818.8862 |  | 0.000036 | 7,072 | 3,402 | 8,439 |
| 74 |  | -940.3357 |  | 0.000017 | 6,516 | 2,843 | 7,404 |
| 75 |  | -817.9728 |  | 0.000793 | 7,104 | 3,453 | 8,472 |
| 76 |  | -817.9728 |  | 0.000071 | 7,104 | 3,453 | 8,472 |
| 77 |  | -940.3357 |  | 0.000147 | 6,516 | 2,843 | 7,404 |
| 78 | NA |  | NA |  | NA | NA | NA |
| 79 |  | -940.3357 |  | 0.000132 | 6,516 | 2,843 | 7,404 |
| 80 |  | -818.8862 |  | 0.000187 | 7,072 | 3,402 | 8,439 |
| 81 | NA |  | NA |  | NA | NA | NA |
| 82 |  | -817.9728 |  | 0.000078 | 7,104 | 3,453 | 8,472 |
| 83 | NA |  | NA |  | NA | NA | NA |
| 84 |  | -817.9728 |  | 0.000141 | 7,104 | 3,453 | 8,472 |
| 85 | NA |  | NA |  | NA | NA | NA |
| 86 |  | -940.3357 |  | 0.000157 | 6,516 | 2,843 | 7,404 |
| 87 |  | -940.3357 |  | 0.000033 | 6,516 | 2,843 | 7,404 |
| 88 | NA |  | NA |  | NA | NA | NA |
| 89 |  | -817.9728 |  | 0.000179 | 7,104 | 3,453 | 8,472 |
| 90 |  | -940.3357 |  | 0.000196 | 6,516 | 2,843 | 7,404 |
| 91 |  | -940.3357 |  | 0.000130 | 6,516 | 2,843 | 7,404 |
| 92 |  | -940.3357 |  | 0.000163 | 6,516 | 2,843 | 7,404 |
| 93 |  | -818.6883 |  | 0.000153 | 7,105 | 3,389 | 8,450 |
| 94 |  | -894.7376 |  | 0.000035 | 6,450 | 2,847 | 7,327 |
| 95 | NA |  | NA |  | NA | NA | NA |
| 96 |  | -817.9728 |  | 0.000187 | 7,104 | 3,453 | 8,472 |
| 97 |  | -940.3357 |  | 0.000073 | 6,516 | 2,843 | 7,404 |


| 98 | -817.9728 | 0.000506 | 7,104 | 3,453 | 8,472 |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 99 | -894.7376 | 0.000226 | 6,450 | 2,847 | 7,327 |
| 100 | -817.9728 | 0.000322 | 7,104 | 3,453 | 8,472 |

Table D.2. Results from 100 jitter runs for scenario 22.1 e 2 for WAG. Jitter run 0 corresponds to the original optimized estimates. NA: model did not converge.

| Jitter Run | Objective Function | Maximum Gradient | B35\% (t) | OFL (t) | Current MMB (t) |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 0 | -1044.2019 | 0.000047 | 4,885 | 1,226 | 4,578 |
| 1 | -808.1501 | 0.000590 | 5,188 | 1,545 | 4,968 |
| 2 | -1044.2020 | 0.000032 | 4,885 | 1,226 | 4,578 |
| 3 | -1044.2020 | 0.000079 | 4,885 | 1,226 | 4,578 |
| 4 | -1044.2020 | 0.000039 | 4,885 | 1,226 | 4,578 |
| 5 | -1044.2020 | 0.000204 | 4,885 | 1,226 | 4,578 |
| 6 | -808.1501 | 0.000154 | 5,188 | 1,545 | 4,968 |
| 7 | -1044.2020 | 0.000258 | 4,885 | 1,226 | 4,578 |
| 8 | -807.3073 | 0.000057 | 5,242 | 1,578 | 5,037 |
| 9 | -813.3011 | 0.000204 | 5,202 | 1,506 | 4,963 |
| 10 | -812.5013 | 0.000169 | 5,168 | 1,487 | 4,944 |
| 11 | -1044.2020 | 0.000105 | 4,885 | 1,226 | 4,578 |
| 12 | -807.3073 | 0.000464 | 5,242 | 1,578 | 5,037 |
| 13 | -807.3073 | 0.000264 | 5,242 | 1,578 | 5,037 |
| 14 | -1044.2020 | 0.000035 | 4,885 | 1,226 | 4,578 |
| 15 | -808.1501 | 0.000144 | 5,188 | 1,545 | 4,968 |
| 16 | NA | NA | NA | NA | NA |
| 17 | -1044.2020 | 0.000049 | 4,885 | 1,226 | 4,578 |
| 18 | -808.1501 | 0.000247 | 5,188 | 1,545 | 4,968 |
| 19 | -808.1501 | 0.000277 | 5,188 | 1,545 | 4,968 |
| 20 | -1044.2020 | 0.000205 | 4,885 | 1,226 | 4,578 |
| 21 | -813.3011 | 0.000441 | 5,202 | 1,506 | 4,963 |
| 22 | -813.3011 | 0.000129 | 5,202 | 1,506 | 4,963 |
| 23 | -1044.2020 | 0.000116 | 4,885 | 1,226 | 4,578 |
| 24 | -807.3073 | 0.000093 | 5,242 | 1,578 | 5,037 |
| 25 | NA | NA | NA | NA | NA |
| 26 | -1044.2020 | 0.000049 | 4,885 | 1,226 | 4,578 |
| 27 | -808.1501 | 0.000675 | 5,188 | 1,545 | 4,968 |
| 28 | -808.1501 | 0.000113 | 5,188 | 1,545 | 4,968 |
| 29 | -808.1501 | 0.000284 | 5,188 | 1,545 | 4,968 |
| 30 | -808.1501 | 0.000162 | 5,188 | 1,545 | 4,968 |
| 31 | -807.3073 | 0.000065 | 5,242 | 1,578 | 5,037 |
| 32 | -807.3073 | 0.000054 | 5,242 | 1,578 | 5,037 |
| 33 | -807.3073 | 0.000394 | 5,242 | 1,578 | 5,037 |
| 34 | -808.1501 | 0.000038 | 5,188 | 1,545 | 4,968 |


| 35 | -1044.2020 | 0.000084 | 4,885 | 1,226 | 4,578 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 36 | -807.3073 | 0.000190 | 5,242 | 1,578 | 5,037 |
| 37 | -1044.2020 | 0.000041 | 4,885 | 1,226 | 4,578 |
| 38 | -808.1501 | 0.000284 | 5,188 | 1,545 | 4,968 |
| 39 | -1044.2020 | 0.000173 | 4,885 | 1,226 | 4,578 |
| 40 | -807.3073 | 0.000570 | 5,242 | 1,578 | 5,037 |
| 41 | -1044.2020 | 0.000192 | 4,885 | 1,226 | 4,578 |
| 42 | -808.1501 | 0.000084 | 5,188 | 1,545 | 4,968 |
| 43 | -813.3011 | 0.000579 | 5,202 | 1,506 | 4,963 |
| 44 | -1044.2020 | 0.000057 | 4,885 | 1,226 | 4,578 |
| 45 | -1044.2020 | 0.000139 | 4,885 | 1,226 | 4,578 |
| 46 | -1044.2020 | 0.000057 | 4,885 | 1,226 | 4,578 |
| 47 | -1044.2020 | 0.000280 | 4,885 | 1,226 | 4,578 |
| 48 | -1044.2020 | 0.000022 | 4,885 | 1,226 | 4,578 |
| 49 | -1044.2020 | 0.000073 | 4,885 | 1,226 | 4,578 |
| 50 | -812.5013 | 0.000366 | 5,168 | 1,487 | 4,944 |
| 51 | -1044.2020 | 0.000094 | 4,885 | 1,226 | 4,578 |
| 52 | -1033.4540 | 0.000017 | 5,187 | 1,180 | 4,660 |
| 53 | -1044.2020 | 0.000065 | 4,885 | 1,226 | 4,578 |
| 54 | NA | NA | NA | NA | NA |
| 55 | -1044.2020 | 0.000270 | 4,885 | 1,226 | 4,578 |
| 56 | -1044.2020 | 0.000022 | 4,885 | 1,226 | 4,578 |
| 57 | -1044.2020 | 0.000211 | 4,885 | 1,226 | 4,578 |
| 58 | -1044.2020 | 0.000058 | 4,885 | 1,226 | 4,578 |
| 59 | -1044.2020 | 0.000062 | 4,885 | 1,226 | 4,578 |
| 60 | -808.1501 | 0.000557 | 5,188 | 1,545 | 4,968 |
| 61 | -807.3073 | 0.000260 | 5,242 | 1,578 | 5,037 |
| 62 | -1044.2020 | 0.000203 | 4,885 | 1,226 | 4,578 |
| 63 | -807.3073 | 0.000091 | 5,242 | 1,578 | 5,037 |
| 64 | -813.3011 | 0.000542 | 5,202 | 1,506 | 4,963 |
| 65 | -807.3073 | 0.000010 | 5,242 | 1,578 | 5,037 |
| 66 | -1044.2020 | 0.000060 | 4,885 | 1,226 | 4,578 |
| 67 | NA | NA | NA | NA | NA |
| 68 | NA | NA | NA | NA | NA |
| 69 | -813.3011 | 0.000104 | 5,202 | 1,506 | 4,963 |
| 70 | -807.3073 | 0.000272 | 5,242 | 1,578 | 5,037 |
| 71 | -1044.2020 | 0.000098 | 4,885 | 1,226 | 4,578 |
| 72 | -807.3073 | 0.000033 | 5,242 | 1,578 | 5,037 |
| 73 | NA | NA | NA | NA | NA |
| 74 | -1044.2020 | 0.000023 | 4,885 | 1,226 | 4,578 |
| 75 | -1044.2020 | 0.000025 | 4,885 | 1,226 | 4,578 |
| 76 | -813.3011 | 0.000399 | 5,202 | 1,506 | 4,963 |
| 77 | -1044.2020 | 0.000007 | 4,885 | 1,226 | 4,578 |
| 78 | -1033.4540 | 0.000133 | 5,187 | 1,180 | 4,660 |


| 79 | -807.3073 | 0.000088 | 5,242 | 1,578 | 5,037 |
| ---: | ---: | :---: | :---: | :---: | :---: |
| 80 | -807.3073 | 0.000287 | 5,242 | 1,578 | 5,037 |
| 81 | -1044.2020 | 0.000014 | 4,885 | 1,226 | 4,578 |
| 82 | -808.1501 | 0.000168 | 5,188 | 1,545 | 4,968 |
| 83 | NA | NA | NA | NA | NA |
| 84 | -1044.2020 | 0.000216 | 4,885 | 1,226 | 4,578 |
| 85 | -807.3073 | 0.000182 | 5,242 | 1,578 | 5,037 |
| 86 | -807.3073 | 0.000226 | 5,242 | 1,578 | 5,037 |
| 87 | -1044.2020 | 0.000185 | 4,885 | 1,226 | 4,578 |
| 88 | -807.3073 | 0.000283 | 5,242 | 1,578 | 5,037 |
| 89 | -808.1501 | 0.000355 | 5,188 | 1,545 | 4,968 |
| 90 | -813.3011 | 0.000077 | 5,202 | 1,506 | 4,963 |
| 91 | -1033.4540 | 0.000024 | 5,187 | 1,180 | 4,660 |
| 92 | -1044.2020 | 0.000011 | 4,885 | 1,226 | 4,578 |
| 93 | -808.1501 | 0.000121 | 5,188 | 1,545 | 4,968 |
| 94 | -813.3011 | 0.000244 | 5,202 | 1,506 | 4,963 |
| 95 | -1044.2020 | 0.000220 | 4,885 | 1,226 | 4,578 |
| 96 | -807.3073 | 0.000124 | 5,242 | 1,578 | 5,037 |
| 97 | -808.1501 | 0.000075 | 5,188 | 1,545 | 4,968 |
| 98 | -808.1501 | 0.000075 | 5,188 | 1,545 | 4,968 |
| 99 | -808.1501 | 0.000186 | 5,188 | 1,545 | 4,968 |
| 100 | -808.1501 | 0.000227 | 5,188 | 1,545 | 4,968 |

## Appendix E: Gmacs

## Introduction

The CPT/SSC in January/February 2023, suggested to implement the Aleutian Islands golden king crab assessment models in GMACS for the May/June 2023 CPT/SSC meetings. Following this suggestion, all models were implemented in GMACS. The modified base model 22.9 c was also run outside GMACS to compare some of its results with that of GMACS.

## Method

The model 22.9 c was a modification of model 22.1 e 2 (with three catchability and three additional SDs) for GMACS implementation. Estimated parameters from model 22.9c for EAG (EAG22.9c) and WAG (WAG22.9c) that were reparametrized for GMACS computational formulas were input to GMACS ctl file. Parallel data and projection files were also created for GMACS runs (e.g., GMACS9cEAG22.1e2CatchNo.ctl, GMACS9cEAG22.1e2CatchNo.dat, and GMACS9cEAG22.1e2CatchNo.prj).

Models 22.1e2, 22.1f (with Year and Area interaction CPUE), 22.1g (with cooperative survey CPUE indices), and 22.1h (with Year and Area interaction CPUE and cooperative survey CPUE indices) were implemented in GMACS. The OFL and ABC estimated by the GMACS models are presented in the executive summary tables. The MMB, CPUE, recruitment, and fishing mortality trends were also compared among the GMACS models and the status quo model, 22.9 c , in the main report. We provide the GMACS input files for model 22.1 e 2 for EAG and WAG, respectively, as examples on how GMACS being implemented for Aleutian Islands golden king crab:

## A. Input data files for model 22.1e2:

1. EAG22.1e2 ctl file
2. \# EAG22_1e2 Update
3. \# Controls for leading parameter vector theta
4. \# LEGEND FOR PRIOR
5. \# $0=$ uniform, $1=$ normal, $2=\operatorname{lognormal}, 3=$ beta, $4=$ gamma
6. \#
7. \# ntheta
8. 9
9. \#ival lb ub phz prior p1 p2
10. $0.22 \quad 0.01 \quad 1.0 \quad 13 \quad 2 \quad 0.18 \quad 0.04$
11. $7.790339494-10.020 .0 \quad 1 \quad 0 \quad-10.0 \quad 20.0$
12. 
13. $12.0 \quad-10.0 \quad 20.0 \quad-3 \quad 0 \quad-10.0 \quad 20.0$
$\begin{array}{lllllllll}\text { 15. } & 8.0 & -10.0 & 20.0 & -1 & 0 & -10.0 & 20.0\end{array}$
14. 

$\begin{array}{lllllllll}\text { 17. } & 110.0 & 103.0 & 165.0 & -2 & 1 & 72.5 & 7.25 & \#\end{array}$
18.
19. $1.6161266570 .001 \quad 20.0 \quad 3 \quad 0 \quad 0.1 \quad 5.0 \quad$ distribution
20.
21. $-0.693147181-10.0 \quad 0.75$-1 $00-10.0 \quad 0.75 \quad \#$
22.
23. $0.7330 .2 \begin{array}{llllll}1.0 & -2 & 3 & 3.0 & 2.0\end{array}$
$\begin{array}{lllllll}0.001 & 0.0 & 1.0 & -3 & 3 & 1.01 & 1.01\end{array}$
25.
26. \# weight-at-length input method $\left(1=\right.$ allometry $\left[\mathrm{w} \_1=\mathrm{a}^{*} \mathrm{l}^{\wedge} \mathrm{b}\right], 2=$ vector by sex $)$
27. 2
28. \#a, in kg
29. \# 1.445E-07
30. \#b
31. \# 3.281126995
32. \# Male weight-at-length
33. $0.581515707 \quad 0.679328169 \quad 0.788032347 \quad 0.908278308 \quad 1.040724257 \quad 1.186036294 \quad 1.344888179 \quad 1.517961114$ $\begin{array}{lllllllll}1.705943543 & 1.90953096 & 2.129425732 & 2.366336933 & 2.620980182 & 2.894077494 & 3.186357141 & 3.498553516\end{array}$ 3.993657581
34. \#
35. \# Proportion mature by sex, males
36. 0. 0. 0. 1. 1. 1. 1. 1. 1. 1. 1. 1. 1. 1. 1.1. 1.
37. \# Proportion legal by sex, males
38. 0. 0. 0. 0. 0. 0. 0. 1. 1. 1. 1. 1. 1. 1. 1.1. 1.
39. \#\# GROWTH PARAM CONTROLS
40. \#\# Two lines for each parameter if split sex, one line if not \#\#
41. \# Use growth transition matrix option (1=read in growth-increment matrix; 2=read in size-transition; 3=gamma distribution for size-increment; 4=gamma distribution for size after increment) ( 1 to 8 options available)
42. \# option 8 is normal distributed growth incrment, size after incrment is normal
43. 8
44. \# growth increment model ( $0=$ prespecified; $1=$ alpha/beta; $2=$ estimated by size-class; $3=$ pre-specified/emprical)
45. 1
46. \# molt probability function ( $0=$ pre-specified; $1=$ flat; $2=$ declining logistic)
47. 2
48. \# maximum size-class (males then females)
49. \# Maximum size-class for recruitment(males then females)
50. 5
51. \#\# number of size-increment periods
52. 1
53. \#\# Year(s) size-increment period changes (blank if no changes)
54.
55. \#\# number of molt periods
56. 1
57. \#\# Year(s) molt period changes (blank if no changes)
58.
59. \#\# Beta parameters are relative to a base level ( $1=$ Yes; $0=$ no $)$
60. 1 \#
61.
62. \# Growth parameters
63. \# ival lb ub phz prior p1 p2 \# parameter \#
64. 25.2890479510 .0
65. 0.090568482 -0.4 $20.0 \quad 7 \quad 0 \quad 0.0$
66. $\begin{array}{lllllllll}3.68086767 & 0.01 & 5.0 & 7 & 0 & 0.0 & 3.0 & \#\end{array}$


```
104.## Selectivity P(capture of all sizes)
105.## gear par sel
phz start end ##
106.# index index par sex ival lb ub prior p1 p2 mirror period period ##
107.## Gear-1
108. 1 1 1 1 0 121.2664438 105.0 180.0 0 100.0 190.0 3 1960 2004 #
109. 1
110. 1 3 1 1 0 136.6281955 105.0 180.0 0 100.0 190.0 3 2005 2022 #
111. 1 4 4 2 0 8.21588572 0.01 20.0}0
112.
113.# Gear-2
114. 2
115.## Retained
116.## gear par sel phz start end
117.# index index par sex ival lb ub prior p1 p2 mirror period period
118.# Gear-1
119.-1
120.-1
121.
122.# Gear-2
123.-2
124.# Number of asyptotic parameters
125.1
126.# Fleet Sex Year ival lb ub phz
127. 1}
128.## PRIORS FOR CATCHABILITY
129.## If a uniform prior is selected for a parameter then the lb and ub are used (p1 ##
130.## and p2 are ignored). ival must be > 0
##
131.## only allowed to use uniform or lognormal prior
132.## if anlytic q estimation step is chosen, turn off estimating q by changing the estimation phase to be -ve
133.## LEGEND
##
134.## prior: 0 = uniform, 1 = normal, 2 = lognormal, 3 = beta, 4 = gamma }\quad#
135.#
136.## SURVEYS/INDICES ONLY
137.## fishery and observer CPUE
138.## Analytic (0=not analytically solved q, use uniform or lognormal prior;
139.## 1= analytic),
140.## Lambda =multilier for iput CV, Emphasis = multiplier for likelihood
```




```
215.## Specific initial values for the natural mortality devs (0-no, 1=yes)
216.## 1
217.## ival 
219.
220.## TAGGING controls CONTROLS
221.1 # emphasis on tagging data (1 =use tag LH, 0=ignore)
222.## Maturity specific natural mortality
223.###
224.# maturity specific natural mortality? (yes = 1; no = 0; only for use if nmature > 1)
225.0
```



```
251.0 0 0.001 0 # Groundfish
252.## EMPHASIS FACTORS (Priors)
253.##
254.# Log_fdevs meanF Mdevs Rec_devs Initial_devs Fst_dif_dev Mean_sex-Ratio Fvecs Fdovs
255.# 0
256.# Log_fdevs meanF Mdevs Rec_devs Initial_devs Fst_dif_dev Mean_sex-Ratio Molt_prob Free selectivity
257. Init_n_at_len Fvecs Fdovs
258. 0
259.## EOF
260.9999
```


## \#\#\#\#\#\#\#\#\#\#\#\#\#\#\#\# <br> EAG22.1e2 dat file



[^0]| 0. | 0.16666667 | 0. | 0.46073059 | 0. | 0.37260274 | $\# 1975$ |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 0. | 0.16666667 | 0. | 0.46073059 | 0. | 0.37260274 | $\# 1976$ |
| 0. | 0.16666667 | 0. | 0.46073059 | 0. | 0.37260274 | $\# 1977$ |
| 0. | 0.16666667 | 0. | 0.46073059 | 0. | 0.37260274 | $\# 1978$ |
| 0. | 0.16666667 | 0. | 0.46073059 | 0. | 0.37260274 | $\# 1979$ |
| 0. | 0.16666667 | 0. | 0.46073059 | 0. | 0.37260274 | $\# 1980$ |
| 0. | 0.44109589 | 0. | 0.18630137 | 0. | 0.37260274 | $\# 1981$ |
| 0. | 0.483561644 | 0. | 0.143835616 | 0. | 0.37260274 | $\# 1982$ |
| 0. | 0.483561644 | 0. | 0.143835616 | 0. | 0.37260274 | $\# 1983$ |
| 0. | 0.315068493 | 0. | 0.312328767 | 0. | 0.37260274 | $\# 1984$ |
| 0. | 0.168493151 | 0. | 0.45890411 | 0. | 0.37260274 | $\# 1985$ |
| 0. | 0.252054795 | 0. | 0.375342466 | 0. | 0.37260274 | $\# 1986$ |
| 0. | 0.087671233 | 0. | 0.539726027 | 0. | 0.37260274 | $\# 1987$ |
| 0. | 0.3 | 0. | 0.32739726 | 0. | 0.37260274 | $\# 1988$ |
| 0. | 0.4 | 0. | 0.22739726 | 0. | 0.37260274 | $\# 1989$ |
| 0. | 0.265753425 | 0. | 0.361643836 | 0. | 0.37260274 | $\# 1990$ |
| 0. | 0.273972603 | 0. | 0.353424658 | 0. | 0.37260274 | $\# 1991$ |
| 0. | 0.276712329 | 0. | 0.350684932 | 0. | 0.37260274 | $\# 1992$ |
| 0. | 0.419178082 | 0. | 0.208219178 | 0. | 0.37260274 | $\# 1993$ |
| 0. | 0.249315068 | 0. | 0.378082192 | 0. | 0.37260274 | $\# 1994$ |
| 0. | 0.223287671 | 0. | 0.404109589 | 0. | 0.37260274 | $\# 1995$ |
| 0. | 0.328767123 | 0. | 0.298630137 | 0. | 0.37260274 | $\# 1996$ |
| 0. | 0.28630137 | 0. | 0.34109589 | 0. | 0.37260274 | $\# 1997$ |
| 0. | 0.263013699 | 0. | 0.364383562 | 0. | 0.37260274 | $\# 1998$ |
| 0. | 0.245205479 | 0. | 0.382191781 | 0. | 0.37260274 | $\# 1999$ |
| 0. | 0.179452055 | 0. | 0.447945205 | 0. | 0.37260274 | $\# 2000$ |
| 0. | 0.160273973 | 0. | 0.467123288 | 0. | 0.37260274 | $\# 2001$ |
| 0. | 0.156164384 | 0. | 0.471232877 | 0. | 0.37260274 | $\# 2002$ |
| 0. | 0.157534247 | 0. | 0.469863014 | 0. | 0.37260274 | $\# 2003$ |
| 0. | 0.143835616 | 0. | 0.483561644 | 0. | 0.37260274 | $\# 2004$ |
| 0. | 0.432876712 | 0. | 0.194520548 | 0. | 0.37260274 | $\# 2005$ |
| 0. | 0.331506849 | 0. | 0.295890411 | 0. | 0.37260274 | $\# 2006$ |
| 0. | 0.368493151 | 0. | 0.25890411 | 0. | 0.37260274 | $\# 2007$ |
| 0. | 0.302739726 | 0. | 0.324657534 | 0. | 0.37260274 | $\# 2008$ |
| 0. | 0.32739726 | 0. | 0.3 | 0. | 0.37260274 | $\# 2009$ |
| 0. | 0.293150685 | 0. | 0.334246575 | 0. | 0.37260274 | $\# 2010$ |
| 0. | 0.263013699 | 0. | 0.364383562 | 0. | 0.37260274 | $\# 2011$ |
| 0 |  | 0.0 |  |  |  |  |


| 0. | 0.275342466 | 0. | 0.352054795 | 0. | 0.37260274 | \#2012 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0. | 0.27260274 | 0. | 0.354794521 | 0. | 0.37260274 | \#2013 |
| 0. | 0.247945205 | 0. | 0.379452055 | 0. | 0.37260274 | \#2014 |
| 0. | 0.228767123 | 0. | 0.398630137 | 0. | 0.37260274 | \#2015 |
| 0. | 0.420547945 | 0. | 0.206849315 | 0. | 0.37260274 | \#2016 |
| 0. | 0.409589041 | 0. | 0.217808219 | 0. | 0.37260274 | \#2017 |
| 0. | 0.349315068 | 0. | 0.278082192 | 0. | 0.37260274 | \#2018 |
| 0. | 0.32739726 | 0. | 0.3 0. |  | 260274 \#2 |  |
| 0. | 0.365753425 | 0. | 0.261643836 | 0. | 0.37260274 | \#2020 |
| 0. | 0.294520548 | 0. | 0.332876712 | 0. | 0.37260274 | \#2021 |
| 0. | 0.38630137 | 0 . | 0.24109589 | 0 . | 0.37260274 | \#2022 |
| \# |  |  |  |  |  |  |
| \# Fishing fleet names (delimited with : no spaces in names) Pot_Fishery Trawl_Bycatch <br> \# Survey names (delimited with : no spaces in names) keep empty |  |  |  |  |  |  |
|  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |
| \# Are the seasons discrete-instantaneous (0) or continuous (1) |  |  |  |  |  |  |
| 111111 |  |  |  |  |  |  |
| \# Number of catch data frames |  |  |  |  |  |  |
| 3 |  |  |  |  |  |  |
| \# Number of rows in each data frame |  |  |  |  |  |  |
| \# 1993 total catch is missing, |  |  |  | up | to 2022/2 | data |
| \# | 1991 ground |  | bycatch is mi | issing |  |  |
|  | retained cat |  | 1981/82-2022/23 |  |  |  |
|  | $42 \quad 32 \quad 33$ |  |  |  |  |  |
| \#\# CATCH DATA in t |  |  |  |  |  |  |
| \#\# Type of catch: $1=$ retained, $2=$ discard, $0=$ total (retained+discard, slide says 3 ) |  |  |  |  |  |  |
| \#\# Units of catch: $1=$ biomass, $2=$ numbers |  |  |  |  |  |  |
| \# Mult: $1=$ use data as thy are, $2=$ multiply by this number (e.g., lbs to kg ) |  |  |  |  |  |  |
| \#\# Retained Catch (numbers from 1981-1984; tonnes from 1985 onwards) |  |  |  |  |  |  |
| \#year seas fleet sex |  |  | obs cv type | unit | mult effor | discard |
| 1981 | 31 |  | 203.9680 .03161 | 2 | 10 |  |
| 1982 | 3 |  | 529.7870 .03161 | 2 | 10 |  |
| 1983 | 311 |  | 662.280 .03161 | 2 | 10 |  |
| 1984 | 3 |  | 801.10 .03161 | 2 | 100 |  |
| 1985 | 311 |  | 2730.320 .03161 | 1 | 10 |  |


| 1986 | 3 | 1 | 1 | 2844.910 .0316 | 1 | 1 | 0 | 0.2 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1987 | 3 | 1 | 1 | 1908.790 .03161 | 1 | 1 | 0 | 0.2 |
| 1988 | 3 | 1 | 1 | 2423.60 .03161 | 1 | 1 | 0 | 0.2 |
| 1989 | 3 | 1 | 1 | 2776.770 .03161 | 1 | 1 | 0 | 0.2 |
| 1990 | 3 | 1 | 1 | 1637.480 .0316 | 1 | 1 | 0 | 0.2 |
| 1991 | 3 | 1 | 1 | 2026.350 .0316 | 1 | 1 | 0 | 0.2 |
| 1992 | 3 | 1 | 1 | 2125.040 .03161 | 1 | 1 | 0 | 0.2 |
| 1993 | 3 | 1 | 1 | 1420.580 .0316 | 1 | 1 | 0 | 0.2 |
| 1994 | 3 | 1 | 1 | 2038.350 .0316 | 1 | 1 | 0 | 0.2 |
| 1995 | 3 | 1 | 1 | 2224.010 .03161 | 1 | 1 | 0 | 0.2 |
| 1996 | 3 | 1 | 1 | 1624.070 .0316 | 1 | 1 | 0 | 0.2 |
| 1997 | 3 | 1 | 1 | 1481.020 .0316 | 1 | 1 | 0 | 0.2 |
| 1998 | 3 | 1 | 1 | 1414.760 .03161 | 1 | 1 | 0 | 0.2 |
| 1999 | 3 | 1 | 1 | 1334.880 .0316 | 1 | 1 | 0 | 0.2 |
| 2000 | 3 | 1 | 1 | 1359.490 .0316 | 1 | 1 | 0 | 0.2 |
| 2001 | 3 | 1 | 1 | 1401.420 .03161 | 1 | 1 | 0 | 0.2 |
| 2002 | 3 | 1 | 1 | 1243.190 .0316 | 1 | 1 | 0 | 0.2 |
| 2003 | 3 | 1 | 1 | 1297.260 .0316 | 1 | 1 | 0 | 0.2 |
| 2004 | 3 | 1 | 1 | 1269.730 .03161 | 1 | 1 | 0 | 0.2 |
| 2005 | 3 | 1 | 1 | 1272.160 .03161 | 1 | 1 | 0 | 0.2 |
| 2006 | 3 | 1 | 1 | 1389.50 .03161 | 1 | 1 | 0 | 0.2 |
| 2007 | 3 | 1 | 1 | 1329.370 .03161 | 1 | 1 | 0 | 0.2 |
| 2008 | 3 | 1 | 1 | 1421.860 .0316 | 1 | 1 | 0 | 0.2 |
| 2009 | 3 | 1 | 1 | 1448.280 .0316 | 1 | 1 | 0 | 0.2 |
| 2010 | 3 | 1 | 1 | 1412.730 .03161 | , | 1 | 0 | 0.2 |
| 2011 | 3 | 1 | 1 | 1444.360 .03161 | 1 | 1 | 0 | 0.2 |
| 2012 | 3 | 1 | 1 | 1499.290 .0316 | 1 | 1 | 0 | 0.2 |
| 2013 | 3 | 1 | 1 | 1546.080 .03161 | 1 | 1 | 0 | 0.2 |
| 2014 | 3 | 1 | 1 | 1553.360 .03161 | 1 | 1 | 0 | 0.2 |
| 2015 | 3 | 1 | 1 | 1692.90 .03161 | 1 | 1 | 0 | 0.2 |
| 2016 | 3 | 1 | 1 | 1658.660 .03161 | 1 | , | 0 | 0.2 |
| 2017 | 3 | 1 | 1 | 1620.860 .03161 | 1 | 1 | 0 | 0.2 |
| 2018 | 3 | 1 | 1 | 1865.110 .03161 | 1 | 1 | 0 | 0.2 |
| 2019 | 3 | 1 | 1 | 2067.470 .03161 | 1 | 1 | 0 | 0.2 |
| 2020 | 3 | 1 | 1 | 1735.370 .03161 | 1 | 1 | 0 | 0.2 |
| 2021 | 3 | 1 | 1 | 1785.440 .03161 | 1 | 1 | 0 | 0.2 |
| 2022 | 3 | 1 | 1 | 1564.320 .03161 | 1 | 1 | 0 | 0.2 |

## \#

\#\# Total Catch (tonnes throughout)

| \#year | seas | fleet | sex | obs cv type |  | its | mult | effort | discard_mortality |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1990 | 3 | 1 | 1 | 3980.730 .358893929 | 0 | 1 | 1 | 0 | 0.2 |
| 1991 | 3 | 1 | 1 | 6596.740 .212951406 | 0 | 1 | 1 | 0 | 0.2 |
| 1992 | 3 | 1 | 1 | 5435.640 .296058703 | 0 | 1 | 1 | 0 | 0.2 |
| 1994 | 3 | 1 | 1 | 3444.230 .375117372 | 0 | 1 | 1 | 0 | 0.2 |
| 1995 | 3 | 1 | 1 | 4640.820 .051194102 | 0 | 1 | 1 | 0 | 0.2 |
| 1996 | 3 | 1 | 1 | 2563.320 .04474373 | 0 | 1 | 1 | 0 | 0.2 |
| 1997 | 3 | 1 | 1 | 2976.80 .059889204 | 0 | 1 | 1 | 0 | 0.2 |
| 1998 | 3 | 1 | 1 | 3140.990 .0680779 | 0 | 1 | 1 | 0 | 0.2 |
| 1999 | 3 | 1 | 1 | 2605.620 .065963387 | 0 | 1 | 1 | 0 | 0.2 |
| 2000 | 3 | 1 | 1 | 2759.910 .057628024 | 0 | 1 | 1 | 0 | 0.2 |
| 2001 | 3 | 1 | 1 | 2237.550 .060173859 | 0 | 1 | 1 | 0 | 0.2 |
| 2002 | 3 | 1 | 1 | 1915.660 .064883292 | 0 | 1 | 1 | 0 | 0.2 |
| 2003 | 3 | 1 | 1 | 1901.610 .065047278 | 0 | 1 | 1 | 0 | 0.2 |
| 2004 | 3 | 1 | 1 | 1694.870 .087224566 | 0 | 1 | 1 | 0 | 0.2 |
| 2005 | 3 | 1 | 1 | 1742.040 .118801346 | 0 | 1 | 1 | 0 | 0.2 |
| 2006 | 3 | 1 | 1 | 1646.830 .123871783 | 0 | 1 | 1 | 0 | 0.2 |
| 2007 | 3 | 1 | 1 | 1819.860 .12997936 | 0 | 1 | 1 | 0 | 0.2 |
| 2008 | 3 | 1 | 1 | 1823.510 .16628614 | 0 | 1 | 1 | 0 | 0.2 |
| 2009 | 3 | 1 | 1 | 1770.080 .204527938 | 0 | 1 | 1 | 0 | 0.2 |
| 2010 | 3 | 1 | 1 | 1756.660 .197720567 | 0 | 1 | 1 | 0 | 0.2 |
| 2011 | 3 | 1 | 1 | 1780.60 .217727165 | 0 | 1 | 1 | 0 | 0.2 |
| 2012 | 3 | 1 | 1 | 1946.590 .197259943 | 0 | 1 | 1 | 0 | 0.2 |
| 2013 | 3 | 1 | 1 | 1851.560 .184593328 | 0 | 1 | 1 | 0 | 0.2 |
| 2014 | 3 | 1 | 1 | 1967.390 .213240733 | 0 | 1 | 1 | 0 | 0.2 |
| 2015 | 3 | 1 | 1 | 2135.810 .188674437 | 0 | 1 | 1 | 0 | 0.2 |
| 2016 | 3 | 1 | 1 | 2234.130 .165738888 | 0 | 1 | 1 | 0 | 0.2 |
| 2017 | 3 | 1 | 1 | 2339.370 .170274949 | 0 | 1 | 1 | 0 | 0.2 |
| 2018 | 3 | 1 | 1 | 2734.630 .189279828 | 0 | 1 | 1 | 0 | 0.2 |
| 2019 | 3 | 1 | 1 | 3032.730 .17733387 | 0 | 1 | 1 | 0 | 0.2 |
| 2020 | 3 | 1 | 1 | 2608.060 .172996036 | 0 | 1 | 1 | 0 | 0.2 |
| 2021 | 3 | 1 | 1 | 2426.950 .188674437 | 0 | 1 | 1 | 0 | 0.2 |
| 2022 | 3 | 1 | 1 | 1731.770 .317585088 | 0 | 1 | 1 | 0 | 0.2 |
| \# |  |  |  |  |  |  |  |  |  |
| \#\# | Trawl |  | fishery discards (in tonnes) |  |  |  |  |  |  |


| 1989 | 3 | 2 | 1 | 0.826511 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1990 | 3 | 2 | 1 | 2.59394 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 1992 | 3 | 2 | 1 | 1.22658 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 1993 | 3 | 2 | 1 | 1.15375 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 1994 | 3 | 2 | 1 | 0.357445 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 1995 | 3 | 2 | 1 | 1.01804 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 1996 | 3 | 2 | 1 | 0.265799 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 1997 | 3 | 2 | 1 | 0.106796 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 1998 | 3 | 2 | 1 | 1.06278 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 1999 | 3 | 2 | 1 | 0.642352 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 2000 | 3 | 2 | 1 | 1.12817 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 2001 | 3 | 2 | 1 | 1.66704 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 2002 | 3 | 2 | 1 | 2.38549 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 2003 | 3 | 2 | 1 | 1.31099 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 2004 | 3 | 2 | 1 | 0.297833 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 2005 | 3 | 2 | 1 | 1.83486 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 2006 | 3 | 2 | 1 | 3.3144 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 2007 | 3 | 2 | 1 | 1.92908 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 2008 | 3 | 2 | 1 | 4.30175 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 2009 | 3 | 2 | 1 | 2.05905 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 2010 | 3 | 2 | 1 | 6.27075 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 2011 | 3 | 2 | 1 | 5.2775 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 2012 | 3 | 2 | 1 | 6.17064 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 2013 | 3 | 2 | 1 | 3.13431 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 2014 | 3 | 2 | 1 | 2.86222 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 2015 | 3 | 2 | 1 | 0.696906 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 2016 | 3 | 2 | 1 | 2.072410 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 2017 | 3 | 2 | 1 | 1.240610 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 2018 | 3 | 2 | 1 | 1.249180 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 2019 | 3 | 2 | 1 | 3.927930 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 2020 | 3 | 2 | 1 | 0.726643 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 2021 | 3 | 2 | 1 | 1.243970 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 2022 | 3 | 2 | 1 | 2.148710 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| \# |  |  |  |  |  |  |  |  |  |
| \#\# RELATIVE ABUNDANCE DATA |  |  |  |  |  |  |  |  |  |
| \#\# Units of abundance: $1=$ biomass, $2=$ numbers |  |  |  |  |  |  |  |  |  |
| \#\# Number of relative abundance indicies |  |  |  |  |  |  |  |  |  |


| \#\# sex:1=male;2=female; 0=both <br> \#\# maturity: $1=$ immature; $2=$ mature; $0=$ both $)$ |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| \# Fishery CPUE index, Observer CPUE index2 |  |  |  |  |  |  |  |  |  |
| $3$ |  |  |  |  |  |  |  |  |  |
| \# Index Type (1=Selecivity; 2=retention) |  |  |  |  |  |  |  |  |  |
| \# AEPAEP |  |  |  |  |  |  |  |  |  |
| 222 |  |  |  |  |  |  |  |  |  |
| \#\# Number of rows in each index |  |  |  |  |  |  |  |  |  |
| 42 |  |  |  |  |  |  |  |  |  |
| ry |  |  |  |  |  |  |  |  |  |
| \# Sex: 1 = male, 2 = female, $0=$ both" << endl; |  |  |  |  |  |  |  |  |  |
| \# Maturity: $1=$ immature, $2=$ mature, $0=$ both |  |  |  |  |  |  |  |  |  |
| \# Units of survey: $1=$ biomass, $2=$ numbers |  |  |  |  |  |  |  |  |  |
| \# Indices are in numbers |  |  |  |  |  |  |  |  |  |
| \# |  | ser |  |  | CPUE | index non int | action |  |  |
| 1 | 1995 | 3 | 1 | 1 | 0 | 1.003366325 | 0.031764858 | 2 | 0.5 |
| 1 | 1996 | 3 | 1 | 1 | 0 | 0.944433433 | 0.021762674 | 2 | 0.5 |
| 1 | 1997 | 3 | 1 | 1 | 0 | 0.874195678 | 0.023687839 | 2 | 0.5 |
| 1 | 1998 | 3 | 1 | 1 | 0 | 1.000445312 | 0.019428072 | 2 | 0.5 |
| 1 | 1999 | 3 | 1 | 1 | 0 | 0.915413606 | 0.020204677 | 2 | 0.5 |
| 1 | 2000 | 3 | 1 | 1 | 0 | 0.819562414 | 0.019360965 | 2 | 0.5 |
| 1 | 2001 | 3 | 1 | 1 | 0 | 1.042909239 | 0.017502339 | 2 | 0.5 |
| 1 | 2002 | 3 | 1 | 1 | 0 | 1.102925652 | 0.019080923 | 2 | 0.5 |
| 1 | 2003 | 3 | 1 | 1 | 0 | 0.971420869 | 0.019998649 | 2 | 0.5 |
| 1 | 2004 | 3 | 1 | 1 | 0 | 1.439353738 | 0.019084714 | 2 | 0.5 |
| 2 | 2005 | 3 | 1 | 1 | 0 | 0.968350994 | 0.027044329 | 2 | 0.5 |
| 2 | 2006 | 3 | 1 | 1 | 0 | 0.776625585 | 0.031272167 | 2 | 0.5 |
| 2 | 2007 | 3 | 1 | 1 | 0 | 0.871246208 | 0.025708974 | 2 | 0.5 |
| 2 | 2008 | 3 | 1 | 1 | 0 | 0.866988479 | 0.030195942 | 2 | 0.5 |
| 2 | 2009 | 3 | 1 | 1 | 0 | 0.715265818 | 0.044121831 | 2 | 0.5 |
| 2 | 2010 | 3 | 1 | 1 | 0 | 0.741946999 | 0.042159637 | 2 | 0.5 |
| 2 | 2011 | 3 | 1 | 1 | 0 | 1.062417451 | 0.030695665 | 2 | 0.5 |
| 2 | 2012 | 3 | 1 | 1 | 0 | 1.017615153 | 0.029396069 | 2 | 0.5 |
| 2 | 2013 | , | 1 | 1 | 0 | 0.997683607 | 0.028390886 | 2 | 0.5 |
| 2 | 2014 | 3 | 1 | 1 | 0 | 1.323666322 | 0.024088322 | 2 | 0.5 |
| 2 | 2015 | 3 | 1 | 1 | 0 | 1.275905076 | 0.022845987 | 2 |  |


| 2 | 2016 | 3 | 1 | 1 | 0 | 1.053788888 | 0.02 | 5396432 | 0.5 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2 | 2017 | 3 | 1 | 1 | 0 | 0.996858932 | 0.02 | 9855682 | 0.5 |  |
| 2 | 2018 | 3 | 1 | 1 | 0 | 1.223204376 | 0.02 | 68502242 | 0.5 |  |
| 2 | 2019 | 3 | 1 | 1 | 0 | 1.14646792 | 0.023 | 2 | 0.5 |  |
| 2 | 2020 | 3 | 1 | 1 | 0 | 1.042800763 | 0.02 | 8576992 | 0.5 |  |
| 2 | 2021 | 3 | 1 | 1 | 0 | 0.930732399 | 0.03 | 2219225 | 0.5 |  |
| 2 | 2022 | 3 | 1 | 1 | 0 | 1.273555912 | 0.03 | 9856689 | 0.5 |  |
| \# |  |  |  |  |  |  |  |  |  |  |
| \# | Cooperative survey 22_1g |  |  |  |  | If substitued for observer 2015 to 2022, identify one less record |  |  |  |  |
| \#2 | 2015 | 3 | 1 | 1 | 0 | 1.240344388 |  | 256823882 | 0.5 |  |
| \#2 | 2016 | 3 | 1 | 1 | 0 | 0.976402759 | 0.02 | 297941012 | 0.5 |  |
| \#2 | 2017 | 3 | 1 | 1 | 0 | 1.171793928 |  | 2 | 0.5 |  |
| \#2 | 2018 | 3 | 1 | 1 | 0 | 1.150314407 |  | 287750152 | 0.5 |  |
| \#2 | 2019 | 3 | 1 | 1 | 0 | 0.693938794 | 0.05 | 2946842 2 | 0.5 |  |
| \#2 | 2021 | 3 | 1 | 1 | 0 | 0.761920262 |  | 421154862 | 0.5 |  |
| \#2 | 2022 | 3 | 1 | 1 | 0 | 1.158591194 |  | 3 | 0.5 |  |
| \# |  |  |  |  |  |  |  |  |  |  |
| \# | Year:Area <br> Observer |  | interaction for model 22_1f |  |  |  |  |  |  |  |
| \# |  |  |  |  | index |  |  |  |  |  |
| \# | 1 | 1995 | 3 | 1 | 1 | $0 \quad 0.770819$ | 5563 | 0.139118217 | 2 | 0.5 |
| \# | 1 | 1996 | 3 | 1 | 1 | 00.779457 | 7772 | 0.135878349 | 2 | 0.5 |
| \# | 1 | 1997 | 3 | 1 | 1 | 00.776463 | 3237 | 0.127739124 | 2 | 0.5 |
| \# | 1 | 1998 | 3 | 1 | 1 | 00.930787 | 7475 | 0.091241829 | 2 | 0.5 |
| \# | 1 | 1999 | 3 | 1 | 1 | 00.886416 | 6941 | 0.098097874 | 2 | 0.5 |
| \# | 1 | 2000 | 3 | 1 | 1 | 00.859286 | 6215 | 0.139616133 | 2 | 0.5 |
| \# | 1 | 2001 | 3 | 1 | 1 | $0 \quad 1.241040$ | 792 | 0.12167127 | 2 | 0.5 |
| \# | 1 | 2002 | 3 | 1 | 1 | $0 \quad 1.260192$ | 2236 | 0.054214095 | 2 | 0.5 |
| \# | 1 | 2003 | 3 | 1 | 1 | $0 \quad 1.127970$ | 0188 | 0.067798981 | 2 | 0.5 |
| \# | 1 | 2004 | 3 | 1 | 1 | 01.713905 | 5535 | 0.054433946 | 2 | 0.5 |
| \# | 2 | 2005 | 3 | 1 | 1 | 01.046292 | 2434 | 0.070760737 | 2 | 0.5 |
| \# | 2 | 2006 | 3 | 1 | 1 | 00.767160 | 3355 | 0.130029491 | 2 | 0.5 |
| \# | 2 | 2007 | 3 | 1 | 1 | $0 \quad 0.857740$ | 745 | 0.103098468 | 2 | 0.5 |
| \# | 2 | 2008 | 3 | 1 | 1 | 00.866440 |  | 0.103461996 | 2 | 0.5 |
| \# | 2 | 2009 | 3 | 1 | 1 | $0 \quad 0.759251$ | 1496 | 0.140273096 | 2 | 0.5 |
| \# | 2 | 2010 | 3 | 1 | 1 | 00.773862 | 2187 | 0.134006209 | 2 | 0.5 |
| \# | 2 | 2011 | 3 | 1 | 1 | $0 \quad 1.055689$ | 9343 | 0.072813727 | 2 | 0.5 |
| \# | 2 | 2012 | 3 | 1 | 1 | 00.998007 | 7774 | 0.08009003 | 2 | 0.5 |


| \# | 2 | 2013 | 3 | 1 | 1 | 0 | 1.041574963 | 0.076681214 | 2 | 0.5 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| \# | 2 | 2014 | 3 | 1 | 1 | 0 | 1.299069357 | 0.047484405 | 2 | 0.5 |
| \# | 2 | 2015 | 3 | 1 | 1 | 0 | 1.23557422 | 0.05090457 | 2 | 0.5 |
| \# | 2 | 2016 | 3 | 1 | 1 | 0 | 1.057224682 | 0.06893231 | 2 | 0.5 |
| \# | 2 | 2017 | 3 | 1 | 1 | 0 | 0.948934847 | 0.08704369 | 2 | 0.5 |
| \# | 2 | 2018 | 3 | 1 | 1 | 0 | 1.179166576 | 0.056386774 | 2 | 0.5 |
| \# | 2 | 2019 | 3 | 1 | 1 | 0 | 1.100964664 | 0.083027709 | 2 | 0.5 |
| \# | 2 | 2020 | 3 | 1 | 1 | 0 | 1.008465927 | 0.07838537 | 2 | 0.5 |
| \# | 2 | 2021 | 3 | 1 | 1 | 0 | 0.956585096 | 0.085527924 | 2 | 0.5 |
| \# | 2 | 2022 | 3 | 1 | 1 | 0 | 1.289193041 | 0.053482771 | 2 | 0.5 |
| \# ${ }^{\text {a }}$ |  |  |  |  |  |  |  |  |  |  |
|  | dex | Year S | eas | fleet | Sex |  | urity index | cv abundanc |  | unit |
|  | 3 | 1985 | 3 | 1 | 1 | 0 | 1.628685686 | 0.031256541 | 2 | 0.5 |
|  | 3 | 1986 | 3 | 1 | 1 | 0 | 1.228858309 | 0.03860399 | 2 | 0.5 |
|  | 3 | 1987 | 3 | 1 | 1 | 0 | 0.955170913 | 0.051223518 | 2 | 0.5 |
|  | 3 | 1988 | 3 | 1 | 1 | 0 | 1.035770885 | 0.03950348 | 2 | 0.5 |
|  | 3 | 1989 | 3 | 1 | 1 | 0 | 1.076478459 | 0.03179462 | 2 | 0.5 |
|  | 3 | 1990 | 3 | 1 | 1 | 0 | 0.986817549 | 0.045649076 | 2 | 0.5 |
|  | 3 | 1991 | 3 | 1 | 1 | 0 | 0.904618567 | 0.047224701 | 2 | 0.5 |
|  | 3 | 1992 | 3 | 1 | 1 | 0 | 0.917176073 | 0.047355474 | 2 | 0.5 |
|  | 3 | 1993 | 3 | 1 | 1 | 0 | 0.914494509 | 0.05332578 | 2 | 0.5 |
|  | 3 | 1994 | 3 | 1 | 1 | 0 | 0.808572288 | 0.051417951 | 2 | 0.5 |
|  | 3 | 1995 | 3 | 1 | 1 | 0 | 0.77981996 | 0.055409819 | 2 | 0.5 |
|  | 3 | 1996 | 3 | 1 | 1 | 0 | 0.779120743 | 0.055920142 | 2 | 0.5 |
|  | 3 | 1997 | 3 | 1 | 1 | 0 | 1.050514781 | 0.042865274 | 2 | 0.5 |
|  | 3 | 1998 | 3 | 1 | 1 | 0 | 1.214100014 | 0.042009806 | 2 | 0.5 |
| \# |  |  |  |  |  |  |  |  |  |  |
| \#\#\# Number of length frequency matrices |  |  |  |  |  |  |  |  |  |  |
| \#3 |  |  |  |  |  |  |  |  |  |  |
| 2 |  |  |  |  |  |  |  |  |  |  |
| \#\# Number of rows in each matrix |  |  |  |  |  |  |  |  |  |  |
| 3832 |  |  |  |  |  |  |  |  |  |  |
| \#33 |  |  |  |  |  |  |  |  |  |  |
| \#\# Number of bins in each matrix (columns of size data) |  |  |  |  |  |  |  |  |  |  |
| 1717 |  |  |  |  |  |  |  |  |  |  |
| \#17 |  |  |  |  |  |  |  |  |  |  |
| \#\# SIZE COMPOSITION DATA FOR ALL FLEETS |  |  |  |  |  |  |  |  |  |  |

## \#\# SIZE COMP LEGEND

\#\# Sex: $1=$ male, $2=$ female, $0=$ both sexes combined
\#\# Type of composition: $1=$ retained, $2=$ discard, $0=$ total composition
\#\# Maturity state: $1=$ immature, $2=$ mature, $0=$ both states combined
\#\# Shell condition: $1=$ new shell, $2=$ old shell, $0=$ both shell types combined
\#\# Type 1 effective sample: Nsamp
\#\# Retain catch size comp
\#\# AEP updated the effective Ns
\#\#Year, Seas, Fleet, Sex, Type, Shell, Maturity, Nsamp, DataVec

| 19853 | 11 | 1 | $0 \quad 0$ | $57 \quad 0.000000$ | 0.000000 | 0.000000 | 0.000000 | 0.000000 | 0.002122 | 0.034669 | 0.103747 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0.158923 | 0.156292 |  | 0.157127 | 0.133423 | 0.108521 | 0.0615450 | 0.038431 | 0.020136 | 25065 |  |  |
| 19863 | 11 | 1 | 00 | 110.000000 | 0.000000 | 0.000000 | 0.000000 | 0.000000 | 0.000635 | 0.030377 | 0.143149 |
| 0.183126 | 0.212534 |  | 0.136044 | 0.114523 | 0.075306 | 0.038519 | 0.039528 | 0.016971 | 9288 |  |  |
| 19873 |  | 1 | 00 | 610.000000 | 0.000000 | 0.003518 | 0.000000 | 0.000550 | 0.003212 | 0.070524 | 0.162974 |
| 0.240875 | 0.168335 |  | 0.132893 | 0.076020 | 0.050479 | 0.037065 | 0.026783 | 0.011753 | 5022 |  |  |
| 19883 | 1 | 1 | $0 \quad 0$ | 3520.000000 | 0.000000 | 0.0 .000000 | - 0.000000 | 0.000250 | 0.004988 | 0.043836 | 0.121611 |
| 0.173481 | 0.179156 |  | 0.161137 | 0.132840 | 0.073217 | 0.043037 | 0.025108 | 0.020902 | 0437 |  |  |
| 19893 | 11 | 1 | 00 | 7920.000000 | 0.000000 | 0.0 .000000 | 0.000066 | 0.000195 | 0.008435 | 0.108452 | 0.234714 |
| 0.191637 | 0.123151 |  | 0.094370 | 0.075312 | 0.057163 | $0.038218 \quad 0.0$ | 0.0262850 | 0.019802 | 2201 |  |  |
| 19903 | 11 | 1 | 00 | 1630.000000 | 0.000052 | 20.000052 | 20.000000 | 0.000340 | 0.005531 | 0.079874 | 0.226018 |
| 0.260315 | 0.183031 |  | 0.112587 | 0.066439 | 0.038093 | 0.016649 | 0.005442 | 0.002781 | 2796 |  |  |
| 19913 | 1 | 1 | 00 | 140 0.000000 | 0.000000 | 0.0 .000000 | 0.000000 | 0.000287 | 0.006172 | 0.074641 | 0.201726 |
| 0.233318 | 0.206834 |  | 0.127877 | 0.072609 | 0.040713 | 0.018307 | 0.009776 | 0.004928 | 2812 |  |  |
| 19923 | 11 | 1 | 00 | 490.000000 | 0.000000 | 0.000056 | 0.000120 | 0.000452 | 0.005204 | 0.074976 | 0.188394 |
| 0.240279 | 0.192046 |  | 0.126742 | 0.085203 | 0.048454 | 0.024934 | 0.008597 | 0.002697 | 1846 |  |  |
| 19933 | 1 | 1 | 00 | $340 \quad 0.000000$ | 0.000000 | 0.0 .000000 | - 0.000000 | 0.001271 | 0.006339 | 0.057846 | 0.227652 |
| 0.263149 | 0.193126 |  | 0.115423 | 0.061702 | 0.041289 | 0.019439 | 0.008024 | 0.001523 | 03216 |  |  |
| 19943 | 11 | 1 | 00 | 3190.000000 | 0.000000 | 0.0 .000000 | 0.0 .000000 | 0.000000 | 0.005146 | 0.056488 | 0.187163 |
| 0.253136 | 0.241073 |  | 0.112635 | 0.071796 | 0.038426 | 0.016716 | 0.011135 | 0.003629 | 2656 |  |  |
| 19953 | 11 | 1 | 00 | 8790.000000 | 0.000000 | 0.0 .000367 | 0.000000 | 0.000132 | 0.002554 | 0.053244 | 0.174310 |
| 0.237169 | 0.205691 |  | 0.131577 | 0.086227 | 0.054200 | 0.029541 | 0.014691 | 0.006267 0 | 4031 |  |  |
| 19963 | 1.1 | 1 | 00 | 5470.000000 | 0.000509 | 9.000000 | ) 0.002673 | 0.004458 | 0.010646 | 0.076046 | 0.176767 |
| 0.219822 | 0.183488 |  | 0.129821 | 0.083593 | 0.049809 | 0.029215 | 0.022160 0. | 0.009716 | 01277 |  |  |
| 19973 | 1 | 1 | 00 | 5380.000165 | 0.000000 | 0.000000 | 0.000000 | 0.000546 | 0.005501 | 0.067013 | 0.195912 |
| 0.241333 | 0.187580 |  | 0.126671 | 0.078708 | 0.047831 | 0.025562 | 0.0149750 | 0.006349 | 01855 |  |  |


| 19983 | 11 | 1 | 00 |
| :---: | :---: | :---: | :---: |
| 0.237512 | 0.195717 |  | 0.131940 |
| 19993 | 1 | 1 | 00 |
| 0.256172 | 0.191463 |  | 0.123275 |
| 20003 | 11 | 1 | 00 |
| 0.243407 | 0.197233 |  | 0.140484 |
| 20013 | 1 | 1 | 00 |
| 0.254416 | 0.209148 |  | 0.150723 |
| 20023 | 1 | 1 | 00 |
| 0.231895 | 0.215249 |  | 0.146064 |
| 20033 | 1 | 1 | 00 |
| 0.198660 | 0.225076 |  | 0.168816 |
| 20043 | 11 | 1 | 00 |
| 0.215850 | 0.214303 |  | 0.163649 |
| 20053 | 1 | 1 | 00 |
| 0.199591 | 0.218250 |  | 0.176555 |
| 20063 | 11 | 1 | 00 |
| 0.179791 | 0.184967 |  | 0.175434 |
| 20073 | 1 | 1 | 00 |
| 0.188152 | 0.182646 |  | 0.168733 |
| 20083 | 1 | 1 | 00 |
| 0.201275 | 0.170907 |  | 0.164015 |
| 20093 | 11 | 1 | 00 |
| 0.178121 | 0.204593 |  | 0.179856 |
| 20103 | 11 | 1 | 00 |
| 0.190138 | 0.186530 |  | 0.154632 |
| 20113 | 11 | 1 | 00 |
| 0.169147 | 0.214179 |  | 0.181341 |
| 20123 | 11 | 1 | 00 |
| 0.167506 | 0.197858 |  | 0.161194 |
| 20133 | 11 | 1 | 00 |
| 0.166734 | 0.180076 |  | 0.184391 |
| 20143 | 1 | 1 | 00 |
| 0.171561 | 0.183012 |  | 0.168880 |
| 20153 | 11 | 1 | 00 |
| 0.169842 | 0.197348 |  | 0.152410 |


| 5410.000000 | 0.000000 | $0 \quad 0.000000$ | 0.000000 | 0.000153 | 0.001613 | 0.058033 | 0.195363 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0.079974 | 0.046411 | 0.030546 | 0.015402 | 0.004854 0, | 0.002485 |  |  |
| 4630.000000 | $0 \quad 0.000000$ | 0.0 .000000 | 0.000000 | 0.000000 | 0.002647 | 0.056968 | 0.209816 |
| 0.073622 | 0.044721 | 0.023946 | $0.011020 \quad 0$. | 0.005430 0 | 0.000921 |  |  |
| 4360.000481 | 10.000000 | 0.000000 | 0.0 .000000 | 0.000000 | 0.002408 | 0.038199 | 0.187100 |
| 0.088336 | 0.054458 | 0.027952 | 0.012388 0. | 0.005379 0. | 0.002176 |  |  |
| 4880.000000 | 0.0 .000040 | 0.0 .000000 | 00.000000 | 0.000000 | 0.002185 | 0.043398 | 0.166360 |
| 0.084320 | 0.049034 | 0.024928 | 0.010970 0. | 0.0024530 | 0.002028 |  |  |
| 4060.000692 | 20.000000 | 0.0000000 | 0.000000 | 0.000000 | 0.001140 | 0.042702 | 0.173724 |
| 0.090496 | 0.052512 | 0.029190 | 0.012247 | 0.0028090 | 0.001280 |  |  |
| 4050.000000 | 0.0 .000000 | 0.000000 | 0.0 .000000 | - 0.000104 | 0.000939 | 0.025425 | 0.128996 |
| 0.127193 | 0.062420 | 0.035472 | 0.017291 | 0.0057260 | 0.003883 |  |  |
| $280 \quad 0.000000$ | 0.0 .000000 | 0.0 .000000 | 0.0 .000000 | 0.000000 | 0.000153 | 0.036696 | 0.127904 |
| 0.120783 | 0.069026 | 0.033788 | $0.016064 \quad 0$. | $0.001630 \quad 0$ | 0.000154 |  |  |
| 2660.000000 | $0 \quad 0.000000$ | $0 \quad 0.000000$ | 0.000000 | 0.000000 | 0.000885 | 0.018795 | 0.118321 |
| 0.132109 | 0.068852 | 0.035158 | 0.023218 | 0.004347 0, | 0.003920 |  |  |
| 2340.000000 | $0 \quad 0.000000$ | 00.000000 | 0.000000 | 0.000000 | 0.000266 | 0.016116 | 0.084749 |
| 0.156561 | 0.101305 | 0.053838 | 0.027473 | 0.011261 0 | 0.008238 |  |  |
| 1990.000317 | $7 \quad 0.000000$ | 00.000000 | 0.000000 | 0.000616 | 0.000000 | 0.023977 | 0.115069 |
| 0.124654 | 0.089646 | 0.056234 | 0.027344 | 0.0154020 | 0.007211 |  |  |
| 1970.000000 | 0.000000 | $0 \quad 0.000000$ | 0.0 .000000 | 0.000000 | 0.000886 | 0.012873 | 0.104580 |
| 0.131524 | 0.089417 | 0.069199 | 0.030247 | 0.0132940 | 0.011783 |  |  |
| $170 \quad 0.000000$ | 0.0000000 | 0.0 .000000 | 0.0 .000000 | 0.000000 | 0.000000 | 0.012998 | 0.085646 |
| 0.132916 | 0.096605 | 0.064687 | 0.026752 | $0.012521 \quad 0$ | 0.005305 |  |  |
| 1830.000424 | $4 \quad 0.000000$ | 0.0 .000000 | 0.000000 | 0.000000 | 0.000497 | 0.019071 | 0.124157 |
| 0.124061 | 0.080623 | 0.064508 | $0.031903 \quad 0$. | 0.012549 0. | 0.010908 |  |  |
| $160 \quad 0.000000$ | $0 \quad 0.000000$ | 0.0 .000000 | 0.0 .000000 | 0.000000 | 0.000000 | 0.006553 | 0.080423 |
| 0.118590 | 0.107631 | 0.063368 | $0.033478 \quad 0$. | 0.017831 0. | 0.007460 |  |  |
| 1870.000000 | 0.000000 | 00.000000 | 0.0 .000000 | 0.000000 | 0.000924 | 0.011670 | 0.080888 |
| 0.133335 | 0.105248 | 0.071755 | 0.041681 | 0.019324 0, | 0.008617 |  |  |
| 1930.000000 | $0 \quad 0.000000$ | 00.000000 | 0.0 .000000 | 0.000000 | 0.001621 | 0.015499 | 0.104071 |
| 0.127462 | 0.095836 | 0.060360 | $0.035295 \quad 0$. | 0.0189790 | 0.009676 |  |  |
| 1680.000000 | $0 \quad 0.000000$ | 00.000000 | 0.000000 | 0.000000 | 0.001431 | 0.022137 | 0.091465 |
| 0.121834 | 0.102642 | 0.069861 | $0.035479 \quad 0$. | 0.022149 0. | 0.009550 |  |  |
| $190 \quad 0.000000$ | 0.000000 | 0.0 .000000 | 0.0 .000000 | 0.000000 | 0.000000 | 0.011420 | 0.072221 |
| 0.136227 | 0.095458 | 0.076222 | 0.042626 0. | 0.025670 0. | 0.020557 |  |  |






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|  | 7 | 1 | 8 | 3 | 1 | 1 | 2004 | 1 |
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|  | 7 | 1 | 9 | 3 | 1 | 1 | 2004 | 2 |
|  | 7 | 1 | 10 | 3 | 1 | 1 | 2004 | 1 |
|  | 7 | 1 | 11 | 3 | 1 | 1 | 2004 | 3 |
|  | 7 | 1 | 13 | 3 | 1 | 1 | 2004 | 6 |
|  | 7 | 1 | 14 | 3 | 1 | 1 | 2004 | 3 |
|  | 8 | 1 | 10 | 3 | 1 | 1 | 2004 | 1 |
|  | 8 | 1 | 11 | 3 | 1 | 1 | 2004 | 7 |
|  | 8 | 1 | 12 | 3 | 1 | 1 | 2004 | 2 |
|  | 8 | 1 | 14 | 3 | 1 | 1 | 2004 | 1 |
|  | 9 | 1 | 11 | 3 | 1 | 1 | 2004 | 1 |
|  | 9 | 1 | 12 | 3 | 1 | 1 | 2004 | 4 |
|  | 9 | 1 | 13 | 3 | 1 | 1 | 2004 | 1 |
|  | 9 | 1 | 15 | 3 | 1 | 1 | 2004 | 1 |
|  | 9 | 1 | 16 | 3 | 1 | 1 | 2004 | 1 |
|  | 10 | 1 | 13 | 3 | 1 | 1 | 2004 | 1 |
|  | 13 | 1 | 15 | 3 | 1 | 1 | 2004 | 1 |
| \#Year4 |  |  |  |  |  |  |  |  |
|  | 1 | 1 | 10 | 4 | 1 | 1 | 2004 | 6 |
|  | 1 | 1 | 11 | 4 | 1 | 1 | 2004 | 1 |
|  | 2 | 1 | 8 | 4 | 1 | 1 | 2004 | 1 |
|  | 2 | 1 | 10 | 4 | 1 | 1 | 2004 | 1 |
|  | 2 | 1 | 11 | 4 | 1 | 1 | 2004 | 5 |
|  | 3 | 1 | 8 | 4 | 1 | 1 | 2004 | 3 |
|  | 3 | 1 | 9 | 4 | 1 | 1 | 2004 | 3 |
|  | 3 | 1 | 10 | 4 | 1 | 1 | 2004 | 3 |
|  | 3 | 1 | 11 | 4 | 1 | 1 | 2004 | 1 |
|  | 3 | 1 | 12 | 4 | 1 | 1 | 2004 | 1 |
|  | 3 | 1 | 14 | 4 | 1 | 1 | 2004 | 1 |
|  | 4 | 1 | 9 | 4 | 1 | 1 | 2004 | 1 |
|  | 4 | 1 | 10 | 4 | 1 | 1 | 2004 | 2 |
|  | 4 | 1 | 12 | 4 | 1 | 1 | 2004 | 1 |
|  | 4 | 1 | 14 | 4 | 1 | 1 | 2004 | 1 |
|  | 5 | 1 | 8 | 4 | 1 | 1 | 2004 | 1 |
|  | 5 | 1 | 10 | 4 | 1 | 1 | 2004 | 1 |
|  | 5 | 1 | 11 | 4 | 1 | 1 | 2004 | 4 |
|  | 5 | 1 | 12 | 4 | 1 | 1 | 2004 | 4 |


|  | 6 | 1 | 11 | 4 | 1 | 1 | 2004 | 1 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | 6 | 1 | 12 | 4 | 1 | 1 | 2004 | 2 |
| \#Year5 | 7 | 1 | 10 | 4 | 1 | 1 | 2004 | 2 |
|  | 7 | 1 | 12 | 4 | 1 | 1 | 2004 | 1 |
|  | 7 | 1 | 14 | 4 | 1 | 1 | 2004 | 2 |
|  | 8 | 1 | 14 | 4 | 1 | 1 | 2004 | 2 |
|  | 1 | 1 | 10 | 5 | 1 | 1 | 2004 | 2 |
|  | 2 | 1 | 11 | 5 | 1 | 1 | 2004 | 1 |
|  | 2 | 1 | 12 | 5 | 1 | 1 | 2004 | 1 |
|  | 2 | 1 | 16 | 5 | 1 | 1 | 2004 | 1 |
|  | 3 | 1 | 11 | 5 | 1 | 1 | 2004 | 1 |
|  | 3 | 1 | 13 | 5 | 1 | 1 | 2004 | 3 |
|  | 3 | 1 | 14 | 5 | 1 | 1 | 2004 | 1 |
|  | 5 | 1 | 14 | 5 | 1 | 1 | 2004 | 1 |
|  | 7 | 1 | 13 | 5 | 1 | 1 | 2004 | 1 |
|  | 7 | 1 | 14 | 5 | 1 | 1 | 2004 | 1 |
|  |  |  |  |  |  |  |  |  |
|  | 1 | 1 | 8 | 6 | 1 | 1 | 2004 | 1 |
|  | 1 | 1 | 9 | 6 | 1 | 1 | 2004 | 1 |
|  | 1 | 1 | 11 | 6 | 1 | 1 | 2004 | 1 |
|  | 1 | 1 | 12 | 6 | 1 | 1 | 2004 | 1 |
|  | 1 | 1 | 13 | 6 | 1 | 1 | 2004 | 2 |
|  | 2 | 1 | 11 | 6 | 1 | 1 | 2004 | 2 |
|  | 2 | 1 | 14 | 6 | 1 | 1 | 2004 | 1 |
|  | 3 | 1 | 9 | 6 | 1 | 1 | 2004 | 1 |
|  | 4 | 1 | 10 | 6 | 1 | 1 | 2004 | 2 |

\#\# eof
9999
\#\#\#\#\#\#\#\#\#\#\#\#\#\#\#\#
EAG22.1e2 prj file

References




## B. Input data files for model 22.1e2:

## WAG22.1e2 ctl file

\# WAG22_1e2 Update
\# Controls for leading parameter vector theta
\# LEGEND FOR PRIOR:
\# 0 = uniform, $1=$ normal, $2=$ lognormal, $3=$ beta, $4=$ gamma
\# ntheta
9
\# ival lb ub phz prior p1 p2 \# parameter \#
$\begin{array}{llllllll}0.22 & 0.01 & 1.0 & -3 & 2 & 0.18 & 0.04 & \text { \# M }\end{array}$
7.607526807-10.0 20.0 $\quad 1 \quad 0 \quad-10.0 \quad 20.0 \quad$ \# ln R0, logarithm of unfished recruits, from my model
$\begin{array}{cccccccc}12.0 & -10.0 & 20.0 & -3 & 0 & -10.0 & 20.0 \quad \text { \# } \ln \text { Rini, logarithm of initial recruitment(syr) }\end{array}$
$8.0 \quad-10.0 \quad 20.0 \quad-1 \quad 0 \quad-10.0 \quad 20.0 \quad$ \# One par freed, ln Rbar, logarithm of average recruits(syr+1,nyr)

```
110.0 103.0}165.0 -2 1 2 7 72.5 7.25 # recruitment expected value, ra, Expected value of recruitment distributio
0.928697601 0.001 20.0 3 0 0 0.1 5.0 ###1 (any run change after first gmacs run0 recruitment scale (variance component)
-0.693147181-10.0 0.75-1 0 - -10.0 0.75 # ln (SigmaR),
0.73 0.2 1.0 
0.001 0.0 1.0 - -3 3 1.01 1.01 # recruitment autocorrelation (only used if R is constrained by a S-R relationship)
# weight-at-length input method (1 = allometry [w_1 = a*1^b], 2 = vector by sex)
2
#a, in kg
# 1.445E-07
#b
# 3.281126995
# Male weight-at-length
0.581515707 
1.90953096 2.129425732 2.366336933 2.620980182 
#
# Proportion mature by sex, males
0. 0. 0. 1. 1. 1. 1. 1. 1. 1. 1. 1. 1. 1. 1. 1.1
# Proportion legal by sex, males
0.0.0.0.0. 0. 0. 1. 1. 1. 1. 1. 1. 1. 1. 1. 1.
## GROWTH PARAM CONTROLS ##
## Two lines for each parameter if split sex, one line if not ##
# Use growth transition matrix option (1=read in growth-increment matrix; 2=read in size-transition; 3=gamma distribution for size-increment; 4=gamma
distribution for size after increment) (1 to 8 options available)
# option 8 is normal distributed growth incrment, size after incrment is normal
8
# growth increment model (0=prespecified; 1=alpha/beta; 2=estimated by size-class;3=pre-specified/emprical)
1
# molt probability function (0=pre-specified; 1=flat;2=declining logistic)
2
\# maximum size-class (males then females)
\# Maximum size-class for recruitment(males then females)
```

```
5
## number of size-increment periods
1
## Year(s) size-increment period changes (blank if no changes)
## number of molt periods
1
## Year(s) molt period changes (blank if no changes)
## Beta parameters are relative to a base level (1=Yes;0=no)
1 #
# Growth parameters
# ival lb ub phz prior p1 p2 # parameter #
    25.41312879
```



```
    3.6720539}00.0
    141.25406}655.0 165.0 7 7 0 0 0.0 999.0 
    0.105950464 -0.1 2.0 7 7 0 0.0
\# The custom growth-increment matrix
\# custom molt probability matrix
\#\# SELECTIVITY CONTROLS
\#\# Selectivity P(capture of all sizes). Each gear must have a selectivity and a \#\#
\#\# retention selectivity. If a uniform prior is selected for a parameter then the \#\#
\#\# lb and ub are used (p1 and p2 are ignored)
\#\#
\#\# LEGEND
\#\#
\#\# sel type: \(0=\) parametric (nclass), \(1=\) indiviudal parameter for each class(nclass), \#\#
\#\# \(\quad 2=\) logistic ( 2 , inflection point and slope), \(3=\operatorname{logistic} 95(2,50 \%\) and \(95 \%\) selection), \(4=\) double normal ( 3 parameters),
\#\#
\#\# 5: Flat equal to zero (1 parameter; phase must be negative), UNIFORM1
\#\# 6: Flat equal to one (1 parameter; phase must be negative), UNIFORM0
\#\# 7: Flat-topped double normal selectivity (4 parameters)
\#\# 8: Declining logistic selectivity with initial values ( \(50 \%\) and \(95 \%\) selection plus extra)
\#\# Extra (type 1): number of selectivity parameters to be estimated
\#\# gear index: use + ve for selectivity, -ve for retention \#\#
```

```
## sex dep: 0 for sex-independent, 1 for sex-dependent
##
## ivector for number of year blocks or nodes ##
## Gear-1 Gear-2
## PotFishery Trawl Byc
    2 1 # selectivity time periods
    0 0 # set 0 for male only fishery, sex specific selectivity, 0 for sex independent selectivity
    2 % # male selectivity type model (flat equal to zero, 1 parameter) or logistic or double normal etc.
    0
    0 0 # extra paramters for each pattern
## Gear-1 Gear-2
    1}10\mathrm{ # retention time periods
    0 0 # set 0 for male only fishery, sex specific retention
    2 % male retention type model (flat equal to one, 1 parameter)
    1 0 # male retention flag (0 = no, 1 = yes)
    0 0 # extra
    1 # determines if maximum selectivity at size is forced to equal 1 or not
```

\#\# Selectivity P(capture of all sizes)
\#\# gear par sel phz start end \#\#
\# index index par sex ival lb ub prior p1 p2 mirror period period \#\#
\#\# Gear-1
$\begin{array}{lllllllllllll}1 & 1 & 1 & 0 & 135.5311804 & 105.0 & 180.0 & 0 & 100.0 & 190.0 & 3 & 1960 & 2004 \text { \#set sex } 0 \text { for male only fishery }\end{array}$
$\begin{array}{llllllllllll}1 & 2 & 2 & 0 & 18.14843141 & 0.01 & 20.0 & 0 & 0.1 & 50.0 & 3 & 1960 \\ 2004\end{array}$
$1 \begin{array}{lllllllllllll}1 & 3 & 1 & 0 & 134.8771653 & 105.0 & 180.0 & 0 & 100.0 & 190.0 & 3 & 2005 & 2022 \text { \# }\end{array}$
$1 \begin{array}{lllllllllll} & 4 & 2 & 0 & 7.309517947 & 0.01 & 20.0 & 0 & 0.1 & 50.0 & 3\end{array} 2005 \quad 2022$ \#
\# Gear-2
$\begin{array}{lllllllllllll}2 & 5 & 1 & 0 & 1.00 & 0.99 & 1.02 & 0 & 10.0 & 200.0 & -3 & 1960 & 2022\end{array}$
\#\# Retained
\#\# gear par sel phz start end
\# index index par sex ival lb ub prior p1 p2 mirror period period
\# Gear-1
$\begin{array}{lllllllllllll}-1 & 6 & 1 & 0 & 136.0107883 & 105.0 & 180.0 & 0 & 100.0 & 190.0 & 3 & 1960 & 2022 \text { \# }\end{array}$
$\begin{array}{llllllllllll}-1 & 7 & 2 & 0 & 1.873573427 & 0.0001 & 20.0 & 0 & 0.1 & 50.0 & 3 & 1960 \\ 2022 \#\end{array}$
\# Gear-2
$\begin{array}{lllllllllllll}-2 & 8 & 1 & 0 & 1.00 & 0.99 & 1.01 & 0 & 10.0 & 200.0 & -3 & 1960 & 2022\end{array}$
\# Number of asyptotic parameters
1
\# Fleet Sex Year ival lb ub phz
$1 \quad 1 \quad 1960 \quad 0.000001 \quad 0 \quad 1 \quad-3$
\#\# PRIORS FOR CATCHABILITY
\#\# If a uniform prior is selected for a parameter then the lb and ub are used (p1 \#\#
\#\# and p2 are ignored). ival must be >0
\#\# only allowed to use uniform or lognormal prior
\#\# if anlytic q estimation step is chosen, turn off estimating q by changing the estimation phase to be -ve
\#\# LEGEND \#\#
\#\# prior: $0=$ uniform, $1=$ normal, $2=$ lognormal, $3=$ beta, $4=$ gamma $\quad$ \#\#
\#\# SURVEYS/INDICES ONLY
\#\# fishery and observer CPUE
\#\# Analytic ( $0=$ not analytically solved q, use uniform or lognormal prior;
\#\# 1= analytic),
\#\# Lambda =multilier for iput CV, Emphasis = multiplier for likelihood
\#\# ival lb ub phz prior p1 p2 Analytic? LAMBDA Emphasis
$0.0010410670 .00000010 .01 \quad 1 \quad 0 \quad 0.0 \quad 1.0 \quad 0 \quad 1 \quad 1 \quad 1 \quad$ \# observer cpue index 1995-2004
$0.0010929650 .0000001 \quad 0.01 \quad 1 \quad 0 \quad 0.0 \quad 1.0 \quad 0 \quad 1 \quad 1 \quad 1 \quad$ \# observer cpue index 2005-2022
$\begin{array}{lllllllllll}0.000729841 & 0.0000001 & 0.01 & 1 & 0 & 0.0 & 1.0 & 0 & 1 & 1 & \text { \# fishery cpue index 1985-1998 }\end{array}$
\#\# if a uniform prior is specified then use lb and ub rather than p 1 and p 2
\#\# ADDITIONAL CV FOR SURVEYS/INDICES
\#\# If a uniform prior is selected for a parameter then the lb and ub are used (p1 \#\#
\#\# and p2 are ignored). ival must be $>0$, lb should be>0
\#\# LEGEND \#\#
\#\# prior type: $0=$ uniform, $1=$ normal, $2=$ lognormal, $3=$ beta, $4=$ gamma $\quad \# \#$
\#\# ival lb ub phz prior p1 p2
$0.0001439930 .0000001 \quad 0.5 \quad 6 \quad 0 \quad 0.5 \quad 100$ \# obs CPUE additional CV adjusted for abundance in 1000s
$\begin{array}{lllllll}0.000223956 & 0.0000001 & 0.5 & 6 & 0 & 0.5 & 100 \text { \# obs CPUE additional CV adjusted for abundance in 1000s }\end{array}$
$0.0001182240 .0000001 \quad 0.5 \quad 6 \quad 0 \quad 0.5 \quad 100$ \# fishery CPUE additional CV adjusted for abundance in 1000s
\#\#\# Pointers to how the the additional CVs are used ( 0 ignore; $>0$ link to one of the parameters
123
\#\#\#\#
\#\# if a uniform prior is specified then use lb and ub rather than p 1 and p 2
\#\#PENALTIES FOR AVERAGE FISHING MORTALITY RATE FOR EACH GEAR
\#\#
\#\# Trap Trawl
\#\# Male F, Female F, early_phasepenalty_sd, later_phasepenalty_sd, meanmaleF_phase, meanfemaleF_phase,
\#\# lb meanF, ub meanF,lbannualmaleF(F_dev), ubannual maleF(F_dev), lbannualfemaleF(F_dev), ubannual femaleF(F_dev)

```
##
## Mean_F Fema-Offset STD_PHZ1 STD_PHZ2 PHZ_M PHZ_F Lb Ub Lb Ub Lb Ub
    0.499150563
    0.000281948
## OPTIONS FOR SIZE COMPOSTION DATA
## One column for each data matrix
## LEGEND
##
## Likelihood: 1 = Multinomial with estimated/fixed sample size
##
##
## 2 = Robust approximation to multinomial
## 3 = logistic normal (NIY) ##
## 4 = multivariate-t (NIY)
## 5 = Dirichlet
##
## AUTO TAIL COMPRESSION
## pmin is the cumulative proportion used in tail compression
##
# ret to
#
    1 1 # Type of likelihood
    0}00\mathrm{ # Auto tail compression (pmin)
    1 1 # Initial value for effective sample size multiplier
    -4 -4 # Phz for estimating effective sample size (if appl.)
    12 # Composition aggregator if you put 1 for each it will merge, do not merge (why merge)
    #
    1 1 # Set to 2 for survey-like predictions; 1 for catch-like predictions
    #
    0.447178761356032 0.568777746325553 # Emphasis AEP for Dritchlet (Ret, Tot, multiplier of stage1 ESS)
    1 1 #AMBDA 0 to ignore the length comp
## TIME VARYING NATURAL MORTALIIY RATES ##
## Type: 0 = constant natural mortality 
##
## 1 = Random walk (deviates constrained by variance in M) ##
## 2 = Cubic Spline (deviates constrained by nodes & node-placement) ##
## 3 = Blocked changes (deviates constrained by variance at specific knots) ##
## 4 = Changes in pre-specified blocks
    ##
## 5 = Changes in some knots ##
## 6 = Changes in Time blocks ##
0 # M type
## M is relative (YES=1;NO=0)
```

```
## Phase of estimation
3
## STDEV in m_dev for Random walk
0.25
## Number of nodes for cubic spline or number of step-changes for option 3
1
#0
## Year position of the knots (vector must be equal to the number of nodes)
1960
## number of breakpoints in M by size (keep it at 0)
0
# line groups for breakpoint
8
## Specific initial values for the natural mortality devs (0-no, 1=yes)
## 1
## ival 
## TAGGING controls CONTROLS
1
# emphasis on tagging data (1 =use tag LH, 0=ignore)
## Maturity specific natural mortality
# maturity specific natural mortality? (yes = 1; no = 0; only for use if nmature > 1)
0
\begin{tabular}{llllllllll} 
\#\# & ival & lb & ub & phz & prior & p 1 & p 2 & \# parameter & \#\# \\
0 & -1 & 1 & -1 & 0 & 1 & 1 & &
\end{tabular}
## OTHER CONTROLS
#
1960 # First year of recruitment estimation,rec_dev.
2022 # last year of recruitment estimation, rec_dev
1 # phase for recruitment estimation,
-2 # phase for recruitment sex-ratio estimation
0.5 # Initial value for Expected sex-ratio
-3 # Phase for initial recruitment estimation, rec_ini phase
# # VERBOSE FLAG (0 = off, 1 = on, 2 = objective func; 3 diagnostics)
0 # Initial conditions (0 = Unfished, 1 = Steady-state fished, 2 = Free parameters, 3 = Free parameters (revised))
1 # Lambda (proportion of mature male biomass for SPR reference points).
0 # Stock-Recruit-Relationship (0 = none, 1 = Beverton-Holt)
```

10 \# Maximum phase (stop the estimation after this phase), 10 Maximum phase. If you put 1 it will stop after phase 1
-1 \# Maximum number of function calls, if 1 , stop at fn 1 call; if -1 , run as long as it takes
1 \# Calculate reference points ( $0=$ no)
200 \# Year to compute equilibria
\#\# EMPHASIS FACTORS (CATCH)
\#ret_male tot_male Groundfish
$4 \quad 2 \quad 1$
\#\# EMPHASIS FACTORS (Priors) by fleet: fdev_total, Fdov_total, Fdev_year, Fdov_year
000.0010 \# Pot fishery
000.0010 \# Groundfish
\#\# EMPHASIS FACTORS (Priors)
\#\#


## WAG22.1e2 dat file

1. \#======================================================================1
2. \# Gmacs Main Data File Version 1.1: WAGModel 22.1e2Update up to 2022/23 data
3. \# data from WAG22_9C are used
4. \# GEAR_ INDEX DESCRIPTION
5. \# 1 : Pot fishery Retained catch
6. \# 2 : Pot fishery total catch
7. \# 3 : Trawl bycatch
8. \# 4 : Observer CPUE
9. \# 5 : Fishery CPUE
10. 
11. \# Fisheries: 1 Pot Fishery, 2 Pot Total
12. \# Cooperative Survey:
13. 1960 \# initial (start year)
14. 2022 \# terminal (end year)
15. \#2023 \# Projection year (for forecast, OFL and ABC calculation)
16. 6 \# Number of seasons: season 1 for N est, season 2 for Jul 1 to MidFishing, season 3 for inst.remove C, season 4 for to spawning time,Feb15, season 5 for inst remove byc\&estimate MMB, season 6 for remaining time to June 30 and R enter
17. 2 \# Number of distinct data groups or number of fleets (pot fishing, groundfish fishing)
18. 1 \# Number of sexes (males)
19. 1 \# Number of shell condition types
20. 1 \# Number of maturity types
21. 17 \# Number of size-classes in the model
22. 6 \# Season when recruitment occurs,end of year before growth
23. 6 \# Season when molting and growth occur, end of year after recruitment
24. 5 \# Season to calculate MMB
25. 1 \# Season for N output
26. \# maximum size-class (males then females)
27. 17
28. \# size_breaks (a vector giving the break points between size intervals with dimension nclass+1, lower limits of bins)
29. $100.5 \quad 105.5110 .5115 .5120 .5125 .5130 .5135 .5140 .5145 .5150 .5155 .5160 .5165 .5170 .5175 .5180 .5185 .5$
30. \# Natural mortality per season input type ( $1=$ vector by season, $2=$ matrix by season/year)
31. 2
32. \# Proportion of the total natural mortality to be applied each season (each row must add to 1 )
33. \# 1 Start biological year (Jul 1) instantaneous N estimation
34. \# 2 to mid fishing time
35. \# 3 instantanous C removal
36. \# 4 to spawning time
37. \# 5 instantaneous byc removal and estimate MMB
38. \# 6 Rest of the period of non fishing from Feb 15 to June 30
39. \#
40. \#Ins N Jul1-MidFish Inst C MidFish-15Feb Ins byc Rest upto end

| 41. | 0. | 0.16666667 | 0. | 0.46073059 | 0. | 0.37260274 | $\# 1960$ |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 42. | 0. | 0.16666667 | 0. | 0.46073059 | 0. | 0.37260274 | $\# 1961$ |
| 43. | 0. | 0.16666667 | 0. | 0.46073059 | 0. | 0.37260274 | $\# 1962$ |
| 44. | 0. | 0.16666667 | 0. | 0.46073059 | 0. | 0.37260274 | $\# 1963$ |
| 45. | 0. | 0.16666667 | 0. | 0.46073059 | 0. | 0.37260274 | $\# 1964$ |
| 46. | 0. | 0.16666667 | 0. | 0.46073059 | 0. | 0.37260274 | $\# 1965$ |
| 47. | 0. | 0.16666667 | 0. | 0.46073059 | 0. | 0.37260274 | $\# 1966$ |
| 48. | 0. | 0.16666667 | 0. | 0.46073059 | 0. | 0.37260274 | $\# 1967$ |
| 49. | 0. | 0.16666667 | 0. | 0.46073059 | 0. | 0.37260274 | $\# 1968$ |
| 50. | 0. | 0.16666667 | 0. | 0.46073059 | 0. | 0.37260274 | $\# 1969$ |
| 51. | 0. | 0.16666667 | 0. | 0.46073059 | 0. | 0.37260274 | $\# 1970$ |
| 52. | 0. | 0.16666667 | 0. | 0.46073059 | 0. | 0.37260274 | $\# 1971$ |
| 53. | 0. | 0.16666667 | 0. | 0.46073059 | 0. | 0.37260274 | $\# 1972$ |


| 54. 0. | 0.16666667 | 0. | 0.46073059 | 0. 0.37260274 | \#1973 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 55. 0. | 0.16666667 | 0. | 0.46073059 | 0.0 .37260274 | \#1974 |
| 56. 0 . | 0.16666667 | 0. | 0.46073059 | 0.0 .37260274 | \#1975 |
| 57. 0. | 0.16666667 | 0. | 0.46073059 | 0.0 .37260274 | \#1976 |
| 58. 0. | 0.16666667 | 0. | 0.46073059 | 0.0 .37260274 | \#1977 |
| 59. 0. | 0.16666667 | 0. | 0.46073059 | 0.0 .37260274 | \#1978 |
| 60. 0. | 0.16666667 | 0. | 0.46073059 | 0.0 .37260274 | \#1979 |
| 61. 0. | 0.16666667 | 0. | 0.46073059 | 0.0 .37260274 | \#1980 |
| 62. 0. | 0.62739726 | 0. | 0.0 | 0.37260274 \#1 |  |
| 63. 0. | 0.564383562 | 0. | 0.063013698 | 0.0 .37260274 | \#1982 |
| 64. 0. | 0.578082192 | 0. | 0.049315068 | 0.0 .37260274 | \#1983 |
| 65. 0. | 0.62739726 | 0. | 0.0 | 0.37260274 \# |  |
| 66. 0 . | 0.62739726 | 0. | 0.0 | 0.37260274 |  |
| 67. 0. | 0.62739726 | 0. | 0.0 | 0.37260274 \#1986 |  |
| 68. 0. | 0.62739726 | 0. | 0.0 | 0.37260274 \#19 |  |
| 69. 0. | 0.62739726 | 0. | 0.0 | 0.37260274 |  |
| 70. 0. | 0.62739726 | 0. | 0.0 | 0.37260274 \#19 |  |
| 71. 0. | 0.62739726 | 0. | 0.0 | 0.37260274 \#19 |  |
| 72. 0. | 0.62739726 | 0. | 0.0 | 0.37260274 |  |
| 73. 0. | 0.62739726 | 0. | 0.0 | 0.37260274 |  |
| 74. 0. | 0.62739726 | 0. | 0.0 . | 0.37260274 \#19 |  |
| 75. 0. | 0.62739726 | 0. | $0 . \quad 0$. | 0.37260274 \#19 |  |
| 76. 0. | 0.62739726 | 0. | 0.0 | 0.37260274 |  |
| 77. 0. | 0.62739726 | 0. | 0.0 . | 0.37260274 \#19 |  |
| 78. 0. | 0.62739726 | 0. | 0.0 | 0.37260274 \#1997 |  |
| 79. 0. | 0.62739726 | 0. | 0.0 | 0.37260274 \#19 |  |
| 80. 0. | 0.62739726 | 0. | 0.0 | 0.37260274 \#19 |  |
| 81. 0. | 0.516438356 | 0. | 0.110958904 | 0.0 .37260274 | \#2000 |
| 82. 0. | 0.435616438 | 0. | 0.191780822 | 0.0 .37260274 | \#2001 |
| 83. 0. | 0.405479452 | 0. | 0.221917808 | 0.0 .37260274 | \#2002 |
| 84. 0. | 0.364383562 | 0. | 0.263013698 | 0.0 .37260274 | \#2003 |
| 85. 0. | 0.317808219 | 0. | 0.309589041 | 0.0 .37260274 | \#2004 |
| 86. 0. | 0.430136986 | 0. | 0.197260274 | 0.0 .37260274 | \#2005 |
| 87. 0. | 0.494520548 | 0. | 0.132876712 | 0.0 .37260274 | \#2006 |
| 88. 0. | 0.530136986 | 0. | 0.097260274 | 0.0 .37260274 | \#2007 |
| 89. 0. | 0.505479452 | 0. | 0.121917808 | 0.0 .37260274 | \#2008 |
| 90. 0. | 0.524657534 | 0. | 0.102739726 | 0.0 .37260274 | \#2009 |


| 91. | 0. | 0.436986301 | 0. | 0.190410959 | 0. | 0.37260274 | $\# 2010$ |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 92. | 0. | 0.471232877 | 0. | 0.156164383 | 0. | 0.37260274 | $\# 2011$ |
| 93. | 0. | 0.504109589 | 0. | 0.123287671 | 0. | 0.37260274 | $\# 2012$ |
| 94. | 0. | 0.509589041 | 0. | 0.117808219 | 0. | 0.37260274 | $\# 2013$ |
| 95. | 0. | 0.501369863 | 0. | 0.126027397 | 0. | 0.37260274 | $\# 2014$ |
| 96. | 0. | 0.463013699 | 0. | 0.164383561 | 0. | 0.37260274 | $\# 2015$ |
| 97. | 0. | 0.42739726 | 0. | 0.2 | 0. | 0.37260274 | $\# 2016$ |
| 98. | 0. | 0.383561644 | 0. | 0.243835616 | 0. | 0.37260274 | $\# 2017$ |
| 99. 0. | 0.394520548 | 0. | 0.232876712 | 0. | 0.37260274 | $\# 2018$ |  |
| 100. 0. | 0.456164384 | 0. | 0.171232876 | 0. | 0.37260274 | $\# 2019$ |  |
| 101. 0. | 0.490410959 | 0. | 0.136986301 | 0. | 0.37260274 | $\# 2020$ |  |
| 102. 0. | 0.493150685 | 0. | 0.134246575 | 0. | 0.37260274 | $\# 2021$ |  |
| 103. 0. | 0.41369863 | 0. | 0.21369863 | 0. | 0.37260274 | $\# 2022$ |  |


| 128. 1986 | 3 | 1 | 1 | 4271.830 .03161 | 1 | 1 | 0 | 0.2 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 129. 1987 | 3 | 1 | 1 | 2535.350 .03161 | 1 | 1 | 0 | 0.2 |
| 130. 1988 | 3 | 1 | 1 | 2471.070 .03161 | 1 | 1 | 0 | 0.2 |
| 131. 1989 | 3 | 1 | 1 | 3062.630 .03161 | 1 | 1 | 0 | 0.2 |
| 132. 1990 | 3 | 1 | 1 | 1636.50 .03161 | 1 | 1 | 0 | 0.2 |
| 133. 1991 | 3 | 1 | 1 | 1359.10 .03161 | 1 | 1 | 0 | 0.2 |
| 134. 1992 | 3 | 1 | 1 | 1030.340 .03161 | 1 | 1 | 0 | 0.2 |
| 135. 1993 | 3 | 1 | 1 | 668.560 .03161 | 1 | 1 | 0 | 0.2 |
| 136. 1994 | 3 | 1 | 1 | 1625.550 .03161 | 1 | 1 | 0 | 0.2 |
| 137. 1995 | 3 | 1 | 1 | 1192.20 .03161 | 1 | 1 | 0 | 0.2 |
| 138. 1996 | 3 | 1 | 1 | 1236.740 .03161 | 1 | 1 | 0 | 0.2 |
| 139. 1997 | 3 | 1 | 1 | 1067.440 .03161 | 1 | 1 | 0 | 0.2 |
| 140. 1998 | 3 | 1 | 1 | 934.6310 .03161 | 1 | 1 | 0 | 0.2 |
| 141. 1999 | 3 | 1 | 1 | 1240.370 .03161 | 1 | 1 | 0 | 0.2 |
| 142. 2000 | 3 | 1 | 1 | 1385.080 .03161 | 1 | 1 | 0 | 0.2 |
| 143. 2001 | 3 | 1 | 1 | 1287.630 .03161 | 1 | 1 | 0 | 0.2 |
| 144. 2002 | 3 | 1 | 1 | 1217.190 .03161 | 1 | 1 | 0 | 0.2 |
| 145. 2003 | 3 | 1 | 1 | 1249.260 .03161 | , | 1 | 0 | 0.2 |
| 146. 2004 | 3 | 1 | 1 | 1265.90 .03161 | 1 | , | 0 | 0.2 |
| 147. 2005 | 3 | 1 | 1 | 1237.60 .03161 | 1 | 1 | 0 | 0.2 |
| 148. 2006 | 3 | 1 | 1 | 1055.970 .03161 | 1 | 1 | 0 | 0.2 |
| 149. 2007 | 3 | 1 | 1 | 1241.730 .03161 | 1 | 1 | 0 | 0.2 |
| 150. 2008 | 3 | 1 | 1 | 1218.890 .03161 | 1 | 1 | 0 | 0.2 |
| 151. 2009 | 3 | 1 | 1 | 1348.460 .03161 | 1 | 1 | 0 | 0.2 |
| 152. 2010 | 3 | 1 | 1 | 1353.820 .03161 | 1 | 1 | 0 | 0.2 |
| 153. 2011 | 3 | 1 | 1 | 1349.850 .03161 | 1 | 1 | 0 | 0.2 |
| 154. 2012 | 3 | 1 | 1 | 1420.140 .03161 | , | 1 | 0 | 0.2 |
| 155. 2013 | 3 | 1 | 1 | 1456.440 .03161 | 1 | 1 | 0 | 0.2 |
| 156. 2014 | 3 | 1 | 1 | 1266.430 .03161 | 1 | 1 | 0 | 0.2 |
| 157. 2015 | 3 | 1 | 1 | 1180.310 .03161 | , | 1 | 0 | 0.2 |
| 158. 2016 | 3 | 1 | 1 | 1050.050 .03161 | , | 1 | 0 | 0.2 |
| 159. 2017 | 3 | 1 | 1 | 1054.430 .03161 | I | 1 | 0 | 0.2 |
| 160. 2018 | 3 | 1 | 1 | 1184.490 .03161 | 1 | 1 | 0 | 0.2 |
| 161. 2019 | 3 | 1 | 1 | 1309.40 .03161 | 1 | 1 | 0 | 0.2 |
| 162. 2020 | 3 | 1 | 1 | 1358.320 .03161 | 1 | 1 | 0 | 0.2 |
| 163. 2021 | 3 | 1 | 1 | 1045.710 .03161 | 1 | 1 | 0 | 0.2 |
| 164. 2022 | 3 | 1 | 1 | 822.6170 .03161 | 1 | 1 | 0 | 0.2 |

165.\#
166.\#\# Total Catch (no mortality applied) in tonnes

| 167.\#year |  | fleet | sex | obs cv type | units |  |  | effort | discard_mortality |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 168.\#1990 | 3 | 1 | 1 | 3981.870 .207908162 | 0 | 1 | 1 | 0 | 0.2 |
| 169.\#1991 | 3 | 1 | 1 | 2118.230 .130117339 | 0 | 1 | 1 | 0 | 0.2 |
| 170.\#1992 | 3 | 1 | 1 | 1039.240 .145158877 | 0 | 1 | 1 | 0 | 0.2 |
| 171.\#1993 | 3 | 1 | 1 | 3601.260 .293606298 | 0 | 1 | 1 | 0 | 0.2 |
| 172.\#1994 | 3 | 1 | 1 | 5053.580 .106740581 | 0 | 1 | 1 | 0 | 0.2 |
| 173.\#1995 | 3 | 1 | 1 | 2618.760 .050729769 | 0 | 1 | 1 | 0 | 0.2 |
| 174.\#1996 | 3 | 1 | 1 | 1972.190 .04474373 | 0 | 1 | 1 | 0 | 0.2 |
| 175.\#1997 | 3 | 1 | 1 | 1891.860 .060142043 | 0 | 1 | 1 | 0 | 0.2 |
| 176.\#1998 | 3 | 1 | 1 | 1106.870 .087743921 | 0 | 1 | 1 | 0 | 0.2 |
| 177.\#1999 | 3 | 1 | 1 | 2178.040 .056445899 | 0 | 1 | 1 | 0 | 0.2 |
| 178.\#2000 | 3 | 1 | 1 | 2272.720 .055136691 | 0 | 1 | 1 | 0 | 0.2 |
| 179.\#2001 | 3 | 1 | 1 | 2154.960 .056882141 | 0 | 1 | 1 | 0 | 0.2 |
| 180.\#2002 | 3 | 1 | 1 | 1900.340 .075835543 | 0 | 1 | 1 | 0 | 0.2 |
| 181.\#2003 | 3 | 1 | 1 | 1867.220 .065763685 | 0 | 1 | 1 | 0 | 0.2 |
| 182.\#2004 | 3 | 1 | 1 | 1886.020 .074221412 | 0 | 1 | 1 | 0 | 0.2 |
| 183.\#2005 | 3 | 1 | 1 | 1796.250 .102938855 | 0 | 1 | 1 | 0 | 0.2 |
| 184.\#2006 | 3 | 1 | 1 | 1551.240 .110618989 | 0 | 1 | 1 | 0 | 0.2 |
| 185.\#2007 | 3 | 1 | 1 | 1614.060 .115699625 | 0 | 1 | 1 | 0 | 0.2 |
| 186.\#2008 | 3 | 1 | 1 | 1733.150 .121676438 | 0 | 1 | 1 | 0 | 0.2 |
| 187.\#2009 | 3 | 1 | 1 | 1689.850 .127517984 | 0 | 1 | 1 | 0 | 0.2 |
| 188.\#2010 | 3 | 1 | 1 | 1604.740 .129358495 | 0 | 1 | 1 | 0 | 0.2 |
| 189.\#2011 | 3 | 1 | 1 | 1516.820 .131675973 | 0 | 1 | , | 0 | 0.2 |
| 190.\#2012 | 3 | 1 | 1 | 1839.080 .114273251 | 0 | 1 | 1 | 0 | 0.2 |
| 191.\#2013 | 3 | 1 | 1 | 1919.090 .108784133 | 0 | 1 | 1 | 0 | 0.2 |
| 192.\#2014 | 3 | 1 | 1 | 1592.180 .112848388 | 0 | 1 | 1 | 0 | 0.2 |
| 193.\#2015 | 3 | 1 | 1 | 1565.390 .105658441 | 0 | 1 | 1 | 0 | 0.2 |
| 194.\#2016 | 3 | 1 | 1 | 1569.990 .116902199 | 0 | 1 | 1 | 0 | 0.2 |
| 195.\#2017 | 3 | 1 | 1 | 1436.590 .138245859 | 0 | 1 | 1 | 0 | 0.2 |
| 196.\#2018 | 3 | 1 | 1 | 1637.420 .145371921 | 0 | 1 | 1 | 0 | 0.2 |
| 197.\#2019 | 3 | 1 | 1 | 1714.120 .122434679 | 0 | 1 | 1 | 0 | 0.2 |
| 198.\#2020 | 3 | 1 | 1 | 1844.190 .112848388 | 0 | 1 | 1 | 0 | 0.2 |
| 199.\#2021 | 3 | 1 | 1 | 1611.550 .130040853 | 0 | 1 | 1 | 0 | 0.2 |
| 200.\#2022 | 3 | 1 | 1 | 1067.950 .134290637 | 0 | 1 | 1 | 0 | 0.2 |



| 238. 1991 | 3 | 2 | 1 | 0.0285788 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 239. 1992 | 3 | 2 | 1 | 0.442401 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 240. 1994 | 3 | 2 | 1 | 0.115027 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 241. 1995 | 3 | 2 | 1 | 0.793348 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 242. 1996 | 3 | 2 | 1 | 2.59512 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 243. 1997 | 3 | 2 | 1 | 0.41718 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 244. 1998 | 3 | 2 | 1 | 1.88422 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 245. 1999 | 3 | 2 | 1 | 1.80473 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 246. 2000 | 3 | 2 | 1 | 1.09494 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 247. 2001 | 3 | 2 | 1 | 0.440945 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 248. 2002 | 3 | 2 | 1 | 1.2808 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 249. 2003 | 3 | 2 | 1 | 0.314029 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 250. 2004 | 3 | 2 | 1 | 1.00981 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 251. 2005 | 3 | 2 | 1 | 3.74606 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 252. 2006 | 3 | 2 | 1 | 2.37268 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 253. 2007 | 3 | 2 | 1 | 1.54793 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 254. 2008 | 3 | 2 | 1 | 9.3034 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 255. 2009 | 3 | 2 | 1 | 4.86086 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 256. 2010 | 3 | 2 | 1 | 2.66153 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 257. 2011 | 3 | 2 | 1 | 2.70594 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 258. 2012 | 3 | 2 | 1 | 4.339 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 259. 2013 | 3 | 2 | 1 | 3.7373 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 260. 2014 | 3 | 2 | 1 | 2.65644 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 261. 2015 | 3 | 2 | 1 | 1.54775 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 262. 2016 | 3 | 2 | 1 | 2.36316 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 263. 2017 | 3 | 2 | 1 | 1.47562 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 264. 2018 | 3 | 2 | 1 | 1.35828 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 265. 2019 | 3 | 2 | 1 | 4.1994 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 266. 2020 | 3 | 2 | 1 | 0.807148 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 267. 2021 | 3 | 2 | 1 | 0.736219 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 268. 2022 | 3 | 2 | 1 | 1.57222 | 1.31082 | 1 | 1.538461538 | 0 | 0.65 |
| 269.\# |  |  |  |  |  |  |  |  |  |
| 270.\#\# RELATIVE ABUNDANCE DATA |  |  |  |  |  |  |  |  |  |
| 271.\#\# Units of abundance: $1=$ biomass, $2=$ numbers |  |  |  |  |  |  |  |  |  |
| 272.\#\# Number of relative abundance indicies |  |  |  |  |  |  |  |  |  |
| 273.\#\# sex: $1=$ male; $2=$ female; $0=$ both |  |  |  |  |  |  |  |  |  |
| 274.\#\# maturity: $1=$ immature; $2=$ mature; $0=$ both) |  |  |  |  |  |  |  |  |  |


| 275. |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 276.\# Fishery CPUE index, Observer CPUE index2 |  |  |  |  |  |  |  |  |  |
| 277.3 |  |  |  |  |  |  |  |  |  |
| 278.\# Index Type (1=Selecivity; 2=retention) |  |  |  |  |  |  |  |  |  |
| 279.\# |  |  |  |  |  |  |  |  |  |
| 280.222 |  |  |  |  |  |  |  |  |  |
| 281.\#\# Number of rows in each index |  |  |  |  |  |  |  |  |  |
| 282.42 ( |  |  |  |  |  |  |  |  |  |
| 283.\# Fishery CPUE index NB error in GLM fits on Observer and Fish Tick data |  |  |  |  |  |  |  |  |  |
| 284. \# Sex: $1=$ male, $2=$ female, $0=$ both" << endl; |  |  |  |  |  |  |  |  |  |
| 285.\# Maturity: $1=$ immature, $2=$ mature, $0=$ both |  |  |  |  |  |  |  |  |  |
| 286.\# Units of survey: $1=$ biomass, $2=$ numbers |  |  |  |  |  |  |  |  |  |
| 287.\# |  | server |  | CPUE | index non i | nteraction |  |  |  |
| 288.\# | Index | Year | Seas | fleet | Sex maturity | index cv |  | danceunit | timing |
| 289. 1 | 1995 | 3 | 11 | 0 | 1.163450543 | 0.025857704 | 2 | 0.5 |  |
| 290. 1 | 1996 | 3 | 11 | 0 | 0.97648904 | 0.021511415 | 2 | 0.5 |  |
| 291. 1 | 1997 | 3 | 11 | 0 | 0.981304103 | 0.022469717 | 2 | 0.5 |  |
| 292. 1 | 1998 | 3 | 11 | 0 | 1.086551568 | 0.025374771 | 2 | 0.5 |  |
| 293. 1 | 1999 | 3 | 11 | 0 | 0.907505347 | 0.023769584 | 2 | 0.5 |  |
| 294. 1 | 2000 | 3 | 11 | 0 | 0.839863867 | 0.024411304 | 2 | 0.5 |  |
| 295. 1 | 2001 | 3 | 11 | 0 | 0.816336291 | 0.027530338 | 2 | 0.5 |  |
| 296. 1 | 2002 | 3 | 11 | 0 | 0.914179816 | 0.026441808 | 2 | 0.5 |  |
| 297. 1 | 2003 | 3 | 11 | 0 | 1.157880838 | 0.020575936 | 2 | 0.5 |  |
| 298. 1 | 2004 | 3 | 11 | 0 | 1.253453001 | 0.018745536 | 2 | 0.5 |  |
| 299. 2 | 2005 | 3 | 11 | 0 | 1.244515552 | 0.027573854 | 2 | 0.5 |  |
| 300. 2 | 2006 | 3 | 11 | 0 | 1.16435816 | 0.029929613 | 2 | 0.5 |  |
| 301.2 | 2007 | 3 | 11 | 0 | 1.036778129 | 0.039818759 | 2 | 0.5 |  |
| 302. 2 | 2008 | 3 | 11 | 0 | 1.195645823 | 0.028202204 | 2 | 0.5 |  |
| 303. 2 | 2009 | 3 | 11 | 0 | 1.262396438 | 0.028421542 | 2 | 0.5 |  |
| 304. 2 | 2010 | 3 | 1 | 0 | 1.084726207 | 0.031824727 | 2 | 0.5 |  |
| 305. 2 | 2011 | 3 | 11 | 0 | 1.146452662 | 0.030897779 | 2 | 0.5 |  |
| 306. 2 | 2012 | 3 | 11 | 0 | 1.192790764 | 0.027363345 | 2 | 0.5 |  |
| 307. 2 | 2013 | 3 | 11 | 0 | 0.847917296 | 0.033762195 | 2 | 0.5 |  |
| 308. 2 | 2014 | 3 | 11 | 0 | 0.806940513 | 0.038700423 | 2 | 0.5 |  |
| 309. 2 | 2015 | 3 | 11 | 0 | 0.788264297 | 0.037107615 | 2 | 0.5 |  |
| 310. 2 | 2016 | 3 | 1 | 0 | 0.897549409 | 0.035148646 | 2 | 0.5 |  |
| 311. 2 | 2017 | 3 | 11 | 0 | 1.036609029 | 0.035289449 | 2 | 0.5 |  |


| 312. 2 | 2018 | 3 | 1 | 1 | 0 | 1.299461476 | 0.029798198 | 2 | 0.5 |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 313.2 | 2019 | 3 | 1 | 1 | 0 | 0.991911968 | 0.032201659 | 2 | 0.5 |  |  |
| 314.2 | 2020 | 3 | 1 | 1 | 0 | 0.916059291 | 0.032947352 | 2 | 0.5 |  |  |
| 315.2 | 2021 | 3 | 1 | 1 | 0 | 0.716004296 | 0.048492746 | 2 | 0.5 |  |  |
| 31.2 | 2022 | 3 | 1 | 1 | 0 | 0.700765061 | 0.061628428 | 2 | 0.5 |  |  |
| 317.\# |  |  |  |  |  |  |  |  |  |  |  |
| 318.\# | Observer | CPUE | index | Year:Area | interaction |  |  |  |  |  |  |
| 319.\# | Index | Year | Seas | fleet | Sex | maturity | index cv | abundance unit timing |  |  |  |
| 320.\# | 1 | 1995 | 3 | 1 | 1 | 0 | 1.133652946 | 0.069718757 | 2 | 0.5 |  |
| 321.\# | 1 | 1996 | 3 | 1 | 1 | 0 | 1.000484309 | 0.066359544 | 2 | 0.5 |  |
| 32.\# | 1 | 1997 | 3 | 1 | 1 | 0 | 1.027866305 | 0.064635806 | 2 | 0.5 |  |
| 323.\# | 1 | 1998 | 3 | 1 | 1 | 0 | 1.055034595 | 0.07078737 | 2 | 0.5 |  |
| 324.\# | 1 | 1999 | 3 | 1 | 1 | 0 | 0.929131556 | 0.075595223 | 2 | 0.5 |  |
| 325.\# | 1 | 2000 | 3 | 1 | 1 | 0 | 0.878845 | 0.085619486 | 2 | 0.5 |  |
| 326.\# | 1 | 2001 | 3 | 1 | 1 | 0 | 0.800724588 | 0.150681487 | 2 | 0.5 |  |
| 327.\# | 1 | 2002 | 3 | 1 | 1 | 0 | 0.927286702 | 0.080120543 | 2 | 0.5 |  |
| 328.\# | 1 | 2003 | 3 | 1 | 1 | 0 | 1.124685361 | 0.051779174 | 2 | 0.5 |  |
| 329.\# | 1 | 2004 | 3 | 1 | 1 | 0 | 1.192308203 | 0.048654508 | 2 | 0.5 |  |
| 330.\# | 2 | 2005 | 3 | 1 | 1 | 0 | 1.098557248 | 0.031415881 | 2 | 0.5 |  |
| 331.\# | 2 | 2006 | 3 | 1 | 1 | 0 | 1.2001859 | 0.026348646 | 2 | 0.5 |  |
| 332.\# | 2 | 2007 | 3 | 1 | 1 | 0 | 1.117238516 | 0.031991099 | 2 | 0.5 |  |
| 333.\# | 2 | 2008 | 3 | 1 | 1 | 0 | 1.266812491 | 0.026825118 | 2 | 0.5 |  |
| 334.\# | 2 | 2009 | 3 | 1 | 1 | 0 | 1.942018671 | 0.010706643 | 2 | 0.5 |  |
| 335.\# | 2 | 2010 | 3 | 1 | 1 | 0 | 1.114826837 | 0.032702002 | 2 | 0.5 |  |
| 336.\# | 2 | 2011 | 3 | 1 | 1 | 0 | 1.137488788 | 0.035867277 | 2 | 0.5 |  |
| 337.\# | 2 | 2012 | 3 | 1 | 1 | 0 | 1.158742564 | 0.029153051 | 2 | 0.5 |  |
| 338.\# | 2 | 2013 | 3 | 1 | 1 | 0 | 0.829556069 | 0.054498673 | 2 | 0.5 |  |
| 339.\# | 2 | 2014 | 3 | 1 | 1 | 0 | 0.824706908 | 0.056750517 | 2 | 0.5 |  |
| 340.\# | 2 | 2015 | 3 | 1 | 1 | 0 | 0.779538685 | 0.065705281 | 2 | 0.5 |  |
| 341.\# | 2 | 2016 | 3 | 1 | 1 | 0 | 0.943096351 | 0.042029573 | 2 | 0.5 |  |
| 342.\# | 2 | 2017 | 3 | 1 | 1 | 0 | 1.019779755 | 0.036455146 | 2 | 0.5 |  |
| 343.\# | 2 | 2018 | 3 | 1 | 1 | 0 | 1.213066432 | 0.026018102 | 2 | 0.5 |  |
| 344.\# | 2 | 2019 | 3 | 1 | 1 | 0 | 0.835699771 | 0.061291602 | 2 | 0.5 |  |
| 345.\# | 2 | 2020 | 3 | 1 | 1 | 0 | 0.74582215 | 0.07699689 | 2 | 0.5 |  |
| 346.\# | 2 | 2021 | 3 | 1 | 1 | 0 | 0.660993184 | 0.103407235 | 2 | 0.5 |  |
| 347.\# | 2 | 2022 | 3 | 1 | 1 | 0 | 0.732593435 | 0.083117711 | 2 | 0.5 |  |
| 348.\# |  |  |  |  |  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |  |  |  |  |


$\begin{array}{llllllllllllllll}380.1985 & 3 & 1 & 1 & 1 & 0 & 0 & 45 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.001779 & 0.044090\end{array}$ $\begin{array}{llllllllll}0.139800 & 0.188711 & 0.163604 & 0.152817 & 0.127810 & 0.078532 & 0.048539 & 0.033765 & 0.010007 & 0.010546\end{array}$ $\begin{array}{lllllllllllllll}381.1986 & 3 & 1 & 1 & 1 & 0 & 0 & 23 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.007269 & 0.065322\end{array}$ $\begin{array}{llllllllll}0.139635 & 0.160400 & 0.149577 & 0.152965 & 0.117009 & 0.088900 & 0.048740 & 0.041813 & 0.019731 & 0.008640\end{array}$

| 382. 19873 | 11 | 1 | $0 \quad 0$ | 80.000000 | 0.000000 | 0.000000 | 0.000000 | 0.000000 | 0.005423 | 0.061702 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0.159142 | 0.191295 |  | 0.150047 | 0.1460960 | 0.098283 | 0.081854 | 0.058236 | 0.032244 | 0.011334 | 0.004346 |
| 383. 19883 | 11 | 1 | $0 \quad 0$ | 2860.000000 | 0.000000 | 0.000000 | 0.000000 | 0.000222 | 0.007071 | 0.071516 |
| 0.192850 | 0.244429 |  | 0.184985 | 0.113489 | 0.074476 | 0.043376 | 0.0279010 | 0.014870 | 0.009961 | 0.014855 |
| 384. 19893 | 11 | 1 | 00 | 5130.000014 | 40.000000 | 0.000000 | 0.000013 | 30.000101 | 10.003677 | 0.050292 |
| 0.210811 | 0.266781 |  | 0.169976 | 0.115880 | 0.079394 | 0.050048 | 0.030497 | 0.013281 | 0.0063760 | 0.002859 |
| 385. 19903 | 11 | 1 | 00 | 2050.000017 | 70.000000 | 0.000000 | 0.000000 | 0.000745 | 5.011654 | 0.100154 |
| 0.240280 | 0.225196 |  | 0.161285 | 0.107812 | 0.071549 | 0.041358 | $0.019198 \quad 0$ | 0.0099040 | 0.0063020 | 0.004545 |
| 386. 19913 | 11 | 1 | 00 | 1020.000000 | 0.000020 | 0.000060 | 0.000325 | 50.001316 | 60.007526 | 0.093017 |
| 0.244868 | 0.245519 |  | 0.173825 | 0.105867 | 0.061824 | 0.034887 | $0.017634 \quad 0$ | 0.008044 | $0.003539 \quad 0$ | 0.001730 |
| 387. 19923 | 11 | 1 | $0 \quad 0$ | $76 \quad 0.000000$ | 0.000000 | 0.000032 | 0.000000 | 0.000293 | 0.003852 | 0.060350 |
| 0.197759 | 0.237800 |  | 0.199371 | 0.131348 | 0.080404 | 0.045223 | $0.024627 \quad 0$ | 0.012611 | $0.004512 \quad 0$ | 0.001816 |
| 388. 19933 | 11 | 1 | $0 \quad 0$ | 3780.000000 | 0.000000 | 0.000000 | 0.000000 | 0.000162 | 20.001564 | 0.030740 |
| 0.173187 | 0.272635 |  | 0.232143 | 0.144654 | 0.070916 | 0.043269 | 0.019418 | 0.0082950 | 0.0021750 | 0.000843 |
| 389. 19943 | 11 | 1 | $0 \quad 0$ | 3670.000000 | 0.000000 | 0.000127 | 70.000000 | 0.000024 | $4 \quad 0.001784$ | 0.046992 |
| 0.177933 | 0.253944 |  | 0.198165 | 0.133629 | 0.084930 | 0.053384 | 0.0286820 | 0.013116 | 0.0058720 | 0.001417 |
| 390. 19953 | 11 | 1 | $0 \quad 0$ | 7050.000044 | 0.000000 | 0.000 | 40.000000 | 0.000238 | 0.001792 | 0.047010 |
| 0.189629 | 0.247586 |  | 0.194753 | 0.130339 0 | 0.082564 | 0.051882 | $0.030110 \quad 0$ | $0.014359 \quad 0$ | $0.006334 \quad 0$ | 0.003027 |
| 391. 19963 | 11 | 1 | 00 | 8170.000000 | 0.000000 | 0.000000 | 0.000000 | 0.000000 | 0.003134 | 0.068075 |
| 0.206839 | 0.229461 |  | 0.192799 | 0.127056 | 0.075589 | 0.048745 | 0.028763 | 0.012721 | $0.005090 \quad 0$. | 0.001728 |
| 392. 19973 | 11 | 1 | $0 \quad 0$ | 9840.000000 | 0.000000 | 0.000000 | 0.000000 | 0.000078 | 8.003000 | 0.059292 |
| 0.204909 | 0.240721 |  | 0.186437 | 0.125003 | 0.075450 | 0.047468 | 0.029456 | $0.015703 \quad 0$ | 0.0079070 | 0.004577 |
| 393. 19983 | 11 | 1 | $0 \quad 0$ | 6130.000000 | 0.000000 | 0.000000 | 0.000108 | $8 \quad 0.000000$ | 0.002820 | 0.057800 |
| 0.218632 | 0.266244 |  | 0.203012 | 0.122690 | 0.067337 | 0.033865 | 0.016304 | 0.006846 | $0.002719 \quad 0$ | 0.001625 |
| 394. 19993 | 11 | 1 | $0 \quad 0$ | 9150.000000 | 0.000000 | 0.000000 | 0.000000 | 0.000000 | 0.002046 | 0.043548 |
| 0.178562 | 0.266933 |  | 0.224152 | 0.1399250 | 0.077925 | 0.041227 | 0.017160 | 0.005705 | 0.001947 | 0.000868 |
| 395. 20003 | 11 | 1 | $0 \quad 0$ | 10290.000000 | 0.000000 | 00.000000 | 0.000055 | 50.000000 | 0.002742 | 20.054185 |
| 0.194604 | 0.239411 |  | 0.192948 | 0.137152 | 0.080274 | 0.048645 | 0.026673 | 0.014585 | $0.005697 \quad 0$. | 0.003029 |
| 396. 20013 | 11 | 1 | $0 \quad 0$ | 8980.000000 | 0.000000 | 0.000028 | 0.000000 | 0.000103 | 30.001500 | 0.045630 |
| 0.199267 | 0.265957 |  | 0.214346 | 0.135542 | 0.067145 | 0.038974 | 0.019746 | 0.007561 | 0.0029420 | 0.001260 |
| 397. 20023 | 11 | 1 | $0 \quad 0$ | 6280.000000 | 0.000000 | 0.000000 | 0.000000 | 0.000071 | 0.001663 | 0.058986 |
| 0.218218 | 0.276414 |  | 0.210594 | 0.128255 | 0.057707 | 0.027606 | 0.012857 | 0.005319 | $0.001734 \quad 0$ | 0.000576 |
| 398. 20033 | 11 | 1 | $0 \quad 0$ | 6880.000000 | 0.000000 | 0.000000 | 0.000000 | 0.000000 | 0.001895 | 0.049682 |
| 0.218629 | 0.275161 |  | 0.207658 | 0.127200 | 0.062539 | 0.033554 | 0.014795 | 0.005876 | 0.0019160 | 0.001094 |
| 399. 20043 | 11 | 1 | $0 \quad 0$ | 4490.000000 | 0.000000 | 0.000000 | 0.000000 | 0.000066 | 60.001215 | 0.051327 |
| 0.215030 | 0.283751 |  | 0.206811 | 0.118840 | 0.062679 | 0.034830 | 0.015537 | 0.006658 | 0.0020260 | 0.001232 |

400. 2005
0.172732
401. 20063
0.136486
402. 20073
0.129360
403. 20083
0.124737
404. 20093
0.111051
405. 2010
0.105370
406. 2011
0.112176
407. 20123
0.124402
408. 2013
0.109425
409. 20143
0.180855
410. 20153
0.200801
411. 20163
0.166920
412. 2017
0.132743
413. 20183
0.131101
414. 20193
0.156779
415. 20203
0.152912
416. 20213
0.141352
417. $2022 \quad 3 \quad 1 \quad 1 \quad 1 \quad 0 \quad 0$
$0.173348 \quad 0.247915 \quad 0.215028$
$\begin{array}{lllllllll}337 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.000485 & 0.034904\end{array}$ $\begin{array}{lllllll}0.147408 & 0.082432 & 0.046486 & 0.023695 & 0.016908 & 0.007266 & 0.003203\end{array}$ $\begin{array}{llllllll}337 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.000450 & 0.027936\end{array}$ $\begin{array}{lllllll}0.161660 & 0.101178 & 0.070427 & 0.041574 & 0.022708 & 0.007786 & 0.004229\end{array}$ $\begin{array}{llllllll}276 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.000573 & 0.025731\end{array}$ $\begin{array}{llllllll}0.150598 & 0.112913 & 0.084343 & 0.052940 & 0.026878 & 0.013391 & 0.006285\end{array}$ $\begin{array}{llllllll}318 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.000348 & 0.002243 & 0.027529\end{array}$ $\begin{array}{lllllll}0.164574 & 0.103519 & 0.086701 & 0.053975 & 0.030228 & 0.014680 & 0.005122\end{array}$ $\begin{array}{llllllll}362 & 0.000000 & 0.000000 & 0.000000 & 0.000147 & 0.000000 & 0.000396 & 0.019110\end{array}$ $\begin{array}{lllllll}0.154035 & 0.115423 & 0.089833 & 0.066136 & 0.043036 & 0.021639 & 0.011753\end{array}$ $\begin{array}{llllllll}328 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.000358 & 0.017161\end{array}$ $\begin{array}{lllllll}0.144264 & 0.119206 & 0.095711 & 0.066710 & 0.043825 & 0.021922 & 0.014333\end{array}$ $\begin{array}{lllllllll}295 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.000810 & 0.020228\end{array}$ $\begin{array}{lllllll}0.144850 & 0.118773 & 0.084588 & 0.072303 & 0.044293 & 0.029024 & 0.024552\end{array}$ $\begin{array}{llllllll}288 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.000395 & 0.024683\end{array}$ $\begin{array}{lllllll}0.147574 & 0.100955 & 0.078569 & 0.053528 & 0.034691 & 0.021014 & 0.019392\end{array}$ $\begin{array}{llllllll}327 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.000770 & 0.010442\end{array}$ $\begin{array}{llllllll}0.161229 & 0.106223 & 0.083860 & 0.044186 & 0.032530 & 0.022399 & 0.017097\end{array}$ $\begin{array}{llllllll}305 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.004066 & 0.037110\end{array}$ $\begin{array}{lllllll}0.134732 & 0.098282 & 0.060790 & 0.042639 & 0.025869 & 0.010910 & 0.005522\end{array}$ $\begin{array}{llllllll}287 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.002504 & 0.047044\end{array}$ $\begin{array}{lllllll}0.132490 & 0.077399 & 0.050622 & 0.029030 & 0.014841 & 0.006719 & 0.002275\end{array}$ $\begin{array}{llllllll}408 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.001173 & 0.034989\end{array}$ $\begin{array}{lllllll}0.141610 & 0.084137 & 0.044617 & 0.024752 & 0.012436 & 0.004111 & 0.003987\end{array}$ $\begin{array}{llllllll}309 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.026198\end{array}$ $\begin{array}{lllllll}0.147671 & 0.085251 & 0.066069 & 0.037995 & 0.028973 & 0.013134 & 0.011450\end{array}$ $\begin{array}{llllllll}291 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.000420 & 0.017810\end{array}$ $\begin{array}{lllllll}0.150588 & 0.102787 & 0.078363 & 0.040693 & 0.024314 & 0.011658 & 0.009653\end{array}$ $\begin{array}{lllllllll}363 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.000307 & 0.035200\end{array}$ $\begin{array}{lllllll}0.148367 & 0.099119 & 0.068591 & 0.044726 & 0.021362 & 0.009767 & 0.007780\end{array}$ $\begin{array}{llllllll}462 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.000927 & 0.030215\end{array}$ $\begin{array}{lllllll}0.156945 & 0.094998 & 0.056883 & 0.032933 & 0.015328 & 0.010220 & 0.008659\end{array}$ $\begin{array}{lllllllll}446 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.000295 & 0.037810\end{array}$ $\begin{array}{lllllll}0.144480 & 0.075795 & 0.056631 & 0.026119 & 0.012881 & 0.005879 & 0.004141\end{array}$ $\begin{array}{llllllll}296 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.000000 & 0.00179081 & 0.034329\end{array}$ $\begin{array}{lllllll}0.149247 & 0.080187 & 0.047234 & 0.0279027 & 0.0140501 & 0.00642464 & 0.00254371\end{array}$


418. 3
457.Same as the EAG22.1e2 tagging data......
458........
419. 

460.\#\# eof
461.9999

## WAG22.1e2 prj file

\# References

| \# | = Do not compute MSY (1=Yes) |
| :--- | :--- |
| 0 | \# Set to 0 if F35\% applied to this fleet; 1 if future F is to be fixed |
| 19862016 | \# for Rbar calc, First and last year for average recruitment,/MMB for Bspr calculation (Tier 3 or Tier 4) |
| 19852022 | \# First and last years for average sex ratio |
| 20122022 | \# First and last years for average F for discards |
| 20222022 | \# First and last years for M (0=last year) |
| 20222022 | \# First and last years for proportion of the season |
| 0 | \# Year for specifying growth (0=last year) |
| 20120 | \# First and last years for average selex and discard (0=last year) |

\# OFL specifications
0.35 \# Target SPR ratio for Bmsy proxy.

3 \# Tier
0.10 \# Alpha (cut-off)
0.25 \# Beta (limit)
1.00 \# Gamma
0.75 \# ABC-OFL buffer

0 \# If compute MSY selection is zero, yield function compute slection should be set to zero. Produce a yield curve ( $1=\mathrm{yes} ; 2=\mathrm{no}$ )
\# Projection material
Last year of projection from the terminal (last year data) year
1
01.2

1
2
-3423.8
19862016
\# Number of strategies (0 for no projections)
\# Range of $F$ values
\# 0 for no mortality for non-directed fleets (see input \#1 above); 1=Yes
\# Mcmc replicates per draw
\# Fixed BMSY (negative number for replicate-specific)
\# for Rbar calc, First and last year for average recruitment

| 19852022 | \# First and last years for average sex ratio |
| :---: | :---: |
| 20122022 | \# First and last years for average F for discards |
| 20222022 | \# First and last years for M (0=last year) |
| 20222022 | \# First and last years for proportion of the season |
| 0 | \# Year for specifying growth for projections ( $0=$ last year) |
| 20120 | \# First and last years for average selex and discard (0=last year) |
| 1 | \# Stock-recruitment option (1=Mean Rec;2=Ricker;3=BH;4=Mean and CV) |
| 8 | \# age-at-recruitment |
| \# |  |
| 19602022 | \# First and last years for generating future recruitment (only used if Stock_recruitment option = 1) |
| 2013.292673 | \# Mean recruitment for projections in 1000s |
| 0.35 | \# SigmaR (only used if Stock_recruitment option = 2) |
| 0.2 | \# ProwR |
| -999 | \# Initial eps (first rec_dev, set to -999 to generate it) |
| \#\#\# State strategy |  |
| 0 | \# Apply strategies [OFL, ABC] (1=yes;0=no) |
| 0.001397568 | \# Mean weight (1985-2022) to use (mature in t) |
| 0.001928397 | \# Mean weight (1985-2022) to use (legal in t) |
| \# Stop after XX mcdraws |  |
| 10000 |  |
| \# Full diag |  |
| 0 |  |
| \#\# eof 9999 |  |


[^0]:    \# Number of maturity types
    7 \# Number of size-classes in the model
    \# Season when recruitment occurs,end of year before growth
    \# Season when molting and growth occur, end of year after recruitment
    \# Season to calculate MMB
    \# Season for N output
    \# maximum size-class (males then females)
    17
    \# size_breaks (a vector giving the break points between size intervals with dimension nclass+1, lower limits of bins) 100.5105 .5110 .5115 .5120 .5125 .5130 .5135 .5140 .5145 .5150 .5155 .5160 .5165 .5170 .5175 .5180 .5185 .5
    \# Natural mortality per season input type ( $1=$ vector by season, $2=$ matrix by season/year)
    2
    \# Proportion of the total natural mortality to be applied each season (each row must add to 1)
    \# 1 Start biological year (Jul 1) instantaneous N estimation
    \# 2 to mid fishing time
    \# 3 instantanous C removal
    \# 4 to spawning time
    \# 5 instantaneous byc removal and estimate MMB
    \# 6 Rest of the period of non fishing from Feb 15 to June 30
    \#
    \#
    \#Ins N Jul1-MidFish Inst C MidFish-15Feb Ins byc Rest upto end
    $0.0 .166666670 .0 .46073059 \quad 0.0 .37260274$ \#1960
    0. 0.16666667 0. 0.46073059 0. 0.37260274 \#1961
    0. $0.16666667 \quad 0.00 .46073059 \quad 0.0 .37260274$ \#1962
    0. 0.16666667 0. 0.46073059 0. 0.37260274 \#1963
    0. $0.16666667 \quad 0.00 .46073059 \quad 0.0 .37260274$ \#1964
    0. 0.16666667 0. 0.46073059 0. 0.37260274 \#1965
    0.0 .16666667 0. $0.46073059 \quad 0.0 .37260274$ \#1966
    0. $0.16666667 \quad 0.0 .46073059 \quad 0.0 .37260274$ \#1967
    0. $0.16666667 \quad 0.0 .46073059 \quad 0.0 .37260274$ \#1968
    $\begin{array}{lllllll}0 . & 0.16666667 & 0 . & 0.46073059 & 0 . & 0.37260274 & \# 1969\end{array}$
    0. 0.16666667 0. 0.46073059 0. 0.37260274 \#1970
    0. $0.16666667 \quad 0.0 .46073059 \quad 0.0 .37260274$ \#1971
    0. 0.16666667 0. $0.46073059 \quad 0.0 .37260274$ \#1972
    0. $0.16666667 \quad 0.0 .46073059 \quad 0.0 .37260274$ \#1973
    0. $0.16666667 \quad 0.0 .46073059 \quad 0.0 .37260274$ \#1974

